

# MultivariateMicrostructureModels

## *Empirical Market Microstructure*

(2006, Oxford University Press)

Companion *Mathematica* notebook

Joel Hasbrouck

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This notebook covers material in Chapters 9 and 10

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### ■ Preliminaries

---

#### ■ Notations

```
<< Notation`
```

The following commands define symbolizations that are convenient for labeling things.

```
Symbolize[Anything_Rule]; Symbolize[Anything_Rules];  
Symbolize[Anything_Solution]; Symbolize[Anything_Solutions];
```

```
Symbolize[ $\sigma^2$ ]
```

---

#### ■ Additional initializations specific to this notebook

```
SetAttributes[c, Constant]
```

#### □ *Permutation routine*

The following routine is useful for permuting the order of variables in coefficient and covariance matrices.

```
Permute[m_, i_] := Transpose[Transpose[m[[i]]][[i]]]
```

For example:

```
(vTest = Table[σMin[i,j],Max[i,j], {i, 4}, {j, 4}]) // MatrixForm
```

$$\begin{pmatrix} \sigma_{1,1} & \sigma_{1,2} & \sigma_{1,3} & \sigma_{1,4} \\ \sigma_{1,2} & \sigma_{2,2} & \sigma_{2,3} & \sigma_{2,4} \\ \sigma_{1,3} & \sigma_{2,3} & \sigma_{3,3} & \sigma_{3,4} \\ \sigma_{1,4} & \sigma_{2,4} & \sigma_{3,4} & \sigma_{4,4} \end{pmatrix}$$

```
Permute[vTest, {2, 4, 1, 3}] // MatrixForm
```

$$\begin{pmatrix} \sigma_{2,2} & \sigma_{2,4} & \sigma_{1,2} & \sigma_{2,3} \\ \sigma_{2,4} & \sigma_{4,4} & \sigma_{1,4} & \sigma_{3,4} \\ \sigma_{1,2} & \sigma_{1,4} & \sigma_{1,1} & \sigma_{1,3} \\ \sigma_{2,3} & \sigma_{3,4} & \sigma_{1,3} & \sigma_{3,3} \end{pmatrix}$$

This reverses the permutation:

```
Permute[%, {3, 1, 4, 2}] // MatrixForm
```

$$\begin{pmatrix} \sigma_{1,1} & \sigma_{1,2} & \sigma_{1,3} & \sigma_{1,4} \\ \sigma_{1,2} & \sigma_{2,2} & \sigma_{2,3} & \sigma_{2,4} \\ \sigma_{1,3} & \sigma_{2,3} & \sigma_{3,3} & \sigma_{3,4} \\ \sigma_{1,4} & \sigma_{2,4} & \sigma_{3,4} & \sigma_{4,4} \end{pmatrix}$$

```
(vTest = Table[σMin[i,j],Max[i,j], {i, 3}, {j, 3}]) // MatrixForm
```

$$\begin{pmatrix} \sigma_{1,1} & \sigma_{1,2} & \sigma_{1,3} \\ \sigma_{1,2} & \sigma_{2,2} & \sigma_{2,3} \\ \sigma_{1,3} & \sigma_{2,3} & \sigma_{3,3} \end{pmatrix}$$

```
Permute[%, {3, 1, 2}] // MatrixForm
```

$$\begin{pmatrix} \sigma_{3,3} & \sigma_{1,3} & \sigma_{2,3} \\ \sigma_{1,3} & \sigma_{1,1} & \sigma_{1,2} \\ \sigma_{2,3} & \sigma_{1,2} & \sigma_{2,2} \end{pmatrix}$$

#### □ Matrix polynomial formatting routines

```
PolyForm[m_, L_ : L] :=  
  Plus @@ Table[MatrixForm[Coefficient[m, L, i]] If[i != 0, Li, ""],  
    {i, -Max[Exponent[Flatten[m /. L → 1/L], L]], Max[Exponent[Flatten[m], L]]}];
```

#### □ Vector autocovariance generating function (from VMA representation);

In the following,  $\Theta$  is assumed to be a matrix polynomial in  $L$ ;  $\Omega$  is the covariance matrix. For more detail, see the last section of this notebook.

```
G[θ_, Ω_, z_] := (θ /. L → z) . Ω . Transpose[θ /. L → z-1];
```

## ■ Extended example with autocorrelated trades (Chapter 9)

### ■ Structural Model (Section 9.2)

```
qRule = qt := vt + β vt-1; mRule = mt := mt-1 + wt;
wRule = wt := λ vt + ut; pRule = pt := mt + c qt;
TableForm @ {qRule, mRule, wRule, pRule}
```

```
qt := vt + β vt-1
mt := mt-1 + wt
wt := λ vt + ut
pt := mt + c qt
```

```
ΔpRule = Δpt := (pt /. pRule /. mRule /. wRule /. qRule) - (pt-1 /. pRule /. qRule);
Simplify[Δpt /. ΔpRule]
```

```
ut - c (β v-2+t + v-1+t) + λ vt + c (β v-1+t + vt)
```

```
nValues = {σv2 → .8738, β → 0.381, c → 1, λ → 2}
```

```
{σv2 → 0.8738, β → 0.381, c → 1, λ → 2}
```

□ VMA (not normalized)

The coefficients of current and lagged v<sub>t</sub> in Δp<sub>t</sub> may be tabulated as:

```
Coefficient[Δpt /. ΔpRule, vt-#] & /@ {0, 1, 2}
```

```
{c + λ, -c + c β, -c β}
```

This enables us to write:  $\begin{pmatrix} \Delta p_t \\ q_t \end{pmatrix} = B_0 \begin{pmatrix} u_t \\ v_t \end{pmatrix} + B_1 \begin{pmatrix} u_{t-1} \\ v_{t-1} \end{pmatrix} + B_2 \begin{pmatrix} u_{t-2} \\ v_{t-2} \end{pmatrix}$  where

```
BRules =
{B0 → {{1, λ + c}, {0, 1}}, B1 → {{0, c (β - 1)}, {0, β}}, B2 → {{0, -c β}, {0, 0}}};
{Map[MatrixForm, BRules, {2}]}
```

```
{ {B0 →  $\begin{pmatrix} 1 & c + \lambda \\ 0 & 1 \end{pmatrix}$ , B1 →  $\begin{pmatrix} 0 & c(-1 + \beta) \\ 0 & \beta \end{pmatrix}$ , B2 →  $\begin{pmatrix} 0 & -c\beta \\ 0 & 0 \end{pmatrix}$  } }
```

□ VMA (normalized so that the coefficient of the leading disturbance is the identity matrix)

```
(B0i = Inverse[B0 /. BRules]) // MatrixForm
```

$$\begin{pmatrix} 1 & -c - \lambda \\ 0 & 1 \end{pmatrix}$$

Coefficients:

```
 $\theta_{\text{Rules}} = \text{Table}[\theta_i \rightarrow \text{Simplify}[(B_i /. BRules) . B0i], \{i, 0, 2\}];$   
 $\{\text{Map}[\text{MatrixForm}, \theta_{\text{Rules}}, 2]\}$ 
```

$$\left\{ \left\{ \theta_0 \rightarrow \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}, \theta_1 \rightarrow \begin{pmatrix} 0 & c(-1 + \beta) \\ 0 & \beta \end{pmatrix}, \theta_2 \rightarrow \begin{pmatrix} 0 & -c\beta \\ 0 & 0 \end{pmatrix} \right\} \right\}$$

NOTE: There is a typo in the first printing of the book. The correct expression for  $\theta_1$  is the one given above.

```
 $\theta[L_] := \theta_0 + \theta_1 L + \theta_2 L^2;$ 
```

Disturbances:

```
 $\epsilon_{\text{Rule}} = \epsilon_{t\_} \rightarrow \text{Evaluate}[B_0 . \{\{u_t\}, \{v_t\}\} /. BRules];$   
 $\text{Map}[\text{MatrixForm}, \epsilon_{\text{Rule}}, 1]$ 
```

$$\epsilon_{t\_} \rightarrow \begin{pmatrix} u_t + (c + \lambda) v_t \\ v_t \end{pmatrix}$$

Covariance matrix:

```
 $\Omega_{\text{Rule}} = \Omega \rightarrow \text{Simplify} /@ B_0 . \{\{\sigma_u^2, 0\}, \{0, \sigma_v^2\}\} . \text{Transpose}[B_0] /. BRules;$   
 $\Omega /. \Omega_{\text{Rule}} // \text{MatrixForm}$ 
```

$$\begin{pmatrix} \sigma_u^2 + (c + \lambda)^2 \sigma_v^2 & (c + \lambda) \sigma_v^2 \\ (c + \lambda) \sigma_v^2 & \sigma_v^2 \end{pmatrix}$$

Random walk variance is in the upper left hand corner of

```
 $(\theta[1] . \Omega . \text{Transpose}[\theta[1]]) /. \Omega_{\text{Rule}} /. \theta_{\text{Rules}} // \text{Simplify} // \text{MatrixForm}$ 
```

$$\begin{pmatrix} \sigma_u^2 + \lambda^2 \sigma_v^2 & (1 + \beta) \lambda \sigma_v^2 \\ (1 + \beta) \lambda \sigma_v^2 & (1 + \beta)^2 \sigma_v^2 \end{pmatrix}$$

□ VAR:

Starting from the VMA  $\begin{pmatrix} \Delta p_t \\ q_t \end{pmatrix} = \theta(L) \epsilon_t$ , compute the VAR by inversion, i.e.,  $\theta(L)^{-1} \begin{pmatrix} \Delta p_t \\ q_t \end{pmatrix} = \epsilon_t$ :

```
θInv = Inverse[θ[L] /. θRules] // Simplify; θInv // MatrixForm
```

$$\begin{pmatrix} 1 & \frac{c L (1 + (-1 + L) \beta)}{1 + L \beta} \\ 0 & \frac{1}{1 + L \beta} \end{pmatrix}$$

Series expansion to the fourth order:

```
s = Series[θInv, {L, 0, 4}]; s // MatrixForm
```

$$\begin{pmatrix} 1 & (c - c \beta) L + c \beta^2 L^2 - c \beta^3 L^3 + c \beta^4 L^4 + O[L]^5 \\ 0 & 1 - \beta L + \beta^2 L^2 - \beta^3 L^3 + \beta^4 L^4 + O[L]^5 \end{pmatrix}$$

Rules for VAR coefficient matrices:

```
φRules = Table[φi → (Coefficient[#, L, i] & /@ Normal[s]), {i, 0, 4}];  
{Map[MatrixForm, φRules, {2}]}
```

$$\left\{ \left\{ \phi_0 \rightarrow \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}, \phi_1 \rightarrow \begin{pmatrix} 0 & c - c \beta \\ 0 & -\beta \end{pmatrix}, \phi_2 \rightarrow \begin{pmatrix} 0 & c \beta^2 \\ 0 & \beta^2 \end{pmatrix}, \phi_3 \rightarrow \begin{pmatrix} 0 & -c \beta^3 \\ 0 & -\beta^3 \end{pmatrix}, \phi_4 \rightarrow \begin{pmatrix} 0 & c \beta^4 \\ 0 & \beta^4 \end{pmatrix} \right\} \right\}$$

This VAR is of the form  $(I + \phi_1 L + \phi_2 L^2 + \dots) \begin{pmatrix} \Delta p_t \\ q_t \end{pmatrix} = \epsilon_t$ . Alternatively,

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \dots + \epsilon_t \text{ where } Y_t \equiv \begin{pmatrix} \Delta p_t \\ q_t \end{pmatrix}.$$

```
φRules = Table[φi → -φi /. φRules, {i, 1, 4}];  
{Map[MatrixForm, φRules, {2}]}
```

$$\left\{ \left\{ \phi_1 \rightarrow \begin{pmatrix} 0 & -c + c \beta \\ 0 & \beta \end{pmatrix}, \phi_2 \rightarrow \begin{pmatrix} 0 & -c \beta^2 \\ 0 & -\beta^2 \end{pmatrix}, \phi_3 \rightarrow \begin{pmatrix} 0 & c \beta^3 \\ 0 & \beta^3 \end{pmatrix}, \phi_4 \rightarrow \begin{pmatrix} 0 & -c \beta^4 \\ 0 & -\beta^4 \end{pmatrix} \right\} \right\}$$

```
{Map[MatrixForm, φRules /. nValues, {2}]}
```

$$\left\{ \left\{ \phi_1 \rightarrow \begin{pmatrix} 0 & -0.619 \\ 0 & 0.381 \end{pmatrix}, \phi_2 \rightarrow \begin{pmatrix} 0 & -0.145161 \\ 0 & -0.145161 \end{pmatrix}, \right. \right. \\ \left. \left. \phi_3 \rightarrow \begin{pmatrix} 0 & 0.0553063 \\ 0 & 0.0553063 \end{pmatrix}, \phi_4 \rightarrow \begin{pmatrix} 0 & -0.0210717 \\ 0 & -0.0210717 \end{pmatrix} \right\} \right\}$$

#### □ Impulse response functions

The VAR is of infinite order. We'll only be looking at the leading four terms:

$$\mathbf{VAR}_{\text{Rule}} = Y_t \Rightarrow \sum_{k=1}^4 \phi_k \cdot Y_{t-k} + \epsilon_t; Y_t /. \mathbf{VAR}_{\text{Rule}}$$

$$\phi_1 \cdot Y_{-1+t} + \phi_2 \cdot Y_{-2+t} + \phi_3 \cdot Y_{-3+t} + \phi_4 \cdot Y_{-4+t} + \epsilon_t$$

(The "." indicates matrix multiplication.)

The following rules set  $\epsilon_t = 0$  for  $t \neq 0$ ,  $y_t = 0$  for  $t < 0$ , and zero out matrix products involving zero.

```
CleanupRules = {y0 -> e0, yk_ -> 0 /; k < 0, ek_ -> 0 /; k != 0, Dot[_ , 0] -> 0};
```

Subsequent to an initial disturbance of  $\epsilon_0$ , the initial realization of  $y$  is:

```
y0 /. CleanupRules
```

```
e0
```

One period later ...

```
y1 /. VARRule /. CleanupRules
```

```
phi1.e0
```

Two periods later ...

```
y2 /. VARRule /. VARRule /. CleanupRules
```

```
phi2.e0 + phi1.phi1.e0
```

The following rule obtains the impulse response function by automating the recursion

```
psiRule = psi_t_ -> (FixedPoint[# /. VARRule &, y_t, t] /. CleanupRules);
```

```
psi2 /. psiRule
```

```
phi2.e0 + phi1.phi1.e0
```

```
psi2 /. psiRule /. phiRules /. eRule // Simplify
```

```
{{-c beta v0}, {0}}
```

The cumulative IR function is:

$$\Psi_{\text{Rule}} = \Psi_{t\_} \rightarrow \sum_{i=0}^t \psi_i;$$

If we substitute in the "correct" structural specification for  $\epsilon_0$ , we get the "correct" long-run price change.

```
psi2 /. PsiRule /. psiRule /. phiRules /. eRule // Simplify // MatrixForm
```

$$\begin{pmatrix} u_0 + \lambda v_0 \\ (1 + \beta) v_0 \end{pmatrix}$$

In practice, we can't identify  $u_0$  and  $v_0$ , we we're left with:

```
εStatisticalRule = εt → {{εΔp,t}, {εq,t}}; ε0 /. εStatisticalRule // MatrixForm
```

$$\begin{pmatrix} \epsilon_{\Delta p,0} \\ \epsilon_{q,0} \end{pmatrix}$$

And the cumulative impulse response function becomes:

```
Ψ2 /. ΨRule /. ψRule /. φRules /. εStatisticalRule // Simplify // MatrixForm
```

$$\begin{pmatrix} -c \epsilon_{q,0} + \epsilon_{\Delta p,0} \\ (1 + \beta) \epsilon_{q,0} \end{pmatrix}$$

## ■ Contemporaneous effects and Cholesky decompositions (Section 9.4)

A note on notation and conventions. For a symmetric positive definite matrix  $V$ , the Cholesky decomposition is usually expressed as  $V = L L'$  where  $L$  is lower triangular. In some references, though, and in *Mathematica*, the convention is  $V = U' U$  where  $U$  is upper triangular. The book follows the latter (*Mathematica*) convention.

As an example, here's a 2 x2 covariance matrix

```
V = {{σ12, σ12}, {σ12, σ22}};  
V // MatrixForm
```

$$\begin{pmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{pmatrix}$$

```
F = Simplify[CholeskyDecomposition[V], {σ12 > 0, σ22 > 0, -√σ12 σ22 < σ12 < √σ12 σ22}] ;  
F // MatrixForm
```

$$\begin{pmatrix} \sqrt{\sigma_1^2} & \frac{\sigma_{12}}{\sqrt{\sigma_1^2}} \\ 0 & \sqrt{\sigma_2^2 - \frac{\sigma_{12}^2}{\sigma_1^2}} \end{pmatrix}$$

Verify that this recreates the original covariance matrix:

```
Transpose[F].F // Simplify // MatrixForm
```

$$\begin{pmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{pmatrix}$$

Here is the structural covariance matrix:

```
Map[MatrixForm,  $\Omega_{\text{Rule}}$ , 1]
```

$$\Omega \rightarrow \begin{pmatrix} \sigma_u^2 + (c + \lambda)^2 \sigma_v^2 & (c + \lambda) \sigma_v^2 \\ (c + \lambda) \sigma_v^2 & \sigma_v^2 \end{pmatrix}$$

Its Cholesky decomposition with price changes "first" is:

```
F = Simplify[CholeskyDecomposition[ $\Omega$  /.  $\Omega_{\text{Rule}}$ ], {c > 0,  $\lambda$  > 0,  $\sigma_u^2$  > 0,  $\sigma_v^2$  > 0}];  
Transpose[F] // MatrixForm
```

$$\begin{pmatrix} \sqrt{\sigma_u^2 + (c + \lambda)^2 \sigma_v^2} & 0 \\ \frac{(c + \lambda) \sigma_v^2}{\sqrt{\sigma_u^2 + (c + \lambda)^2 \sigma_v^2}} & \sqrt{\frac{\sigma_u^2 \sigma_v^2}{\sigma_u^2 + (c + \lambda)^2 \sigma_v^2}} \end{pmatrix}$$

This implies that a one unit shock to  $\epsilon_{\Delta p, t}$  will have a contemporaneous effect on trades. This is inconsistent with the structural model. If we compute the Cholesky decomposition with "trades" first (by permutation) we have:

```
F = Permute[Simplify[CholeskyDecomposition[Permute[ $\Omega$  /.  $\Omega_{\text{Rule}}$ , {2, 1}]],  
{c > 0,  $\lambda$  > 0,  $\sigma_u^2$  > 0,  $\sigma_v^2$  > 0}], {2, 1}]; Transpose[F] // MatrixForm
```

$$\begin{pmatrix} \sqrt{\sigma_u^2} & (c + \lambda) \sqrt{\sigma_v^2} \\ 0 & \sqrt{\sigma_v^2} \end{pmatrix}$$

```
Transpose[F].{{0}, {1}} // MatrixForm
```

$$\begin{pmatrix} (c + \lambda) \sqrt{\sigma_v^2} \\ \sqrt{\sigma_v^2} \end{pmatrix}$$

## ■ Random-walk variance

Starting from the VMA  $\begin{pmatrix} \Delta p_t \\ q_t \end{pmatrix} = \theta(L) \epsilon_t$ , compute the VAR by inversion, i.e.,  $\theta(L)^{-1} \begin{pmatrix} \Delta p_t \\ q_t \end{pmatrix} = \epsilon_t$ :

$$\begin{pmatrix} \sigma_w^2 & \dots \\ \vdots & \square \end{pmatrix} = \theta(1) \Omega \theta(1)'$$

$\theta$  is identified (or get  $\theta(1) = \varphi(1)^{-1}$ ).

```
 $\theta[1]$  /.  $\theta_{\text{Rules}}$  // Simplify // MatrixForm
```

$$\begin{pmatrix} 1 & -c \\ 0 & 1 + \beta \end{pmatrix}$$



```
{{1, -c}}.(Ω /. ΩRule).Transpose[{{1, -c}}] // Simplify // MatrixForm
```

$$\left( \sigma_u^2 + \lambda^2 \sigma_v^2 \right)$$

In practice, we'd be working from an estimated covariance matrix:

```
{{1, -c}}.V.Transpose[{{1, -c}}] // Simplify // MatrixForm
```

$$\left( \sigma_1^2 + c^2 \sigma_2^2 - 2 c \sigma_{12} \right)$$

Resolve contemporaneous effects with Cholesky decompositions.

□ "Prices first"

```
F = Simplify[CholeskyDecomposition[Ω /. ΩRule], {c > 0, λ > 0, σu2 > 0, σv2 > 0}];  
Transpose[F] // MatrixForm
```

$$\begin{pmatrix} \sqrt{\sigma_u^2 + (c + \lambda)^2 \sigma_v^2} & 0 \\ \frac{(c + \lambda) \sigma_v^2}{\sqrt{\sigma_u^2 + (c + \lambda)^2 \sigma_v^2}} & \sqrt{\frac{\sigma_u^2 \sigma_v^2}{\sigma_u^2 + (c + \lambda)^2 \sigma_v^2}} \end{pmatrix}$$

```
d = First @ {{1, -c}}.Transpose[F] // Simplify
```

$$\left\{ \frac{\sigma_u^2 + \lambda (c + \lambda) \sigma_v^2}{\sqrt{\sigma_u^2 + (c + \lambda)^2 \sigma_v^2}}, -c \sqrt{\frac{\sigma_u^2 \sigma_v^2}{\sigma_u^2 + (c + \lambda)^2 \sigma_v^2}} \right\}$$

```
Plus @@ (d^2) // Simplify
```

$$\sigma_u^2 + \lambda^2 \sigma_v^2$$

□ Trades "first"

```
F = Permute[Simplify[CholeskyDecomposition[Permute[Ω /. ΩRule, {2, 1}]],  
{c > 0, λ > 0, σu2 > 0, σv2 > 0}], {2, 1}]; Transpose[F] // MatrixForm
```

$$\begin{pmatrix} \sqrt{\sigma_u^2} & (c + \lambda) \sqrt{\sigma_v^2} \\ 0 & \sqrt{\sigma_v^2} \end{pmatrix}$$

```
d = First @ {{1, -c}}.Transpose[F] // Simplify
```

$$\left\{ \sqrt{\sigma_u^2}, \lambda \sqrt{\sigma_v^2} \right\}$$

```
Plus @@ (d^2) // Simplify
```

$$\sigma_u^2 + \lambda^2 \sigma_v^2$$

## ■ Pricing error

Structural model:

```
P_t - m_t /. PRule
```

```
c q_t
```

$$\Rightarrow \text{Var}(s_t) = c^2 \text{Var}(q_t) = c^2 (1 + \beta)^2 \sigma_v^2$$

Lower bound

```
b = (theta_1 + theta_2) + theta_2
```

```
theta_1 + 2 theta_2
```

```
b /. thetaRules // Simplify // MatrixForm
```

$$\begin{pmatrix} 0 & -c(1+\beta) \\ 0 & \beta \end{pmatrix}$$

```
(sv = b.O.Transpose[b] /. OmegaRule /. thetaRules // Simplify) // MatrixForm
```

$$\begin{pmatrix} (c + c\beta)^2 \sigma_v^2 & -c\beta(1+\beta) \sigma_v^2 \\ -c\beta(1+\beta) \sigma_v^2 & \beta^2 \sigma_v^2 \end{pmatrix}$$

```
sv /. nValues // MatrixForm
```

$$\begin{pmatrix} 1.66648 & -0.459759 \\ -0.459759 & 0.126842 \end{pmatrix}$$

## ■ Problems

### ■ Exercise 9.1 (Glosten and Harris)

Definitions:

```
mRule = m_t_ -> m_{t-1} + w_t ;
```

```
wRule = w_t_ -> lambda_0 q_t + lambda_1 Q_t + u_t ;
```

```
PRule = p_t_ -> m_t + c_1 q_t + c_2 Q_t ;
```

With these definitions,  $m_t$ ,  $w_t$  and  $p_t$  are (respectively):

```
TableForm[{ m_t /. m_Rule, w_t /. w_Rule, p_t /. p_Rule}]
```

$$\begin{aligned} & m_{-1+t} + w_t \\ & u_t + q_t \lambda_0 + Q_t \lambda_1 \\ & m_t + c_1 q_t + c_2 Q_t \end{aligned}$$

The price change at time  $t$  is  $\Delta p_t =$

```
 $\Delta p_{Rule} = \Delta p_t \rightarrow (p_t /. p_{Rule} /. m_{Rule} /. w_{Rule}) - (p_{t-1} /. p_{Rule});$   
 $\text{Simplify}[\Delta p_t /. \Delta p_{Rule}]$ 
```

$$c_1 (-q_{-1+t} + q_t) + c_2 (-Q_{-1+t} + Q_t) + u_t + q_t \lambda_0 + Q_t \lambda_1$$

The vector of variables is:

```
y_Rule = y_t_ -> Transpose[{{Delta p_t, q_t, Q_t}}]; MatrixForm[y_t /. y_Rule]
```

$$\begin{pmatrix} \Delta p_t \\ q_t \\ Q_t \end{pmatrix}$$

and the disturbances are  $\epsilon_t =$

```
 $\epsilon_{Rule} = \epsilon_t \rightarrow \text{Transpose}[\{u_t, q_t, Q_t\}]; \text{MatrixForm}[\epsilon_t /. \epsilon_{Rule}]$ 
```

$$\begin{pmatrix} u_t \\ q_t \\ Q_t \end{pmatrix}$$

With substitutions, the vector of system variables is:

```
(y_t /. y_Rule /. Delta p_Rule) // MatrixForm // Simplify
```

$$\begin{pmatrix} c_1 (-q_{-1+t} + q_t) + c_2 (-Q_{-1+t} + Q_t) + u_t + q_t \lambda_0 + Q_t \lambda_1 \\ q_t \\ Q_t \end{pmatrix}$$

The MA model is  $y_t = \theta_0 \epsilon_t + \theta_1 \epsilon_{t-1}$  where the coefficient matrices are:

```
ClearAll[theta];  
thetaRules = Table[Rule[theta_s, Table[Coefficient[(y_t /. y_Rule /. Delta p_Rule)[[i, 1]],  
      (epsilon_t - s /. epsilon_Rule)[[j, 1]], {i, 3}, {j, 3}], {s, 0, 1}];  
{Map[MatrixForm, thetaRules, {2}]}]
```

$$\left\{ \left\{ \theta_0 \rightarrow \begin{pmatrix} 1 & c_1 + \lambda_0 & c_2 + \lambda_1 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix}, \theta_1 \rightarrow \begin{pmatrix} 0 & -c_1 & -c_2 \\ 0 & 0 & 0 \\ 0 & 0 & 0 \end{pmatrix} \right\} \right\}$$

To build the covariance matrix, we need to work out  $\text{Cov}(q_t, Q_t) = E[Q_t \text{Sign}(Q_t)] = E|Q_t|$

```
ClearAll[Omega]
```

If  $Q_t \sim N(0, \sigma_Q^2)$ , then  $E[|Q|] =$

$$\text{Simplify}\left[\int_{-\infty}^{\infty} \text{Abs}[x] \text{PDF}[\text{NormalDistribution}[0, \sigma_Q], x] dx, \{\sigma_Q \in \text{Reals}, \text{Re}[(\sigma_Q)^2] > 0\}\right]$$

$$\sqrt{\frac{2}{\pi}} \sigma_Q$$

So the disturbance covariance matrix is:

$$\Omega_{\text{Rule}} = \Omega \rightarrow \left( \left\{ \{\sigma_u^2, 0, 0\}, \{0, 1, \sigma_{q,Q}\}, \{0, \sigma_{q,Q}, \sigma_Q^2\} \right\} / . \sigma_{q,Q} \rightarrow \sqrt{2 \sigma_Q^2 / \pi} \right);$$

$$\text{Map}[\text{MatrixForm}, \Omega_{\text{Rule}}, 1]$$

$$\Omega \rightarrow \begin{pmatrix} \sigma_u^2 & 0 & 0 \\ 0 & 1 & \sqrt{\frac{2}{\pi}} \sqrt{\sigma_Q^2} \\ 0 & \sqrt{\frac{2}{\pi}} \sqrt{\sigma_Q^2} & \sigma_Q^2 \end{pmatrix}$$

The random-walk variance is  $\sigma_w^2 =$

$$\text{Simplify}\left[(\theta_0 + \theta_1) . \Omega . \text{Transpose}[\theta_0 + \theta_1] / . \theta_{\text{Rules}} / . \Omega_{\text{Rule}}, \{\sigma_Q > 0, \sigma_- \in \text{Reals}\}\right][[1, 1]]$$

$$\sigma_u^2 + \lambda_0^2 + 2 \sqrt{\frac{2}{\pi}} \sqrt{\sigma_Q^2} \lambda_0 \lambda_1 + \sigma_Q^2 \lambda_1^2$$

□ *Decomposition with  $q_t$  first:*

$$\text{F1} = \text{Simplify}\left[\text{Permute}[\text{CholeskyDecomposition}[\text{Permute}[\Omega / . \Omega_{\text{Rule}}, \{2, 3, 1\}]], \{3, 1, 2\}], \{\sigma_Q^2 > 0, \sigma_u^2 > 0\}\right];$$

$$\text{F1} // \text{MatrixForm}$$

$$\begin{pmatrix} \sqrt{\sigma_u^2} & 0 & 0 \\ 0 & 1 & \sqrt{\frac{2}{\pi}} \sqrt{\sigma_Q^2} \\ 0 & 0 & \sqrt{\frac{-2+\pi}{\pi}} \sqrt{\sigma_Q^2} \end{pmatrix}$$

$$\text{VarDecomp1} = ((\theta_0 + \theta_1) . \text{Transpose}[\text{F1}] / . \theta_{\text{Rules}} // \text{Simplify})[[1]]$$

$$\left\{ \sqrt{\sigma_u^2}, \lambda_0 + \sqrt{\frac{2}{\pi}} \sqrt{\sigma_Q^2} \lambda_1, \sqrt{\frac{-2+\pi}{\pi}} \sqrt{\sigma_Q^2} \lambda_1 \right\}$$

The variance components corresponding to  $u_t$ ,  $q_t$  and  $Q_t$  are:

```
{VarDecomp12 // Simplify} // TableForm
```

$$\sigma_u^2 \left( \lambda_0 + \sqrt{\frac{2}{\pi}} \sqrt{\sigma_Q^2} \lambda_1 \right)^2 - \frac{(-2+\pi) \sigma_Q^2 \lambda_1^2}{\pi}$$

Verify that they add up to the correct  $\sigma_w^2$ :

```
Plus @@ VarDecomp12 // Simplify
```

$$\sigma_u^2 + \lambda_0^2 + 2 \sqrt{\frac{2}{\pi}} \sqrt{\sigma_Q^2} \lambda_0 \lambda_1 + \sigma_Q^2 \lambda_1^2$$

□ *Decomposition with  $Q_t$  first:*

```
F2 = Simplify[Permute[CholeskyDecomposition[Permute[Ω /. ΩRule, {3, 2, 1}]],  
  {3, 2, 1}], {σ_Q > 0, σ_u > 0}];  
F2 // MatrixForm
```

$$\begin{pmatrix} \sqrt{\sigma_u^2} & 0 & 0 \\ 0 & \sqrt{\frac{-2+\pi}{\pi}} & 0 \\ 0 & \sqrt{\frac{2}{\pi}} & \sqrt{\sigma_Q^2} \end{pmatrix}$$

```
VarDecomp2 = ((θ0 + θ1).Transpose[F2] /. θRules // Simplify)[[1]]
```

$$\left\{ \sqrt{\sigma_u^2}, \sqrt{\frac{-2+\pi}{\pi}} \lambda_0, \sqrt{\frac{2}{\pi}} \lambda_0 + \sqrt{\sigma_Q^2} \lambda_1 \right\}$$

```
{VarDecomp22} // Simplify // TableForm
```

$$\sigma_u^2 - \frac{(-2+\pi) \lambda_0^2}{\pi} \left( \sqrt{\frac{2}{\pi}} \lambda_0 + \sqrt{\sigma_Q^2} \lambda_1 \right)^2$$

... and verify:

```
Plus @@ VarDecomp22 // Simplify
```

$$\sigma_u^2 + \lambda_0^2 + 2 \sqrt{\frac{2}{\pi}} \sqrt{\sigma_Q^2} \lambda_0 \lambda_1 + \sigma_Q^2 \lambda_1^2$$

## ■ Exercise 9.2 (Madhavan, Richardson and Roomans)

The model is:

```
qRule = q_t_ -> v_t + beta q_{t-1};
mRule = m_t_ -> m_{t-1} + w_t;
wRule = w_t_ -> lambda v_t + u_t;
pRule = p_t_ -> m_t + c q_t;
```

With these rules,  $q_t$ ,  $m_t$ ,  $w_t$ , and  $p_t$  are:

```
{q_t /. qRule, m_t /. mRule, w_t /. wRule, p_t /. pRule} // TableForm

beta q_{-1+t} + v_t
m_{-1+t} + w_t
u_t + lambda v_t
m_t + c q_t
```

The price change is  $\Delta p_t =$

```
Delta pRule = Delta p_t_ -> Delta p_t -> (p_t /. pRule /. mRule /. wRule) - (p_{t-1} /. pRule);
Simplify[Delta p_t /. Delta pRule]

Delta p_t -> -c q_{-1+t} + c q_t + u_t + lambda v_t
```

The system variables are  $y_t =$

```
yRule = y_t_ -> Transpose[{{Delta p_t, q_t}}]; y_t /. yRule // MatrixForm

( Delta p_t
  q_t )
```

The vector of disturbances is:

```
eRule = e_t_ -> Transpose[{{u_t, v_t}}]; e_t /. eRule // MatrixForm

( u_t
  v_t )

ClearAll[0]
```

With substitutions, the vector of system variables becomes  $y_t =$

```
{{Delta p_t}, {q_t}} /. Delta p_t -> (p_t /. pRule /. mRule /. wRule) - (p_{t-1} /. pRule) /. q_t -> (q_t /. qRule) //
Simplify // MatrixForm

( c (-1 + beta) q_{-1+t} + u_t + (c + lambda) v_t
  beta q_{-1+t} + v_t )
```

This is a first-order vector autoregressive process:  $Y_t = \phi Y_{t-1} + \theta \epsilon_t$  where

```
 $\phi_{Rule} = \phi \rightarrow \{\{0, c(-1 + \beta)\}, \{0, \beta\}\}; MatrixForm /@ \phi_{Rule}$ 
```

$$\phi \rightarrow \begin{pmatrix} 0 & c(-1 + \beta) \\ 0 & \beta \end{pmatrix}$$

and

```
 $\theta_{Rule} = \theta \rightarrow \{\{1, (c + \lambda)\}, \{0, 1\}\}; MatrixForm /@ \theta_{Rule}$ 
```

$$\theta \rightarrow \begin{pmatrix} 1 & c + \lambda \\ 0 & 1 \end{pmatrix}$$

It may be put in vector moving average (VMA) form as:  $Y_t = \underbrace{(\mathbf{I} - \phi L)^{-1} \theta}_{\text{VMA coefficients}} \epsilon_t$ . We could obtain the VMA

coefficient matrices by doing the series expansion. Here, though, to compute the random-walk variance, we just need the sum of the moving average coefficients, and  $(\mathbf{I} - \phi)^{-1} \theta =$

```
maSum = Inverse[IdentityMatrix[2] - (\phi /. \phi_{Rule})].(\theta /. \theta_{Rule}) // Simplify;
maSum // MatrixForm
```

$$\begin{pmatrix} 1 & \lambda \\ 0 & \frac{1}{1-\beta} \end{pmatrix}$$

To compute the random-walk variance  $\sigma_w^2$ , take the upper left hand entry of  $\left[ (\mathbf{I} - \phi)^{-1} \theta \right] \Omega \left[ (\mathbf{I} - \phi)^{-1} \theta \right]'$

```
maSum.{\{\sigma_u^2, 0\}, \{0, \sigma_v^2\}}.Transpose[maSum] // Simplify // MatrixForm
```

$$\begin{pmatrix} \sigma_u^2 + \lambda^2 \sigma_v^2 & \frac{\lambda \sigma_v^2}{1-\beta} \\ \frac{\lambda \sigma_v^2}{1-\beta} & \frac{\sigma_v^2}{(-1+\beta)^2} \end{pmatrix}$$

The quantity  $\sigma_u^2 + \lambda^2 \sigma_v^2$  summarizes the public information and trade-related components of the random-walk.

## ■ Multiple Securities (Chapter 10)

### ■ Two securities in a Roll model, not cointegrated, prices only. (Chapter 10, Section 10.1)

#### □ Model

```

PRule = Pi,t := mi,t + c qi,t;
mRule = mi,t := mi,t-1 + wi,t;
wRule = wi,t := ui,t;
ΔPRule = ΔPi,t := (Pi,t / PRule / mRule / wRule) - (Pi,t-1 / PRule);
ΔPt := { {ΔP1,t}, {ΔP2,t} };

```

```

ΔPt / ΔPRule

```

```

{ {-c q1,-1+t + c q1,t + u1,t}, {-c q2,-1+t + c q2,t + u2,t} }

```

#### □ Compute autocovariances

We will obtain  $\Gamma_k = E\Delta P_t \Delta P_{t-k}$ , where the following rules are used to eliminate terms in the expectation with zero cross-products. (Note: these are general rules; some of them are left over from the univariate case.)

```

εRules = { ε[q2] → 1, ε[u2] → σu2, ε[q-u-] → 0, ε[qtqs] := 0 /; t != s,
  ε[utus] := 0 /; t != s, ε[q-u-] → 0, ε[u2-] := σu2, ε[u-us] → 0,
  ε[q-us] := 0 /; t == s, ε[q-us] := 0 /; t != s, ε[q2-] → 1,
  ε[a- + b-] := ε[a] + ε[b],
  ε[c-] := c /; NumberQ[c],
  ε[c_Symbol] := c /; MemberQ[Attributes[c], Constant],
  ε[c a-] := c ε[a] /; NumberQ[c],
  ε[c_Symbol a-] := c ε[a] /; MemberQ[Attributes[c], Constant],
  ε[c_Symboln_Integer a-] := cn ε[a] /; MemberQ[Attributes[c], Constant] };

```



$\mathcal{E}_{\text{Rules}}$  // TableForm

```

 $\mathcal{E}[q_-^2] \rightarrow 1$ 
 $\mathcal{E}[u_-^2] \rightarrow \sigma_u^2$ 
 $\mathcal{E}[q_- u_-] \rightarrow 0$ 
 $\mathcal{E}[q_{s_-} q_{t_-}] \rightarrow 0 \text{ ; } t \neq s$ 
 $\mathcal{E}[u_{s_-} u_{t_-}] \rightarrow 0 \text{ ; } t \neq s$ 
 $\mathcal{E}[q_{-,s_-} u_{-,t_-}] \rightarrow 0$ 
 $\mathcal{E}[u_{-,s_-}^2] \rightarrow \sigma_u^2$ 
 $\mathcal{E}[u_{-,s_-} u_{-,t_-}] \rightarrow 0$ 
 $\mathcal{E}[q_{-,s_-} q_{-,t_-}] \rightarrow \rho \text{ ; } t == s$ 
 $\mathcal{E}[q_{-,s_-} q_{-,t_-}] \rightarrow 0 \text{ ; } t \neq s$ 
 $\mathcal{E}[q_{-,s_-}^2] \rightarrow 1$ 
 $\mathcal{E}[a_- + b_-] \rightarrow \mathcal{E}[a_-] + \mathcal{E}[b_-]$ 
 $\mathcal{E}[c_-] \rightarrow c \text{ ; } \text{NumberQ}[c]$ 
 $\mathcal{E}[c\_Symbol] \rightarrow c \text{ ; } \text{MemberQ}[\text{Attributes}[c], \text{Constant}]$ 
 $\mathcal{E}[a_- c_-] \rightarrow c \mathcal{E}[a_-] \text{ ; } \text{NumberQ}[c]$ 
 $\mathcal{E}[a_- c\_Symbol] \rightarrow c \mathcal{E}[a_-] \text{ ; } \text{MemberQ}[\text{Attributes}[c], \text{Constant}]$ 
 $\mathcal{E}[a_- c\_Symbol^{n\_Integer}] \rightarrow c^n \mathcal{E}[a_-] \text{ ; } \text{MemberQ}[\text{Attributes}[c], \text{Constant}]$ 

```

SetAttributes[ $\mathcal{E}$ , Listable]

$\Gamma_{\text{Rule}} = \Gamma_{k_-} \rightarrow (\mathcal{E}[\text{Expand}[(\Delta p_t / \Delta p_{\text{Rule}}) \cdot \text{Transpose}[\Delta p_{t-k} / \Delta p_{\text{Rule}}]]]) // \mathcal{E}_{\text{Rules}};$

$\Gamma_0 / \Gamma_{\text{Rule}}$  // MatrixForm

$$\begin{pmatrix} 2c^2 + \sigma_u^2 & 2c^2\rho \\ 2c^2\rho & 2c^2 + \sigma_u^2 \end{pmatrix}$$

$\Gamma_0$ ,  $\Gamma_1$  and  $\Gamma_2$  are:

Map[MatrixForm, Table[ $\Gamma_k / \Gamma_{\text{Rule}}$ , {k, 0, 2}], 1]

$$\left\{ \begin{pmatrix} 2c^2 + \sigma_u^2 & 2c^2\rho \\ 2c^2\rho & 2c^2 + \sigma_u^2 \end{pmatrix}, \begin{pmatrix} -c^2 & -c^2\rho \\ -c^2\rho & -c^2 \end{pmatrix}, \begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix} \right\}$$

Note that  $\Gamma_2 = 0$ . We can verify that  $\Gamma_k = 0$  for  $|k| \geq 2$ .

Here are some trial numerical values:

nValues = { $\sigma_u^2 \rightarrow 1$ ,  $c \rightarrow 10$ ,  $\rho \rightarrow .9$ }

{ $\sigma_u^2 \rightarrow 1$ ,  $c \rightarrow 10$ ,  $\rho \rightarrow 0.9$ }

Map[MatrixForm, Table[ $\Gamma_k / \Gamma_{\text{Rule}}$ , {k, 0, 1}], 1] /. nValues

$$\left\{ \begin{pmatrix} 201 & 180. \\ 180. & 201 \end{pmatrix}, \begin{pmatrix} -100 & -90. \\ -90. & -100 \end{pmatrix} \right\}$$

### Set up VMA form

Since the autocovariances vanish after lag 1, we have a VMA(1) structure:

```
θ[L_] := IdentityMatrix[2] + Table[θi,j, {i, 2}, {j, 2}] L;  
θ[L] // PolyForm
```

$$\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} + L \begin{pmatrix} \theta_{1,1} & \theta_{1,2} \\ \theta_{2,1} & \theta_{2,2} \end{pmatrix}$$

```
Ω = {{σ12, σ1,2}, {σ1,2, σ22}}; Ω // MatrixForm
```

$$\begin{pmatrix} \sigma_1^2 & \sigma_{1,2} \\ \sigma_{1,2} & \sigma_2^2 \end{pmatrix}$$

Autocovariance generating function:

```
g[z_] := θ[z].Ω.Transpose[θ[z-1]]
```

We'll obtain the  $\Gamma_k$  for the VMA representation by picking off the appropriate coefficient in the AGF.

```
ΓVMARule = Γk -> Simplify[Coefficient[g[z], z, k]];  
Map[MatrixForm, Table[Γk /. ΓVMARule, {k, 0, 1}]] // MatrixForm
```

$$\begin{pmatrix} \sigma_1^2 + \sigma_1^2 \theta_{1,1}^2 + \sigma_2^2 \theta_{1,2}^2 + 2 \theta_{1,1} \theta_{1,2} \sigma_{1,2} & \sigma_{1,2} + \theta_{1,2} (\sigma_2^2 \theta_{2,2} + \theta_{2,1} \sigma_{1,2}) + \theta_{1,1} \\ \sigma_{1,2} + \theta_{1,2} (\sigma_2^2 \theta_{2,2} + \theta_{2,1} \sigma_{1,2}) + \theta_{1,1} (\sigma_1^2 \theta_{2,1} + \theta_{2,2} \sigma_{1,2}) & \sigma_2^2 + \sigma_1^2 \theta_{2,1}^2 + \sigma_2^2 \theta_{2,2}^2 + 2 \theta_{2,1} \theta_{2,2} \sigma_{1,2} \\ \sigma_1^2 \theta_{1,1} + \theta_{1,2} \sigma_{1,2} & \sigma_2^2 \theta_{1,2} + \theta_{1,1} \sigma_{1,2} \\ \sigma_1^2 \theta_{2,1} + \theta_{2,2} \sigma_{1,2} & \sigma_2^2 \theta_{2,2} + \theta_{2,1} \sigma_{1,2} \end{pmatrix}$$

Solve for VMA parameters by equating the two representations:

```
eq = Flatten @ MapThread[Equal, {{Γ0, Γ1} /. ΓRule, {Γ0, Γ1} /. ΓVMARule}, 3];  
eq // TableForm
```

$$\begin{aligned} 2 c^2 + \sigma_u^2 &= \sigma_1^2 + \sigma_1^2 \theta_{1,1}^2 + \sigma_2^2 \theta_{1,2}^2 + 2 \theta_{1,1} \theta_{1,2} \sigma_{1,2} \\ 2 c^2 \rho &= \sigma_{1,2} + \theta_{1,2} (\sigma_2^2 \theta_{2,2} + \theta_{2,1} \sigma_{1,2}) + \theta_{1,1} (\sigma_1^2 \theta_{2,1} + \theta_{2,2} \sigma_{1,2}) \\ 2 c^2 \rho &= \sigma_{1,2} + \theta_{1,2} (\sigma_2^2 \theta_{2,2} + \theta_{2,1} \sigma_{1,2}) + \theta_{1,1} (\sigma_1^2 \theta_{2,1} + \theta_{2,2} \sigma_{1,2}) \\ 2 c^2 + \sigma_u^2 &= \sigma_2^2 + \sigma_1^2 \theta_{2,1}^2 + \sigma_2^2 \theta_{2,2}^2 + 2 \theta_{2,1} \theta_{2,2} \sigma_{1,2} \\ -c^2 &= \sigma_1^2 \theta_{1,1} + \theta_{1,2} \sigma_{1,2} \\ -c^2 \rho &= \sigma_2^2 \theta_{1,2} + \theta_{1,1} \sigma_{1,2} \\ -c^2 \rho &= \sigma_1^2 \theta_{2,1} + \theta_{2,2} \sigma_{1,2} \\ -c^2 &= \sigma_2^2 \theta_{2,2} + \theta_{2,1} \sigma_{1,2} \end{aligned}$$

The third equation is redundant, so ...

```
e2 = Drop[eq, {3}]
```

$$\begin{aligned} \{2c^2 + \sigma_u^2 &= \sigma_1^2 + \sigma_1^2 \theta_{1,1}^2 + \sigma_2^2 \theta_{1,2}^2 + 2\theta_{1,1}\theta_{1,2}\sigma_{1,2}, \\ 2c^2\rho &= \sigma_{1,2} + \theta_{1,2}(\sigma_2^2\theta_{2,2} + \theta_{2,1}\sigma_{1,2}) + \theta_{1,1}(\sigma_1^2\theta_{2,1} + \theta_{2,2}\sigma_{1,2}), \\ 2c^2 + \sigma_u^2 &= \sigma_2^2 + \sigma_1^2\theta_{2,1}^2 + \sigma_2^2\theta_{2,2}^2 + 2\theta_{2,1}\theta_{2,2}\sigma_{1,2}, \quad -c^2 = \sigma_1^2\theta_{1,1} + \theta_{1,2}\sigma_{1,2}, \\ -c^2\rho &= \sigma_2^2\theta_{1,2} + \theta_{1,1}\sigma_{1,2}, \quad -c^2\rho = \sigma_1^2\theta_{2,1} + \theta_{2,2}\sigma_{1,2}, \quad -c^2 = \sigma_2^2\theta_{2,2} + \theta_{2,1}\sigma_{1,2} \end{aligned}$$

```
parms = {θ1,1, θ1,2, θ2,1, θ2,2, σ12, σ22, σ1,2};
```

We'll attempt only a numerical solution here:

```
s = NSolve[e2 /. nValues, parms]; s // TableForm
```

$\theta_{1,1} \rightarrow -1.1501$	$\theta_{1,2} \rightarrow 0.22006$	$\theta_{2,1} \rightarrow 0.22006$	$\theta_{2,2} \rightarrow -1.1501$	$\sigma_1^2 \rightarrow 105.796$	$\sigma_2^2 \rightarrow 1$
$\theta_{1,1} \rightarrow -0.82994$	$\theta_{1,2} \rightarrow -0.100096$	$\theta_{2,1} \rightarrow -0.100096$	$\theta_{2,2} \rightarrow -0.82994$	$\sigma_1^2 \rightarrow 108.997$	$\sigma_2^2 \rightarrow 1$
$\theta_{1,1} \rightarrow -1.22269$	$\theta_{1,2} \rightarrow 0.147465$	$\theta_{2,1} \rightarrow 0.147465$	$\theta_{2,2} \rightarrow -1.22269$	$\sigma_1^2 \rightarrow 92.0027$	$\sigma_2^2 \rightarrow 9$
$\theta_{1,1} \rightarrow -0.902535$	$\theta_{1,2} \rightarrow -0.172692$	$\theta_{2,1} \rightarrow -0.172692$	$\theta_{2,2} \rightarrow -0.902535$	$\sigma_1^2 \rightarrow 95.2042$	$\sigma_2^2 \rightarrow 9$

Consider the matrix

```
DD = (θ[L] /. L → 1/L) * L // Simplify;
DD // MatrixForm
```

$$\begin{pmatrix} L + \theta_{1,1} & \theta_{1,2} \\ \theta_{2,1} & L + \theta_{2,2} \end{pmatrix}$$

For the VMA to be invertible, the roots of the determinantal equation  $\text{Det}[DD]=0$  must lie inside the unit circle. We check:

```
Table[Solve[Det[DD] == 0 /. s[[i]], L] // Flatten, {i, 4}] // MatrixForm
```

$$\begin{pmatrix} L \rightarrow 0.930036 & L \rightarrow 1.37016 \\ L \rightarrow 0.729844 & L \rightarrow 0.930036 \\ L \rightarrow 1.07523 & L \rightarrow 1.37016 \\ L \rightarrow 0.729844 & L \rightarrow 1.07523 \end{pmatrix}$$

Only the second solution is invertible:

```
sInvertible = s[[2]]
```

$$\{\theta_{1,1} \rightarrow -0.82994, \theta_{1,2} \rightarrow -0.100096, \theta_{2,1} \rightarrow -0.100096, \\ \theta_{2,2} \rightarrow -0.82994, \sigma_1^2 \rightarrow 108.997, \sigma_2^2 \rightarrow 108.997, \sigma_{1,2} \rightarrow 95.2958\}$$

```
Table[θi,j, {i, 2}, {j, 2}] /. sInvertible // MatrixForm
```

$$\begin{pmatrix} -0.82994 & -0.100096 \\ -0.100096 & -0.82994 \end{pmatrix}$$

```
 $\Omega /. sInvertible // MatrixForm$ 
```

```
 $\begin{pmatrix} 108.997 & 95.2958 \\ 95.2958 & 108.997 \end{pmatrix}$ 
```

Verify the computation of the random-walk variance:

```
 $\theta[1].\Omega.Transpose[\theta[1]] /. sInvertible // Chop // MatrixForm$ 
```

```
 $\begin{pmatrix} 1. & 0 \\ 0 & 1. \end{pmatrix}$ 
```

Variance decomposition.

```
 $(F = CholeskyDecomposition[\Omega /. sInvertible]) // MatrixForm$ 
```

```
 $\begin{pmatrix} 10.4402 & 9.12779 \\ 0. & 5.06762 \end{pmatrix}$ 
```

```
 $NumberForm[Transpose[F] // MatrixForm, \{5, 2\}]$ 
```

```
 $\begin{pmatrix} 10.44 & 0.00 \\ 9.13 & 5.07 \end{pmatrix}$ 
```

```
 $NumberForm[\theta[1].Transpose[F] /. sInvertible // MatrixForm, \{5, 3\}]$ 
```

```
 $\begin{pmatrix} 0.862 & -0.507 \\ 0.507 & 0.862 \end{pmatrix}$ 
```

```
 $r = (\theta[1].Transpose[F] /. sInvertible)[[1]]$ 
```

```
 $\{0.861799, -0.50725\}$ 
```

```
 $r^2$ 
```

```
 $\{0.742698, 0.257302\}$ 
```

## ■ Cointegrated model (Chapter 10, section 10.2)

### □ Model

```

$$\begin{aligned} P_{Rule} &= \{ p_{1,t} \mapsto m_t + c q_t, p_{2,t} \mapsto m_{t-1} \}; \\ m_{Rule} &= m_t \mapsto m_{t-1} + u_t; \\ \Delta P_{Rule} &= \Delta p_{i,t} \mapsto (p_{i,t} /. P_{Rule} /. m_{Rule}) - (p_{i,t-1} /. P_{Rule}); \\ \Delta P_t &:= \{ \{ \Delta p_{1,t} \}, \{ \Delta p_{2,t} \} \}; \end{aligned}$$

```

The vector of price differences is:

$$\Delta p_t / \cdot \Delta p_{\text{Rule}} // \text{MatrixForm}$$

$$\begin{pmatrix} -c q_{-1+t} + c q_t + u_t \\ u_{-1+t} \end{pmatrix}$$

The difference between the two prices is:

$$\left( (p_{1,t} / \cdot p_{\text{Rule}} / \cdot m_{\text{Rule}}) - (p_{2,t} / \cdot p_{\text{Rule}}) \right)$$

$$c q_t + u_t$$

□ *Structural VMA representation: VMA I*

$$\theta[L_] := \theta_0 + \theta_1 L$$

$$\epsilon_{\text{Rule}} = \epsilon_{t-} \rightarrow \{\{u_t\}, \{q_t\}\};$$

```

 $\theta_{\text{Rules}} =$ 
  Table[Rule[ $\theta_s$ , Table[Coefficient[( $\Delta p_t / \cdot \Delta p_{\text{Rule}}$ )[[i, 1]], ( $\epsilon_{t-s} / \cdot \epsilon_{\text{Rule}}$ )[[j, 1]],
    {i, 2}, {j, 2}]], {s, 0, 1}];
  Map[MatrixForm,  $\theta_{\text{Rules}}$ , {2}] // TableForm

```

$$\theta_0 \rightarrow \begin{pmatrix} 1 & c \\ 0 & 0 \end{pmatrix} \quad \theta_1 \rightarrow \begin{pmatrix} 0 & -c \\ 1 & 0 \end{pmatrix}$$

```
Clear[ $\Omega$ ];
```

$$\Omega_{\text{Rule}} = \Omega \rightarrow \{\{\sigma_u^2, 0\}, \{0, 1\}\}$$

$$\Omega \rightarrow \{\{\sigma_u^2, 0\}, \{0, 1\}\}$$

Autocovariance generating function:

```

agfI = G[ $\theta[L] / \cdot \theta_{\text{Rules}}$ ,  $\Omega / \cdot \Omega_{\text{Rule}}$ , z];
PolyForm[agfI, z]

```

$$z \begin{pmatrix} -c^2 & 0 \\ \sigma_u^2 & 0 \end{pmatrix} + \frac{\begin{pmatrix} -c^2 & \sigma_u^2 \\ 0 & 0 \end{pmatrix}}{z} + \begin{pmatrix} 2c^2 + \sigma_u^2 & 0 \\ 0 & \sigma_u^2 \end{pmatrix}$$

This can't be inverted:

```
Series[Inverse[ $\theta[L] / \cdot \theta_{\text{Rules}}$ ], {L, 0, 5}] // PolyForm
```

$$\begin{pmatrix} 0 & 0 \\ \frac{1}{c} & -\frac{1}{c} \end{pmatrix} + L \begin{pmatrix} 0 & 0 \\ \frac{1}{c} & -\frac{1}{c} \end{pmatrix} + L^2 \begin{pmatrix} 0 & 0 \\ \frac{1}{c} & -\frac{1}{c} \end{pmatrix} + L^3 \begin{pmatrix} 0 & 0 \\ \frac{1}{c} & -\frac{1}{c} \end{pmatrix} + L^4 \begin{pmatrix} 0 & 0 \\ \frac{1}{c} & -\frac{1}{c} \end{pmatrix} + L^5 \begin{pmatrix} 0 & 0 \\ \frac{1}{c} & -\frac{1}{c} \end{pmatrix} + \frac{\begin{pmatrix} 0 & 1 \\ 0 & -\frac{1}{c} \end{pmatrix}}{L}$$

But it gives the correct random-walk variance:

```
Θ[1].Ω.Transpose[Θ[1]] /. ΘRules /. ΩRule
```

```
{ {σu2, σu2}, {σu2, σu2}}
```

(There is one random-walk component underlying both securities.)

□ VMA representation computed from autocovariances: VMA II

```
ΘII[L_] := IdentityMatrix[2] + Table[Θi,j, {i, 2}, {j, 2}] L;  
ΘII[L] // PolyForm
```

```
( 1  0 ) + L ( Θ1,1  Θ1,2 )  
  0  1       ( Θ2,1  Θ2,2 )
```

```
ΩIIRule = Ω → { {σ12, σ1,2}, {σ1,2, σ22}};  
Ω /. ΩIIRule // MatrixForm
```

```
( σ12    σ1,2 )  
 ( σ1,2  σ22 )
```

```
agfII = G[ΘII[L], Ω /. ΩIIRule, z];  
PolyForm[agfII, z, vert]
```

```
PolyForm[ { {  $\frac{\Theta_{1,2} (z \sigma_2^2 \Theta_{1,2} + (1 + z \Theta_{1,1}) \sigma_{1,2})}{z} + \left(1 + \frac{\Theta_{1,1}}{z}\right) (\sigma_1^2 (1 + z \Theta_{1,1}) + z \Theta_{1,2} \sigma_{1,2})$ ,  

 $\left(1 + \frac{\Theta_{2,2}}{z}\right) (z \sigma_2^2 \Theta_{1,2} + (1 + z \Theta_{1,1}) \sigma_{1,2}) + \frac{\Theta_{2,1} (\sigma_1^2 (1 + z \Theta_{1,1}) + z \Theta_{1,2} \sigma_{1,2})}{z}$  },  

 $\left\{ \frac{\Theta_{1,2} (\sigma_2^2 (1 + z \Theta_{2,2}) + z \Theta_{2,1} \sigma_{1,2})}{z} + \left(1 + \frac{\Theta_{1,1}}{z}\right) (z \sigma_1^2 \Theta_{2,1} + (1 + z \Theta_{2,2}) \sigma_{1,2})$ ,  

 $\left(1 + \frac{\Theta_{2,2}}{z}\right) (\sigma_2^2 (1 + z \Theta_{2,2}) + z \Theta_{2,1} \sigma_{1,2}) + \frac{\Theta_{2,1} (z \sigma_1^2 \Theta_{2,1} + (1 + z \Theta_{2,2}) \sigma_{1,2})}{z} \right\}$ , z, vert ]
```

```
e = Union @ Flatten @ Table[MapThread[Equal,  

    {Coefficient[agfI, z, k], Coefficient[agfII, z, k]}, 2], {k, 0, 1}];  
e //  
TableForm
```

```
0 = σ22 Θ1,2 + Θ1,1 σ1,2  

0 = σ22 Θ2,2 + Θ2,1 σ1,2  

0 = σ12 Θ1,1 Θ2,1 + σ22 Θ1,2 Θ2,2 + σ1,2 + Θ1,2 Θ2,1 σ1,2 + Θ1,1 Θ2,2 σ1,2  

-c2 = σ12 Θ1,1 + Θ1,2 σ1,2  

σu2 = σ12 Θ2,1 + Θ2,2 σ1,2  

σu2 = σ22 + σ12 Θ2,1 + σ22 Θ2,2 + 2 Θ2,1 Θ2,2 σ1,2  

2 c2 + σu2 = σ12 + σ12 Θ1,1 + σ22 Θ1,2 + 2 Θ1,1 Θ1,2 σ1,2
```

```
parms = {θ1,1, θ1,2, θ2,1, θ2,2, σ12, σ22, σ1,2};
```

```
solution = Solve[e, parms];  
solution // Transpose // TableForm
```

$$\begin{array}{ll} \sigma_1^2 \rightarrow \frac{c^4 + 3 c^2 \sigma_u^2 + (\sigma_u^2)^2}{c^2 + \sigma_u^2} & \sigma_1^2 \rightarrow \frac{c^4 + 3 c^2 \sigma_u^2 + (\sigma_u^2)^2}{c^2 + \sigma_u^2} \\ \sigma_2^2 \rightarrow \frac{c^2 \sigma_u^2}{c^2 + \sigma_u^2} & \sigma_2^2 \rightarrow \frac{c^2 \sigma_u^2}{c^2 + \sigma_u^2} \\ \theta_{1,2} \rightarrow \frac{c^2}{c^2 + \sigma_u^2} & \theta_{1,2} \rightarrow \frac{c^2}{c^2 + \sigma_u^2} \\ \theta_{2,2} \rightarrow -\frac{\sigma_u^2}{c^2 + \sigma_u^2} & \theta_{2,2} \rightarrow -\frac{\sigma_u^2}{c^2 + \sigma_u^2} \\ \theta_{1,1} \rightarrow -\frac{c^2}{c^2 + \sigma_u^2} & \theta_{1,1} \rightarrow -\frac{c^2}{c^2 + \sigma_u^2} \\ \theta_{2,1} \rightarrow \frac{\sigma_u^2}{c^2 + \sigma_u^2} & \theta_{2,1} \rightarrow \frac{\sigma_u^2}{c^2 + \sigma_u^2} \\ \sigma_{1,2} \rightarrow \frac{c^2 \sigma_u^2}{c^2 + \sigma_u^2} & \sigma_{1,2} \rightarrow \frac{c^2 \sigma_u^2}{c^2 + \sigma_u^2} \end{array}$$

There are two solutions, but they're the same:

```
solution[[1]] === solution[[2]]
```

```
True
```

So

```
solution = solution[[1]]
```

$$\left\{ \sigma_1^2 \rightarrow \frac{c^4 + 3 c^2 \sigma_u^2 + (\sigma_u^2)^2}{c^2 + \sigma_u^2}, \sigma_2^2 \rightarrow \frac{c^2 \sigma_u^2}{c^2 + \sigma_u^2}, \theta_{1,2} \rightarrow \frac{c^2}{c^2 + \sigma_u^2}, \right. \\ \left. \theta_{2,2} \rightarrow -\frac{\sigma_u^2}{c^2 + \sigma_u^2}, \theta_{1,1} \rightarrow -\frac{c^2}{c^2 + \sigma_u^2}, \theta_{2,1} \rightarrow \frac{\sigma_u^2}{c^2 + \sigma_u^2}, \sigma_{1,2} \rightarrow \frac{c^2 \sigma_u^2}{c^2 + \sigma_u^2} \right\}$$

```
θII[1] /. solution // Simplify // MatrixForm
```

$$\begin{pmatrix} \frac{\sigma_u^2}{c^2 + \sigma_u^2} & \frac{c^2}{c^2 + \sigma_u^2} \\ \frac{\sigma_u^2}{c^2 + \sigma_u^2} & \frac{c^2}{c^2 + \sigma_u^2} \end{pmatrix}$$

Consider the matrix

```
DD = (θII[L] /. L → 1/L) * L // Simplify;  
DD // MatrixForm
```

$$\begin{pmatrix} L + \theta_{1,1} & \theta_{1,2} \\ \theta_{2,1} & L + \theta_{2,2} \end{pmatrix}$$

For the VMA to be invertible, the roots of the determinantal equation  $\text{Det}[DD]=0$  must lie inside the unit circle.

We check:

```
Table[Solve[Det[DD] == 0 /. solution, L] // Flatten, {1, 2}] // MatrixForm
```

$$\begin{pmatrix} L \rightarrow 0 & L \rightarrow 1 \\ L \rightarrow 0 & L \rightarrow 1 \end{pmatrix}$$

Uh Oh. Let's try to compute the VAR by polynomial expansion:

```
Series[Inverse[ΘII[L]] /. solution, {L, 0, 4}] // Simplify // PolyForm
```

$$\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} + L \begin{pmatrix} \frac{c^2}{c^2 + \sigma_u^2} & -\frac{c^2}{c^2 + \sigma_u^2} \\ -\frac{\sigma_u^2}{c^2 + \sigma_u^2} & \frac{\sigma_u^2}{c^2 + \sigma_u^2} \end{pmatrix} + L^2 \begin{pmatrix} \frac{c^2}{c^2 + \sigma_u^2} & -\frac{c^2}{c^2 + \sigma_u^2} \\ -\frac{\sigma_u^2}{c^2 + \sigma_u^2} & \frac{\sigma_u^2}{c^2 + \sigma_u^2} \end{pmatrix} + L^3 \begin{pmatrix} \frac{c^2}{c^2 + \sigma_u^2} & -\frac{c^2}{c^2 + \sigma_u^2} \\ -\frac{\sigma_u^2}{c^2 + \sigma_u^2} & \frac{\sigma_u^2}{c^2 + \sigma_u^2} \end{pmatrix} + L^4 \begin{pmatrix} \frac{c^2}{c^2 + \sigma_u^2} & -\frac{c^2}{c^2 + \sigma_u^2} \\ -\frac{\sigma_u^2}{c^2 + \sigma_u^2} & \frac{\sigma_u^2}{c^2 + \sigma_u^2} \end{pmatrix}$$

No convergence here, either.

□ *Exercise 10.1*

```
solution
```

$$\left\{ \sigma_1^2 \rightarrow \frac{c^4 + 3 c^2 \sigma_u^2 + (\sigma_u^2)^2}{c^2 + \sigma_u^2}, \sigma_2^2 \rightarrow \frac{c^2 \sigma_u^2}{c^2 + \sigma_u^2}, \theta_{1,2} \rightarrow \frac{c^2}{c^2 + \sigma_u^2}, \right. \\ \left. \theta_{2,2} \rightarrow -\frac{\sigma_u^2}{c^2 + \sigma_u^2}, \theta_{1,1} \rightarrow -\frac{c^2}{c^2 + \sigma_u^2}, \theta_{2,1} \rightarrow \frac{\sigma_u^2}{c^2 + \sigma_u^2}, \sigma_{1,2} \rightarrow \frac{c^2 \sigma_u^2}{c^2 + \sigma_u^2} \right\}$$

```
ClearAll[VarianceDecomposition]
```

```
Ω /. ΩIIRule
```

$$\left\{ \left\{ \sigma_1^2, \sigma_{1,2} \right\}, \left\{ \sigma_{1,2}, \sigma_2^2 \right\} \right\}$$



```

VarianceDecomposition[s_, perm_ : {1, 2}] := Module[{},
  Print["Ordering of prices is: ", TableForm[{perm}]];
  F =
    Permute[CholeskyDecomposition[Permute[Ω /. ΩIIRule /. s // N, perm]], perm];
  Print["Cholesky factorization (in correct order) F=", MatrixForm[F]];
  θSum = θII[1] /. s // N;
  a = θSum.Transpose[F];
  Print["θ(1)F'=", MatrixForm[a]];
  v = a[[1]]^2;
  Print["Variance contributions:", MatrixForm[{v}]];
  tv = Total[v];
  Print["Total variance=", tv];
  Print["Information shares= ", MatrixForm[{v} / tv]];
];
VarianceDecomposition[Join[solution, {σu2 → 1, c → 2}], {1, 2}]

```

Ordering of prices is: 1 2

Cholesky factorization (in correct order)  $F = \begin{pmatrix} 2.40832 & 0.332182 \\ 0. & 0.830455 \end{pmatrix}$

$\theta(1)F' = \begin{pmatrix} 0.747409 & 0.664364 \\ 0.747409 & 0.664364 \end{pmatrix}$

Variance contributions: (0.558621 0.441379)

Total variance=1.

Information shares= (0.558621 0.441379)

```

VarianceDecomposition[Join[solution, {σu2 → 1, c → 2}], {2, 1}]

```

Ordering of prices is: 2 1

Cholesky factorization (in correct order)  $F = \begin{pmatrix} 2.23607 & 0. \\ 0.894427 & 0.894427 \end{pmatrix}$

$\theta(1)F' = \begin{pmatrix} 0.447214 & 0.894427 \\ 0.447214 & 0.894427 \end{pmatrix}$

Variance contributions: (0.2 0.8)

Total variance=1.

Information shares= (0.2 0.8)

□ *VECM representations.*

Here is the VAR representation for the price *levels*:

```
ϕ[L_] := Evaluate[Simplify[Inverse[θII[L]] (1 - L) /. solution] // Normal];
ϕ[L] // PolyForm
```

$$\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} + L \begin{pmatrix} -\frac{\sigma_u^2}{c^2 + \sigma_u^2} & -\frac{c^2}{c^2 + \sigma_u^2} \\ -\frac{\sigma_u^2}{c^2 + \sigma_u^2} & -\frac{c^2}{c^2 + \sigma_u^2} \end{pmatrix}$$

Verify (19.1.34) in Hamilton

```
Simplify[ϕ[1].θII[1] /. solution] // MatrixForm
```

$$\begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix}$$

Rearrange  $\phi(L) p_t = \epsilon_t$  as  $p_t + \phi_1 p_{t-1} = \epsilon_t \Leftrightarrow p_t - p_{t-1} = (-I - \phi_1) p_{t-1} + \epsilon_t$ . The coefficient of  $p_{t-1}$  on the r.h.s. is:

```
-IdentityMatrix[2] - Coefficient[ϕ[L], L, 1] // Together // MatrixForm
```

$$\begin{pmatrix} -\frac{c^2}{c^2 + \sigma_u^2} & \frac{c^2}{c^2 + \sigma_u^2} \\ \frac{\sigma_u^2}{c^2 + \sigma_u^2} & -\frac{\sigma_u^2}{c^2 + \sigma_u^2} \end{pmatrix}$$

This can be written as the error correction term  $\gamma [1 \ -1] p_{t-1}$  where

$$\gamma = \left\{ \left\{ \frac{-c^2}{\sigma_u^2 + c_1^2} \right\}, \left\{ \frac{\sigma_u^2}{\sigma_u^2 + c^2} \right\} \right\};$$

```
γ.{1, -1} // MatrixForm
```

$$\begin{pmatrix} -\frac{c^2}{\sigma_u^2 + c_1^2} & \frac{c^2}{\sigma_u^2 + c_1^2} \\ \frac{\sigma_u^2}{c^2 + \sigma_u^2} & -\frac{\sigma_u^2}{c^2 + \sigma_u^2} \end{pmatrix}$$