

# THE SOCIAL INTEGRATION OF INTERNATIONAL MIGRANTS: EVIDENCE FROM THE NETWORKS OF SYRIANS IN GERMANY \*

MICHAEL BAILEY<sup>†</sup>  
DREW JOHNSTON<sup>‡</sup>  
MARTIN KOENEN<sup>§</sup>  
THERESA KUCHLER<sup>¶</sup>  
DOMINIC RUSSEL<sup>||</sup>  
JOHANNES STROEBEL<sup>\*\*</sup>

## Abstract

We use de-identified data from Facebook to study the social integration of Syrian migrants in Germany, a country that received a large influx of refugees during the Syrian Civil War. We construct measures of migrants' social integration based on Syrians' friendship links to Germans, their use of the German language, and their participation in local social groups. We find large variation in Syrians' social integration across German counties, and use a movers' research design to document that these differences are largely due to causal effects of place. Regional differences in the social integration of Syrians are shaped both by the rate at which German natives befriend other locals in general (*general friendliness*) and the relative rate at which they befriend local Syrian migrants versus German natives (*relative friending*). We follow the friending behavior of Germans that move across locations to show that both general friendliness and relative friending are more strongly affected by place-based effects such as local institutions than by persistent individual characteristics of natives (e.g., attitudes toward neighbors or migrants). Relative friending is higher in areas with lower unemployment and more completed government-sponsored integration courses. Using variation in teacher availability as an instrument, we find that integration courses had a substantial causal effect on the social integration of Syrian migrants. We also use fluctuations in the presence of Syrian migrants across high school cohorts to show that natives with quasi-random exposure to Syrians in school are more likely to befriend other Syrian migrants in other settings, suggesting that contact between groups can shape subsequent attitudes towards migrants.

**JEL Codes:** F22, J15, K37, D85

**Keywords:** Integration, immigration, social networks, place effects

---

\*This draft: March 3, 2022. This research was facilitated through a research consulting agreement between the academic authors and Facebook. Bailey is an employee at Facebook. We thank Raj Chetty, Ed Glaeser, Nathan Hendren, Nathan Nunn, Larry Katz, Jesse Shapiro and seminar participants at Harvard, HKS, and NYU for valuable feedback.

<sup>†</sup>Facebook, Inc: mcbailey@fb.com.

<sup>‡</sup>Harvard University: drewjohnston@g.harvard.edu

<sup>§</sup>Harvard University: martin\_koenen@g.harvard.edu

<sup>¶</sup>Stern School of Business, New York University, NBER, CEPR, cesifo: tkuchler@stern.nyu.edu

<sup>||</sup>Harvard University: drussel@g.harvard.edu

<sup>\*\*</sup>Stern School of Business, New York University, NBER, CEPR, cesifo: johannes.stroebel@nyu.edu

In 2019, there were 272 million international migrants comprising 3.5% of the world’s population (United Nations, 2019). The challenge of fostering communities that harmoniously integrate new arrivals with natives has therefore become of increasing importance to policymakers around the globe (e.g., European Commission, 2020). Yet, because of difficulties with measuring social networks using traditional data sources, researchers have long struggled to understand the determinants of the social integration of migrants in their host communities.

In this paper, we use de-identified data from Facebook, a global online social networking service, to study the factors that shape the social integration of newly arriving migrants and refugees. We focus on individuals who recently migrated from Syria to Germany. In the wake of the Syrian Civil War, millions of Syrians fled conflict and economic collapse in their home country. Germany has been the largest recipient country in the EU, with around 800,000 Syrians settling in Germany since 2014, equivalent to about one percent of the German population. The social and economic integration of these migrants has been a dominant political issue in Germany in the years since, with policy makers attempting to facilitate this integration through a variety of programs (see Bundesregierung, 2021). In 2018 alone, for instance, the German government invested two billion Euros in integration courses intended to teach migrants the German language and provide information on the country’s culture, legal system, and more.

While some studies have investigated the *economic* integration of Syrians in Germany—with a particular focus on attempts to bring refugees into the labor force through apprenticeship programs—substantial data challenges have meant that little work has been done to understand the *social* integration of these migrants: how frequently and intensely do Syrian migrants interact with local Germans?<sup>1</sup> How does this differ across demographic characteristics and locations? And why? Our unique data and research design allow us to provide answers to these questions.

We begin by constructing a sample of Facebook users who are currently living in Germany, but who specified a hometown or high school in Syria in their Facebook profile, or who previously had a predicted home region in Syria.<sup>2</sup> This simple methodology generates spatial variation in Syrian migrant population shares across German counties that closely resembles German administrative data. We also construct a group of users that we call “German natives” based on self-reported profile information, home region predictions, and German language usage.<sup>3</sup> We use these data to measure Syrian migrants’ social integration along three key dimensions: (i) friendships between migrants and German natives; (ii) migrants’ German language usage; and (iii) migrants’ participation in local social groups.

Syrian migrant users have five local German native friends on average, and 30% of them produce German content such as posts or comments on Facebook. Controlling for Facebook usage patterns,

---

<sup>1</sup>While there is no single definition of social integration, the concept is often defined by the frequency of interactions of individuals of different groups (e.g., Phillips et al., 2019). We also recognize the important distinction between social integration and *assimilation* (Berry, 1997). The latter, which is defined in terms of cultural identity, is not the focus of our work.

<sup>2</sup>Estimated home region is determined by a person’s information on Facebook, including the stated city on their Facebook profile, and device and connection information (also see Herdağdelen et al., 2016; Chi et al., 2019).

<sup>3</sup>We describe these criteria in detail in Appendix B. When constructing both the “Syrian migrant” and “German native” samples, we do not make any inference about citizenship status or race and ethnicity. Our intent is instead to create samples of users that appear to have lived in Syria but now live in Germany (Syrian migrant sample), or that have lived in Germany for a substantial amount of time and exclusively or primarily use the German language (German native sample). For example, the former sample will include a German citizen who self-reports a Syrian hometown; the latter will include individuals of Syrian descent who report a German hometown and primarily use the German language on Facebook.

younger and male Syrian migrants have higher levels of social integration than others. We also find large spatial heterogeneity in Syrian migrants' social integration across the 401 German counties (*Kreise*): an average Syrian migrant living in a 10th percentile county has more than twice as many native German friends as an average Syrian migrant living in a 90th percentile county. These spatial patterns are highly correlated across our three measures of social integration. We show that these measures pick up true differences in integration rather than sampling variation or differences in Facebook usage; they also align with survey-based measures of integration available at higher levels of geographic aggregation.

We then analyze whether the observed geographic differences in social integration are the result of selection (e.g., migrants with a higher propensity to integrate select to live in certain regions) or whether they correspond to *causal* effects of location. We first argue that the initial allocation of Syrian refugees to different locations—a process that is largely random—suggests that the geographic differences in refugee integration can likely be interpreted as causal effects of place. To confirm this, we use a mover research design that follows the (relatively few) Syrian migrants who move across German counties. These Syrian movers' social integration patterns quickly adjust from those of their origin towards those of their destination county, confirming that most of the observed regional differences in social integration are indeed due to causal place-based factors rather than migrant characteristics. We estimate that at least 74% of the observed spatial differences are due to causal effects of place. A mover design does not capture the effects of places on long-run friending behaviors (e.g., if a migrant learns the German language or takes an integration course in one place and then uses this knowledge to make more friends in another), and our estimate is therefore likely to *understate* the full effects of place.

We next ask: “What makes migrants more likely to integrate in one place versus another?” Motivated by research that characterizes communities' overall connectedness (e.g., Coleman, 1988; Putnam, 1995a; Chetty et al., 2022a,b), we first study general patterns of local friending. Put simply, if natives in a given location are more likely to befriend *any* of their German neighbors they may also be more likely to befriend newly arriving migrants. We call this measure *general friendliness*, the rate at which native Germans befriend their neighbors in general. The second determinant of Syrian migrants' social integration is a region's *relative friending*, the relative probability of natives befriending local Syrians versus local natives. Both general friendliness and relative friending vary across counties, with relative friending explaining about two-thirds of the spatial variation in the social integration of Syrian migrants.

We explore the extent to which county-level variation in relative friending and general friendliness are driven by the characteristics and preferences of the population of local natives versus the structure of local institutions. Studying the behavior of *native* movers, we show that place-based effects play a dominant role, but that individual characteristics and preferences become more important for older natives. Native movers under the age of 30 adjust their general friendliness about three-quarters of the way to comparable destination natives within a year of moving; their relative friending adjusts fully to that of comparable destination natives. For individuals above the age of 30, we find general friendliness and relative friending to adjust about half the way to those of comparable destination natives within a year of moving. This finding is consistent with work in a number of other settings that finds that new places exert stronger effects on younger individuals (e.g., Kling, Liebman and Katz, 2007; Chetty, Hendren and Katz, 2016; Chetty and Hendren, 2018a; Chyn, 2018).

Next, we study the determinants of these two drivers of migrants' social integration by exploring their relationships with county-level characteristics. General friendliness is decreasing with population density: Germans are less likely to befriend their neighbors in cities than in the countryside. Relative friending decreases with a county's population share that was Syrian in 2019, but increases with the share that was Syrian in 2010. This finding speaks to the "ethnic enclaves" literature that finds migrant networks support integration in some settings and hinder it in others (e.g. Lazear, 1999; Edin, Fredriksson and Åslund, 2003; Cutler, Glaeser and Vigdor, 2008; Beaman, 2012; Sale, 2021; Martén, Hainmueller and Hangartner, 2019). Here, earlier migrants may provide information or connections to support new arrivals' integration, and their long-term presence could have positively shaped local natives' views towards Syrians. On the other hand, many migrants arriving at the same time may lead to fewer migrant-native connections in part because the presence of many others with similar backgrounds facilitates the formation of cliques (Chetty et al., 2022b). Relative friending is also lower in places with stronger support for the Alternative for German (AfD), a far-right political party in favor of limiting immigration.

For policymakers, language and integration courses are one of the few direct tools to foster the social integration of migrants and such courses have been a central component of the German government's integration policy. We next explore the effectiveness of such integration courses, thereby contributing to a literature that has studied various government policies intended to assimilate minority groups or to improve their labor market prospects (Abdelgadir and Fouka, 2020; Abramitzky, Boustan and Eriksson, 2020; Arendt et al., 2020; Bandiera et al., 2019; Battisti, Giesing and Laurentsyeva, 2019; Fouka, 2020; Lleras-Muney and Shertzer, 2015). We find that in counties with more completed integration courses, relative friending is higher. We use an instrumental variables (IV) approach to study whether the provision of local integration courses has a *causal* effect on integration outcomes. Our instrument, the local unemployment rate of qualified teachers at the start of the large influx of Syrian migrants, is strongly correlated with the completion of courses, even after controlling for the general local unemployment rate. This is consistent with anecdotal evidence suggesting the unavailability of teachers substantially limited the government's initial ability to offer integration courses. Our estimates suggest that a 10% increase in 2015-19 integration course completion per Syrian (driven by an increase in the availability of such courses) raised relative friending by 14% and friending integration by 17%. At the regional level, integration courses also improve employment outcomes among Syrians.

In the final section of the paper we return to the determinants of natives' friending behaviors that do not change after a move, and study the longer-term effects of childhood exposure to Syrian migrants on subsequent friending patterns. Specifically, we use fluctuations in the presence of Syrian migrants across high school cohorts as a quasi-random source of variation of exposure to such migrants. We find that exposure to Syrian migrants in high school leads to higher probabilities of German natives befriending Syrians even outside the high school setting, consistent with the contact hypothesis, which outlines the circumstances in which social contact between members of different groups can help to reduce prejudice and animosity (Allport, Clark and Pettigrew, 1954; Paluck, Green and Green, 2019).<sup>4</sup>

---

<sup>4</sup>Existing work has found that interactions with members of other social groups such as minorities reduce bias against that group according to self-reported measures (Boisjoly et al., 2006) or short-run decisions in a similar context to the initial exposure (Rao, 2019; Carrell, Hoekstra and West, 2015). Perhaps most closely related to this analysis is Bursztyn et al. (2021), who show that exposure to Arab-Muslims in the U.S. translates into greater self-reported personal contact with Arab-Muslims.

The concept of social integration has long been important in social science research (e.g., Srole, 1956; Coleman, 1988; Putnam, 1995a; Alesina, Baqir and Easterly, 1999). Within this literature, our work relates most closely to studies that use surveys or assimilation-related measures to proxy for immigrants' social integration. Laurentsyeva and Venturini (2017) provide one overview (see also Niehues, Rother and Siegert, 2021; Schmidt, Jacobsen and Krieger, 2020; Cheung and Phillimore, 2014). In contrast to these studies, we are able to directly measure and study key elements of migrants' social integration, allowing us to provide a number of novel insights on the determinants of this integration.

Our work also relates to a literature studying the economic integration of refugees in high-income countries. Becker and Ferrara (2019) and Brell, Dustmann and Preston (2020) provide overviews. We complement this literature by showing that migrants' social integration in a region increases with their economic integration and decreases with the overall unemployment rate. We also provide evidence that integration courses had a substantial causal effect on migrants' economic integration.

We also add to a literature that uses experimental and quasi-experimental methods to study the causal effects of local environments on a variety of economic, social, and health outcomes (see Chyn and Katz, 2021, for one overview). We believe we are the first to use a mover-based research design to study the effects of place on migrants' social integration, adding to existing evidence that is observational or relies on quasi-random refugee settlements (e.g. Åslund and Rooth, 2007; Damm, 2014; Braun and Dwenger, 2017; Aksoy, Poutvaara and Schikora, 2020; Jaschke, Sardoschau and Tabellini, 2021; Sale, 2021). We also introduce the use of movers to study the effect of places on *native* rates of befriending migrants, highlighting that place-based effects are not primarily picking up preferences of local natives.

The remainder of this paper is structured as follows. In Section 1 we describe our data, sample, and outcomes of interest, before documenting overall patterns of social integration. Section 2 explores the relationship between individual-level Syrian migrant characteristics and integration outcomes. In Section 3 we generate regional measures of social integration and use movers to study the extent to which they reflect place-based effects. Section 4 focuses on natives and local institutions, exploring the forces that make migrants more likely to integrate in one place versus another. Section 5 looks at how quasi-random exposure to migrants shapes natives' long-term behavior. We conclude in Section 6.

## 1 Data and Descriptive Statistics

We work with de-identified data from the online social networking site Facebook. Created in 2004, Facebook had over 2.8 billion monthly active users by March 2021, including 423 million in Europe (Facebook, 2021). Facebook is used widely by Syrian migrants in Germany to share information and communicate with friends and family in Syria and elsewhere (Scheibe, Zimmer and Stock, 2019). Many individuals opened their Facebook accounts prior to arriving in Germany, while others likely created accounts during their migration, as Facebook was frequently cited as a tool used by refugees fleeing to Europe to share information (e.g., Dekker et al., 2018; Mall et al., 2015; Ritscher, 2016).<sup>5</sup>

Establishing a "friendship" connection on Facebook requires the consent of both individuals, and a person can have at most 5,000 connections. As a result, Facebook connections are primarily between

---

<sup>5</sup>For example, the introductory passage of Mustafa and Lamb (2017), a memoir of a young Syrian refugee in Germany, highlights that: "To be a successful migrant you need to know the law. You need to be resourceful. You need a smartphone and to be on Facebook and WhatsApp."

individuals who interact in person (Jones et al., 2013). Facebook networks therefore resemble real-world social networks more closely than networks on other online platforms where uni-directional links to non-acquaintances (e.g., celebrities) are common. Indeed, prior studies have used Facebook data to study the relationship between real world social connections and a wide variety of economic and social outcomes. For example, Facebook connections predict patterns of trade (Bailey et al., 2021), patent citations (Bailey et al., 2018b), travel flows (Bailey et al., 2020c,b), disease transmission (Kuchler, Russel and Stroebel, 2021), bank lending (Rehbein and Rother, 2020), social program participation (Wilson, 2019), and investment decisions (Kuchler et al., 2020). Information on individuals’ Facebook networks can also help explain their product adoption decisions (Bailey et al., 2019a), housing choices (Bailey et al., 2018a, 2019b), and beliefs and behaviors surrounding public health issues (Bailey et al., 2020a).

## 1.1 Sample Construction

We construct our primary sample from a sub-population of Facebook users who had active accounts in October 2021, were 18 or older, lived in Germany, and had 25 or more friends. Each user is predicted to have a home region in one of 401 German districts (*Kreis*, *Landkreis*, or *Stadtkreis*), with an average population of just over 200,000.<sup>6</sup> We will refer to these geographies as “counties” throughout. These counties fall into 16 federal states. From this primary sample, we define two sub-samples.

**Syrian Migrant Sample.** For many analyses, we focus on a set of users who specify a hometown or high school in Syria in their Facebook profile, or who previously had a predicted home region in Syria (see footnotes 2 and 3 for more information). There are about 350,000 such users, which we refer to as “Syrian migrants.” The median county has 464 Syrian migrant users; the 10th-90th percentile range is 167 to 1,705 users. Figure 1 plots shares of our Facebook sample that are Syrian migrants against analogous shares in the full population constructed using administrative data from the Federal Statistical Office of Germany. Panel (a) plots these across state-by-age-by-gender groups. Panel (b) plots these across county-by-gender groups. The correlation (weighted by the true population size) in each panel is greater than 0.9, highlighting that we observe users in our sample consistently across geography, age, and gender groups. The level of the shares is close to the 45-degree line, suggesting that we pick up Syrian and non-Syrian users at roughly the same rates in the data, and at similar rates across locations.

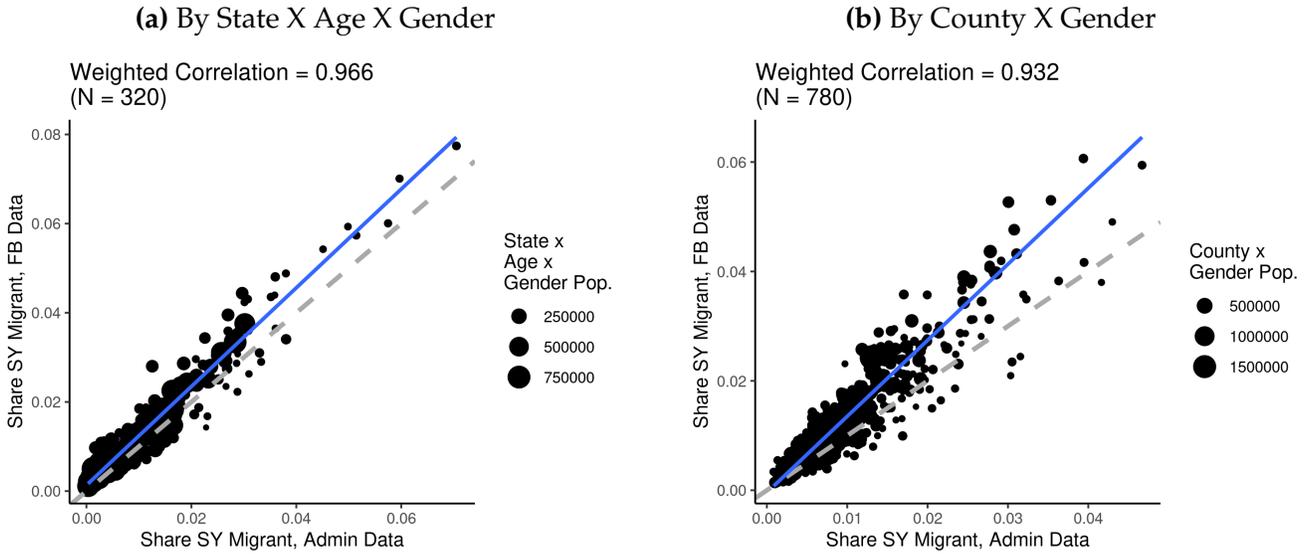
Integration levels generally increase with time in a new place (see, e.g., Niehues, Rother and Siegert, 2021, and Figure 2). To account for this fact, many of our analyses will condition on information about when Syrian migrant users first appeared in Germany. Appendix Figure A2 shows that the number of users in our Syrian migrant Facebook sample in Germany at the end of each year from 2012 to 2019 is generally well aligned with the Syrian migrant population at that time as measured in administrative data. Within the Syrian migrant sample, 62% of individuals first used Facebook *outside* of Germany. For some analyses, we further focus on this sub-sample of users for which we have a better ability to observe the timing of their arrival in Germany.

**German Native Sample.** We also construct a group of users, which we refer to as “German natives”, who meet the criteria described in Appendix B based on self-reported profile information, home region

---

<sup>6</sup>For this paper, we use the set of counties that correspond the European Union NUTS 2016 classification, level 3. Locations are assigned to users based on their information and activity on Facebook, including their self-reported profile information, and device and connection information.

**Figure 1: Syrian Migrant Sample vs. Admin Data**



**Note:** Figures show the shares of the primary sample of Facebook users that are also in the Syrian migrant sample (on the y-axis), against shares of the population that are Syrian from administrative data (on the x-axis). The size of each dot is proportional to the true population it represents. The solid blue lines are from weighted linear regressions. The dashed grey line is the line  $y = x$ . Panel (a) plots these shares by state, age, and gender. The age groups are 18-24, 25-29, 30-34, 35-39, 40-44, 45-49, 50-54, 55-59, 60-64, and 65+. There are 16 states X 10 age groups X 2 genders = 320 observations. Panel (b) plots these shares by county and gender. Admin data is unavailable for 11 counties. There are 390 counties X 2 genders = 780 observations.

predictions, and German language usage. We identify 18 million such users. The median county has 34,063 German native users; the 10th-90th percentile range is 17,057 to 74,651 German native users. Appendix Figure A3 benchmarks this sample against administrative data. The share of users in the primary Facebook sample that are German native is somewhat lower than the true population share, a result of our relatively strict assignment criteria. However, the German population shares are consistent across space and gender, with population-weighted correlations between county-by-gender shares of German natives in the Facebook sample versus the actual population of 0.94.

## 1.2 Measures of Integration

When studying the social integration of migrants, a central challenge is directly measuring outcomes of interest at scale. We overcome this difficulty by constructing measures from de-identified Facebook data. We capture the social integration of Syrian migrant users using three primary measures (see Appendix C for a detailed definitions): (1) The number of native German friends a Syrian migrant user has in the same county or a bordering county; (2) an indicator for whether the Syrian migrant user produces content such as posts and comments in German; and (3) the number of local native groups a Syrian migrant user is in, where we consider groups with 5-5,000 users, of which  $\geq 50\%$  are identified as German natives and  $\geq 75\%$  are located in the Syrian migrant's NUTS2 region. We will also focus certain analyses on German natives, measuring the number of Syrian migrant friends each of these users has.

## 1.3 Sample Summary Statistics

Panel (a) of Table 1 summarizes the Syrian migrant sample. The median Syrian migrant user is 31 years old, with a 10th-90th percentile range of 22 to 48 years. The sample is 32% female, somewhat lower

than 40% in the administrative data.<sup>7</sup> The median number of Facebook friends and groups joined is 226 and 56, respectively. The median user in the Syrian migrant sample first used Facebook in Germany 23 quarters ago. About 8% list a German college on their profile.

**Table 1: Syrian Migrant and German Native Sample Summary Characteristics**

**Panel (a): Syrian Migrant Sample**

	Mean	SD	P10	P25	P50	P75	P90	P99
Age	32.90	10.26	22	25	31	38	48	66
Female (0/100)	32.07	46.68	0	0	0	100	100	100
DE College (0/100)	7.92	27.00	0	0	0	0	0	100
N Friends	347.89	385.84	62	117	226	423	751	2431
N Groups	104.55	137.09	8	22	56	129	256	831
Qs Since 1st on FB in DE	20.30	8.04	7	15	23	25	28	36
N Local Native Friends	5.03	12.24	0	0	1	4	13	87
N Local Syrian Friends	14.99	17.43	1	4	9	20	36	103
Produces DE Content (0/100)	30.40	46.00	0	0	0	100	100	100
N Local Native Groups	0.55	1.41	0	0	0	0	2	9

**Panel (b): German Native Sample**

	Mean	SD	P10	P25	P50	P75	P90	P99
Age	40.23	13.79	24	29	38	51	60	77
Female (0/100)	51.74	49.97	0	0	100	100	100	100
DE College (0/100)	32.93	47.00	0	0	0	100	100	100
N Friends	253.72	243.28	51	93	181	327	535	1535
N Groups	25.22	34.52	2	6	14	30	59	231
Qs Since 1st on FB in DE	31.87	8.26	18	33	36	36	36	36
N Local Native Friends	122.52	128.88	12	32	79	168	295	687
N Local Syrian Friends	0.09	0.34	0	0	0	0	0	2
Produces DE Content (0/100)	100.00	0.00	100	100	100	100	100	100
N Local Native Groups	3.98	4.92	0	1	2	5	10	26

**Note:** Table presents summary statistics describing users in our samples. Panel (a) shows users in the Syrian migrant sample. Panel (b) shows users in the German native sample. Each measure is winsorized at the 99% level. Section 1.1 describes sample construction. Appendix C provides more information on how individual-level outcomes are defined. Appendix Table A1 provides additional summary statistics.

Syrian migrant users have five native local friends on average. This magnitude is consistent with data from the German Socio-Economic Panel (SOEP), a longitudinal survey of German households that regularly includes modules that over-sample refugees or other groups. In the 2016 wave of the *IAB-BAMF-SOEP Survey of Refugees in Germany*, the average recent Syrian migrant in Germany reported to have 6.2 acquaintances, a number very close to the number of Facebook friends observed in our data.<sup>8</sup>

<sup>7</sup>Appendix Figure A1 colors observations in Figure 1 by gender and age, thereby benchmarking our sample against the true population as reported in the administrative data.

<sup>8</sup>The exact question asked by the SOEP is: "How many German people have you met since your arrival in Germany with whom you have regular contact?" The average responses reported in the text is based on responses from 1,095 survey respondents. The advantage of our data relative to the SOEP is that our sample size is nearly 320 times as large, allowing us to precisely estimate and explore regional differences in integration outcomes. Our data also allow us to use a mover research design to understand the determinants of this geographic variation.

By contrast, Syrian migrant users have 15 Facebook friendships with other Syrian migrants in the same location. About 30% of Syrian migrant users produce content on Facebook in German. At the median and 90th percentiles, Syrian migrant users are members of zero and two local native groups, respectively.

Appendix Figure A4 presents binned scatter plots showing relationships between our three primary integration outcomes—local native friends, German content production, and local native groups—at the individual level. There are strong positive relationships, both with and without controls for individual-level demographics and Facebook usage, providing evidence that our measures are capturing related aspects of social integration. The underlying correlations are high for individual-level measures (between 0.2 and 0.3), as shown in Appendix Tables A2 and A3.

Panel (b) of Table 1 summarizes characteristics of the German natives sample. The median user is 38 years old, with a 10th-90th percentile range of 24 to 60. The sample is 52% female and 33% of users list a German college on their profile. The median German native has a total of 181 Facebook friends, 79 local native friends, and 0.1 local Syrian migrant friends (users at the 99th percentile have two local Syrian migrant friends), highlighting that most German native users are not Facebook friends with a single Syrian migrant. German natives are members of four local native groups on average.

## 2 Individual-Level Characteristics and Integration Outcomes

In this section, we analyze the relationship between individual-level characteristics and integration outcomes of Syrian migrants in Germany. We begin by exploring average integration outcomes across migrant characteristics, before shifting to multivariate analyses.

Figure 2 focuses on the number of friendships with German natives and the share of Facebook content produced in German among the cohort of Syrian migrants with an “observed arrival” in 2015-2016.<sup>9</sup> The top row shows that these migrants become increasingly socially integrated as they spend more time in Germany. For example, after their first quarter in Germany, Syrian migrant users on average had 1.4 native friends and produced 1.7% of their Facebook content in German; three years later, these numbers were 7.3 friends and 4.2% of content, respectively. The bottom row shows considerable heterogeneity in the degree of integration across age and gender groups, with younger and male migrants integrating more quickly than older and female migrants. Three years after arrival, male Syrian migrants who moved to Germany between ages 13-18 had 14.4 native German friends, compared to 4.3 such friends for similarly aged females, and 3.1 such friends for males who arrived after age 45.

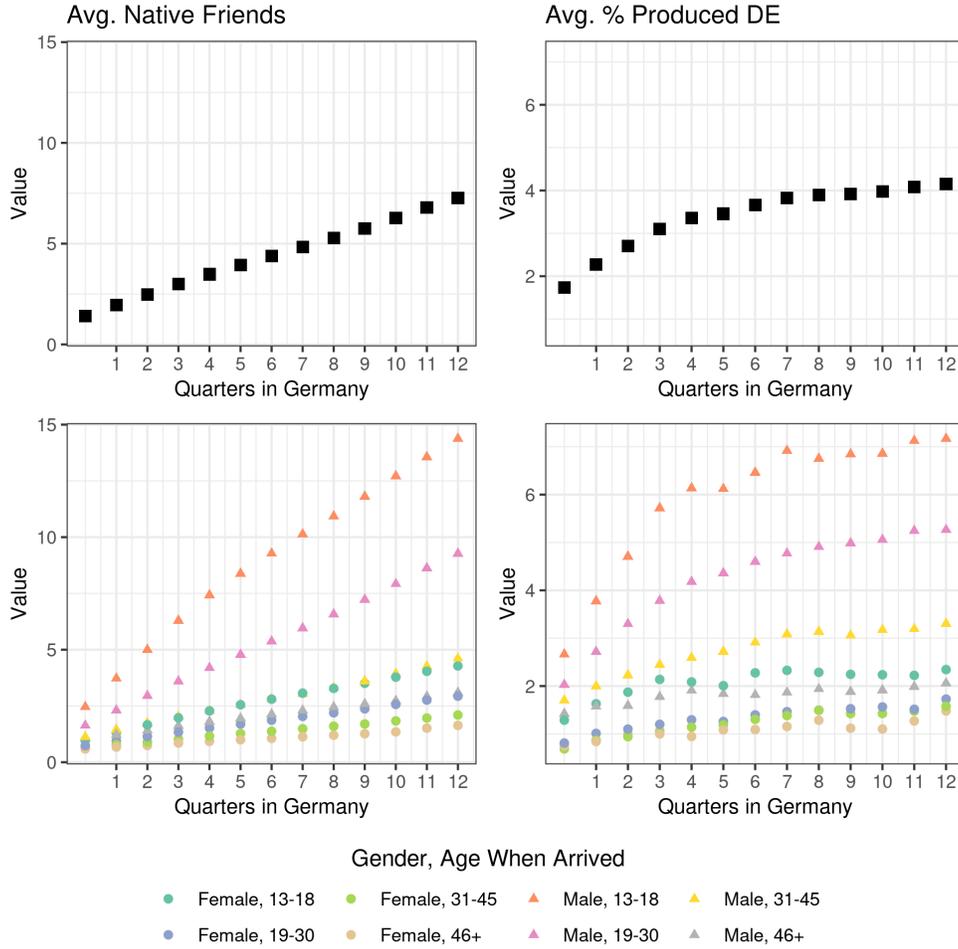
We next further explore this heterogeneity in integration outcomes across Syrian migrants using the following multivariate regression model:

$$Y_{i,j} = \alpha_0 + \alpha_1 Z_i + \psi_j + \epsilon_i. \quad (1)$$

For the results in columns 1-4 of Table 2,  $Y_{i,j}$  is the number of native local friends that individual  $i$  has in their current county of residence,  $j$ .  $Z_i$  is a vector of controls. All specifications include various controls for the amount of time users spend on Facebook, ensuring that differences in observed integration outcomes are not driven by variation in the intensity of Facebook usage. We also include fixed effects

<sup>9</sup>As described in Section 1.1, these are Syrian migrant users who first used Facebook *outside* Germany, then began using Facebook inside Germany in 2015 or 2016. Appendix Figure A5 reproduces this plot with additional integration measures.

**Figure 2: Integration Over Time For 2015-16 Arrival Cohort**



**Note:** Figures show the average values, by quarter, of integration measures for users in the Syrian migrant sample with an observed arrival in 2015 or 2016. The measures are total native friends (left column) and the share of content produced in German (right column). Appendix C provides more details on each measure. The top row shows overall trends. In the bottom row each observation’s shape and color represents a gender-by-age group.

for the user’s number of quarters since arrival in Germany and the number of quarters living in their current county. Some specifications also include fixed effects for the migrant’s county of residence,  $\psi_j$ .

In column 1,  $Z_i$  also includes dummies for age, gender, and whether the user has another Syrian migrant household member or non-household family member who was in Germany more than a year prior to their arrival.<sup>10</sup> Consistent with the univariate patterns in Figure 2, we find that younger and male Syrians befriend disproportionately many local German natives. All else equal, a female Syrian migrant has 3.7 fewer local native friends than a male does. Similarly, a Syrian migrant aged 55 or older has 4.6 fewer native local friends than a comparable individual under the age of 25. Column 1 also shows that, while migrants with a family member who arrived earlier in Germany *outside* of the household have more local native friends, individuals with an earlier arriving Syrian migrant *inside* their household have fewer local native friends. This result adds to prior findings that connections to other migrants

<sup>10</sup>Family and household information is determined through self-reports and model-based imputations. Similar data are used in Bailey et al. (2019a) and Chetty et al. (2022a,b).

**Table 2: Syrian Migrant Integration by Demographics - Friending to Natives**

	Facebook Sample				SOEP Sample	
	N Local Native Friends				N German Acquaintances	
Age 25 - 34	-1.012*** (0.053)	-0.894*** (0.052)	-0.873*** (0.052)	-1.148*** (0.129)	-0.839* (0.47)	-1.089** (0.47)
Age 35 - 44	-2.963*** (0.062)	-3.019*** (0.061)	-2.941*** (0.061)	-2.375*** (0.158)	-1.116* (0.58)	-1.070* (0.58)
Age 45 - 54	-4.012*** (0.080)	-4.102*** (0.079)	-4.147*** (0.079)	-4.765*** (0.184)	-2.362*** (0.78)	-2.238*** (0.77)
Age 55+	-4.548*** (0.100)	-4.531*** (0.098)	-4.586*** (0.099)	-7.226*** (0.241)	-3.378*** (1.24)	-3.594*** (1.23)
Female	-3.676*** (0.043)	-3.610*** (0.042)	-3.225*** (0.045)	-3.267*** (0.090)	-1.421*** (0.47)	-1.512*** (0.48)
Household Member in DE 1+ Year Prior	-0.377*** (0.100)	-0.290** (0.099)	-0.352*** (0.099)			
Non-Household Family in DE 1+ Year Prior	0.524*** (0.091)	0.621*** (0.089)	0.421*** (0.089)			
Quarters Since DE FEs	X	X	X	X	X	X
Prev Quarters in NUTS3 FEs	X	X	X	X		
Personal Usage Controls	X	X	X	X		
County / State FEs		X	X	X		X
Log (1 + Total Outside Germany Friends)			X	X		
Log (1 + Total Other Groups)			X	X		
Log (1 + Total Content Produced Past Year)			X	X		
Household FE				X		
N	349,072	349,072	349,072	84,216	1,095	1,095
R-Squared	0.132	0.160	0.165	0.658	0.048	0.093
Sample Mean	5.029	5.029	5.029	4.195	6.232	6.232

**Note:** Table shows results from regressing various measures on friend-based measures of integration. Each observation in columns 1-4 is a user in the Syrian migrant Facebook sample. Column 1 includes controls for age and gender, as well as fixed effects for the number of quarters on Facebook in their current county and the number of quarters since arrival in Germany. For the latter fixed effect, we use a single dummy value for those for which we do not observe arrival, but obtain nearly identical results if we instead drop these users. It includes dummies for whether the user has another Syrian migrant household member or non-household family member in Germany more than year prior to their arrival. For all users not in the “observe arrival timing” sample, these two dummies are set to 0. It also includes linear controls for log(0.5 + minutes on FB in the last 28 days), log(91 - days on Facebook out of the last 90), log(1081 - days on Facebook out of the last 1080) as measures of the Facebook usage intensity of each Syrian migrant. Column 2 adds county fixed effects. Column 3 adds controls for each user’s total number of friends outside Germany, total number of non-local/native groups joined, and total amount of content produced in the last year. Column 4 adds a household fixed effect, limiting to households for which we observe more than one Syrian migrant. Columns 5 and 6 use data from the Socio-Economic Panel in 2016. The dependent variable in these columns is the number of new acquaintances made in Germany. The exact survey question is, “How many German people have you met since your arrival in Germany with whom you have regular contact?” Each observation is a recent migrant from Syria living in Germany as of the date of the survey. Both columns 5 and 6 include controls for the number of quarters in Germany. Column 6 also controls for state fixed-effects. Significance levels: \*(p<0.10), \*\*(p<0.05), \*\*\*(p<0.01).

support integration in some settings and hinder it in others (e.g., Lazear, 1999; Edin, Fredriksson and Åslund, 2003; Cutler, Glaeser and Vigdor, 2008; Damm, 2009; Beaman, 2012; Martén, Hainmueller and Hangartner, 2019). In our context, the results suggest that somewhat-distant familial connections might provide support and guidance to help the social integration of newly arriving migrants, whereas the presence of close household connections might reduce the need to form connections with local natives.

Column 2 adds fixed effects for the Syrian migrants' current county of residence to the regression. The  $R^2$  increases by 21% from 0.132 to 0.160, consistent with there being important systematic regional differences in the social integration of Syrian migrants. We will explore this role of place in shaping integration in the following sections. The coefficients on the demographic characteristics in  $Z_i$  are largely unaffected by the addition of county fixed effects, suggesting there is a little selection based on these characteristics into more or less integrated places.

Column 3 adds controls for each user's total number of friends outside Germany, total number of groups joined, and total amount of recent content produced. These controls absorb additional variation in individuals' Facebook usage patterns beyond those in column 1, but could also remove variation in the true sociability of individuals that might influence their ability and desire to socially integrate with natives. While most coefficients remain largely unchanged, the gender coefficient falls somewhat in absolute terms, from -3.6 to -3.2. A possible interpretation is that Syrian migrant men generally have larger social networks, but, even conditional on overall network size, also make more German friends.

In column 4 of Table 2 we add household fixed effects while dropping individuals without additional household members from the sample. Even within the same household, and conditional on general Facebook usage patterns, younger and male Syrian migrants are better socially integrated.

Appendix Table A4 presents results analogous to column 1-4 of Table 2 for our key language- and group-based measures of social integration, and Table A5 presents results analogous to column 3 of Table 2 for a number of other outcomes. Across all measures, we find highly consistent relationships between age, gender, and family connections and the social integration of Syrian migrants.

One concern with this analysis may be that, despite our strict controls for Facebook usage and the consistency of our results across outcome, the observed differences in integration outcomes across demographic groups may still be driven by patterns of Facebook usage, rather than reflecting true demographic variation in social integration. To address this concern, we also look at related outcomes in the Socio-Economic Panel data, namely the number of native acquaintances made in Germany among a sample of recent Syrian migrants. In 2016, the SOEP administered a survey specifically targeted at recent migrants to Germany. We focus on the 1,095 Syrian migrants in the data that are 18+ years old.

Columns 5 and 6 show that the patterns of friending across demographics in the SOEP data mirror those we observe in the Facebook data in columns 1-4. Female and older migrants have fewer local acquaintances than male and younger migrants, respectively, on average. This holds with state fixed effects in column 6. Indeed, even the coefficient estimates using the Facebook and SOEP data are generally quite similar. We interpret this as reassuring as it shows that the patterns of social integration we identify in the Facebook data align closely with available survey evidence. The Facebook data, however, is much larger and more detailed, allowing us to more precisely explore the spatial variation in integration and to better understand the determinants of this variation. We do so in the following sections.

### 3 Regional Variation in Integration

In this section, we explore how places shape integration outcomes for Syrian migrants. We first construct measures of social integration at the county level, and document substantial regional heterogeneity in social integration outcomes that is not driven by sampling error or patterns of Facebook usage. Different

measures of social integration are also substantially correlated with each other across counties. We then use a movers research design to show that these spatial differences are, to a large extent, driven by causal place-based effects rather than selection in the type of migrants in each location. In subsequent sections we decompose local measures of integration to better understand their determinants.

### 3.1 County-Level Measures

In Table 2 we saw that county-level fixed effects explain a sizeable amount of variation in Syrian migrants' integration outcomes. To further explore this regional variation, we first estimate county-level averages for our different baseline measures of Syrian migrants' integration outcomes. These estimates have high reliability, suggesting that the observed differences in social integration do not arise from sampling error (see Appendix D).<sup>11</sup>

We want our county-level measures to reflect true differences in the social integration of Syrian migrants, rather than capture spatial patterns of Facebook usage. While we find no spatial differences in Facebook usage among Syrian migrants, there are small spatial differences in Facebook usage patterns of German natives, which could influence some of our measures of Syrian migrants' integration. For example, in a region where a somewhat smaller part of the native population uses Facebook, it might look as if local Syrians were relatively less well integrated according to the "local native friends" measure because we observe a smaller share of their actual friendships in the data. To account for such concerns, we residualize the observed average integration outcomes on county-level measures of the intensive and extensive Facebook usage of German natives. In the rest of this paper, we focus on exploring these residualized regional measures of the social integration outcomes of Syrian migrants.<sup>12</sup>

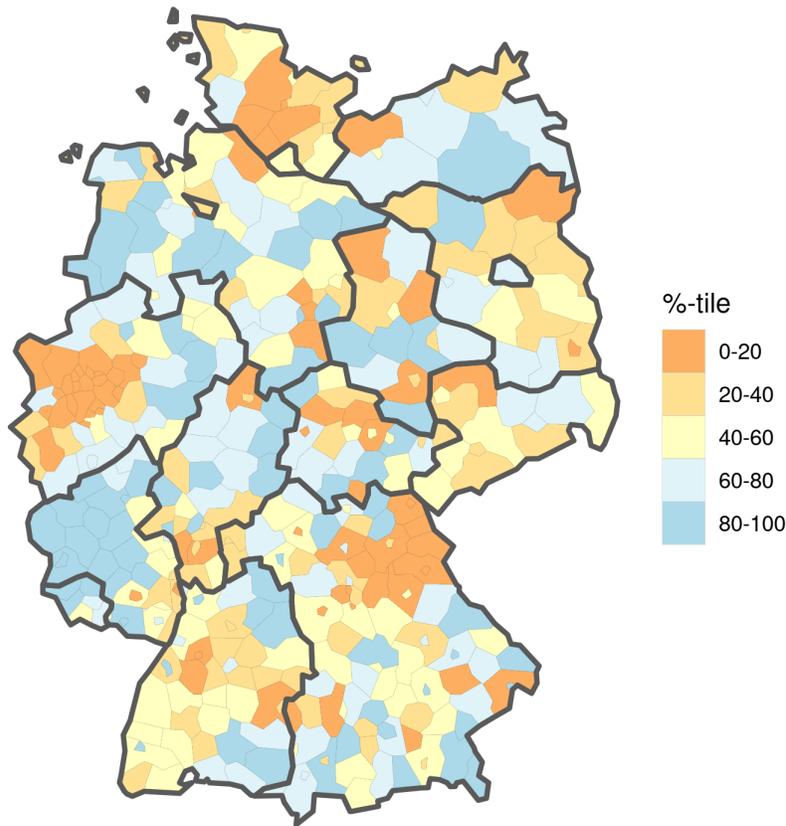
To validate these regional measures of the social integration of migrants, we compare them against the average number of native acquaintances made by recent Syrian migrants in Germany in the SOEP. These survey data are only available at less granular geographic levels. We can thus only compare the two data sources at the state (and state by age-group) levels. We find that, despite different definitions of friendships and small sample sizes in the SOEP data, the regional measures of social integration are correlated with  $\rho \approx 0.5$  across the two data sets (see Appendix Figure A7).

Figure 3 maps our county-level measure of friending integration (Appendix Figures A8 and A9 plot analogous maps for our language-based and group-based measures of social integration). Syrian migrants in a 90th percentile county make more than twice as many local native friends as Syrian migrants in a 10th percentile county (7.9 vs. 3.9). Consistent with anecdotal evidence in Nawras (2017), the social integration of migrants tends to be highest in rural areas: migrants living in counties along the southern border, in Rhineland-Palatinate (along the western border), in Lower Saxony (in the northwest), and in Mecklenburg-Western Pomerania (near the Baltic Sea in the northeast) each have particularly high levels of social integration. By contrast, many mid-sized cities—such as Ansbach, Kaiserslautern, and Cottbus—rank among the bottom 20% of places in terms of the average integration of migrants living there. Migrants living in larger cities, including Berlin, Munich, and Cologne, often have intermediate

<sup>11</sup>For example, Appendix Table A9 shows that if we randomly split the individual-level data into two halves and estimate our regional average of native friending in each half, the two are highly correlated with a correlation coefficient of 0.94.

<sup>12</sup>Due to Facebook business restrictions, we are unable to publicly characterize the spatial distribution of natives' Facebook usage patterns. We verify that the high reliability estimates documented above are not driven by usage differences: in Appendix Table A9, we show that the split-sample reliability after residualizing on usage remains 0.94.

**Figure 3:** Regional Estimates of Integration - Friending to Native Locals



**Note:** Figure shows county-level estimates of Syrian migrant integration based on the average number of local native friends among Syrian migrants in each county (residualized on regional patterns of German natives' Facebook usage). Colors correspond to measure ventiles. Darker red and blue areas indicate the lowest and highest integration counties, respectively.

levels of social integration. Interestingly, there do not appear to be systematic differences between East and West Germany, despite their history as distinct countries. We explore the determinants of these regional differences in migrant's integration in greater detail in Section 4.3.

Panel A of Table 3 shows population-weighted county-level correlations between our various integration measures. The different integration outcomes are positively correlated across counties: those counties where Syrian migrants make more German friends are also the counties where they are more likely to use the German language, and more likely to participate in local social groups.

The observed regional variation in integration outcomes for Syrian migrants could be explained by at least two different forces. One possibility is that places have causal effects on integration, either because of characteristics of the German natives living there, or because of institutional factors associated with the location. A second possibility is that there exist systematic differences in observable or unobservable characteristics of Syrian migrants by place that shape their propensity to integrate (for example, if migrants with knowledge of the German language are more likely to choose to live in certain areas). In the next section, we assess the relative importance of the characteristics of migrants vs. the features of locations by studying changes in the integration outcomes of migrants who move between counties.

**Table 3: Correlation Between Outcomes, Regional Level**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
<b>Panel A: Baseline Integration Measures</b>							
(1) SY Migrants - N Local Native Friends	1.00						
(2) SY Migrants - Produced Content in DE	0.59	1.00					
(3) SY Migrants - N Local Native Groups	0.25	0.49	1.00				
(4) SY Migrants - N Local SY Friends	-0.03	-0.51	-0.41	1.00			
<b>Panel B: Decomposition of Integration Measures</b>							
(5) General Friendliness	0.62	0.29	-0.04	0.11	1.00		
(6) Relative Friending	0.73	0.51	0.40	-0.16	-0.05	1.00	
<b>Panel C: Labor Market Integration Measure</b>							
(7) Share Syrians in Employment or Training	0.45	0.59	0.13	-0.36	0.29	0.33	1.00

**Note:** Table presents correlations between county-level estimates. The outcomes in panel A are the regional averages of Syrian migrants after residualizing on local German natives' Facebook usage, as described in Section 3.1. The outcomes in panel B are the regional decomposition measures described in Section 4.1. Row 5 is general friendliness, generated as the regional average of German natives local native friends after residualizing on local patterns of Facebook usage. Row 6 is relative friending, generated as the quotient from dividing the measure in row 1 by the measure in row 5. The outcome in panel C is an external county-level measure of the share of all Syrians that are employed or in training programs according to data from the federal employment agency (see Appendix J). Correlations are weighted by the number of Syrian migrant users in each county. Appendix Table A6 presents analogous signal correlations, which use our reliability estimates to remove noise due to sampling error.

Before doing so, however, it is important to note that asylum seekers in Germany are initially dispersed throughout the country in a quasi-random way and according to a formula based on local population and tax revenues (the *Königsteiner Schlüssel*), and that there are restrictions on resettlement. This institutional framework suggests migrants' current locations are largely determined by a process independent of any prior propensity to integrate. Indeed, in Appendix Section E we compare the distribution of refugees across place to the official assignment key and find that the two line up very closely, hence indicating that the assignment key has been followed relatively strictly even during these years of increased migration.

In addition, we can directly rule out that *observable* Syrian migrant demographics are driving the observed regional differences in average integration outcomes. In particular, regressing migrant's age, gender, and number of quarters since first using Facebook in Germany on county fixed effects results in  $R^2$ s of 0.005, 0.003, and 0.005, respectively, highlighting that these characteristics vary little across counties. This is consistent with the fact that adding county fixed effects in column 2 of Table 2 had little effect on the demographic coefficients in column 1, and the fact that regional integration measures with and without individual-level observable controls are highly correlated (see Appendix Figure A6).

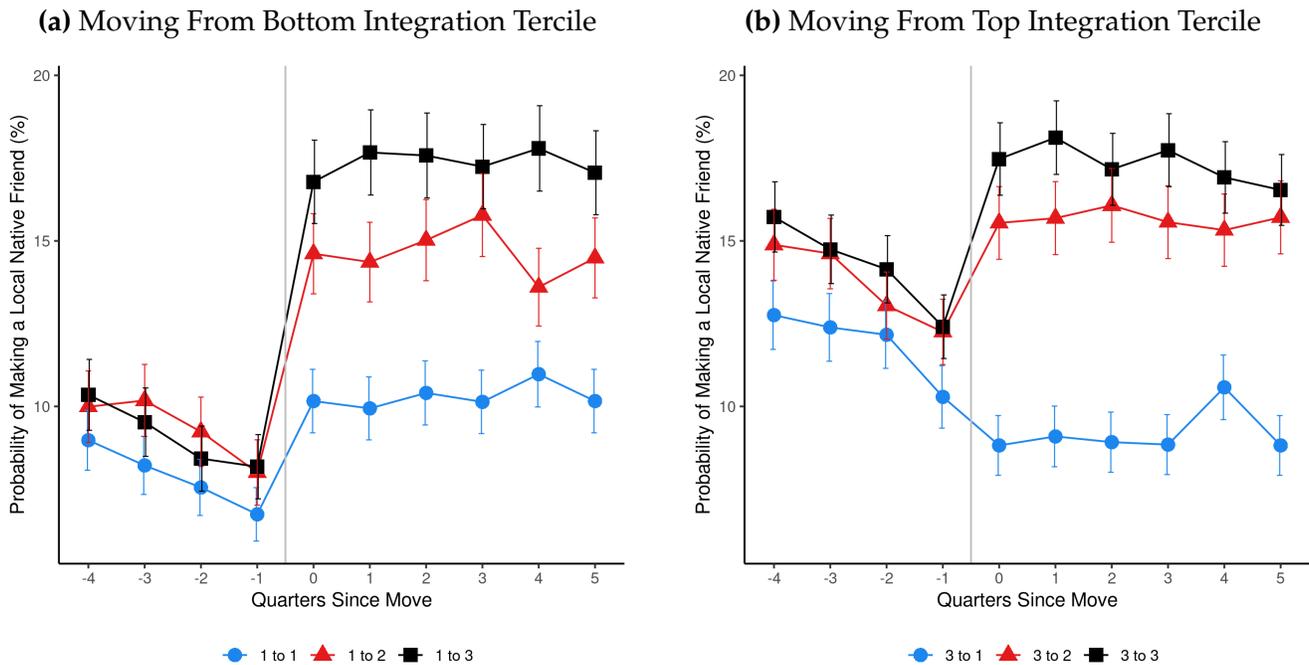
However, while highly suggestive, this evidence does not yet fully rule out that selection on unobservable characteristics might drive some of the observed regional variation in the integration of Syrian migrants. For example, while some restrictions exist on asylum seekers movement after settlement, these are less restrictive for individuals who arrived prior to August 2016 or who have been in Germany for more than three years (see the discussion in Hilbig and Riaz, 2020). Selection from this movement, or other factors that might limit the degree to which the initial place assignment is random, could confound our interpretation of the average regional integration outcomes as the causal effect of place. The next section thus presents an analysis that explores such unobservable selection.

### 3.2 Establishing Causal Place-Based Effects Using Movers

We next study the drivers of regional variation in integration outcomes using an empirical strategy that exploits movers to separate the role of place-based and non-place-based factors. This approach builds on recent work using similar designs to study place-based effects in different contexts (e.g., Card, Heining and Kline, 2013; Finkelstein, Gentzkow and Williams, 2016, 2019; Chetty and Hendren, 2018a,b).

For this analysis, we focus on Syrian migrants who move between German counties, and study changes around these moves in the moving migrants' propensity to befriend local natives. To see the intuition behind this research design, consider a Syrian migrant who moves from Ansbach, where we observe Syrians generally making few native friends, to Saarlouis, where they make more native friends. If the observed differences in the friending behavior of Syrians in Ansbach and Saarlouis were due to characteristics of the Syrians living in those places, we would expect the moving migrant's likelihood of befriending local natives to remain largely unchanged after the move. By contrast, if the observed geographic differences in the social integration of Syrians were primarily due to a causal effect of place, we would expect the moving migrant's likelihood of befriending native locals to increase by the average difference in this likelihood between the two locations. The magnitude of the change in the rate of befriending local Germans around a migrant's move thus identifies the importance of each explanation.

**Figure 4: Change in Syrian Migrants' Friending of Local Natives Around a Move**



**Note:** Figures show the quarterly probability that a moving Syrian migrant befriends a local German native, relative to the timing of the migrant's move within Germany. The population is Syrian migrant users who moved between non-neighboring counties and were in the first and second county for 4+ and 6+ consecutive quarters, respectively. The counties are grouped into terciles (weighted by the overall number of Syrian migrant users) of the regional friending-based measures of integration (mapped in Figure 3). Panels (a) and (b) limit to users who move from a county in the bottom and top tercile of integration, respectively. The different lines in each graph show movers to counties in each of the three terciles of social integration. The individual-level outcomes are residualized by the regional measures of Facebook usage described in Section 3.1. Bars display 95% confidence intervals of the estimates.

To begin, we construct a sample of Syrian migrant users who were in one county for four or more consecutive quarters followed by a different, non-neighborhood county for six or more consecutive quarters. We allow a user to be included for multiple moves so long as each move meets these criteria. Our sample includes 33,772 moves and 31,721 unique movers. In Appendix Figure A10, we compare the number of moves between counties observed in the FB data to admin data: we find that these measures of moves are extremely highly correlated. We interpret this as very reassuring as it indicates that our measures of moves reflect "true moves" rather than picking up Facebook usage behavior.

Figure 4 plots Syrian migrants' probabilities of befriending local natives around a move, where quarter = 0 is the first quarter we observe the migrant in their new location. Counties are grouped into terciles based on the integration measures mapped in Figure 3. Panels (a) and (b) focus on users who lived in a bottom and top tercile county in the year prior to moving, respectively. In each panel, the different lines correspond to movers to counties in different integration terciles. The vertical axis plots the probability that a migrant makes at least one local German friend in a given quarter, a flow measure of social integration that allows us to study changes in the rate of integration around a move. To avoid picking up differences in natives' Facebook usage across locations, we residualize this measure across user-quarters on the same intensive and extensive regional usage measures described in Section 3.1.

In both panels, the likelihood of migrants making new local German friends is somewhat decreasing prior to the move, consistent with individuals investing less effort in making new friends prior to a move. There is little significant variation in the rate of making local German friends across the destination tercile, suggesting that individuals moving to a high-integration place behaved fairly similar prior to the move to individuals moving to a low-integration places. Following the move, the probability of making local German friends varies systematically by the movers' destination, with higher probabilities for individuals moving to places with higher overall social integration levels. The pattern exists in both panels, which we interpret as first evidence for symmetric place-based effects. There is also an additive increase in the rate of making local friends following a move, independent of integration tercile in the origin and destination, consistent with all movers building new local networks in their destinations.

In Section 3.1 we ruled out migrant *observables* determining regional patterns of integration. Building on the descriptive results of this section, and prior studies of place effects (see Chyn and Katz, 2021, for one overview), we now propose a simple model in which a Syrian migrant's rate of befriending local natives is determined by the sum of place-based effects—which we allow to vary across time and with observable migrant characteristics—and other *unobservable* individual-level factors. Since only place-based factors change at the time of a move, this model will allow us to estimate the share of regional variation in the social integration of migrants that can be attributed to place-based effects rather than individual characteristics. Appendix F provides more details on this model, which is similar to Finkelstein, Gentzkow and Williams (2016).

We bring this model to the data by comparing the rate at which moving migrants make native German friends in the year before and after their move to the average friending rates of otherwise similar non-movers in each location.<sup>13</sup> For each user  $i$  moving in quarter  $t$ , the outcome of interest is the

---

<sup>13</sup>In this analysis we limit to movers who were in their origin and destination counties for four or more consecutive quarters each, less stringent than the prior analysis which required six quarters in the destination. In addition, we only include observations for which there is at least five "matched" non-movers in both the origin and destination.

change in the quarterly probability of making at least one local German friend,  $y_{i,t}^\Delta$ , defined as:

$$y_{i,t}^\Delta = 0.25 \left[ \sum_{\tau=t}^{t+3} Y_{i,\tau} - \sum_{\tau=t-4}^{t-1} Y_{i,\tau} \right] \quad (2)$$

Here,  $Y_{i,t}$  is an indicator for whether Syrian migrant  $i$  makes at least one local German friend in quarter  $t$ , the individual-level measure underlying Figure 4. Similar to before, we first residualize each side of the difference on regional measures of natives' Facebook usage. To compare  $y_{i,t}^\Delta$  to differences in the average integration rates of observably similar non-movers in each place, we construct sets of users who match each mover on the important determinants of social integration in Section 2: gender, age group, and time spent in Germany. Formally, for user  $i$  moving in quarter  $t$ , we let  $O(i, t)$  and  $D(i, t)$  be the sets of similar non-movers in the origin at time  $t - 4$  and in the destination at time  $t$ . We then define the differences in their average outcomes,  $x_{i,t}^\Delta$ , as:

$$x_{i,t}^\Delta = 0.25 \cdot \left[ \frac{1}{|D(i, t)|} \sum_{j \in D(i, t)} \sum_{\tau=t}^{t+3} Y_{j,\tau} - \frac{1}{|O(i, t)|} \sum_{j \in O(i, t)} \sum_{\tau=t-4}^{t-1} Y_{j,\tau} \right] \quad (3)$$

The set cardinalities  $|O(i, t)|$  and  $|D(i, t)|$  are the number of non-movers in the observable matched comparison groups for each mover. Intuitively,  $x_{i,t}^\Delta$  is the difference in the average quarterly probability of a non-mover migrant making a native local friend between the destination location in the year after the move and the origin location in the year before the move. Time-specific measures allow for changes in the differences between regions over time. We find this appealing, as regions implemented policies at different times, possibly impacting their relative levels of integration. Again, we residualize each side of the difference on regional measures of natives' Facebook usage. We then estimate:

$$y_{i,t}^\Delta = \alpha_0 + \alpha_1 x_{i,t}^\Delta + \zeta_t + \epsilon_{i,t} \quad (4)$$

where slope  $\alpha_1$  is our outcome of interest. An estimate of  $\alpha_1$  close to 1 would suggest that, within the first year of moving, migrant movers' friending behavior fully adjusts to the level of local non-movers' friending behavior. An  $\alpha_1$  close to 0 would suggest that migrants do not adjust their friending rates systematically toward the level of local non-movers. Because migrant observables do not differ significantly across space, under the relatively weak identification assumptions discussed below,  $\alpha_1$  estimates the share of the observed differences in the social integration of migrants across locations that are due to causal place-based effects rather than unobservable individual characteristics. The quarter of move fixed effect,  $\zeta_T$ , remove variation in overall time trends in the rates of befriending local natives.

One challenge with our estimation is that we only observe a sample estimate of each mover's  $x_{i,t}^\Delta$ , denoted by  $\hat{x}_{i,t}^\Delta$ . Measurement error in the true differences in friending probabilities of non-movers across locations would thus lead to attenuation bias in  $\alpha_1$ . To account for this sampling error, when estimating equation 4, we randomly split the individual-level data of the friending behavior of non-movers used to construct  $\hat{x}_{i,t}^\Delta$  into two sub-samples and instrument for the value constructed in one sub-sample with the value constructed in the other sub-sample (see Appendix D for details).

**Identification Assumptions.** Our interpretation of  $\alpha_1$  relies on a number of identification assumptions. In particular, our model requires place-based effects to be additive and additively separable from any unobservable individual-level factors. Additivity implies, for example, that a move from place A to place B should have the same effect as a move from place B to place A. This is strongly supported by Figure 4, as well as the results in Figure 5 and Table 5. Additive separability implies that migrants’ friending rates between locations will vary by the same *absolute amount* across unobservables (the model does, however, allow for non-additive relationships between our observables—gender, age, and time in Germany—and migrants’ friending rates). Our identification also relies on there being no systematic shocks to unobservable factors that coincide exactly with the move quarter and affect native friending differentially by origin and destination. We find it unlikely that such shocks exist.

Importantly, the model allows for movers to differ from non-movers on observables and unobservables and for these differences to correlate with origin and destination. It also allows the level of movers’ pre-move friending within an origin county to correlate with destination due to differences in individual characteristics. Movers’ native friending around a move can also differ from the trends of non-movers. This could occur if, as suggested by Figure 4, all movers make fewer local connections in anticipation of a move or more connections immediately after a move. Each of these would increase  $\alpha_0$ , but leave  $\alpha_1$  unaffected. While our model would be affected if, for example, these downward trends in movers’ propensity to make friends before relocating differed systematically by the integration levels in the movers’ destinations, Figure 4 provides evidence that such differential trends do not exist.

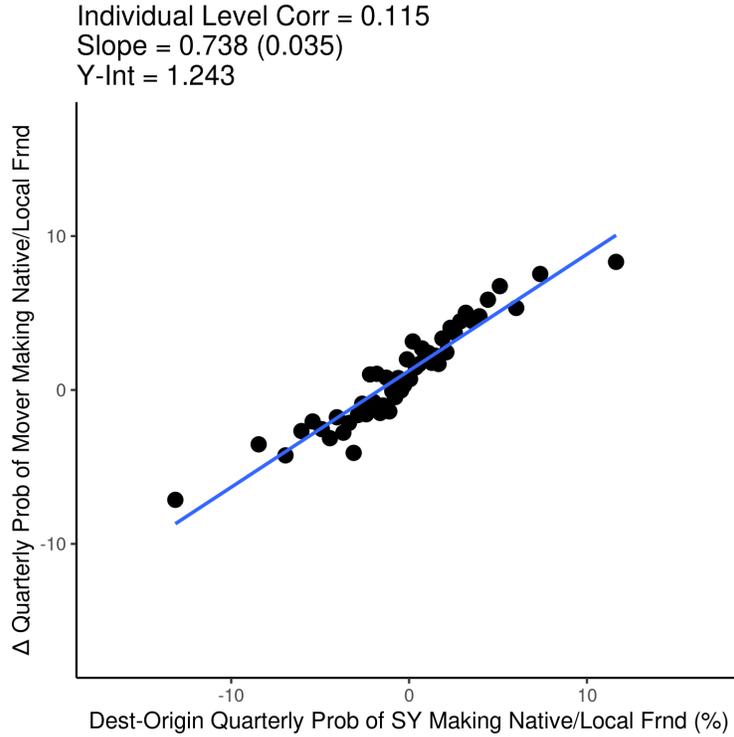
**Table 4: Syrian Migrant Mover and Comparable Non-Mover Sample Summaries**

	All		To Below Median Place		To Above Median Place	
	Movers	Matched	Movers	Matched	Movers	Matched
% Female	18.70	18.70	19.54	19.54	17.95	17.95
Avg Age	27.97	27.49	27.98	27.51	27.97	27.47
Avg Qs in DE	6.47	6.42	6.54	6.50	6.40	6.36
Avg Friends Made (total in year)	44.72	43.97	44.78	44.07	44.66	43.87
% of Qs Produ in DE	45.77	45.01	44.31	44.01	47.09	45.90
% of Qs Makes Native Local Friend	11.80	17.18	10.51	16.72	12.96	17.60

**Note:** Table presents summary statistics describing the movers underlying Figure 5 and their matched non-movers in their origin. Movers are matched to non-movers on county, time, age group (18-29, 30-39, 40+), gender, and the year we first observed the user on Facebook in Germany (i.e., same arrival “cohort”). To be in the final sample, a mover must be matched to five or more non-movers in both the origin and destination. Measures are constructed using the movers’ information in the year prior to the move and their matched users in the origin location and time. Matched non-mover summaries are generated by first constructing measures within each mover’s set of matched movers, then averaging across these measures. “Avg Friends Made” is constructed from summing quarterly measures that are winsorized at the 99% level across all migrant user-by-quarter observations. “% of Qs Makes Native Local Friend” is residualized by local natives’ Facebook usage.

**Comparing Movers and Matched Non-Movers.** Table 4 summarizes our sample of movers and the corresponding matched sample of otherwise similar non-movers in the origin location. The matching is on county, time, age group (18-29, 30-39, 40+), gender, and the year we first observed the user on Facebook in Germany (i.e., same arrival “cohort”). Movers and non-movers are similar or identical on the matched demographics. They also make any friends on Facebook and produce German content at similar rates. Movers are less likely to befriend local natives than non-movers prior to the move—for example, because the anticipation of a move reduces the incentives to make local friends. As discussed above, this will result in a positive estimate of  $\alpha_0$ , but leave our estimate of  $\alpha_1$  unaffected.

**Figure 5:**  $\Delta$  Syrian Migrant Mover Friending Integration vs. Matched Non-Movers



**Note:** Figure shows a binned scatter plot describing the change in the friending of Syrian migrants to German natives before and after a move within Germany. The population is Syrian migrant users who moved between two non-neighboring counties and were in the first and second county for 4+ consecutive quarters each. The y-axis displays  $y_{i,t}^{\Delta}$ , movers' change in the quarterly probability of making a native local friend the year before to after the move. The x-axis displays  $x_{i,t}^{\Delta}$ , the difference in average outcomes for comparable non-movers at the same time. We match each mover to a set of non-movers who lived in the origin location a year before the move and to a set who lived in the destination location at the move. In addition we also match movers to non-movers of the same gender and age bucket (18-29, 30-39, 40+), and whom we first observed on Facebook in Germany in the same year (i.e., same "cohort"). We include observations for which there is at least 5 non-movers in both the origin and destination match group. We control for quarter of move fixed effects. We correct for sampling error in the x-axis measures by randomly splitting the individual-level non-mover data into two halves and instrumenting for one set of averages with the other. See Appendix D for more information this procedure. Standard errors are shown in parentheses.

**Results.** Figure 5 displays a binned scatter plot of  $y_{i,t}^{\Delta}$  against  $x_{i,t}^{\Delta}$ , with the slope corresponding to  $\alpha_1$  in equation 4. The relationship is symmetric around zero and linear, highly consistent with additive effects of place. The slope estimate is 0.738, which suggests that nearly three quarters of the observed regional variation in Syrian migrants' friendship formation with local natives is attributable to place-based effects that occur within the first year of after their move, rather than individual characteristics. In Appendix Figure A11 we plot the slope estimates separately for samples of users that are male, female, younger than 30 years old, 30 to 39 years old, and over 40 years old. For each group, the estimates remain similar, suggesting our results are not driven by any particular demographic group of Syrian migrants.

Table 5 presents further details on the relationship between the friending patterns of movers and stayers. In column 1, we replicate the regression underlying Figure 5. In column 2, we regress  $y_{i,t}^{\Delta}$  separately on the origin and destination components of  $x_{i,t}^{\Delta}$ , resulting in slope magnitudes between 0.71 and 0.78. Columns 3 and 4 show that  $\alpha_1$  remains similar when we include origin fixed effects or destination fixed effects. Jointly, the estimates provide further evidence that places have additive effects.

**Table 5:**  $\Delta$  Syrian Migrant Mover Friending Integration vs. Matched Non-Movers: Robustness

	Change Quarterly Prob of Making Native Local Friend			
Dest-Origin Quarterly Prob of SY Making Native Local Friend	0.738*** (0.036)		0.758*** (0.051)	0.724*** (0.053)
Origin Quarterly Prob of SY Making Native Local Friend		-0.712*** (0.037)		
Dest Quarterly Prob of SY Making Native Local Friend		0.773*** (0.037)		
Quarter FEs	X	X	X	X
Origin County FEs			X	
Dest County FEs				X
N	32,853	32,853	32,849	32,845
Sample Mean	0.934	0.934	0.933	0.938

**Note:** Table shows results from regressions exploring the change in friending of Syrian migrants to German natives, before and after a move within Germany. Column 1 corresponds to the relationship depicted in Figure 5. Column 2 regresses each component of the difference in the right-hand side measure in column 1 separately on the outcome. Columns 3 and 4 repeat column 1 with origin and destination fixed effects, respectively. We correct for sampling error in the right-hand side measures by randomly splitting the individual-level non-mover data into two halves and instrumenting for one set of averages with the other. See Appendix D for more information this procedure. Significance levels: \*( $p < 0.10$ ), \*\*( $p < 0.05$ ), \*\*\*( $p < 0.01$ ).

While this section has focused on measures of social integration based on migrants’ friending patterns, in Appendix G we turn to our language-based measure of integration. Whereas our friending analysis was able to work with panel data in quarterly friending rates, our language outcome—whether the user produces content in German—is only observable at high quality in the cross section. As such, we instead explore how a mover’s language-based integration *today* is shaped by the set of places they have lived, following similar analyses in Chetty and Hendren (2018a) and Finkelstein, Gentzkow and Williams (2019). Our results are consistent with place-based effects driving much of the cross-sectional variation in Syrian migrants’ German language usage.

The prior results have documented that when Syrian migrants move between German counties, their social integration patterns quickly adjust from those of their origin towards those of their destination county. Our results therefore provide evidence that most of the observed regional differences in social integration are explained by the effect of places—either due to institutional factors associated with the location, or due to local native characteristics—rather than by the characteristics of the migrants. In this context, it is important to note, when interpreting our estimates in Figure 5 and Table 5, that a mover design will not even capture the full extent to which individual factors are shaped by place-based effects. For example, Syrian migrants who learn the German language in high integration places (possibly due to having attended integration courses) might then use these skills to make German friends more quickly after moving to a low-integration place. This effect might be considered “place-based” in the sense that it is shaped by features of the mover’s origin location, but will not be captured by our estimates. To the extent that such additional long-term place-based effects are important, our estimates of  $\alpha_1$  will even *understate* the extent to which places truly shape migration outcomes.

Given our strong evidence for large causal effects of locations on migrants’ social integration, in the next section we ask: “What makes migrants more likely to integrate in one place versus another?”

## 4 Determinants of Place-Based Integration

We begin our analysis of the determinants of regions’ integration outcomes by decomposing the regional variation in migrants’ integration outcomes into variation in the rate at which natives make local friends in general (*general friendliness*) and variation in the relative probability of natives befriending Syrians versus other natives (*relative friending*). Both of these objects are important to understanding regional variation in migrants’ integration outcomes. To quantify how these two measures are determined by the institutions of a place versus the preferences of local natives, we study how *natives’* friending behavior changes when they move between counties. Our results suggest that regional differences in Syrians’ social integration are primarily driven by structural or institutional features of locations rather than the preferences of natives living there. Motivated by this finding, we explore the relationship between general friendliness, relative friending and other regional characteristics. We provide evidence that the local provision of integration courses plays an important causal role in shaping integration outcomes through increasing relative friending.

### 4.1 General Friendliness and Relative Friending

Many studies have characterized the *overall* likelihood of community members forming connections with one another (see Coleman, 1988; Putnam, 1995a; Jackson, 2020; Chetty et al., 2022a,b). Motivated by this literature, we distinguish between two forces that can contribute to the regional variation in migrants’ social integration. The first, which we call *general friendliness*, is the overall rate at which natives befriend others in their community: if local natives in a given county are more likely to befriend any neighbor, they might also be more likely to befriend their Syrian migrant neighbors. The second force, which we call *relative friending*, is the relative probability of a German native befriending a local German versus a local Syrian migrant. Our unique data allow us to measure these two components separately, furthering our understanding of the causal effects of place documented in Section 3.

To formalize this, we define a county’s general friendliness as German natives’ average number of local German friends. We define relative friending as Syrian migrants’ average number of local German friends divided by general friendliness. General friendliness and relative friending thus determine friending integration multiplicatively:

$$\underbrace{NLocalFriends_j^{SY \rightarrow DE}}_{\text{Friending Integration}} = \underbrace{NLocalFriends_j^{DE \rightarrow DE}}_{\text{General Friendliness}} \times \underbrace{\frac{NLocalFriends_j^{SY \rightarrow DE}}{NLocalFriends_j^{DE \rightarrow DE}}}_{\text{Relative Friending}} \quad (5)$$

where  $NLocalFriends_j^{DE \rightarrow DE}$  and  $NLocalFriends_j^{SY \rightarrow DE}$  are the average number of local native friends among native and Syrian migrant users in county  $j$ , respectively, after residualizing on regional patterns of Facebook usage in the native population as in Section 3.1.

Intuitively, relative friending captures how much harder it is for a Syrian migrant to make a local native friend than it is for a native German to make that friend. To further build intuition for its determinants, it is possible to re-write county-level relative friending as a function of only natives’ friending behaviors, using the fact that within a county the total number of friendships from local natives to local

Germans must equal the total number of friendships from local Germans to local natives:

$$Rel. \text{ Friending} = \frac{NLocalFriends_j^{SY \rightarrow DE}}{NLocalFriends_j^{DE \rightarrow DE}} = \frac{NLocalFriends_j^{DE \rightarrow SY}}{NLocalFriends_j^{DE \rightarrow DE}} \times \frac{NGer_j}{NSyr_j} = \frac{\frac{NLocalFriends_j^{DE \rightarrow SY}}{NLocalFriends_j^{DE \rightarrow DE}}}{\frac{NSyr_j}{NGer_j}}. \quad (6)$$

Here,  $NGer_j$  and  $NSyr_j$  are the numbers of German native and Syrian migrant Facebook users local to county  $j$ , respectively.  $NLocalFriends_j^{DE \rightarrow SY}$  is the average number of local Syrian friends of German natives in county  $j$ . Relative friending will thus be equal to one if German natives befriend local Syrian migrants and other local German natives in proportion to their population shares.

Panels (a) and (b) of Figure 6 map general friendliness and relative friending by county, respectively, while Panel (c) shows their across-county correlation (with different colors corresponding to different overall levels of integration). General friendliness is higher in South-Western states and lower in Northern Germany, while relative friending is generally higher in Northern Germany. The industrial areas in the Ruhr area of Northrhine-Westphalia—including the cities of Duisburg, Oberhausen, Bottrop, Gelsenkirchen, and Herne—as well as parts of upper Franconia in northern Bavaria have low general friendliness and low relative friending. Accordingly, in Panel (c) these counties are clustered in the bottom-left and colored in red, indicating they have the lowest levels of Syrian migrant social integration. Overall, general friendliness and relative friending are weakly negatively correlated across counties, with a weighted correlation coefficient of -0.05. Section 4.3 explores correlates of this spatial variation in both general friendliness and relative friending.<sup>14</sup>

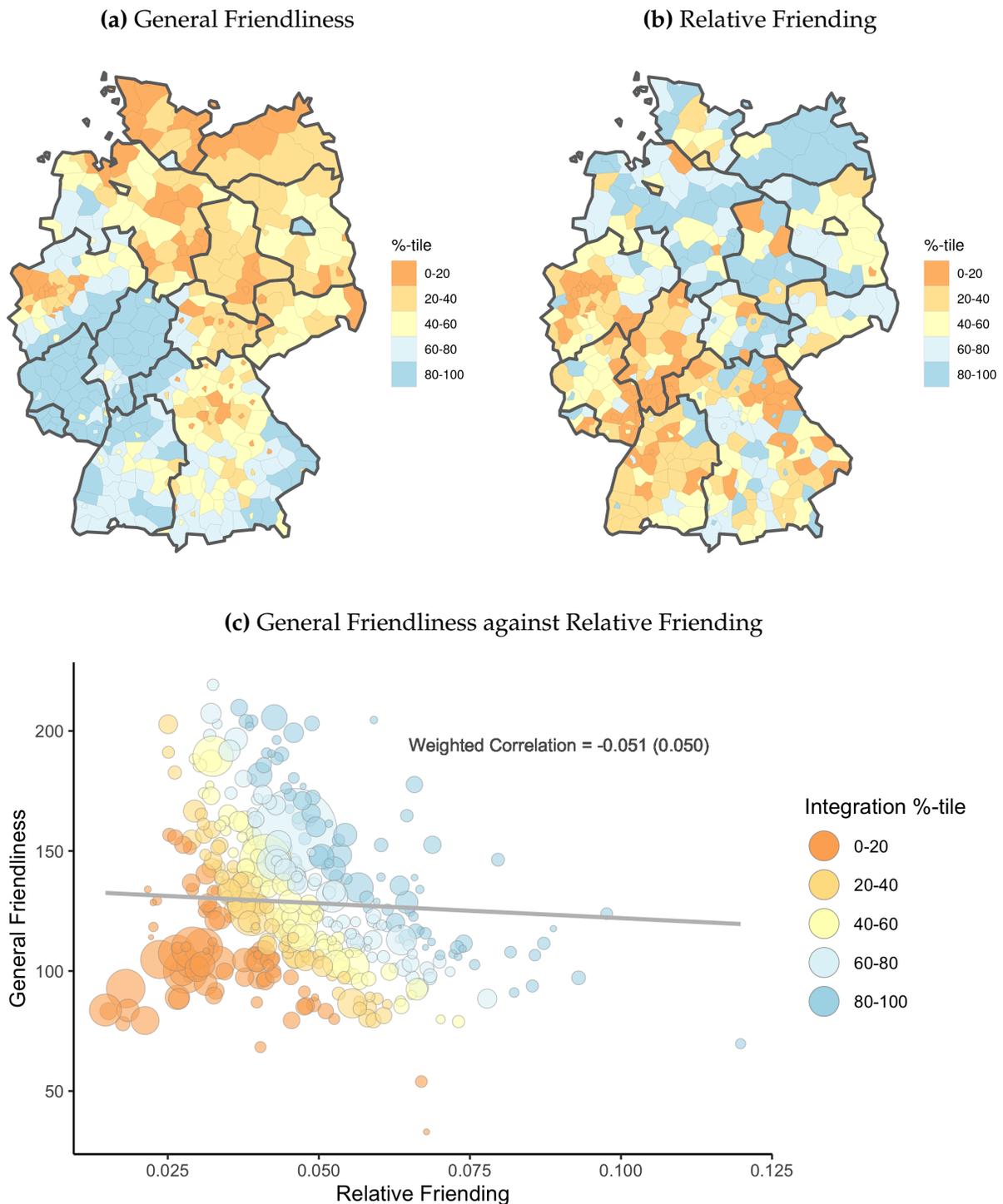
To quantify the relative importance of general friendliness and relative friending in explaining county-level differences in integration outcomes, in columns 1 and 2 of Table 6 we separately regress the log of overall friending integration on the log of each component. The  $R^2$  estimates of 0.41 and 0.66 for general friendliness and relative friending, respectively, suggest that relative friending explains 50% more of the geographic variation in Syrian migrants' integration than general friendliness does.<sup>15</sup>

For some questions, it may not be central to determine whether good integration outcomes in a given place are driven by high general friendliness or high relative friending. For instance, a policy-maker interested in simply assessing the potential of different regions to socially integrate migrants—perhaps because they are interested in determining where to settle new refugees—may be indifferent

<sup>14</sup>In addition to these geographic differences in general friendliness and relative friending, Appendix I shows that individual-level versions of these objects also differ across native demographics. Overall, younger and male natives have more Syrian migrant friends than older and female natives, respectively. These patterns are driven in part by variation in general friendliness (i.e., these natives have more friends in general) as well as by variation in relative friending (i.e., these natives are relatively more likely to befriend a given local migrant vs. a local German). Because Syrian migrants in Germany are more likely to be young and male than the average German native, one possible explanation for this finding is that homophily plays a strong role in shaping which natives befriend Syrian migrants. For example, younger German natives might be more likely to connect with younger Syrian migrants because younger people are more likely to connect in general, rather than because of particular behaviors toward migrants. Consistent with such an interpretation, we show that it is, in fact, older and female natives that are more likely than others to join pro-immigration groups on Facebook, conditional on Facebook usage. These are *opposite* to the relationships we find for relative friending, suggesting that is not necessarily those who are most supportive of pro-immigration groups that are most likely to befriend Syrian migrants.

<sup>15</sup>In Appendix H, we conduct a counterfactual exercise to further explore the degree to which the two components explain the difference between high- and low-integration counties. The results suggest that relative friending explains 1.9 times as much of this difference as general friendliness does.

**Figure 6: Regional Estimates of General Friendliness and Relative Friending**



**Note:** Panel (a) shows county-level estimates of general friendliness, the average number of local native friends among natives in each county (residualized on Facebook usage). Panel (b) shows county-level estimates of relative friending, given by the ratio of the overall friending integration measures and general friendliness (see equation 5, also residualized on Facebook usage). Colors correspond to measure ventiles. Darker red areas indicate the lower values of general friendliness and relative friending, and darker blue areas indicate higher values. Panel (c) shows a county-level scatter plot of relative friending against general friendliness. The size of bubbles corresponds to the number of Syrian migrants in the county. Darker red observations have the lowest friending integration (mapped in Figure 3) and darker blue have the highest.

**Table 6: County-Level Relationship Between Integration Measures**

	<b>Friending Integration</b>		<b>Language</b>		<b>Employment / Training</b>	
General Friendliness	1.098*** (0.13)			0.183*** (0.07)		0.558*** (0.08)
Relative Friending		1.056*** (0.07)		0.255*** (0.03)		0.459*** (0.06)
Friending Integration			0.228*** (0.04)		0.494*** (0.05)	
N	401	401	401	401	385	385
R-Squared	0.408	0.664	0.367	0.374	0.353	0.356

**Note:** Table shows results from multivariate regressions exploring the county-level relationship of integration measures with general friendliness and relative friending. In every specification, the outcomes and all controls are measured in logs. The outcomes are friending integration (columns 1 and 2), the share of Syrian migrants on Facebook who produce German content (columns 3 and 4), and the share of Syrians employed or in training programs (columns 5 and 6) according to data from the federal employment agency (see Appendix J. Regressions are weighted by the number of Syrian migrants in the Facebook data. Significance levels: \*( $p < 0.10$ ), \*\*( $p < 0.05$ ), \*\*\*( $p < 0.01$ ).

to which of the components drive this integration. Indeed, columns 4 and 6 of Table 6 show that both components of social integration have strong and similarly-sized positive associations with language- and economic-based measures of integration that policymakers may care about.<sup>16</sup>

However, the distinction between general friendliness and relative friending might be crucially important for policymakers in other settings. For instance, consider a policymaker seeking to actively improve integration outcomes in a given location. While targeted policies may potentially reduce the gap between natives' rate of befriending local migrants versus other locals (i.e., their relative friending), increasing the overall friending rate of natives (i.e., their general friendliness) is likely more challenging. Consistent with this, in Section 4.4 we show that the causal effects of integration courses on Syrians' integration outcomes is driven entirely by relative friending in a given region, rather than general friendliness. Because general friendliness and relative friending shape overall integration multiplicatively, interventions that raise relative friending will increase integration most in places with high general friendliness. Said differently, increasing relative friending when natives rarely befriend their neighbors will impact integration less than when natives often befriend their neighbors. The ability to measure each component separately therefore allows policymakers to most effectively target interventions.

<sup>16</sup>The results in column 4 and 6 also suggests that the positive correlations between social integration and German language proficiency in column 3 and between social integration and employment outcomes in column 5 are not uniquely driven by language proficiency and employment *causing* integration. In particular, such causal relationships would likely be mediated largely through relative friending: Syrian migrants with better German language skills could raise the relative rate at which natives befriend migrants (relative friending), but would be unlikely to affect the rate at which natives befriend each other (general friendliness). The positive relationship between general friendliness and language skills/employment outcomes is instead more consistent with friending integration impacting language acquisitions (i.e., having more German friends makes it easier to learn German) and employment (i.e., having more German friends makes it easier to find a job), though we also cannot rule out the presence of some omitted variable that is correlated both with general friendliness of natives and migrants' ability language acquisition and labor market outcomes.

## 4.2 To What Extent Are Native Behaviors Place-Based?

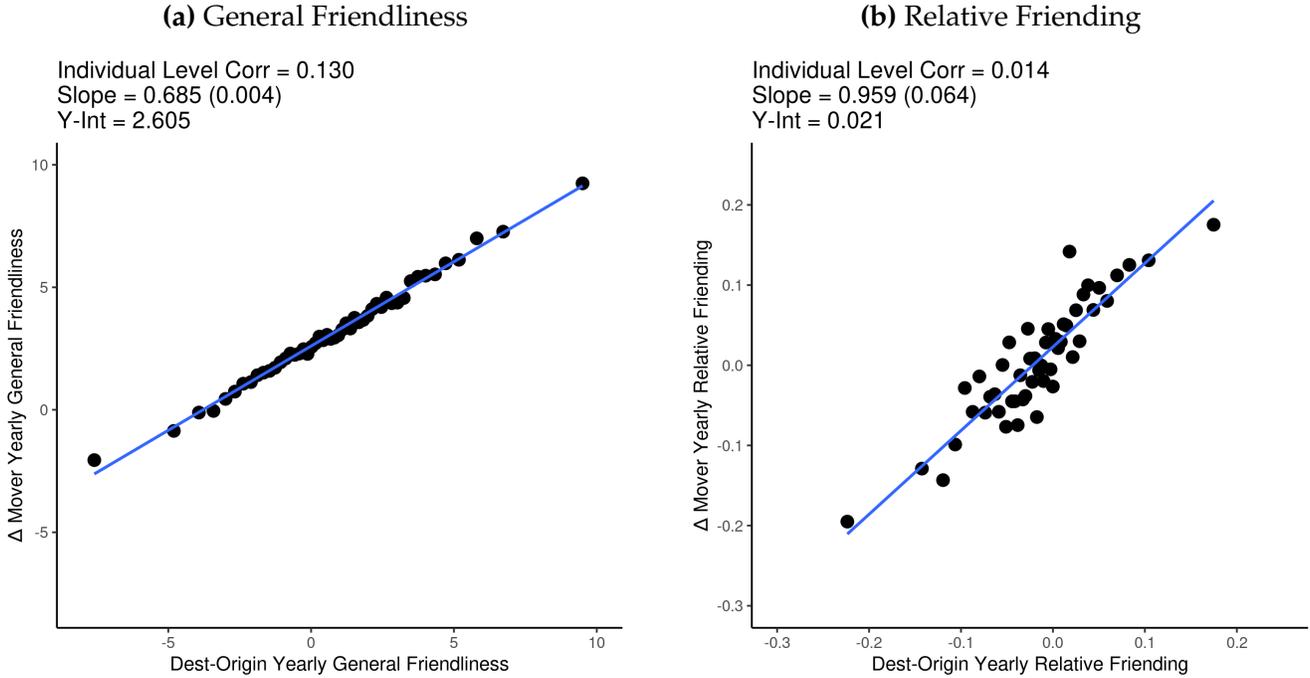
Having decomposed the observed social integration of migrants into the components of general friendliness and relative friending, we next ask what role native characteristics (e.g., attitudes toward neighbors or migrants) versus place-based effects (e.g., the structure of local institutions) play in shaping these determinants. To answer this question, we return to a movers design similar to that introduced in Section 3.2. We focus now on *natives* who move between counties, exploring their local friending patterns before and after the move. The intuition behind our analysis is similar to before. When German natives move between counties with differing levels of general friendliness or relative friending, the degree to which their own behavior changes in their new environment identifies the importance of individual-level versus structural local factors. As in Section 3.2, we focus on users who moved between two non-neighboring counties and who were in the origin and destination counties for at least four consecutive quarters. In this section, we focus on moves that occurred in Q1 2017 or later, when a substantial number of Syrians had already entered Germany, and we can measure relative friending more precisely.

Returning to equation 4, we let  $y_{i,t}^{\Delta}$  be the change in *yearly general friendliness* or the change in *yearly relative friending* around a move. Yearly general friendliness is the number of local native friends a user makes in a given year. Yearly relative friending is the ratio of local Syrian migrant friends to local native friends made by a German native in a given year, compared to the relative population shares of Syrian migrants and natives in that location (i.e., an annualized version of the “ratio of ratios” introduced in equation 6). Similar to before,  $x_{i,t}^{\Delta}$  captures the corresponding average individual-level measures of native stayers in the same place at the same time and in the same gender  $\times$  age group as the mover. Appendix Table A7 summarizes the characteristics of the sample of native movers and matched non-movers. As before, we only observe the sample-based  $\hat{x}_{i,t}^{\Delta}$ . We therefore adjust for the resulting attenuation bias in our estimate  $\hat{\alpha}_1$  in equation 4 by randomly splitting the sample of stayers in half and instrumenting for the estimate of  $x_{i,t}^{\Delta}$  constructed using one half with the estimate constructed using the other (see Appendix D for details on this approach).

Figure 7 shows conditional binned scatter plots of  $y_{i,t}^{\Delta}$  against  $x_{i,t}^{\Delta}$ , with slopes corresponding to  $\alpha_1$  (Appendix Table A8 provides the underlying regressions, as well as some robustness specifications). Panels (a) and (b) show plots for general friendliness and relative friending, respectively. Our interpretation of  $\alpha_1$  here is that it provides an answer to the following question: “within a year of being assigned to a new place, to what extent does a moving native’s friending behavior adjust to that of observably comparable destination non-movers?”<sup>17</sup> An  $\alpha_1$  close to 1 suggests native movers’ behavior completely

<sup>17</sup>There are two reasons why this interpretation is intentionally more narrow than that of the mover design in Section 3.2 (where we interpreted  $\alpha_1$  as the share of across-region variation in integration that is explained by place-based effects). First, whereas regional differences in the observables for which we allow flexibility (gender, age, and arrival cohort) were nearly non-existent for Syrian migrants, regional differences in native demographics do have the potential to shape overall variation in our measures. For example, since older people are less likely to befriend Syrian migrants, regions with older populations on average may have lower levels of integration. Since we match movers to stayers with similar observables, our estimates will not capture observable-driven variation in friending patterns across space. (However in Section 4.3, we will show that relative to other factors, the quantitative importance of these county-level differences in natives’ gender and age is relatively small.) Second, our estimates will still not capture any “long-term effects of place” that do not adjust quickly after a move. These likely play a larger role in our current context. For example, Bursztyn et al. (2021) suggest that natives’ attitudes toward migrants are shaped by their level exposure to foreigners in prior settings, a long-term place effect we will not capture (and something that we will explicitly study in later sections).

**Figure 7:  $\Delta$  Native Mover Behaviors vs. Matched Non-Movers**

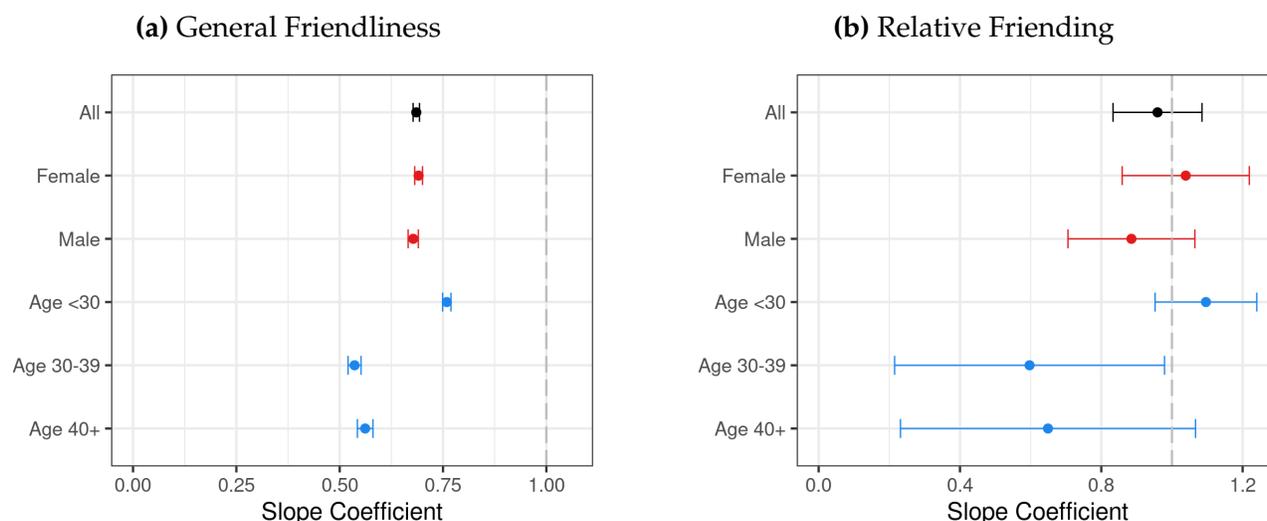


**Note:** Figures show binned scatter plots describing the change in the friending behavior of German natives before and after a move within Germany. The population is German native users who moved between non-neighboring counties and were in the first and second county for 4+ consecutive quarters each. In both panels the y-axis displays  $y_{i,t}^{\Delta}$ , an individual level change in movers' behavior the year before vs. after the move, and the x-axis displays  $\hat{x}_{i,t}^{\Delta}$ , the difference in average outcomes for comparable non-movers at the same time. In panel (a), the outcome is the change in the number of local German native friends made (*yearly general friendliness*) between the years. In panel (b), the outcome is the change in the ratio of the number of local Syrian migrant vs. local native friends, divided by the ratio of the number of local Syrian migrants vs. natives in the Facebook data (*yearly relative friending*) between the years. Panel (b) excludes users who make no local native friends in either the year before or after the move. In both panels we match each mover to a set of non-movers who match on gender and age buckets (18-29, 30-44, 45+). We include observations for which there are at least 1,000 non-movers in both the origin and destination match group. Both panels include quarter of move fixed effects. We correct for sampling error in the x-axis measures by randomly splitting the individual-level non-mover data into two halves and instrumenting for one set of averages with the other. See Appendix D for more information this procedure. Standard errors are shown in parentheses.

adjusts, whereas an  $\alpha_1$  close to 0 suggests it does not adjust at all. In both panels of Figure 7, the plots appear linear and symmetric around zero, providing evidence of additive place-based effects. In panel (a), the resulting slope estimate suggests that, within a year of moving to a new place, a native will adjust their overall friending 69% of the way to the level of comparable destination natives. In panel (b), our estimates suggest that movers' relative friending will adjust nearly fully to that of their destination, though the estimates are somewhat less precise, since few natives make any local Syrian migrant friends (see Table 1). Both panels, therefore, provide evidence that institutional factors and local policies play an important role in shaping various components of natives' friending behaviors. The fact that relative friending adjusts almost fully suggests that time-invariant individual-level characteristics such as attitudes towards migrants appear to play only a small role in explaining this outcome on average.

A number of works studying place-based effects in the United States find that new places exert stronger effects on younger, rather than older, movers (Kling, Liebman and Katz, 2007; Chetty, Hendren and Katz, 2016; Chetty and Hendren, 2018a; Chyn, 2018). Motivated by this, we next test whether place-based effects shape general friendliness and relative friending differentially by age and gender. We do

**Figure 8:**  $\Delta$  Native Mover Behaviors vs. Matched Non-Movers - Slope By Demographics



**Note:** Figures show slopes corresponding to versions of Figure 7. Panels (a) and (b) correspond to panels (a) and (b) of Figure 7, respectively. The coefficients in black are the slopes using the full sample of German native movers; the coefficients in red use samples of only one gender; and the coefficients in blue use samples of only one age group. Bars display 95% confidence intervals. The sample sizes used to generate each coefficient are in panel (a) (from top to bottom): 1,771,041, 920,076, 850,965, 813,375, 427,631, and 530,035. In panel (b) they are: 1,096,874, 578,607, 518,267, 572,880, 245,830, and 278,164.

so by running versions of regression 4 over samples of users with different ages and genders. Figure 8 presents the corresponding estimates of  $\alpha_1$  (corresponding to the slopes in Figure 7). These estimates suggest that native movers under 30 adjust their general friendliness and relative friending around 76% and 110% of the way to the comparable destination users, respectively, within a year. By contrast, native movers 40 or older adjust by only 56% and 65%. Put differently, younger natives' overall friending, and friending to Syrian migrants in particular, is more strongly shaped by the characteristics of place.

One possible interpretation of these findings is that places have cumulative effects on individuals, which become more ingrained over time (a force that would lead our large estimates of place-based effects to *understate* the full role of places on individuals' behaviors). In Section 5, we explore the lasting effects of contact between migrants and natives in one setting on the natives' friending behavior in other settings, which may be part of this unmeasured effect. Throughout the rest of Section 4 we attempt to better understand the institutional and structural factors which, even excluding unmeasured long-term effects, explain the majority of the regional variation in natives' friending with Syrian migrants.

### 4.3 Correlates of Regional Measures of Social Integration

We next explore the relationship of county-level social integration outcomes of migrants—and their determinants of general friendliness and relative friending components—with a variety of regional characteristics. This section explores salient correlations rather than causal relationships. In Section 4.4 we will study the causal effect of one prominent policy intervention, the provision of integration courses.

Figure 9 presents univariate county-level correlations between migrants' social integration and measures of demographic composition, urbanity, economic conditions, attitudes towards migrants, migrant concentration, and integration courses. Red diamonds denote raw correlations, while blue triangles de-

**Figure 9: County-Level Univariate Correlations with Friending Integration**



**Note:** Figure presents correlations between our county-level measure of social integration and various other regional measures. Social integration is based on Syrian migrants' number of native local friends (mapped in Figure 3). Correlations are weighted by the size of the Syrian migrant sample in each county. Red diamonds depict raw, univariate correlations and blue triangles depict correlations after controlling for state fixed effects. The regional measures are average age, log 2018 population density; log average income, log employment rate; the vote share for the Alternative for Germany, demeaned by state, pro-immigration groups per population; log of the shares of the population that were Syrian in 2010 and 2019, and log of the numbers of integration courses completed from 2015-2019 per Syrian. For more information on each measure see Appendix J.

note correlations with state fixed effects.<sup>18</sup> Appendix J describes each measure in detail. Table 7 presents multivariate analyses that explore how these county-level characteristics correlate with social integration, general friendliness, relative friending, and language integration as outcomes.<sup>19</sup> To help with the interpretation of magnitudes, we use the log-form for some of the dependent and explanatory variables, but correlations are very similar with raw magnitudes.

**Demographics & Urbanity.** While the top row of Figure 9 suggests that Syrians tend to be less socially integrated in places with an older population, this relationship becomes insignificant in the multivariate regressions in Table 7. By contrast, in both univariate and multivariate analyses, migrants are better integrated socially in less densely populated areas. The multivariate results in Table 7 show that this is driven by both relative friending and general friendliness being lower in urban areas. These trends are consistent with research that finds that rural areas have higher levels of social capital and lower levels of social isolation relative to more densely populated urban areas (Putnam, 1995b; Rupasingha, Goetz and Freshwater, 2006; The Social Capital Project, 2018; Henning-Smith, Moscovice and Kozhimannil, 2019).

<sup>18</sup> Appendix Figure A12 includes univariate relationships with a number of additional county-level measures. For example, we document that there are no strong relationships between integration and local crime rates.

<sup>19</sup> In Figure 9 and Table 7, we weight all relationships by the county's Syrian migrant sample size, except when we look at general friendliness as outcome variable, in which case we weight by the county's German native sample size.

**Table 7: County-level Multivariate Relationships with Friending Integration**

	Friending Integration		General Friendliness		Relative Friending		Language	
Average Age	-0.032 (0.02)	-0.034* (0.02)	-0.034*** (0.01)	-0.034*** (0.01)	0.015 (0.02)	0.003 (0.02)	-0.005 (0.01)	-0.011* (0.01)
Log Pop. Density 2018	-0.098* (0.05)	-0.136*** (0.03)	-0.029 (0.03)	-0.071*** (0.02)	-0.066** (0.03)	-0.058** (0.03)	-0.034** (0.01)	-0.016 (0.01)
Log Average Income (in EUR)	-0.198 (0.26)	0.140 (0.18)	0.168 (0.14)	0.097 (0.10)	-0.296 (0.20)	0.054 (0.15)	0.070 (0.08)	0.035 (0.06)
Log % Unemployed	-0.056 (0.09)	-0.291*** (0.09)	-0.108*** (0.04)	-0.065* (0.04)	0.015 (0.07)	-0.209*** (0.07)	-0.129*** (0.03)	-0.032 (0.03)
Vote Share AfD European Elections 2014	-8.953*** (2.64)	-6.167*** (1.92)	-1.939** (0.85)	-1.039 (0.69)	-6.917*** (2.29)	-5.091*** (1.55)	-0.569 (0.68)	-1.289** (0.65)
Number of ProAsyl Groups per Pop	4.778* (2.55)	4.286*** (1.40)	-1.381 (1.22)	-0.341 (0.76)	4.876*** (1.69)	3.167** (1.29)	3.557*** (0.85)	1.558** (0.62)
Log Fraction of Syrians 2010	0.105*** (0.02)	0.150*** (0.02)	0.025*** (0.01)	0.030*** (0.01)	0.067*** (0.02)	0.114*** (0.02)	0.019** (0.01)	0.043*** (0.01)
Log Fraction of Syrians 2019	-0.239*** (0.08)	-0.135*** (0.05)	-0.048* (0.03)	-0.065*** (0.02)	-0.117** (0.05)	-0.060 (0.04)	-0.044* (0.02)	-0.103*** (0.02)
Log Int. Courses Completed 2015-19 per Syrian	0.235*** (0.05)	0.200*** (0.04)	0.005 (0.02)	-0.013 (0.02)	0.222*** (0.04)	0.202*** (0.03)	0.076*** (0.02)	0.052*** (0.01)
State FE		x		x		x		x
R-squared	0.487	0.709	0.261	0.665	0.330	0.633	0.519	0.668
N	390	390	390	390	390	390	390	390

**Note:** Table presents results from regressions of various county-level measures on the logs of friending integration (columns 1 and 2), general friendliness (columns 3 and 4), relative friending (columns 5 and 6), and language (columns 7 and 8). The county measures are those discussed in Figure 9. Regressions are weighted by the number of Syrian migrants in the Facebook data in columns 1-2 and 5-8. Regressions in columns 3 and 4 are weighted by the number of natives in the Facebook data. Standard errors are shown in parentheses. Significance levels: \*( $p < 0.10$ ), \*\*( $p < 0.05$ ), \*\*\*( $p < 0.01$ )

**Economic Conditions.** While limited by measurement challenges, a handful of prior works have explored ties between social and economic integration. For example, Laurentsyeve and Venturini (2017) discuss the possibility that employment contributes to migrants’ social integration. Conversely, Cheung and Phillimore (2014) use survey data to highlight the importance of local language proficiency for employment. Figure 9 and Table 7 show that while there is no strong relationship between the average income level in a county and migrants’ social integration, integration does appear to be higher in areas with lower rates of unemployment, in particular when comparing counties within states. For instance, controlling for state fixed effects, we find that a 1% higher unemployment rate is associated with 0.29% lower level of social integration, an effect that is largely driven by lower relative friending rather than general friendliness.

**Attitudes Towards Migrants.** We explore correlations with two measures of local attitudes towards migrants: (i) the vote share for Alternative for Germany or AfD, a political party in favor of limiting immigration, in the 2014 EU Election (predating the main influx of Syrian migrants);<sup>20</sup> and (ii) the number of pro-immigration groups per population. Support for the AfD has a strong negative relationship with social integration and relative friending: a one percentage point increase in AfD vote share relative to state-level averages is associated with a decrease in social integration of nearly 9% and in relative friending of 6.9%. Pro-immigration groups are independent organizations that offer a wide range of services to migrants, including help filing for asylum status, medical attention, and the provision of child care. We

<sup>20</sup>Because political parties in Germany are differentially important across states, and often run with varying policy positions by state, in Figure 9 and Table 7 we always demean AfD vote share by state.

study groups registered with *ProAsyl*, a widely-known pro-immigration organization in the country. In both univariate and multivariate analyses, we find places with relatively more pro-immigration groups per population tend to have higher levels of social integration. Table 7 shows this is driven entirely by variation in relative friending rather than general friendliness.

**Concentration of Migrants.** A number of works study the relationship between local co-ethnic populations and the economic integration of migrants. For example, Edin, Fredriksson and Åslund (2003) and Damm (2009) find a positive effect on earnings for refugees living in areas with many co-ethnic individuals (so-called “ethnic enclaves”), while Cutler, Glaeser and Vigdor (2008) find negative effects if the community has low levels of average education. Our results suggest that migrants do make fewer native friends when there are more Syrian migrants today, driven primarily by relative friending. However, we see that social integration generally increases with the share of the population that was Syrian in 2010, again largely through effects on relative friending. We find similar results when looking at the extent of German language usage. These patterns are consistent with earlier migrants providing important information or connections with natives to new arrivals to aid their social integration, but large communities of migrants arriving at the same time leading to fewer migrant-native connections. It is also possible that local natives more exposed to Syrian migrants in 2010 became more friendly toward Syrians in the future, a notion we explore at the individual level in Section 5.

**Integration Courses.** The German government and other independent organizations have invested heavily in efforts to integrate recent migrants (see, e.g., Bundesregierung, 2021). Integration courses, which are intended to teach migrants the German language and other relevant information, are “at the core of the government’s integration measures” (BAMF, 2015). Indeed, they have been taken by 1.13 million individuals from 2015-2019 (BAMF, 2021). In both the univariate and multivariate analyses, we find strong positive relationships between a county’s social integration outcomes and the number of integration courses completed per Syrian between 2015 and 2019. The effect appears to be entirely driven by a relationship between integration course completion and relative friending. While these results are not causal, they are consistent with integration courses supporting the integration efforts of Syrian migrants. To isolate a possible causal effect of integration courses, we next use an instrumental variables approach that leverages exogenous variation in course availability across regions.

#### 4.4 Causal Effect of Integration Policy: Integration Courses

In this section we study the causal effects of integration courses, a local intervention of central importance to the German government, on integration outcomes. We do so with an instrumental variables (IV) approach, exploiting the impact of quasi-random variation in the presence of qualified teachers on the availability—and in turn completion—of integration courses across counties.

To teach an integration course in Germany, the government requires an individual to have either a college degree in teaching German as a second language or, with a degree in a different pedagogical field, significant experience teaching German as a second language (BAMF, 2018). Because of these substantial requirements, many integration courses are taught by otherwise unemployed teachers. Indeed, in a widely-televised 2016 interview, the federal government’s coordinator of refugee policy (*Flüchtlingskoor-*

**Table 8: Integration Courses and Teacher Unemployment Rates**

	Log Integration Courses per Syrian 2015-19			
Log Unemp. General Schools Teachers 2014 per Syrian	0.088			
	(0.05)			
Log Unemp. Vocat. School Teachers 2014 per Syrian		0.084		
		(0.05)		
Log Unemp. Driving and Sports Teachers 2014 per Syrian			0.052	
			(0.06)	
Log Unemp. Other School Teachers 2014 per Syrian				0.229***
				(0.05)
Control Covariates	x	x	x	x
Control Log General Unemployment Rate	x	x	x	x
F-statistic	2.37	3.67	0.94	20.97
N	390	367	388	390
R-Squared	0.349	0.354	0.347	0.379

**Note:** Table presents results from county-level regressions between various 2014 teacher unemployment rates and integration course completion. The outcome is the log of the number of integration courses completed per Syrian between 2015 and 2019. In all regressions we control linearly for the log of the share of the population unemployed, the number of unemployed people per Syrian (as of 2014) as well as average age, log population density, log average income and log number of open training positions per applicant. Regressions are weighted by the total number of Syrians in each county as of 2019. Standard errors are shown in parentheses. Significance levels: \*(p<0.10), \*\*(p<0.05), \*\*\*(p<0.01)

*dinator*) specifically called on unemployed teachers to meet the rapid demand for integration course instructors (Tagesschau, 2016). The unemployment rate of qualified teachers in a given county at the start of the major influx of migrants, therefore, likely impacted the availability of local integration courses.

We test this story using county-level data on 2014 teacher unemployment from the Federal Employment Agency. Importantly, these data allow us to distinguish between four types of teachers: general, vocational, driving or sports, and other. “Other” teachers are primarily adult educators, often focused on non-native populations, and are much more likely than the other groups of teachers to meet the necessary requirements to teach integration courses. Therefore, if local teacher unemployment affects integration course availability (and therefore completion), it should do so primarily through this particular set of teachers.

Table 8 presents results that are highly consistent with the availability of teachers driving the availability and eventual completion of integration courses. Columns 1-3 show that, after controlling for general unemployment and other county-level covariates, there are no significant relationships between integration course completion and unemployed general, vocational, and driving or sports teachers per Syrian. By contrast, column 4 shows a positive and highly significant relationship for “other” teachers: a 10% increase in their unemployment per Syrian as of 2014 corresponds to a 2.3% increase in integration course completion per Syrian. With an F-statistic of just under 21, this “first stage” relationship for our IV strategy is remarkably strong given the limited number of counties.

While this evidence supports the notion that teacher unemployment meaningfully affects the completion of integration courses, for the measure to serve as a valid instrument it must also satisfy the exclusion restriction. Namely, teacher unemployment must not affect social integration other than through its effect on integration courses. To mitigate concerns that our first stage is driven by general economic conditions or other confounders that might affect integration, we always include a rich set of county-level controls in our regressions: general unemployment, the number of unemployed people per Syrian, average age, population density, average incomes, and open training positions.<sup>21</sup> Moreover, our use of 2014 teacher unemployment, before the large influx of migrants, also allows us to rule out stories in which reverse causality violates the exclusion restriction.

**Table 9:** IV Estimates - Measures of Integration and Integration Courses

	Integration	General Friendliness	Relative Friending	Language	Employ. / Training
Log Integration Courses per Syrian	1.698*** (0.33)	0.204 (0.21)	1.389*** (0.25)	0.193*** (0.07)	0.891*** (0.15)
Control Covariates	x	x	x	x	x
Control Log General Unemployment Rate	x	x	x	x	x
N	390	390	390	390	384

**Note:** Table presents results from county-level IV regressions of various measures related to integration on the completion of integration courses. We instrument for integration courses with the 2014 total number of unemployed “other” per Syrian. In both stages of our estimation we include the same controls as in Table 8. The outcomes are overall friending integration (column 1), general friendliness (column 2), relative friending (columns 3), the share of Syrian migrant Facebook users producing content in German (column 4), and the share of all Syrians employed or in training programs (column 5). All independent and dependent variables are specified in logs. Regressions are weighted by the total number of Syrians as of 2019 except when the outcome variable is general friendliness in which case we weight by the number of German natives in the Facebook data. Standard errors are shown in parentheses. Significance levels: \*( $p < 0.10$ ), \*\*( $p < 0.05$ ), \*\*\*( $p < 0.01$ )

Table 9 presents results for our IV regressions. Column 1 suggests that a 10% increase in completed integration courses increases the social integration of Syrians by nearly 17%. Intuitively, this means that moving a migrant from a 25th to 75th percentile county in terms of the relevant teacher unemployment would result in them making about 1.7 more native friends.

The IV estimates are substantially larger than the OLS estimates in columns 1 and 2 of Table 7. We provide two potential reasons. First, our instrumental variable likely corrects for substantial downward bias due to omitted variables in the OLS estimates. Such downward bias could occur, for example, if integration courses were specifically targeted toward or advertised in areas with the lowest integration levels. We find supporting evidence that this is indeed the case: courses tend to be concentrated in urban places and places with a greater total immigrant share, both factors that are negatively correlated with integration as discussed in Section 4.3. Second, the IV identifies a Local Average Treatment Effect (LATE), rather than an Average Treatment Effect (ATE). Concretely, our instrument makes use of higher teacher unemployment relaxing supply constraints when courses were most scarce. If the marginal course participant aided by these relaxed constraints had higher than average returns from

<sup>21</sup>Note that our set of controls differs somewhat from the variables used in Table 7: in the present part of our analysis, we refrain from controlling for covariates that are potentially endogenous to our outcome of interest, such as the fraction of Syrians in 2019 or the number of pro-immigration groups.

these courses, we would expect the LATE to exceed the ATE. For example, there is suggestive evidence that women were less likely to participate in integration courses when those courses are in short supply (Tissot, 2021); if integration courses had stronger impacts on the integration outcomes of women, the LATE from our IV strategy would thus exceed the ATE. Importantly, the LATE from our IV strategy may be of particular interest to policy makers, whose primary tool to increase the completion of integration courses is to make them more easily accessible.

Columns 2 and 3 of Table 9 present IV estimates of the effect of integration courses on general friendliness and relative friending, respectively. Because friending behavior among natives should not be impacted by integration courses, integration courses should affect overall integration only through relative friending. Highly consistent with this story, we find significant effects for relative friending in column 3, but not for general friendliness in column 2. Our IV estimates suggest that a 10% increase in integration courses completed increases Syrian migrants' relative friending integration by close to 14%.

Columns 4 and 5 measure the causal effect of integration courses on language and economic integration. In particular, our outcomes are the share of Syrian migrant Facebook users producing content in German (in column 4) and the share of all Syrians employed or in training programs (in column 5). For both, we find highly significant and positive effects of integration courses. The IV estimates suggest that a 10% increase in integration courses completed increases language integration by just under 2% and the rate of Syrians in employment or training by about 9%.

## 5 Exposure and Native Behaviors Toward Migrants

In the prior section, we showed that average differences in the propensity of local natives to befriend migrants across locations are largely driven by characteristics of the locations rather than the characteristics of the natives living in those locations. However, individual characteristics still drive a portion of regional variation and, importantly, contribute to the substantial differences in friending behaviors across natives *within* regions. We next explore one possible source of these differences in the propensities of different natives to befriend Syrians: variation natives' prior exposures to Syrians. Our exploration is motivated by a number of studies which—building on the seminal work of Allport, Clark and Pettigrew (1954)—find evidence of intergroup contact reducing negative attitudes or discrimination toward outsiders (Paluck, Green and Green, 2019; Boisjoly et al., 2006; Rao, 2019; Carrell, Hoekstra and West, 2015; Bursztyn et al., 2021). Indeed, in Section 4.3 we showed that a county's relative friending today increases with when there was a larger Syrian population in 2010, consistent with exposure to Syrian migrants at one time leading a local native population to be more friendly towards Syrians in the future.

In this section we use individual-level data to explore whether natives who are exposed to migrants in one setting are more likely to subsequently befriend similar migrants in another setting. To provide quasi-random variation in natives' exposures to migrants, we exploit Germany's relatively strict age cutoffs for school entry, which sort individuals born at similar times into different cohorts at age six. We then use within-school variation in whether a Syrian migrant is assigned to a particular cohort to generate quasi-random variation in natives' exposure to migrants. We find that individuals who are quasi-randomly exposed to a Syrian migrant in their high school are more likely to subsequently make Syrian friends outside of high school.

**Sample Construction.** We generate our primary sample for this analysis by subsetting our German native and Syrian migrant samples into those with a birth date between 1995 and 1999. These individuals were roughly 15 to 19 years old in 2014, at the start of the major influx of Syrian migrants. We observe 26,000 such Syrian migrant users and 2.2 million such German native users.

We match individuals to their high schools using self-reports and friend-based imputations (see Appendix K for details). We assign 63.2% of individuals within this age group to a high school. We then sort individuals into cohorts within a school using the German system of age cutoffs for school entry. In Germany, children are eligible to enroll in school for the first time if they have turned six by a certain date (the exact date varies by state). Though students are allowed to enroll earlier or to defer enrollment at the advice of a doctor, at least 86% of students comply with the entry time suggested by the cutoff date (Schwandt and Wuppermann, 2015). We use data from Schwandt and Wuppermann (2015) to define the school entry cutoff date for each state and year.

**Research Design.** Since students are disproportionately exposed to individuals in their own grade (relative to the years above and below them), variation in cohort composition can generate exogenous differences in the social networks formed by the members of each grade. Similar sources of variation in exposure and network composition have been utilized in other studies (e.g. Chetty et al., 2022b; Billings, Chyn and Haggag, 2021; Sacerdote, 2011). Because Syrian students are relatively uncommon in the German school system overall, we focus on how German natives are affected by having at least one Syrian migrant in their cohort. In particular, we focus on adjacent cohorts within a school where one cohort contains at least one Syrian migrant and the other does not. For instance, if the only Syrian who attends Marie Curie Gymnasium is in the class of 2016, we will study natives who fall on either side of the cutoff that divides the 2015 and 2016 cohorts.<sup>22</sup> We estimate equations of the form:

$$Y_i = \alpha_1 \text{SyrianInCohort}_s + \zeta_{t,L} + \gamma_s + \epsilon_{i,t}. \quad (7)$$

Here,  $Y_i$  is the number of friends of a given type that user  $i$  has today,  $\text{SyrianInCohort}$  is an indicator variable set to one if a user has at least one Syrian in their assigned school cohort,  $\zeta_{t,L}$  is a birth year-by-county fixed effect, and  $\gamma_s$  is a school fixed effect. Under the assumption that it is random whether a student’s birth date places them into a cohort with a Syrian or into an adjacent cohort without one,  $\alpha_1$  identifies the effect of the additional exposure caused by placement into a cohort containing a Syrian. In some specifications, we include an interaction term,  $\text{SyrianInCohort}_c \times \text{CohortSize}_c$ , where  $\text{CohortSize}_c$  is the number of students in cohort  $c$ , normalized to have mean 0 and standard deviation 1. This interaction term will allow us to examine how the effects of exposure differ according to the size of the cohort.

<sup>22</sup>Conceptually, we could also study Germans around the assignment cutoff for the 2016 and 2017 cohorts. However, since many Syrians enter the German school system with low levels of German proficiency, some are assigned to a cohort younger than would be suggested by the assignment rule (though we find that most Syrians have the plurality of their friends in the cohort they would be assigned into under the allocation rules used for Germans). As a result, if we use this second design (where the Syrian is supposed to be in the older cohort), we will swap the treatment and control groups of Germans when the Syrian is assigned to a *younger* cohort. This is not a problem when we consider Germans who are on the cusp of being too old for a cohort with a Syrian, as the treatment and control group will not be swapped if the Syrian is placed in a younger cohort relative to the assignment rule. We also exclude pairs of years where there is a cohort without Syrians that is flanked by cohorts with Syrians. Since Syrians from the older cohort is sometimes mis-assigned, these configurations can lead us to inadvertently compare two cohorts that both contain Syrians, which would attenuate our results.

**Table 10:** Impacts of High School Exposure on Friendship

	Syrian Friends		Syrian Friends (Excluding Classmates)		Syrian Friends (Excluding Syrian Classmates and their Friends)	
Syrian in Cohort	0.020*** (0.002)	0.020*** (0.002)	0.005*** (0.002)	0.005*** (0.002)	0.005*** (0.001)	0.005*** (0.001)
Syrian in Cohort x Standardized Cohort Size		-0.007*** (0.001)		-0.003*** (0.001)		-0.003*** (0.001)
School FE	X	X	X	X	X	X
Birth Year x County FE	X	X	X	X	X	X
N	115,625	115,625	115,625	115,625	115,625	115,625
Mean in Control Cohort	0.054	0.054	0.029	0.029	0.027	0.027

**Note:** Table presents results from regressions of the form outlined in Equation 7. The sample includes Germans who were assigned to one high school cohort where the younger cohort contains a Syrian and the older cohort does not. The treatment years include students who entered kindergarten between 2001 and 2004, while students in the paired control cohorts entered kindergarten between 2002 and 2005. In columns 1-2, we include all Syrian friends that a user makes; in columns 3-4, we only include Syrian friends who did not attend the user’s high school; and in column 5-6 we only include Syrian friends who did not attend the user’s high school and who did not have a prior friendship with a Syrian that attended the user’s high school. In all columns, we include only Syrian friends made in the first 21 years of a person’s life, in order to avoid mechanically calculating larger treatment effects for older users. All users in our sample have already turned 21. In all columns, we cluster standard errors at the school and cohort level. \*( $p < 0.10$ ), \*\*( $p < 0.05$ ), \*\*\*( $p < 0.01$ )

**Effects of Exposure.** In Table 10, we quantify the effects of being randomly assigned to a cohort including a Syrian migrant. The first column presents baseline results: students placed into a cohort containing a Syrian have 0.02 more Syrian friends by age 21, an increase of around 40% relative to the 0.054 Syrian friends that Germans in the adjacent cohort have on average. In the second column, we interact the treatment term with the z-score of cohort size. We find that treated students in a cohort one standard deviation larger than the mean make one-third fewer Syrian friends.

We next turn our attention to the mechanisms through which these friendships can be formed. Broadly speaking, there are three possible mechanisms. First, and most trivially, German natives can befriend the Syrian in their cohort. Second, the Syrian can play a direct role in mediating connections between native Germans and other Syrians by introducing previously disconnected individuals across groups. Third, the presence of the Syrian can play a role in shaping the preferences of native Germans for contact with other Syrians. This last mechanism could play a role in future network formation if stereotypes about individuals outside one’s own group inhibit friendship formation.

In columns 3 and 4, we repeat the regressions in columns 1 and 2, but now include only Syrian friends who did not attend the German’s high school in our outcome measure. This allows us to isolate friends made through the second and third mechanisms above. We find that Germans in the treated cohorts make 0.005 more friends of this type, about 17% more than the average number of such friends in the control group. As in column 2, these effects are larger for students whose cohorts are smaller. These friendships outside of one’s school comprise about one quarter of the overall effect of exposure.

In column 5, we again exclude Syrian friends who attended the German’s high school from our outcome. We now also exclude any friends of individuals who attended the German’s high school,

allowing us to isolate connections made due to third mechanism listed above. We find that the effect size is largely unchanged, indicating that the bulk of the new friendships in columns 3 and 4 are made in new social contexts and not simply connections made by the Syrians in one's school. This positive effect on Syrian friending, even when excluding the Syrian who attended their high school and their friends, suggests quasi-random exposure to Syrian migrants shifts German natives' *preferences* toward making migrant friends.

## 6 Conclusion

The challenge of harmoniously integrating immigrants into new communities has become central for policymakers around the world. In the coming decades, climate change could displace as many as one billion individuals, increasing the flow of international migrants and raising the importance of these challenges further (Kamal, 2017). However, due to the difficulty of measuring social networks using traditional data sources, understanding the drivers of international migrants' social integration has historically proven to be challenging. Are there environments where newly arriving migrants are relatively better integrated, and why? What can governments do to foster the social integration of migrants?

In this paper, we use de-identified data from Facebook to study the social integration of Syrian migrants in Germany along three key dimensions: (i) friendships between migrants and natives; (ii) migrants' German language usage; and (iii) migrants' participation in local social groups. For each dimension we construct German county-level measures of social integration. We provide evidence these measures pick up true differences in social integration rather than sampling variation or geographic differences in Facebook usage. Using a movers design, we show that these differences are largely due to causal place-based factors rather than unobserved characteristics of migrants.

We next ask "why do migrants integrate better in some places than in others?" We show that regional variation in social integration outcomes is shaped by both the rate at which local natives befriend other locals in general (*general friendliness*) and the relative rate at which they form friendships with Syrian migrants in particular (*relative friending*). We show that each force is more strongly determined by place-based effects than individual native characteristics, as natives' friending patterns substantially adjust when they move between locations. Put simply, local institutions and environments appear more important than individual preferences in determining whether a native makes migrant friends (although both play a role).

We then explore the characteristics of more integrated communities, highlighting a number of important patterns that might help policymakers. For example, our results suggest that while large numbers of migrants arriving at the same time may lead to fewer migrant-native connections, when migrants arrive in a place with many *earlier arriving* migrants they make more native connections. We suggest two possible channels through which earlier migrant arrivals might increase integration: (1) earlier arrivals may provide important information or connections to aid new arrivals' social integration; (2) earlier arrivals may increase local natives' openness toward migrants. While both channels likely play a role, we document that natives with quasi-random exposure to Syrian migrants in their high schools are more likely to befriend other Syrian migrants in other settings, thereby providing direct evidence for the latter.

Finally, we ask whether specific policy programs can improve integration outcomes. We find that

government-sponsored integration courses have a substantial positive *causal* effect on relative friending, closing the gap between the rates at which German natives befriend Syrian migrants vs. other natives. This finding highlights that integration outcomes are not immutable, but can be shaped by government policies. We hope that the increasing availability of data sources similar to the ones used in this paper—as well as other digital trace data discussed in Kuchler and Stroebel (2022)—will help researchers better understand the forces that shape social integration and help policymakers develop programs that effectively foster interconnected communities.

## References

- Abdelgadir, Aala, and Vasiliki Fouka.** 2020. “Political Secularism and Muslim Integration in the West: Assessing the Effects of the French Headscarf Ban.” *American Political Science Review*, 114(3): 707–723.
- Abramitzky, Ran, Leah Boustan, and Katherine Eriksson.** 2020. “Do immigrants assimilate more slowly today than in the past?” *American Economic Review: Insights*, 2(1): 125–41.
- Aksoy, Cevat Giray, Panu Poutvaara, and Felicitas Schikora.** 2020. “First Time Around: Local Conditions and Multi-dimensional Integration of Refugees.”
- Alesina, Alberto, Reza Baqir, and William Easterly.** 1999. “Public goods and ethnic divisions.” *The Quarterly journal of economics*, 114(4): 1243–1284.
- Allport, Gordon Willard, Kenneth Clark, and Thomas Pettigrew.** 1954. “The nature of prejudice.”
- Arendt, Jacob Nielsen, Iben Bolvig, Mette Foged, Linea Hasager, and Giovanni Peri.** 2020. “Language Training and Refugees’ Integration.” National Bureau of Economic Research.
- Åslund, Olof, and Dan-Olof Rooth.** 2007. “Do when and where matter? Initial labour market conditions and immigrant earnings.” *The Economic Journal*, 117(518): 422–448.
- Bailey, Michael, Abhinav Gupta, Sebastian Hillenbrand, Theresa Kuchler, Robert Richmond, and Johannes Stroebel.** 2021. “International trade and social connectedness.” *Journal of International Economics*, 129: 103418.
- Bailey, Michael, Drew M Johnston, Martin Koenen, Theresa Kuchler, Dominic Russel, and Johannes Stroebel.** 2020a. “Social Networks Shape Beliefs and Behavior: Evidence from Social Distancing During the COVID-19 Pandemic.” National Bureau of Economic Research.
- Bailey, Michael, Drew M Johnston, Theresa Kuchler, Dominic Russel, Johannes Stroebel, and Bogdan State.** 2020b. “The determinants of social connectedness in europe.” *International Conference on Social Informatics*, 1–14.
- Bailey, Michael, Drew M Johnston, Theresa Kuchler, Johannes Stroebel, and Arlene Wong.** 2019a. “Peer effects in product adoption.” National Bureau of Economic Research.
- Bailey, Michael, Eduardo Dávila, Theresa Kuchler, and Johannes Stroebel.** 2019b. “House price beliefs and mortgage leverage choice.” *The Review of Economic Studies*, 86(6): 2403–2452.

- Bailey, Michael, Patrick Farrell, Theresa Kuchler, and Johannes Stroebel.** 2020c. "Social connectedness in urban areas." *Journal of Urban Economics*, 118: 103264.
- Bailey, Michael, Rachel Cao, Theresa Kuchler, and Johannes Stroebel.** 2018a. "The economic effects of social networks: Evidence from the housing market." *Journal of Political Economy*, 126(6): 2224–2276.
- Bailey, Michael, Rachel Cao, Theresa Kuchler, Johannes Stroebel, and Arlene Wong.** 2018b. "Social connectedness: Measurements, determinants, and effects." *Journal of Economic Perspectives*, 32(3): 259–80.
- BAMF.** 2015. "Konzept für einen bundesweiten Integrationskurs." Available at [https://www.bamf.de/SharedDocs/Anlagen/DE/Integration/Integrationskurse/Kurstraeger/KonzepteLeitfaeden/konz-f-bundesw-integrationskurs.pdf?\\_\\_blob=publicationFile&v=7](https://www.bamf.de/SharedDocs/Anlagen/DE/Integration/Integrationskurse/Kurstraeger/KonzepteLeitfaeden/konz-f-bundesw-integrationskurs.pdf?__blob=publicationFile&v=7), (accessed: 07.30.2021).
- BAMF.** 2018. "Matrix: Zulassungskriterien für Lehrkräfte zum Integrationskurs bis 30.09.2020." Available at <https://www.bamf.de/SharedDocs/Anlagen/DE/Integration/Integrationskurse/Lehrkraefte/matrix-zulassung-lehrkraefte-integrationskurse-bis-30092020.pdf>, (accessed: 07.27.2021).
- BAMF.** 2021. "Geschäftsstatistik zum Integrationskurs 2020." Available at <https://www.bamf.de/DE/Themen/Statistik/Integrationskurszahlen/integrationskurszahlen-node.html;jsessionid=75DC03EF1E61BC40EABEA96CA2FD8930.internet561>, (accessed: 07.30.2021).
- Bandiera, Oriana, Myra Mohnen, Imran Rasul, and Martina Viarengo.** 2019. "Nation-building through compulsory schooling during the age of mass migration." *The Economic Journal*, 129(617): 62–109.
- Battisti, Michele, Yvonne Giesing, and Nadzeya Laurentsyeva.** 2019. "Can job search assistance improve the labour market integration of refugees? Evidence from a field experiment." *Labour Economics*, 61: 101745.
- Beaman, Lori A.** 2012. "Social networks and the dynamics of labour market outcomes: Evidence from refugees resettled in the US." *The Review of Economic Studies*, 79(1): 128–161.
- Becker, Sascha O, and Andreas Ferrara.** 2019. "Consequences of forced migration: A survey of recent findings." *Labour Economics*, 59: 1–16.
- Berry, John W.** 1997. "Immigration, acculturation, and adaptation." *Applied psychology*, 46(1): 5–34.
- Billings, Stephen B., Eric Chyn, and Kareem Haggag.** 2021. "The Long-Run Effects of School Racial Diversity on Political Identity." *American Economic Review: Insights*, 3(3): 267–84.
- Boisjoly, Johanne, Greg J Duncan, Michael Kremer, Dan M Levy, and Jacque Eccles.** 2006. "Empathy or antipathy? The impact of diversity." *American Economic Review*, 96(5): 1890–1905.

- Braun, Sebastian, and Nadja Dwenger.** 2017. "The local environment shapes refugee integration: Evidence from post-war Germany." *Hohenheim Discussion Papers in Business, Economics and Social Sciences*.
- Brell, Courtney, Christian Dustmann, and Ian Preston.** 2020. "The Labor Market Integration of Refugee Migrants in High-Income Countries." *Journal of Economic Perspectives*, 34(1): 94–121.
- Bundesregierung.** 2021. "Migration and integration - what is the German government doing?" Available at <https://www.bundesregierung.de/breg-en/news/migration-und-integration-1659206>, (accessed: 06.18.2021).
- Bursztyn, Leonardo, Thomas Chaney, Tarek Alexander Hassan, and Aakaash Rao.** 2021. "The Immigrant Next Door: Exposure, Prejudice, and Altruism." National Bureau of Economic Research.
- Card, David, Jörg Heining, and Patrick Kline.** 2013. "Workplace heterogeneity and the rise of West German wage inequality." *The Quarterly journal of economics*, 128(3): 967–1015.
- Carrell, Scott E, Mark Hoekstra, and James E West.** 2015. "The impact of intergroup contact on racial attitudes and revealed preferences." National Bureau of Economic Research.
- Chetty, Raj, and Nathaniel Hendren.** 2018a. "The impacts of neighborhoods on intergenerational mobility I: Childhood exposure effects." *The Quarterly Journal of Economics*, 133(3): 1107–1162.
- Chetty, Raj, and Nathaniel Hendren.** 2018b. "The impacts of neighborhoods on intergenerational mobility II: County-level estimates." *The Quarterly Journal of Economics*, 133(3): 1163–1228.
- Chetty, Raj, Nathaniel Hendren, and Lawrence F Katz.** 2016. "The effects of exposure to better neighborhoods on children: New evidence from the Moving to Opportunity experiment." *American Economic Review*, 106(4): 855–902.
- Chetty, Raj, Nathaniel Hendren, Matthew O. Jackson, Theresa Kuchler, Johannes Stroebe, Robert Fluegge, Federico Gonzalez, Matthew Jacob, Drew Johnston, Martin Koenen, Eduardo Laguna-Muggenburg, Florian Mudekereza, Tom Rutter, Nicolaj Thor, Wilbur Townsend, Ruby Zhang, Mike Bailey, Pablo Barbera, Monica Bhole, and The Facebook Core Data Science Team.** 2022a. "Social Capital in the United States I: Measurement and Associations with Economic Mobility."
- Chetty, Raj, Nathaniel Hendren, Matthew O. Jackson, Theresa Kuchler, Johannes Stroebe, Robert Fluegge, Federico Gonzalez, Matthew Jacob, Drew Johnston, Martin Koenen, Eduardo Laguna-Muggenburg, Florian Mudekereza, Tom Rutter, Nicolaj Thor, Wilbur Townsend, Ruby Zhang, Mike Bailey, Pablo Barbera, Monica Bhole, and The Facebook Core Data Science Team.** 2022b. "Social Capital in the United States II: Exposure, Friendings Bias, and the Determinants of Economic Connectedness."
- Cheung, Sin Yi, and Jenny Phillimore.** 2014. "Refugees, social capital, and labour market integration in the UK." *Sociology*, 48(3): 518–536.

- Chi, Guanghua, Joshua E Blumenstock, Lada Adamic, et al.** 2019. "Who Ties the World Together? Evidence from a Large Online Social Network." 451–465, Springer.
- Chyn, Eric.** 2018. "Moved to opportunity: The long-run effects of public housing demolition on children." *American Economic Review*, 108(10): 3028–56.
- Chyn, Eric, and Lawrence F Katz.** 2021. "Neighborhoods Matter: Assessing the Evidence for Place Effects." National Bureau of Economic Research.
- Coleman, James S.** 1988. "Social capital in the creation of human capital." *American journal of sociology*, 94: S95–S120.
- Cutler, David M, Edward L Glaeser, and Jacob L Vigdor.** 2008. "When are ghettos bad? Lessons from immigrant segregation in the United States." *Journal of Urban Economics*, 63(3): 759–774.
- Damm, Anna Piil.** 2009. "Ethnic enclaves and immigrant labor market outcomes: Quasi-experimental evidence." *Journal of Labor Economics*, 27(2): 281–314.
- Damm, Anna Piil.** 2014. "Neighborhood quality and labor market outcomes: Evidence from quasi-random neighborhood assignment of immigrants." *Journal of Urban Economics*, 79: 139–166.
- Dekker, Rianne, Godfried Engbersen, Jeanine Klaver, and Hanna Vonk.** 2018. "Smart refugees: How Syrian asylum migrants use social media information in migration decision-making." *Social Media+ Society*, 4(1): 2056305118764439.
- Deutscher Bundestag.** 2020. "Verteilungsschlüssel bei Bund-Länder-Finanzierungen."
- Edin, Per-Anders, Peter Fredriksson, and Olof Åslund.** 2003. "Ethnic enclaves and the economic success of immigrants—Evidence from a natural experiment." *The quarterly journal of economics*, 118(1): 329–357.
- European Commission.** 2020. "Communication from the Commission to the European Parliament, the Council, the European Economic and Social Committee and the Committee of the Regions: Action plan on Integration and Inclusion 2021-2027."
- Facebook.** 2021. "Facebook Form 10-Q, Q1 2021."
- Finkelstein, Amy, Matthew Gentzkow, and Heidi L Williams.** 2019. "Place-based drivers of mortality: Evidence from migration." National Bureau of Economic Research.
- Finkelstein, Amy, Matthew Gentzkow, and Heidi Williams.** 2016. "Sources of geographic variation in health care: Evidence from patient migration." *The Quarterly Journal of Economics*, 131(4): 1681–1726.
- Fouka, Vasiliki.** 2020. "Backlash: The unintended effects of language prohibition in US schools after World War I." *The Review of Economic Studies*, 87(1): 204–239.
- Henning-Smith, Carrie, Ira Moscovice, and Katy Kozhimannil.** 2019. "Differences in social isolation and its relationship to health by rurality." *The Journal of Rural Health*, 35(4): 540–549.

- Herdağdelen, Amaç, Bogdan State, Lada Adamic, and Winter Mason.** 2016. “The social ties of immigrant communities in the United States.” *Proceedings of the 8th ACM Conference on Web Science*, 78–84.
- Hilbig, Hanno, and Sascha Riaz.** 2020. “Freedom of Movement Restrictions Inhibit the Social Integration of Refugees.”
- Jackson, Matthew O.** 2020. “A typology of social capital and associated network measures.” *Social Choice and Welfare*, 54(2): 311–336.
- Jaschke, Philipp, Sulin Sardoschau, and Marco Tabellini.** 2021. “Scared Straight? Threat and Assimilation of Refugees in Germany.”
- Jones, Jason J., Jaime E. Settle, Robert M. Bond, Christopher J. Fariss, Cameron Marlow, and James H. Fowler.** 2013. “Inferring Tie Strength from Online Directed Behavior.” *PLOS ONE*, 8(1): 1–6.
- Kamal, Baher.** 2017. “Climate Migrants Might Reach One Billion by 2050.” *United Nations OCHA*.
- Kling, Jeffrey R, Jeffrey B Liebman, and Lawrence F Katz.** 2007. “Experimental analysis of neighborhood effects.” *Econometrica*, 75(1): 83–119.
- Kuchler, Theresa, and Johannes Stroebel.** 2022. “Social interactions, resilience, and access to economic opportunity: A research agenda for the field of computational social science.”
- Kuchler, Theresa, Dominic Russel, and Johannes Stroebel.** 2021. “JUE Insight: The geographic spread of COVID-19 correlates with the structure of social networks as measured by Facebook.” *Journal of Urban Economics*, 103314.
- Kuchler, Theresa, Yan Li, Lin Peng, Johannes Stroebel, and Dexin Zhou.** 2020. “Social proximity to capital: Implications for investors and firms.” National Bureau of Economic Research.
- Laurentsyeve, Nadzeya, and Alessandra Venturini.** 2017. “The social integration of immigrants and the role of policy—A literature review.” *Intereconomics*, 52(5): 285–292.
- Lazear, Edward P.** 1999. “Culture and language.” *Journal of political Economy*, 107(S6): S95–S126.
- Lleras-Muney, Adriana, and Allison Shertzer.** 2015. “Did the Americanization movement succeed? An evaluation of the effect of English-only and compulsory schooling laws on immigrants.” *American Economic Journal: Economic Policy*, 7(3): 258–90.
- Mall, Naline, Ayla Jean Yackley, Marton Dunai, and Aleksander Vasovic.** 2015. “For many refugees, journey to Europe begins on Facebook.” *Reuters*.
- Martén, Linna, Jens Hainmueller, and Dominik Hangartner.** 2019. “Ethnic networks can foster the economic integration of refugees.” *Proceedings of the National Academy of Sciences*, 116(33): 16280–16285.
- Mustafa, Nujeen, and Christina Lamb.** 2017. *The Girl from Aleppo: Nujeen’s Escape from War to Freedom*. New York, NY:Harper Wave. Pg. 111.

- Nawras, Nastaran.** 2017. "Heimweh nach der deutschen Kleinstadt." *Die Zeit*.
- Niehues, Wenke, Nina Rother, and Manuel Siegert.** 2021. "Language acquisition and social contacts progress among refugees: fourth wave of the IAB-BAMF-SOEP survey of refugees." Bundesamt für Migration und Flüchtlinge.
- Paluck, Elizabeth Levy, Seth A Green, and Donald P Green.** 2019. "The contact hypothesis re-evaluated." *Behavioural Public Policy*, 3(2): 129–158.
- Phillips, Nolan E, Brian L Levy, Robert J Sampson, Mario L Small, and Ryan Q Wang.** 2019. "The social integration of American cities: Network measures of connectedness based on everyday mobility across neighborhoods." *Sociological Methods & Research*, 0049124119852386.
- Putnam, Robert D.** 1995a. "Bowling Alone: America's Declining Social Capital." *Journal of Democracy*, 6: 65–78.
- Putnam, Robert D.** 1995b. "Tuning in, tuning out: The strange disappearance of social capital in America." *PS: Political science & politics*, 28(4): 664–684.
- Rao, Gautam.** 2019. "Familiarity does not breed contempt: Generosity, discrimination, and diversity in Delhi schools." *American Economic Review*, 109(3): 774–809.
- Rehbein, Oliver, and Simon Rother.** 2020. "Distance in bank lending: The role of social networks." University of Bonn and University of Mannheim, Germany.
- Ritscher, Annia.** 2016. "Mehr als blosser Klickivismus." *Der Bund*.
- Rupasingha, Anil, Stephan J Goetz, and David Freshwater.** 2006. "The production of social capital in US counties." *The journal of socio-economics*, 35(1): 83–101.
- Sacerdote, Bruce.** 2011. "Peer effects in education: How might they work, how big are they and how much do we know thus far?" In *Handbook of the Economics of Education*. Vol. 3, 249–277. Elsevier.
- Sale, Laura.** 2021. "The social determinants of refugee integration and optimal placement: a study in Denmark." *Working Paper*.
- Scheibe, Katrin, Franziska Zimmer, and WG Stock.** 2019. "Social media usage of asylum seekers in Germany." 263–272.
- Schmidt, Katja, Jannes Jacobsen, and Magdalena Krieger.** 2020. "Social integration of refugees is improving." *DIW Weekly Report*, 10(34): 355–363.
- Schwandt, Hannes, and Amelie C Wuppermann.** 2015. "The Youngest Get the Pill: ADHD Misdiagnosis and the production of education in Germany."
- Srole, Leo.** 1956. "Social integration and certain corollaries: An exploratory study." *American sociological review*, 21(6): 709–716.

**Tagesschau.** 2016. "tagesthemen 22:15 Uhr, 14.04.2016." Available at <https://www.youtube.com/watch?v=f9r7ijIr9Tk>, (accessed: 07.27.2021).

**The Social Capital Project.** 2018. "The Geography of Social Capital in America."

**Tissot, Anna.** 2021. "Hürden beim Zugang zum Integrationskurs."

**United Nations.** 2019. "The number of international migrants reaches 272 million, continuing an upward trend in all world regions, says UN." *United Nations Department of Economic and Social Affairs*.

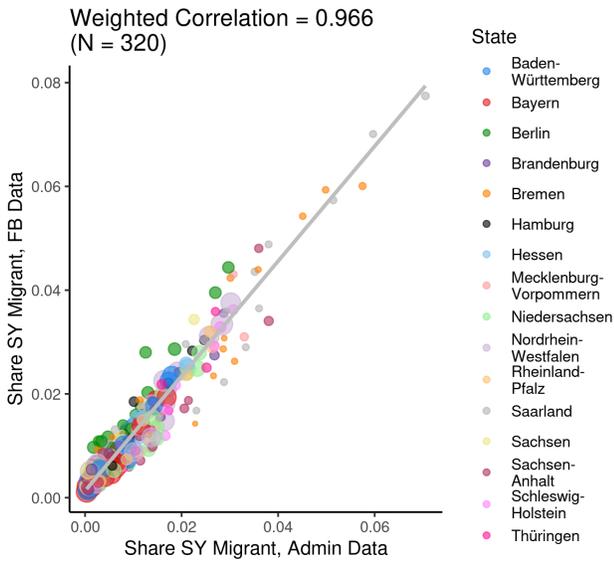
**Wilson, Riley.** 2019. "The impact of social networks on eite claiming behavior." *Review of Economics and Statistics*, 1–45.

# Appendices

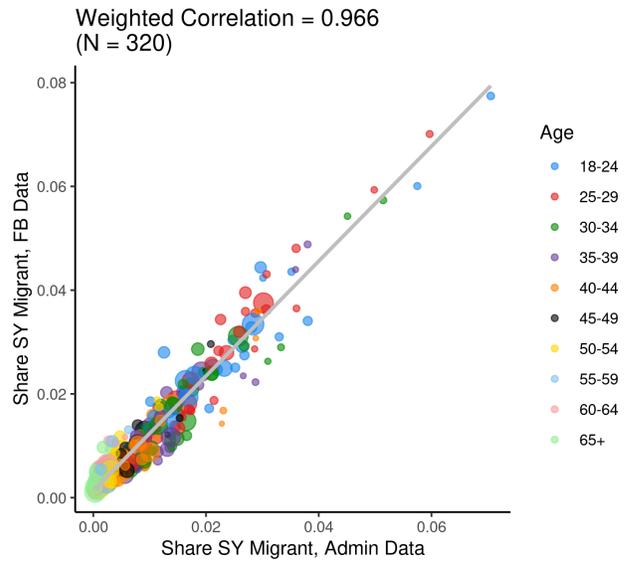
## A Additional Figures and Tables

**Figure A1: Syrian Migrant Sample vs Admin Data**

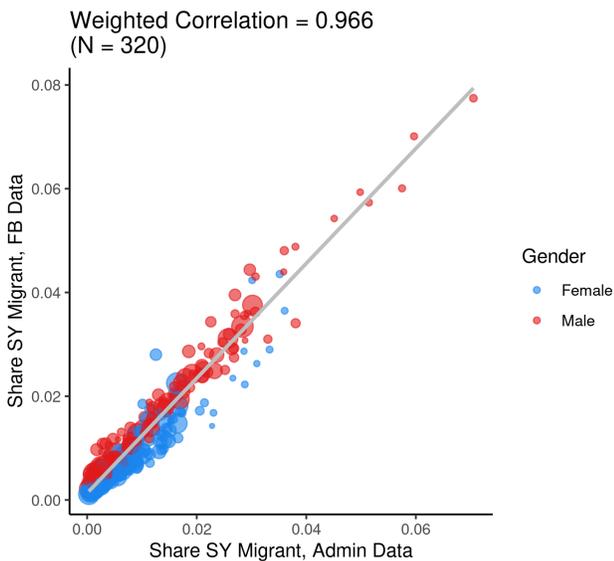
**(a) By State x Age x Gender - Color by State**



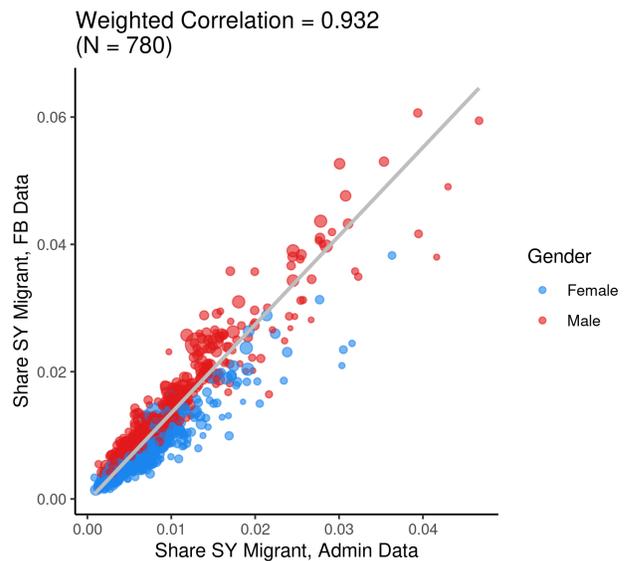
**(b) By State x Age x Gender - Color by Age**



**(c) By State x Age x Gender - Color by Gender**

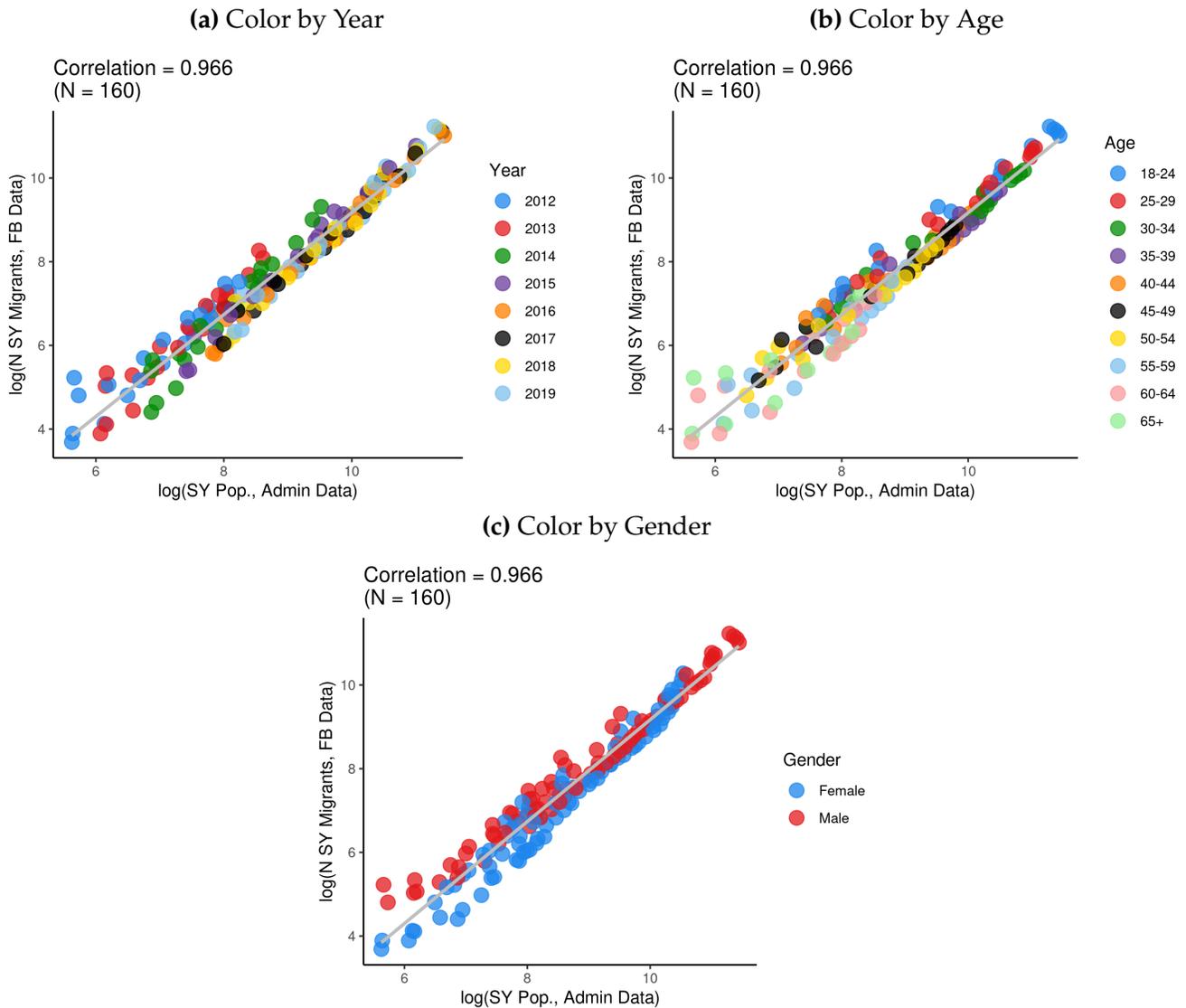


**(d) By County x Gender - Color by Gender**



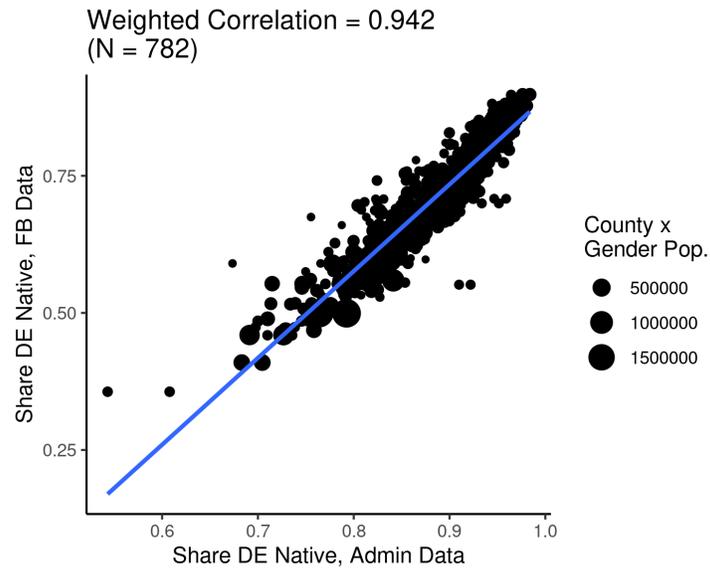
**Note:** Figures show the shares of the primary sample of Facebook users that are also in the Syrian migrant sample (on the y-axis), against shares of the population that are Syrian from administrative data (on the x-axis). The size of each dot is proportional to the size of the population it represents. The solid grey lines are from weighted linear regressions. Panels (a), (b), and (c) plot these shares by state, age, and gender. The age groups are 18-24, 25-29, 30-34, 35-39, 40-44, 45-49, 50-54, 55-59, 60-64, and 65+. There are 16 states X 10 age groups X 2 genders = 320 observations. Panel (d) plots these shares by county and gender. Administrative data is unavailable for 11 counties. There are 390 counties X 2 genders = 780 observations. Panel (a) colors observations by state; panel (b) colors by age; and panels (c) and (d) color by gender.

**Figure A2: Syrian Migrant Sample vs Admin Data - By Age X Gender X Year**



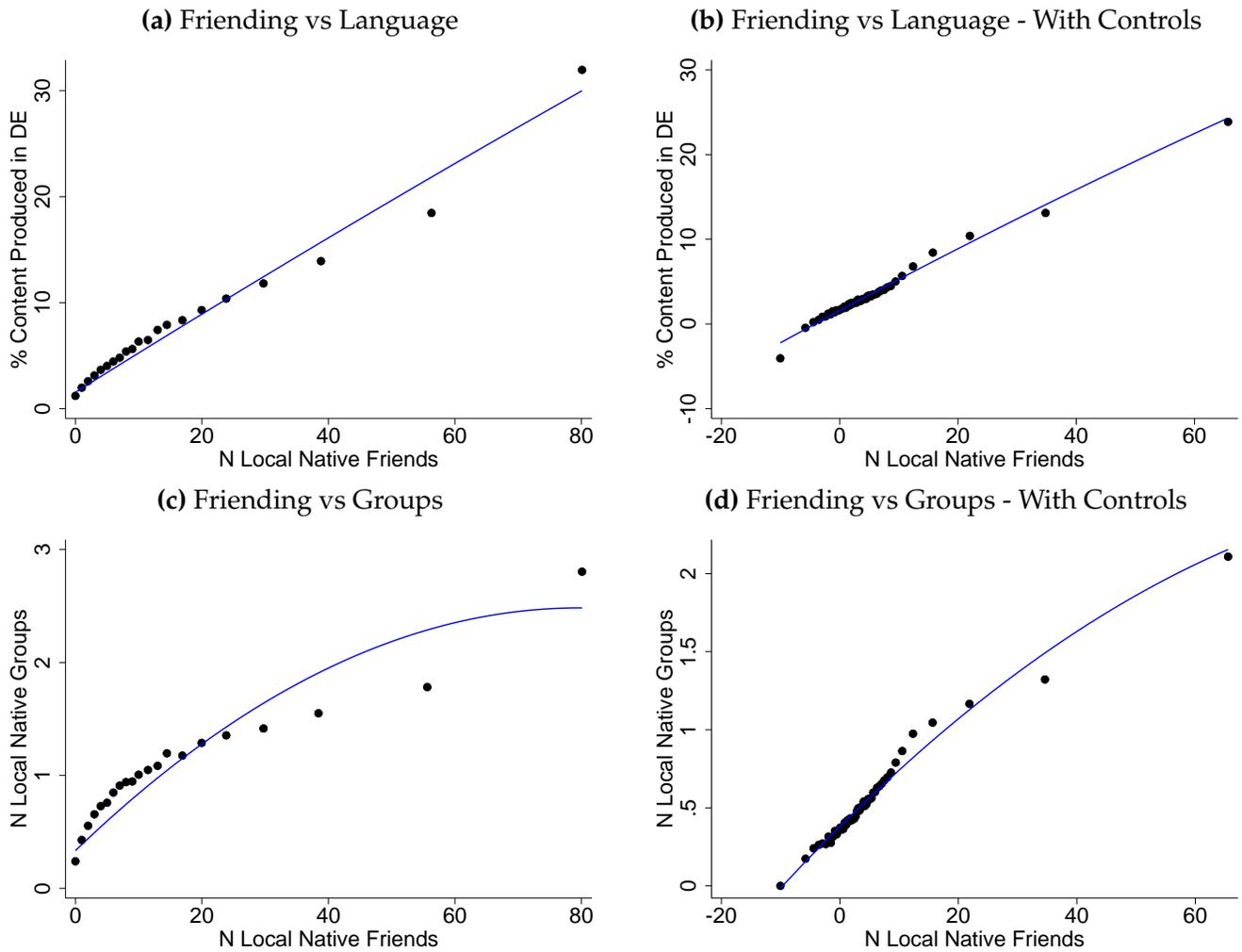
**Note:** Figure shows the number of users in our Syrian migrant sample using Facebook in Germany by the end of each year from 2012 to 2019 (on the y-axis), against analogous measures of Syrian migrant population from German administrative data (on the x-axis). Each observation is an age by gender by year group. The age groups are the same as those used in Figure 1. Both axes are transformed by the natural logarithm. The solid grey line is from a linear regression. Observations are colored by year in panel (a), age in panel (b), and gender in panel (c).

**Figure A3: Native German Sample vs Admin Data**



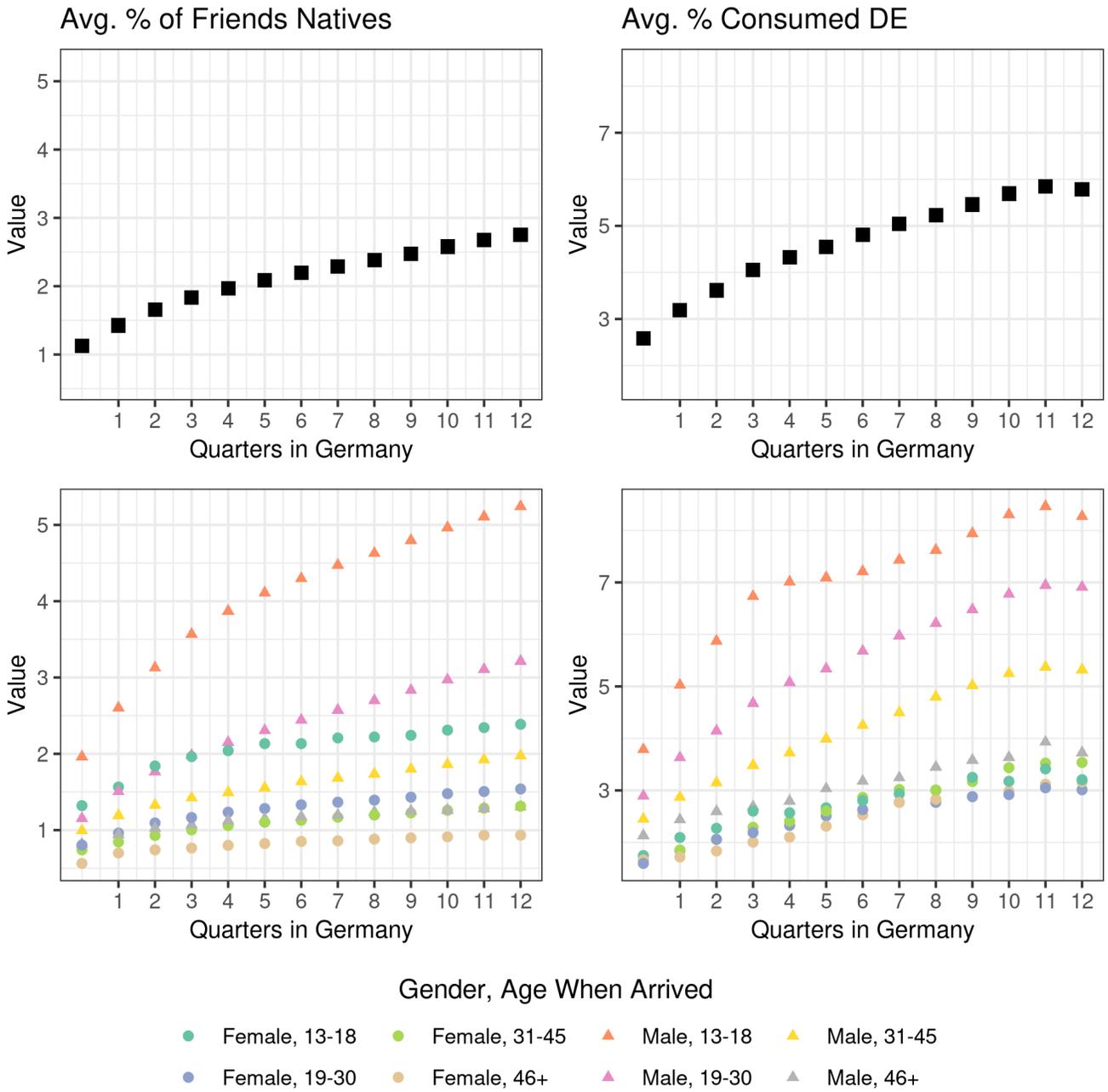
**Note:** Figure shows the shares of the primary sample of Facebook users that are also in the German native sample (on the y-axis), against shares of the population that are native from administrative data (on the x-axis). Each observation is a county by gender group. The size of each dot is proportional to the “true” population it represents. The solid blue lines are from weighted linear regressions. Admin data is unavailable for 10 counties. There are 391 counties X 2 genders = 782 observations.

**Figure A4: Relationship Between Integration Outcomes, Individual Level**



**Note:** Figures show binned scatter plots of individual Syrian migrants' number of local German native friends on the x-axis, against their share of content produced in German in panels (a) and (b), and the number of local native groups they are in panels (c) and (d). Appendix C provides more details on each measure. The measures in panels (b) and (d) are first residualized on the individual-level controls used in column 3 of Table 2. Lines are fit from quadratic regressions.

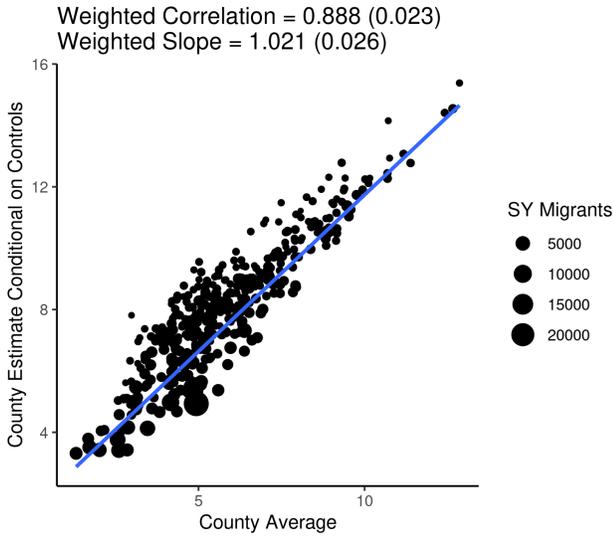
**Figure A5: Integration Over Time For 2015-16 Cohort — Additional Measures**



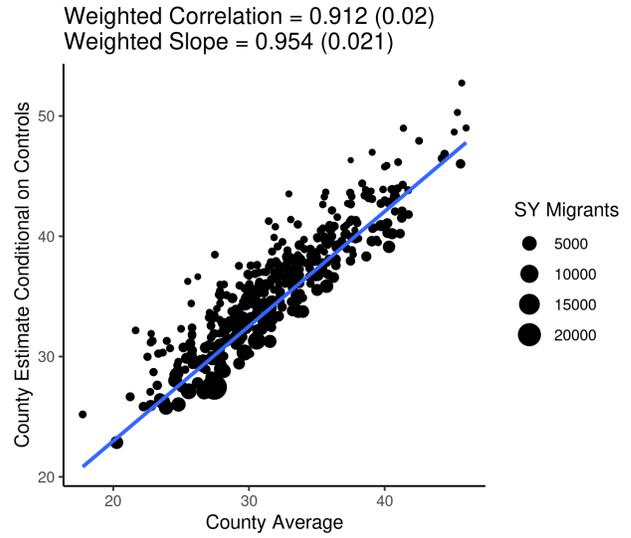
**Note:** Figures show the average values, by quarter, of integration measures for users in the Syrian migrant sample with an observed arrival in 2015 or 2016. The measures are share of friends native (left column) and the share of content consumed in German (right column). Appendix C provides more details on each measure. The top row shows overall trends. In the bottom row each observation's shape and color represents a gender-by-age group.

**Figure A6: Regional Estimates With and Without Controls**

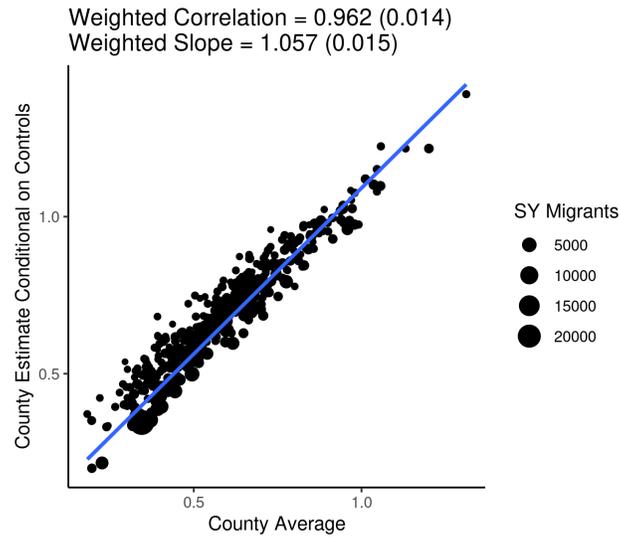
**(a) Friending**



**(b) Language**

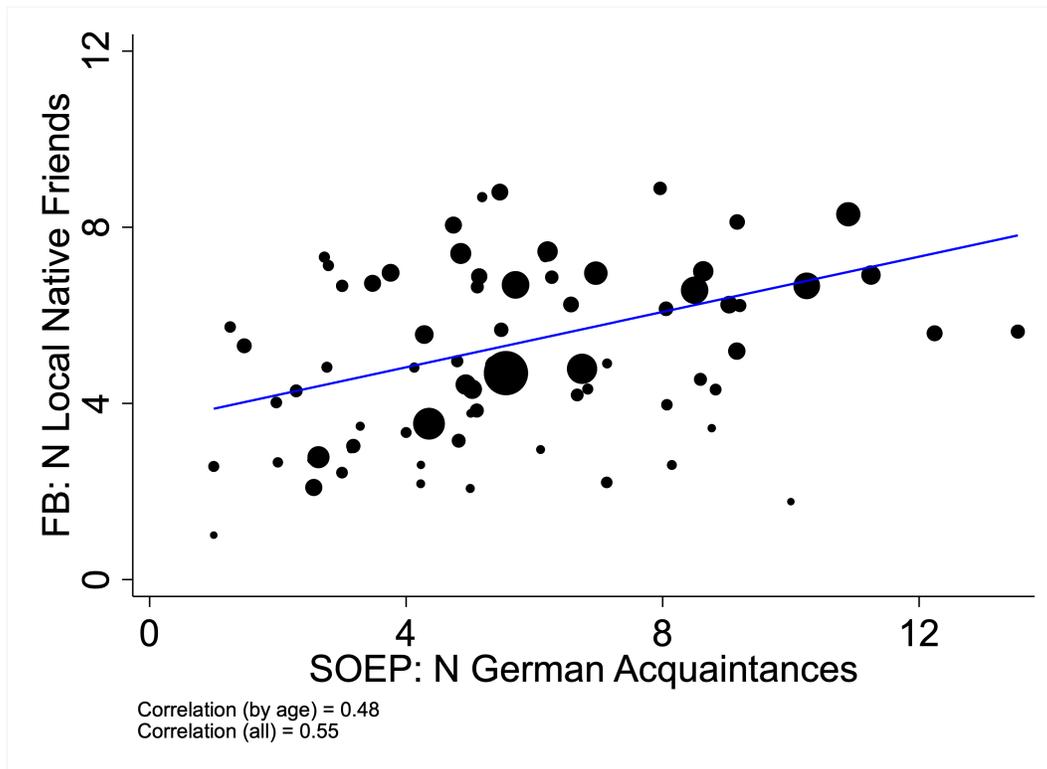


**(c) Groups**



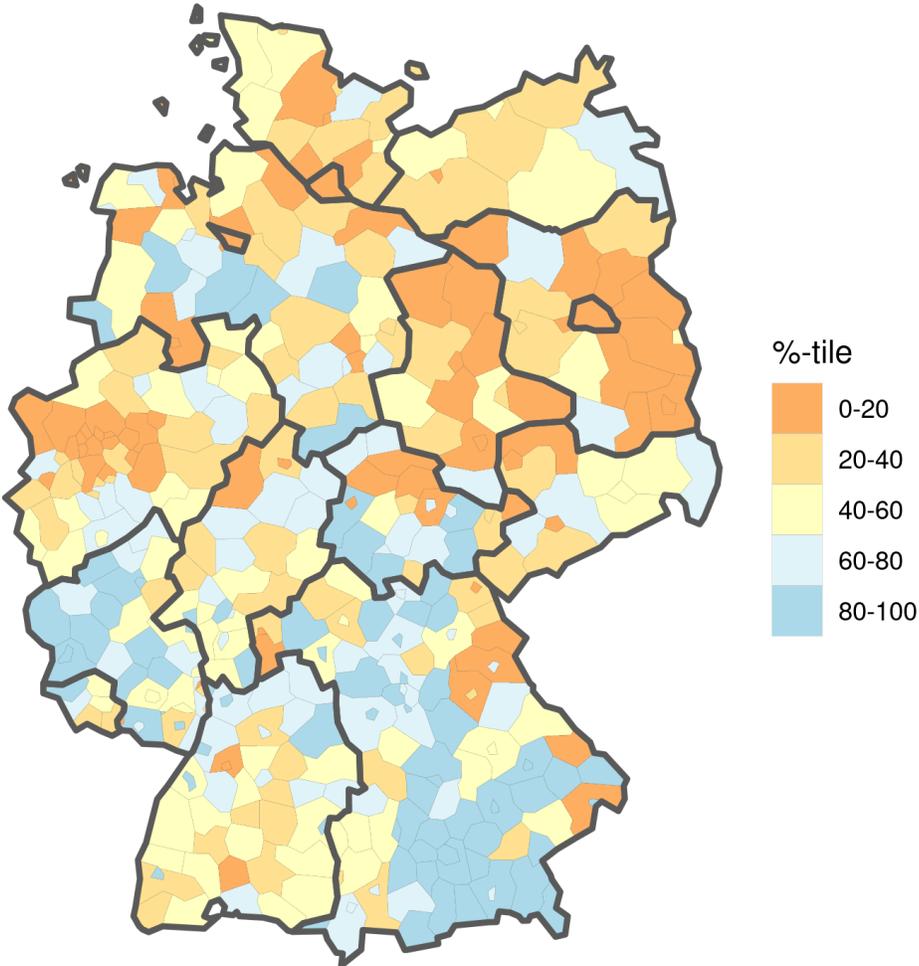
**Note:** Figures show the relationship between county averages of integration outcomes among Syrian migrants vs county-level fixed effect estimates constructed from versions of equation 1. The outcomes are a user's number of local German native friends in panel (a), whether the user produces content in German in panel (b), and the number of local native groups a user is in in panel (c). Appendix C provides more details on each measures. The controls in the fixed effect regressions are those used in column 3 of Table 2.

**Figure A7: Comparing Regional Estimates of Integration - Facebook vs. SOEP**



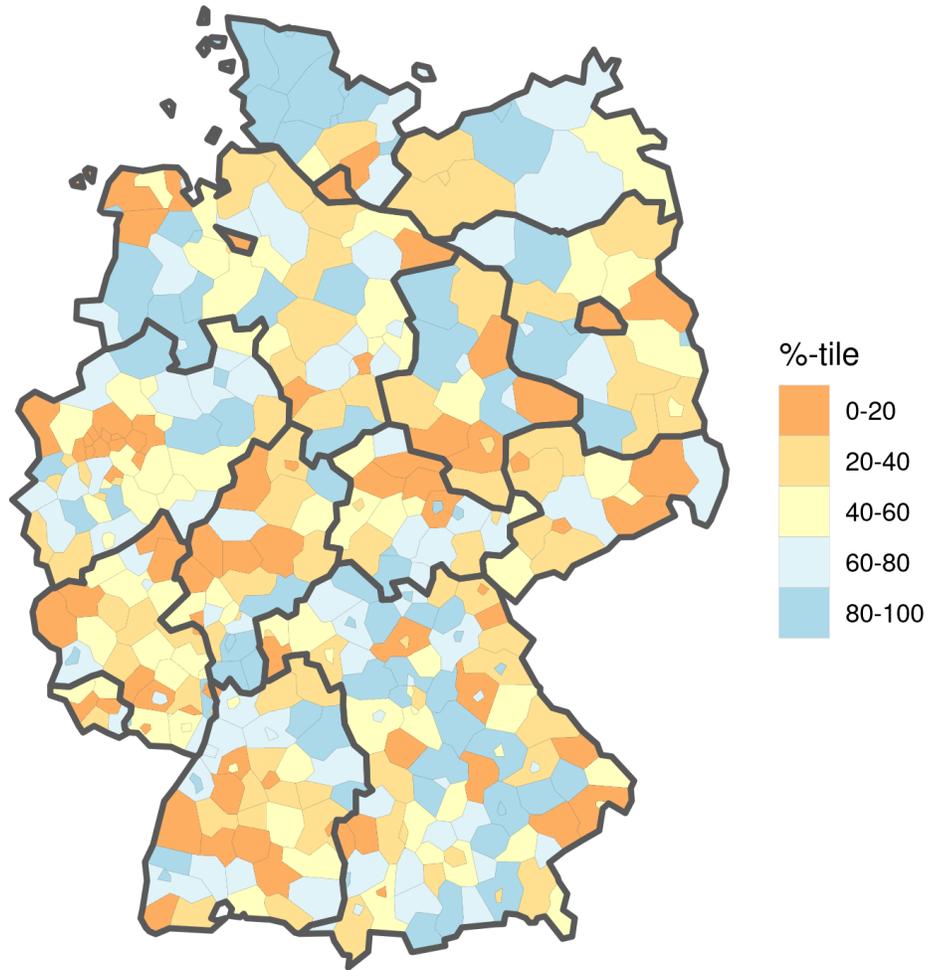
**Note:** Figure compares estimates of social integration based on our Facebook sample with the average number of acquaintances made by recent Syrian migrants in Germany in the SOEP data. The SOEP question is "How many German people have you met since your arrival in Germany with whom you have regular contact?". Each observation in the Figure is a state-by-age-group combination. The size of each dot corresponds to the number of Syrian migrants in the Facebook data. At the bottom of the figure, we report two correlations. The first is a correlation at the state by age-group level, i.e., the same level of aggregation as shown in the plot. The second is a correlation estimated at the state-level, i.e., we further aggregate observations to the state-level and then correlate the two data sources. Both correlations are weighted by the number of Syrian migrants in our Facebook sample.

**Figure A8: Regional Estimates of Integration - German Language Usage**



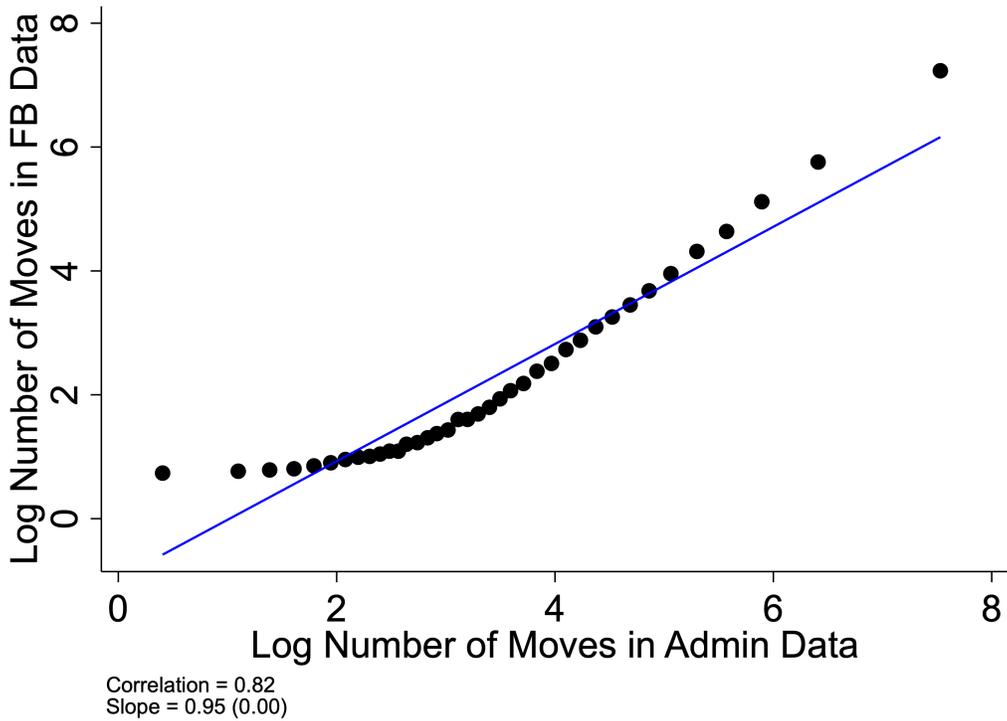
**Note:** Figure shows county-level estimates of Syrian migrant integration based on the share that produce content in the German language (residualized on regional patterns of Facebook usage). Colors correspond to measure ventiles. Darker red and blue areas indicate the lowest and highest integration counties, respectively.

**Figure A9: Regional Estimates of Integration - Local Native Group Joining**



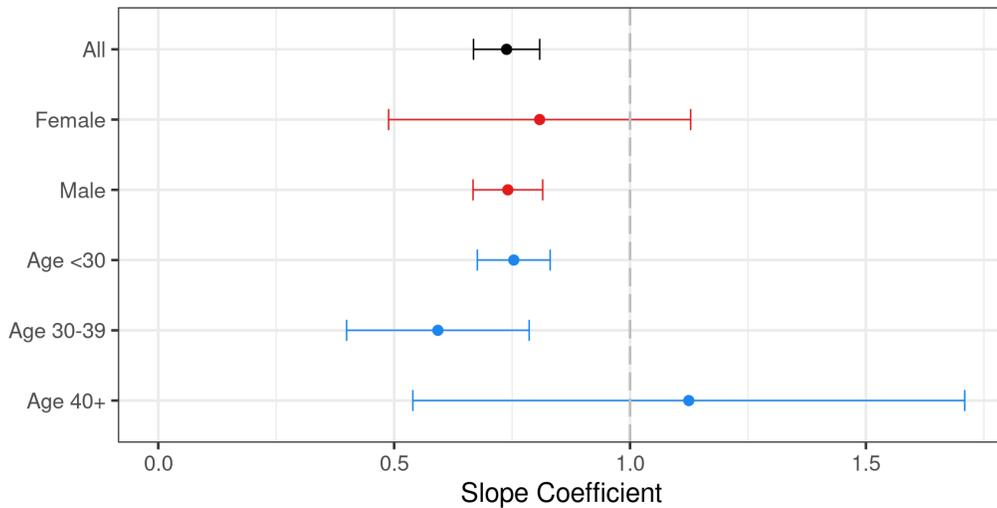
**Note:** Figure shows county-level estimates of Syrian migrant integration based on the average number of native local groups joined (residualized on regional patterns of Facebook usage. This includes the average number of total groups natives in the region have joined, allowing us to account for variation driven by differential usage of the groups feature in general). Colors correspond to measure ventiles. Darker red and blue areas indicate the lowest and highest integration counties, respectively.

**Figure A10: Comparing Movers in Facebook and Administrative Data**



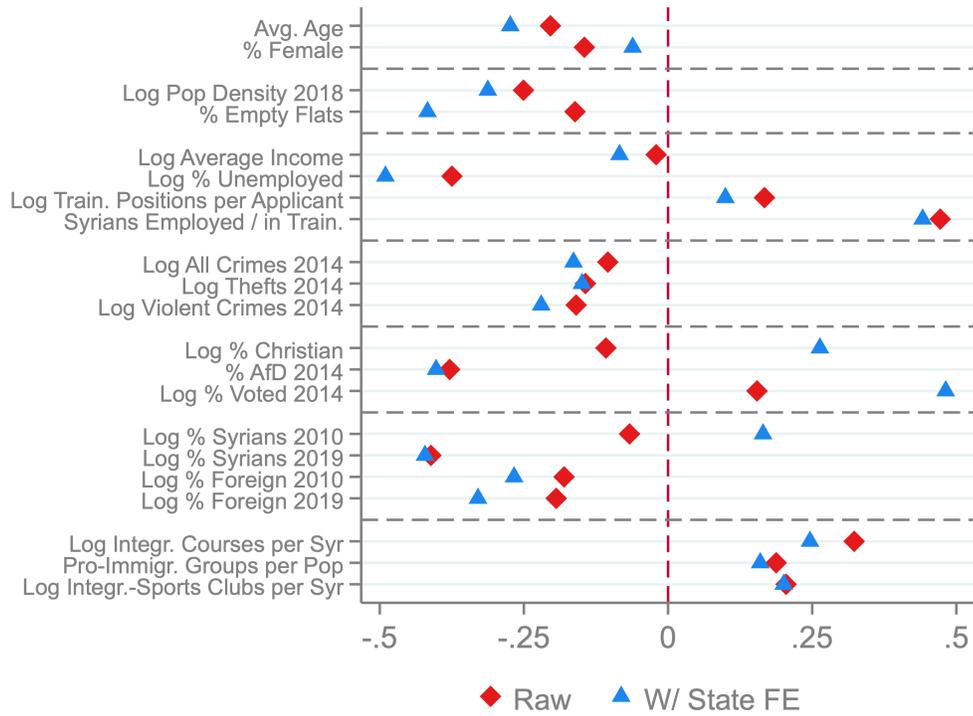
**Note:** Figure compares the number of moves between counties made by individuals between the ages of 18-64 in the years 2016 and 2017 in Facebook and administrative data. We obtained the administrative data from the German Statistical Office. Each observation in this analysis is a county to county combination. The Figure is a binned scatter plot with 40 equally sized bins. The Figure is weighted by the the total number of individuals living in origin and destination county.

**Figure A11: Syrian Migrant Movers - Slope by Demographics**



**Note:** Figure shows slopes corresponding to versions of Figure 5 over certain sub-samples. The coefficient in black corresponds to the slope using the full sample of Syrian migrant movers; the coefficients in red use samples of only one gender; and the coefficients in blue use samples of only one age group. Bars display 95% confidence intervals. The sample sizes used to generate each coefficient are (from top to bottom) 32,853, 6,144, 26,709, 20,796, 8,623, and 3,434.

**Figure A12: County-Level Univariate Correlations with Friending Integration - Long Version**



**Note:** Figure replicates analysis conducted in Figure 9 using an extended set of covariates. For more information regarding covariates, see Appendix J.

**Table A1: Syrian Migrant and German Native Sample Summaries - Additional Measures**

**Panel (a): Syrian Migrant Sample**

	Mean	SD	P10	P25	P50	P75	P90	P99
N Native Friends	9.09	20.54	0	0	2	8	24	151
N Top 50 Native Friends	1.02	2.46	0	0	0	1	3	16
% of Friends Native	3.04	6.19	0.00	0.00	0.80	2.99	8.19	40.25
N Local Other Refugee Country Friends	2.04	3.63	0	0	1	2	6	21
N Local Recent Other Refugee Country Friends	1.04	1.87	0	0	0	1	3	10
% Content Produced in DE	3.39	9.89	0.00	0.00	0.00	2.31	8.48	70.00
% Content Consumed in DE	3.48	8.64	0.00	0.00	0.00	2.91	9.09	60.00
Consumes DE Content (0/100)	41.81	49.32	0	0	0	100	100	100
Account in DE	14.90	35.61	0	0	0	0	100	100
% Groups Local Native	0.88	3.55	0.00	0.00	0.00	0.00	2.22	15.38
Avg. % Native in DE Groups	31.09	30.21	0.15	0.52	25.06	56.44	77.84	92.91

**Panel (b): German Native Sample**

	Mean	SD	P10	P25	P50	P75	P90	P99
N Native Friends	204.73	189.58	40	74	148	269	443	1151
N Top 50 Native Friends	36.87	8.76	25	33	39	43	46	49
% of Friends Native	82.09	14.70	63.75	77.84	86.67	91.61	94.52	98.16
N Local Other Refugee Country Friends	1.12	2.58	0	0	0	1	3	17
N Local Recent Other Refugee Country Friends	0.05	0.22	0	0	0	0	0	1
% Content Produced in DE	94.49	9.70	81.19	92.90	100.00	100.00	100.00	100.00
% Content Consumed in DE	88.60	16.55	65.84	84.06	95.90	100.00	100.00	100.00
Consumes DE Content (0/100)	97.69	15.02	100	100	100	100	100	100
Account in DE	98.61	11.69	100	100	100	100	100	100
% Groups Local Native	22.07	22.34	0.00	4.55	16.67	33.33	50.00	100.00
Avg. % Native in DE Groups	90.42	5.88	83.52	88.16	91.70	94.15	95.95	100.00

**Note:** Table presents summary statistics describing users in our Facebook samples. Panel (a) shows users in the Syrian migrant sample. Panel (b) shows users in the German native sample. Each measure is winsorized at the 99% level. Section 1.1 describes sample construction. Appendix C provides more information on how individual-level outcomes are defined.

**Table A2: Correlation Between Integration Outcomes, Individual Level**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)
(1) N Local Native Friends	1.00														
(2) N Native Friends	0.64	1.00													
(3) N Top 50 Native Friends	0.61	0.54	1.00												
(4) % of Friends Native	0.69	0.61	0.88	1.00											
(5) N Local SY Friends	0.29	0.14	0.01	0.02	1.00										
(6) N Local Other Refugee Country Friends	0.47	0.30	0.19	0.23	0.54	1.00									
(7) N Local Recent Other Refugee Country Friends	0.28	0.15	0.06	0.07	0.53	0.85	1.00								
(8) % Content Produced in DE	0.45	0.40	0.65	0.67	-0.02	0.17	0.03	1.00							
(9) % Content Consumed in DE	0.46	0.40	0.67	0.68	-0.01	0.18	0.05	0.80	1.00						
(10) Produces DE Content	0.24	0.19	0.27	0.31	0.04	0.11	0.03	0.33	0.33	1.00					
(11) Consumes DE Content	0.37	0.25	0.37	0.40	0.09	0.19	0.12	0.52	0.45	0.27	1.00				
(12) Account in DE	0.32	0.21	0.32	0.34	0.11	0.19	0.13	0.34	0.47	0.25	0.57	1.00			
(13) N Local Native Groups	0.29	0.25	0.25	0.27	0.12	0.14	0.09	0.23	0.25	0.14	0.26	0.24	1.00		
(14) % Groups Local Native	0.33	0.26	0.37	0.40	0.03	0.13	0.05	0.36	0.37	0.19	0.27	0.24	0.61	1.00	
(15) Avg. % Native in DE Groups	0.32	0.23	0.32	0.36	0.01	0.14	0.08	0.33	0.35	0.26	0.38	0.36	0.43	0.47	1.00

**Note:** Table presents correlations at the user level across outcome measures for the Syrian migrant sample. Each measure is winsorized at the 99% level. Appendix C provides more information on how outcomes are defined.

**Table A3: Correlation Between Integration Outcomes, Individual Level - With Controls**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)
(1) N Local Native Friends	1.00														
(2) N Native Friends	0.61	1.00													
(3) N Top 50 Native Friends	0.60	0.54	1.00												
(4) % of Friends Native	0.69	0.62	0.86	1.00											
(5) N Local SY Friends	0.20	0.07	0.02	0.03	1.00										
(6) N Local Other Refugee Country Friends	0.39	0.24	0.17	0.20	0.46	1.00									
(7) N Local Recent Other Refugee Country Friends	0.20	0.09	0.05	0.06	0.45	0.83	1.00								
(8) % Content Produced in DE	0.43	0.38	0.61	0.63	-0.01	0.15	0.03	1.00							
(9) % Content Consumed in DE	0.44	0.39	0.63	0.63	0.00	0.16	0.04	0.77	1.00						
(10) Produces DE Content	0.19	0.15	0.21	0.24	0.02	0.06	-0.00	0.27	0.27	1.00					
(11) Consumes DE Content	0.31	0.21	0.32	0.33	0.02	0.12	0.05	0.48	0.40	0.21	1.00				
(12) Account in DE	0.25	0.17	0.26	0.27	0.04	0.12	0.07	0.28	0.42	0.19	0.51	1.00			
(13) N Local Native Groups	0.25	0.22	0.24	0.26	0.03	0.08	0.03	0.23	0.25	0.13	0.22	0.18	1.00		
(14) % Groups Local Native	0.28	0.23	0.33	0.36	0.02	0.09	0.02	0.31	0.32	0.14	0.22	0.18	0.63	1.00	
(15) Avg. % Native in DE Groups	0.23	0.17	0.26	0.29	-0.04	0.05	-0.00	0.27	0.29	0.19	0.30	0.27	0.43	0.42	1.00

**Note:** Table presents correlations at the user level across outcome measures for the Syrian migrant sample. Each measure is first winsorized at the 99% level. Appendix C provides more information on how outcomes are defined. Before constructing the correlations, each measure is residualized on the individual-level controls used in column 3 of Table 2.

**Table A4: Syrian Migrant Integration by Demographics - Language and Groups**

	Produces Content in German (0/100)				N Local Native Groups			
Age 25 - 34	-2.407*** (0.204)	-2.241*** (0.203)	-2.275*** (0.203)	-3.312*** (0.596)	0.167*** (0.006)	0.171*** (0.006)	0.136*** (0.006)	0.140*** (0.019)
Age 35 - 44	-7.133*** (0.238)	-7.161*** (0.237)	-6.875*** (0.237)	-6.615*** (0.733)	-0.002*** (0.007)	-0.007*** (0.007)	0.039* (0.007)	0.072** (0.023)
Age 45 - 54	-13.651*** (0.306)	-13.798*** (0.305)	-12.553*** (0.307)	-16.243*** (0.854)	-0.184*** (0.010)	-0.189*** (0.010)	-0.064*** (0.009)	-0.070*** (0.027)
Age 55+	-18.045*** (0.382)	-18.134*** (0.380)	-16.451*** (0.384)	-24.395*** (1.116)	-0.298*** (0.012)	-0.300*** (0.012)	-0.088*** (0.012)	-0.228*** (0.035)
Female	-15.767*** (0.164)	-15.560*** (0.164)	-16.725*** (0.173)	-18.765*** (0.418)	-0.202*** (0.005)	-0.200*** (0.005)	-0.372*** (0.005)	-0.447*** (0.013)
Household Member in DE 1+ Year Prior	-2.420*** (0.384)	-2.298*** (0.383)	-2.113*** (0.382)		-0.057*** (0.012)	-0.058*** (0.012)	-0.060*** (0.012)	
Non-Household Family in DE 1+ Year Prior	3.418*** (0.347)	3.451*** (0.345)	4.045*** (0.345)		0.023*** (0.011)	0.025*** (0.011)	0.030*** (0.010)	
Quarters Since DE FEs	X	X	X	X	X	X	X	X
Prev Quarters in NUTS3 FEs	X	X	X	X	X	X	X	X
Personal Usage Controls	X	X	X	X	X	X	X	X
County FEs		X	X	X		X	X	X
Log (1 + Total Outside Germany Friends)			X	X			X	X
Log (1 + Total Other Groups)			X	X			X	X
Log (1 + Total Content Produced Past Year)			X	X			X	X
Household FE				X				X
N	349,072	349,072	349,072	84,216	349,072	349,072	349,072	84,216
R-Squared	0.098	0.108	0.113	0.590	0.059	0.076	0.133	0.606
Sample Mean	30.401	30.401	30.401	27.215	0.545	0.545	0.545	0.574

**Note:** Table shows results from regressing various measures on language- and groups-based measures of integration. Each observation in every column is a user in the Syrian migrant Facebook sample. Columns 1 and 5 include controls for age and gender, as well as fixed effects for the number of quarters on Facebook in their current county and the number of quarters since arrival in Germany. For the latter fixed effect, we use a single dummy value for those for which we do not observe arrival, but obtain nearly identical results if we instead drop these users. We also include dummies for whether the user has another Syrian migrant household member or non-household family member in Germany more than year prior to their arrival. For all users not in the “observe arrival timing” sample, these two dummies are set to 0. Columns 2 and 6 add county fixed effects. Columns 3 and 7 add controls for each user’s total number of friends outside Germany, total number of non-local/native groups joined, and total amount of content produced in the last year. Columns 4 and 8 add a household fixed effect, limiting to households for which we observe more than one Syrian migrant. Significance levels: \*(p<0.10), \*\*(p<0.05), \*\*\*(p<0.01).

**Table A5: Syrian Migrant Integration by Demographics - Other Measures**

	N Native Friends	N Top 50 Native Friends	% of Friends Native	% Content Produced in DE	% Content Consumed in DE	Account in DE	% Groups Local Native	Avg. % Native in DE Groups
Age 25 - 34	-0.894*** (0.184)	0.004*** (0.014)	-0.467*** (0.032)	0.076** (0.044)	0.078*** (0.038)	-2.683*** (0.160)	0.197*** (0.010)	-0.136*** (0.160)
Age 35 - 44	-4.728*** (0.216)	-0.263*** (0.016)	-1.446*** (0.038)	-0.694*** (0.051)	-0.749*** (0.044)	-7.099*** (0.187)	0.043 (0.012)	-4.347*** (0.187)
Age 45 - 54	-6.928*** (0.279)	-0.454*** (0.021)	-1.927*** (0.049)	-1.245*** (0.066)	-1.298*** (0.057)	-7.676*** (0.241)	-0.164*** (0.015)	-6.940*** (0.254)
Age 55+	-8.157*** (0.349)	-0.421*** (0.026)	-1.862*** (0.061)	-1.221*** (0.083)	-1.327*** (0.072)	-6.151*** (0.302)	-0.350*** (0.019)	-7.334*** (0.360)
Female	-7.188*** (0.157)	-0.787*** (0.012)	-2.334*** (0.027)	-2.339*** (0.037)	-2.154*** (0.032)	-5.377*** (0.136)	-0.485*** (0.009)	-11.601*** (0.137)
Household Member in DE 1+ Year Prior	-0.610 (0.347)	-0.030 (0.026)	0.013 (0.061)	0.146 (0.082)	-0.057 (0.071)	0.182 (0.300)	-0.014 (0.019)	-0.875*** (0.295)
Non-Household Family in DE 1+ Year Prior	0.667*** (0.314)	0.075*** (0.023)	0.360*** (0.055)	0.535*** (0.074)	0.404*** (0.064)	3.659*** (0.271)	0.098*** (0.017)	2.649*** (0.257)
Quarters Since DE FEs	X	X	X	X	X	X	X	X
Prev Quarters in County FEs	X	X	X	X	X	X	X	X
Personal Usage Controls	X	X	X	X	X	X	X	X
County FEs	X	X	X	X	X	X	X	X
Log (1 + Total Outside Germany Friends)	X	X	X	X	X	X	X	X
Log (1 + Total Other Groups)	X	X	X	X	X	X	X	X
Log (1 + Total Content Produced Past Year)	X	X	X	X	X	X	X	X
N	349,072	349,072	349,072	345,814	346,367	349,072	345,162	237,563
R-Squared	0.064	0.111	0.163	0.121	0.125	0.083	0.077	0.171
Sample Mean	10.592	1.101	3.221	3.388	3.474	14.896	0.754	31.091

**Note:** Table shows results from regressing various measures on outcomes for Syrian migrants in the Facebook sample. All columns include controls for age, gender, time spent on Facebook, number of friends outside Germany, total number of non-local/native groups joined, and total amount of content produced in the last year. They include fixed effects for county, the number of quarters since arrival in Germany (with a single dummy for those for which we do not observe arrival) and the number of quarters on Facebook in their current county. They also include dummies for whether the user has another Syrian migrant household member or non-household family member in Germany more than year prior to their arrival. Column 8 limits to migrants who are members of at least one group of majority users in Germany. Significance levels: \*(p<0.10), \*\*(p<0.05), \*\*\*(p<0.01).

**Table A6: Signal Correlation Between Outcomes, Regional Level**

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
<b>Panel A: Baseline Integration Measures</b>							
(1) SY Migrants - N Local Native Friends	X						
(2) SY Migrants - Produced Content in DE	0.65	X					
(3) SY Migrants - N Local Native Groups	0.27	0.55	X				
(4) SY Migrants - N Local SY Friends	-0.04	-0.55	-0.42	X			
<b>Panel B: Decomposition of Integration Measures</b>							
(5) General Friendliness	0.64	0.31	-0.04	0.11	X		
(6) Relative Friending	0.77	0.56	0.43	-0.16	-0.05	X	
<b>Panel C: Labor Market Integration Measure</b>							
(7) Share Syrians in Employment or Training	0.46	0.63	0.14	-0.36	0.29	0.34	X

**Note:** Table presents signal-adjusted correlations between county-level estimates. The outcomes in panel (a) are the regional averages of Syrian migrants after residualizing on local German natives’ Facebook usage, as described in Section 3.1. The outcomes in panel (b) are the regional decomposition measures described in Section 4.1. Row 5 is general friendliness, generated as a regional average of German natives after residualizing on local German natives’ Facebook usage. Row 6 is relative friending, generated as the quotient from dividing the measure in row 1 by the measure in row 5. The outcome in panel C is an external county-level measure of the share of all Syrians that are employed or in training programs as described in Section 4.4. Correlations are weighted by the number of Syrian migrant users in each county. Our methodology for adjusting correlations to remove sampling error is described in Appendix D.

**Table A7: Native Mover and Comparable Non-Mover Sample Summaries**

**Panel A: Yearly General Friendliness Sample**

	All		To Below Median Place		To Above Median Place	
	Movers	Matched	Movers	Matched	Movers	Matched
% Female	51.95	51.95	51.74	51.74	52.07	52.07
Avg Age	33.70	33.34	34.21	33.87	33.39	33.03
Avg Friends Made (total in year)	21.22	20.11	19.71	19.68	22.12	20.36
Yearly General Friendliness	5.33	9.74	4.81	9.49	5.63	9.89

**Panel B: Yearly Relative Friending Sample**

	All		To Below Median Place		To Above Median Place	
	Movers	Matched	Movers	Matched	Movers	Matched
% Female	52.75	52.75	52.48	52.48	52.90	52.90
Avg Age	31.90	31.86	32.35	32.35	31.65	31.58
Avg Friends Made (total in year)	28.19	20.70	26.41	20.20	29.21	20.99
Yearly Relative Friending	0.20	0.23	0.17	0.22	0.21	0.23

**Note:** Table presents summary statistics describing the users underlying Figure 7. Panels (a) and (b) show summaries for movers and matched non-movers in panels (a) and (b) of Figure 7, respectively. Measures are constructed using movers’ information in the year prior to the move and their matched users in the origin location and time. Matched non-mover summaries are generated by first constructing measures within each mover’s set of matched movers, then averaging across these measures. “Avg Friends Made” is constructed from summing quarterly measures winsorized at the 99% level across all native user-by-quarter observations. The final outcome in each panel is residualized by local natives’ Facebook usage.

**Table A8: Change in Native Mover SY Migrant Friending vs Matched Non-Movers**

	Change in Mover Yearly General Friendliness				Change in Mover Yearly Relative Friending			
Dest-Origin Yearly General Friendliness	0.685*** (0.004)		0.711*** (0.005)	0.602*** (0.005)				
Origin Yearly General Friendliness			-0.636*** (0.005)					
Dest Yearly General Friendliness			0.739*** (0.005)					
Dest-Origin Yearly Relative Friending					0.959*** (0.064)	0.926*** (0.094)	0.988*** (0.086)	
Origin Yearly Relative Friending						-0.988*** (0.071)		
Dest Yearly Relative Friending						0.926*** (0.071)		
Quarter FEs	X	X	X	X	X	X	X	X
Origin County FEs			X				X	
Dest County FEs				X				X
N	1,771,041	1,771,041	1,771,041	1,771,041	1,096,874	1,096,874	1,096,874	1,096,874
Sample Mean	3.160	3.160	3.160	3.160	0.005	0.005	0.005	0.005

**Note:** Table shows results from regressions exploring the change in friending of natives, before and after a move within Germany. Columns 1 and 5 correspond to the relationships depicted in panels (a) and (b) of Figure 7. Columns 2 and 6 regress each component of the difference in the right-hand side measure in columns 1 and 5 separately on the outcome. Columns 3 and 7 repeat columns 1 and 5 with origin fixed effects; columns 4 and 8 repeat columns 1 and 5 with destination fixed effects. We correct for sampling error in the right-hand side measures by randomly splitting the individual-level non-mover data into two halves and instrumenting for one set of averages with the other. See Appendix D for more information this procedure. Significance levels: \*(p<0.10), \*\*(p<0.05), \*\*\*(p<0.01).

## B Construction of “Native German” Sample

For many of our analyses we use a sample of Facebook users, which we refer to as “German natives”, that meet *both* criteria 1 and 2 described below (as well as the primary sample inclusion criteria described in Section 1.1). Our methodology is not intended to proxy for citizenship status or ethnicity; rather it generates a sample of users who generally use the German language and—according to self-reported profile information and home region predictions—appear to have lived in Germany for a substantial amount of time. This will include, for example, individuals of Syrian descent who report a German hometown and primarily use the German language on Facebook. For more details, see footnote 2.

- **Criteria 1:** The user meets *one* of the following
  - The user produces  $\geq 75\%$  of their content in German
  - The user produces  $\geq 50\%$  of their content in German, AND lists a German hometown or high school on their profile
- **Criteria 2:** The user meets *all* of the following
  - Does not list a hometown in a “top migration country”
  - Does not list a high school in a “top migration country”
  - Did not first have a predicted home region in a “top migration country”

The top migration countries are the 15 countries outside of the European Union and within Eastern Europe, the Middle East, or Africa with the most foreign nationals in Germany (for the full list, see: <https://tinyurl.com/8jk2d4yd>).

## C Individual-Level Outcomes

We measure social integration on the three primary dimensions listed below. Within each dimension, we construct a number of specific measures. For brevity, our results primarily focus on a single measure within each dimension, which is noted in **bold**.

### 1. Friendship Measures

- (a) ***N Local Native Friends***: The number of friends a user has in the same county or a bordering county that are in the German native sample.
- (b) *N Native Friends*: The number of friends a user has in the German native sample.
- (c) *N Top 50 Native Friends*: The number of a user's closest 50 friends that are in the German native sample.
- (d) *% of Friends Native*: The percent a user's total friends that are in the German native sample.

### 2. Language Measures

- (a) *% Content Produced in DE*: The share of content a user produces (e.g., in posts, comments) that is in German. "Half-life" of 30 days (i.e., a post 30 days ago is weighted as half a post today).
- (b) *% Content Consumed in DE*: The share of the content a user engages with by using the "react" and "comment" features that is in German. 1 comment = 7 reactions. "Half-life" of 30 days.
- (c) ***Produces Any DE Content*** : An indicator for "**% Content Produced in DE**" is >1%.
- (d) *Consumes Any DE Content*: An indicator for "*% Content Consumed in DE*" is >1%.
- (e) *Account in DE*: Whether a user selected German as their language in their account settings.

### 3. Local Group Participation Measures

- (a) ***N Local Native Groups***: The number of groups a user is in that have 5 - 5,000 users;  $\geq 90\%$  of users in Germany and  $\geq 75\%$  of users in one NUTS2 region; and  $\geq 50\%$  of users in the German native sample.
- (b) *% Groups Local Native*: The share of groups a user is in that match the criteria in "N Local Native Groups."
- (c) *Avg. % Native in DE Groups*: Among groups a user is in which have >90% of users in Germany, the average share of users that are German natives.

We also observe the following additional measures at the individual level:

- *N Local Syrian Friends*: The number of friends a user has in the same county or a bordering county that are in the Syrian migrant sample
- *N Local Other Migrant Country Friends*: The number of friends a user has in the same or bordering county that are migrants (determined by hometown, high school, or past usage) from one of the five countries with the most asylum applicants in Germany in 2020 other than Syria (Turkey, Afghanistan, Iraq, Nigeria, and Iran).

- *N Local Recent Other Migrant Country Friends*: The number of friends a user has matching the “N Local Other Migrant Country Friends” criteria with observed arrival in Germany 2015 or later.<sup>23</sup>

---

<sup>23</sup>As described in Section 1.1, users with an “observed arrival timing” are those who first used Facebook outside of Germany.

## D Assessing the Reliability of Regional Estimates

A potential concern with our regional estimates of integration outcomes is that the differences we observe might be due to sampling error, instead of capturing actual differences in the parameters of interest. In this appendix we explore this concern and describe the methods used to address it.<sup>24</sup>

To assess the degree to which our variation is driven by sampling error, we seek an estimate of:

$$r = \frac{Var(\delta_j)}{Var(\delta_j) + Var(\epsilon_j)} \quad (8)$$

Here  $\delta_j$  is the true (un-observable) parameter for county  $j$ ,  $Var(\delta_j)$  is the variance of that parameter across all counties, and  $Var(\epsilon_j)$  is the variance due to sampling error (noise) when we measure our estimate  $Var(\hat{\delta}_j)$ , such that  $Var(\hat{\delta}_j) = Var(\delta_j) + Var(\epsilon_j)$ . Our outcome of interest is the reliability,  $r$ .

We estimate  $r$  in two ways: (i) a “split sample” estimate generated by randomly splitting the individual-level data in half (within counties) and comparing the resulting estimates; and (ii) a “standard error-based” estimate generated by comparing the magnitudes of the standard error squared of each estimate with the variance of the estimates across counties.

Formally, our “split sample” estimates are given by:

$$\hat{r} = Corr(\hat{\delta}_j^1, \hat{\delta}_j^2) \cdot \frac{\sqrt{Var(\hat{\delta}_j^1)Var(\hat{\delta}_j^2)}}{Var(\hat{\delta}_j)} \quad (9)$$

Where  $\hat{\delta}_j$  is the county-level estimate of  $\delta$  in county  $j$ , the average of individual-level measures across users in the county;  $Var(\hat{\delta}_j^1)$  and  $Var(\hat{\delta}_j^2)$  are the population-weighted variances of these measures in the first and second split samples;  $Var(\hat{\delta}_j)$  is the population-weighted variance in the full sample; and  $Corr(\hat{\delta}_j^1, \hat{\delta}_j^2)$  is the population-weighted correlation.

Our “standard error-based” estimates are given by:

$$\hat{r} = \frac{Var(\hat{\delta}_j) - E[s_{\hat{\delta}_j}^2]}{Var(\hat{\delta}_j)} \quad (10)$$

Where  $s_{\hat{\delta}_j}$  is the standard error of the county level average  $\hat{\delta}_j$  for county  $j$ .

The first two columns of Appendix Table A9 show that the reliability of each of our regional averages is around 0.9 or above regardless of the method used. This suggests that 90% or more of the variance in a given regional measure reflects true latent differences rather than sampling error. As noted in Section 3.1, there are moderate differences in the Facebook usage of natives across space (largely at the intensive margin) which could affect the raw regional averages we measure. To account for this, our estimates in Figure 3 and Appendix Figures A8 and A9 are constructed after residualizing by differences in natives’ Facebook usage. Column 3 of Appendix Table A9 shows split-sample reliability estimates using  $\hat{\delta}_j^1$  and  $\hat{\delta}_j^2$  that have been residualized in this same manner. The reliability estimates are largely unchanged, suggesting our original reliability estimates are not driven by regional differences in usage.

<sup>24</sup>The methods described in this appendix are similar to procedures used in Chetty and Hendren (2018b), Chetty et al. (2022a), and Chetty et al. (2022b).

**Table A9: Reliability of County-Level Measures, Syrian Migrant Sample**

	Reliability		
	Split-Sample	SE-Based	Split-Sample, Usage Control
N Local Native Friends	0.962	0.961	0.938
Produced Any DE Content	0.909	0.901	0.883
N Local Native Groups	0.948	0.946	0.934
N Local Syrian Friends	0.989	0.989	0.989

**Note:** Table shows the reliability of county-level measures. In columns 1 and 2 the measures are averages across Syrian migrant users. In column 3 these measures are residualized on extensive and intensive measures of local natives' Facebook usage, as described in Section 3.1. Reliability is defined by equation 8. The split sample reliability estimates are generated using equation 9. The standard error-based reliability estimates are generated using equation 10.

In Section 4.1, we construct regional measures of *general friendliness* using the German native sample. The sample size for these measures is very large and, accordingly, the reliability estimates using both methods is greater than 0.995. Therefore, essentially all of the sampling error present in our measures of *relative friending* (generated by dividing the Syrian migrant integration outcomes by general friendliness) is driven by the Syrian migrant integration outcomes.

In Table 3 we correlate regional measures against each other across counties. In these cases, the correlations between the estimates may understate the true correlations between parameters because of noise introduced by the sampling error. To recover estimates of the correlation between the true parameters we calculate:

$$\hat{C}orr(\psi_j, \mu_j) = Corr(\hat{\psi}_j, \hat{\mu}_j) \sqrt{\frac{1}{\hat{r}_\psi}} \sqrt{\frac{1}{\hat{r}_\mu}} \quad (11)$$

Where  $Corr(\hat{\psi}_j, \hat{\mu}_j)$  is the correlation between estimates  $\hat{\psi}_j$  and  $\hat{\mu}_j$  (of parameters  $\psi_j$  and  $\mu_j$ ) across all counties  $j$ , and  $\hat{r}_\psi$  and  $\hat{r}_\mu$  are their reliability estimates from equation 10. We present these “signal correlations” in Appendix Table A6.

In Section 3.2 and 4.2, we use certain regional (and region-by-demographics) measures as right-hand side variables in our movers specifications. The sampling error in these estimates will attenuate their regression coefficients. To see this, take the simple regression  $Y = \beta \cdot X + \omega$  where we observe  $\hat{X}$ , an estimate of  $X$  with independent sampling error  $\epsilon$ . Then when estimating  $Y = \hat{\beta} \cdot \hat{X} + \nu$  we have:

$$\begin{aligned} \hat{\beta} &= \frac{Cov(Y, \hat{X})}{Var(\hat{X})} \\ &= \frac{Cov(Y, X + \epsilon)}{Var(X + \epsilon)} \\ &= \frac{Cov(Y, X)}{Var(X) + Var(\epsilon)} < \frac{Cov(Y, X)}{Var(X)} = \beta \end{aligned} \quad (12)$$

To account for this, in our movers analyses we first randomly split the individual-level data used to construct the relevant right-hand side measures in two halves. We then instrument for the value con-

structed by one half with the other. To see the intuition behind this procedure, let  $\hat{X}_1$  and  $\hat{X}_2$  be the split sample estimates. Then the first stage of a two-stage least squares estimate is given by  $\hat{X}_1 = \phi_1 \cdot \hat{X}_2 + \nu_1$ , where  $\phi_1 = \hat{r} = \frac{\text{Var}(X)}{\text{Var}(X) + \text{Var}(e_2)}$ . The reduced form is given by  $Y = \phi_2 \cdot \hat{X}_2 + \nu_2$ , where  $\phi_2 = \frac{\text{Cov}(Y, X)}{\text{Var}(X) + \text{Var}(e_2)}$ . Then the resulting estimate is:

$$\hat{\beta} = \frac{\phi_2}{\phi_1} = \phi_2 \cdot \frac{1}{\hat{r}} \approx \frac{\text{Cov}(Y, X)}{\text{Var}(X)} = \beta. \quad (13)$$

## E Königsteiner Schlüssel and the Assignment of Refugees to Place

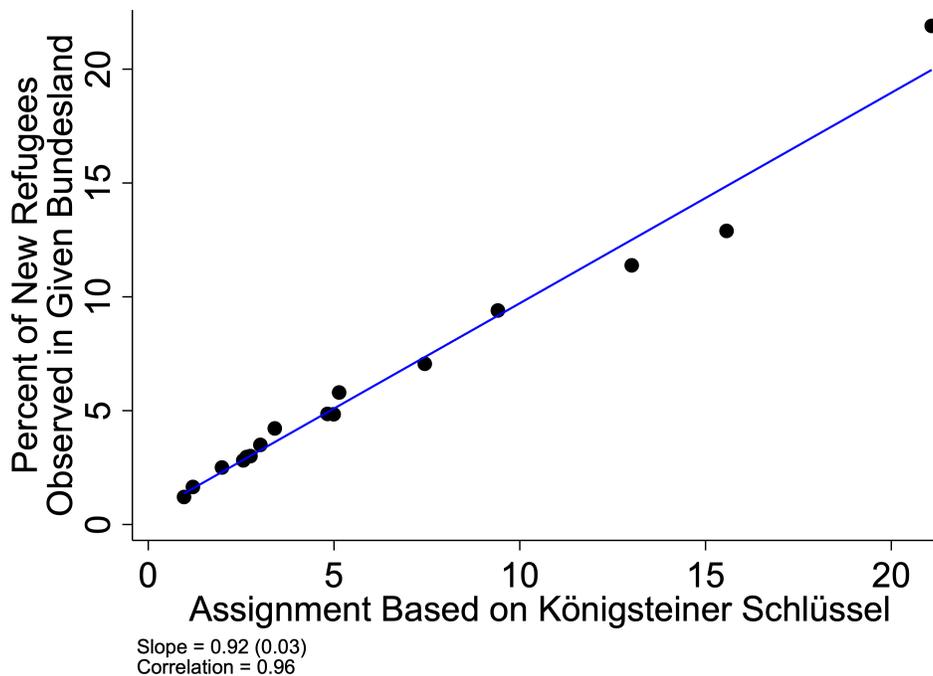
In this section, we attempt to compare the official refugee allocation rule — the so-called Königsteiner Schlüssel (engl. key) — to observed administrative data on refugee assignment.

The Königsteiner Key is an allocation rule which was designed in the 1940s to assign refugees to the sixteen different German states. It takes as input a state's population and tax income and weights these two factors with  $1/3$  and  $2/3$ , respectively (Deutscher Bundestag, 2020). The key is updated annually, but given the slow-moving nature of its inputs, it is stable over time.

To infer to what extent the key has been abided to during the time period of interest for our study, we compare the 2019 assignment key (for data availability reasons) to the percentage of the total number of refugees that live in a given state and have been in Germany for less than 1 year, for each year from 2015 to 2019. The latter measure is intended to approximate for new-arrivals in the absence of direct data on this and the data for this approximate measure is obtained from the German Statistical Office.

Figure A10 shows the result of our comparison. The correlation of 0.96 and a slope of 0.92 indicates that the observed assignment lines up very closely with the official assignment rule. We find this reassuring, as it suggests that despite the large influx of migrants during these year, refugee assignment largely followed the official assignment key. In turn, this is suggestive that once controlling for the Königsteiner key, assignment to place is somewhat random.

**Table A10:** Comparison Königsteiner Key and Assignment of Refugees to Place



**Note:** Figure compares assignment of recent refugees to place with the official assignment key, i.e. the Königsteiner Schlüssel from 2019. The Königsteiner Schlüssel is comprised of a state's total population and a state's tax income where the former is weighted with one third and the latter is weighted two thirds. Assignment of recent refugees is approximated by the percentage of the total number of refugees that live in a given state and have been in Germany for less than 1 year, for each year from 2015 to 2019. The data comes from the German Statistical Office.

## F Additional Information on Primary Movers Model

In Section 3.2 we study the drivers of regional variation in friending outcomes for Syrian migrants. We let each individual's friending outcome be the sum of their county's effect ( $\text{PlaceEffect}^{(p)}$ ) and their personal individual effect ( $\text{IndivEffect}_i$ ). Let  $\text{AvgIndivEffect}^{(p)}$  be the average of  $\text{IndivEffects}$  for individuals in county  $p$ . Then the difference between the average outcomes,  $x$ , in two regions, (2) and (1), is the sum of differences between the place-based effect and the average of individual-effects.

$$x^{(2)} - x^{(1)} = (\text{PlaceEffect}^{(2)} - \text{PlaceEffect}^{(1)}) + (\text{AvgIndivEffect}^{(2)} - \text{AvgIndivEffect}^{(1)}) \quad (14)$$

We want to know the share of  $x^{(2)} - x^{(1)}$  that is due to place-based effects, formally:

$$\frac{\text{PlaceEffect}^{(2)} - \text{PlaceEffect}^{(1)}}{(\text{PlaceEffect}^{(2)} - \text{PlaceEffect}^{(1)}) + (\text{AvgIndivEffect}^{(2)} - \text{AvgIndivEffect}^{(1)})} \quad (15)$$

We cannot observe any of these parameters directly, but we know that when a mover moves from (1) to (2), only the place-based factors should change. So for mover  $i$  who moves from (1) to (2) at time  $t$ :

$$y_{i,t}^{\Delta} = (\text{PlaceEffect}^{(2)} - \text{PlaceEffect}^{(1)}) \quad (16)$$

Where  $y_{i,t}^{\Delta}$  is the change in outcome before and after the move for mover  $i$ . Then  $\alpha$ , below, is equivalent to equation 15, our outcome of interest.

$$y_{i,t}^{\Delta} = \alpha \cdot (x^{(2)} - x^{(1)}) \quad (17)$$

In our final model we allow for separate place effects across certain observable demographics such as age and gender, as well as time. The  $\text{AvgIndivEffect}$  is then the average of the remaining unobservable individual effects. When estimating  $\alpha$  we remove the variation in  $y_{i,t}^{\Delta}$  explained by overall time trends (e.g., if throughout Germany Syrian migrants make more native friends over time) by adding quarter of move fixed effects,  $\zeta_t$ .

## G Cross-Sectional Analysis of Movers and German Language Usage

We assess the degree to which selection drives our regional estimates of German language integration using a cross-sectional movers design. This follows similar designs in Chetty and Hendren (2018a) and Finkelstein, Gentzkow and Williams (2019), and differs from the design used in Sections 3.2 and 4.2 which utilize panel data on movers' friending. In particular, we model German language usage as a linear combination of the outcomes of non-movers in each of the mover's locations. Then, using the same mover criteria as in Figure 5, we estimate:

$$y_i = \alpha_0 + \alpha_1 \sum_p q(i, p) * x_{p,d(i)} + \kappa_{d(i)} + \epsilon_i \quad (18)$$

Here,  $y_i$  is an indicator for whether individual  $i$  produces German content on Facebook and  $q(i, p)$  is the share of their quarters in Germany spent in place  $p$ . The notation  $d(i)$  represents a set of demographics used to match movers to similarly situated non-movers.  $x_{p,d}$  is the share of users in place  $p$  and demographic group  $d$  that produces German content, and  $\kappa_{d(i)}$  are demographic group fixed effects.<sup>25</sup> In our strictest specifications, we also add fixed effects for users' first and current county in Germany.

**Table A11:** Syrian Migrant Mover Language Integration vs Weighted Average of Places

	Produces Content in German (0/100)				
Predicted Prob. Of Using German (Weighted Avg. of Places Lived)	0.863*** (0.037)	0.857*** (0.043)	0.863*** (0.058)	0.813*** (0.042)	0.816*** (0.058)
FEs	Cohort	Cohort	Cohort	Cohort X Curr. Cnty.	Cohort X Curr. Cnty. X First Cnty.
Sample		< 75% in Max County	< 60% in Max County		
N	23,249	18,233	10,172	23,069	14,474
Sample Mean	38.075	37.959	38.252	38.099	36.977

**Note:** Table shows results for comparisons between the German language usage of Syrian migrants who moved between counties and their predicted language usage based on the outcomes of non-movers in the places they lived. For each location, movers are matched non-movers by age, gender, and the first year they used Facebook in Germany (cohort). Column 1 shows our baseline specification from equation 18, which includes cohort fixed effects. Column 2 limits to only users who spent < 75% of their quarters in Germany in one county. Column 3 limits to those who spent < 60%. Column 4 repeats column 1 with cohort-by-current county fixed effects; column 5 repeats column 1 with cohort-by-current county-by-first county in Germany fixed effects. We correct for sampling error in the right-hand side measures by randomly splitting the individual-level non-mover data into two halves and instrumenting for one set of averages with the other. See Appendix D for more information this procedure. Significance levels: \*(p<0.10), \*\*(p<0.05), \*\*\*(p<0.01).

In contrast to equation 4, our unit of observation is a mover, not a move, and we use movers' location for every quarter they have been in Germany. As in our panel analyses, we cannot observe  $x_{p,c(i)}$ , but instead account for sampling error by constructing estimates  $\hat{x}_{p,c(i)}$  from random halves of the data and instrumenting for one with the other. We also again relax the assumption of fully additive-seperability between individual-level factors and place-based effects by matching movers to similarly situated non-

<sup>25</sup>These fixed effects remove variation driven by the demographic matching from our slope estimates.

movers on gender, age group, and year of arrival in Germany. This allows for non-additive interactions with these demographics. We enforce that each mover must have 20 matched non-movers.<sup>26</sup>

Table A11 presents results from our analysis. In column 1, an estimate of  $\alpha_1$  close to 1 would suggest that a Syrian migrant's likelihood of using German on Facebook is close to the averages of migrants in each location they have lived, weighted by the amount of time they lived in each location. The resulting slope estimate of 0.86 shows that this is the case. While this evidence is consistent with places having an *effect* on migrants' German language integration, it does not rule out alternative explanations. For example, it is possible that our sample includes many users who have spent a long time in a single location, and that the right hand side weighted averages are often dominated by a single region. If this were the case, our estimates could be largely driven by movers behaving similarly to local non-movers in general, rather than by place-based effects in particular. Columns 2 and 3 provide evidence that this story does not drive our overall results, as our estimates of  $\alpha_1$  remain similar when limiting our sample to users who spent <75% or <60% of their time in Germany in one county, respectively.

In column 4 we take another approach to testing whether our results are indicative of causal effects of place. In particular, we control for each user's current county, thereby identifying our slope estimates from variation in the user's origin counties. The slope estimate decrease slightly, but remains around 0.81. This suggests that much of the variation in language outcomes amongst movers across regions today is determined by where they *originally* lived in Germany, providing evidence against selection effects. In the final column, we control for both first county and final county fixed effects. Our identification, therefore, comes from the amount of *time* users' spend in each particular place. The slope estimates remains at 0.82, providing more evidence that a migrant's probability of using the German language scales linearly in proportion to the time they spend in high- and low-integration places.

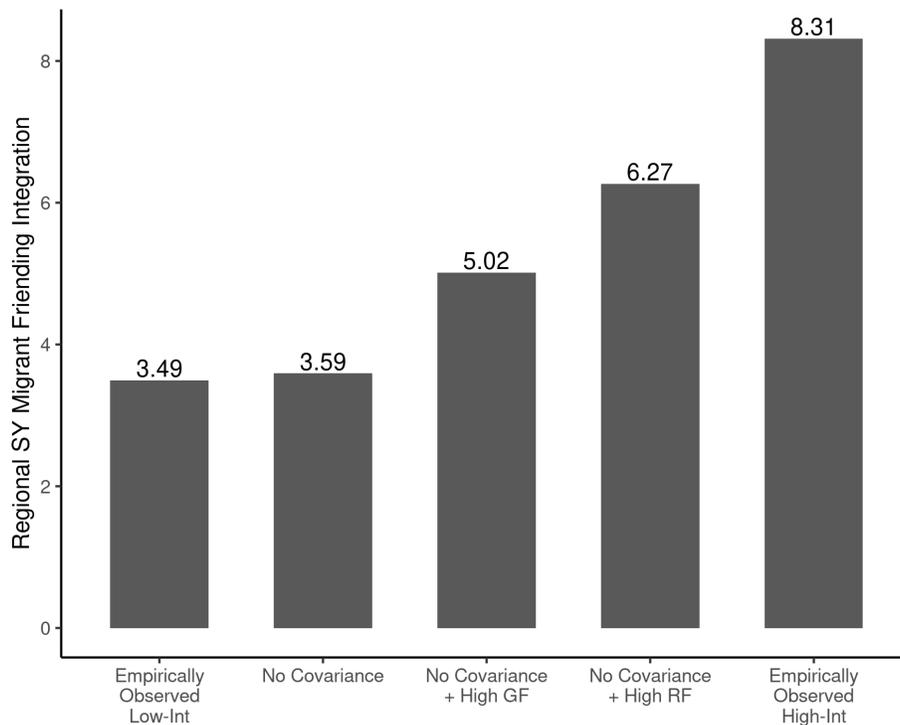
---

<sup>26</sup>This threshold is higher than the five user minimum in Section 3.2. Our sample in this analysis, however, will remain larger because we (mechanically) do not enforce temporal matching.

## H Decomposition of High- vs Low-Integration Regional Differences

In Figure A13, we conduct counterfactual exercises to explore the degree to which each of our two components explain the differences between counties with high- and low-friending integration. This follows a similar exercise in Chetty et al. (2022b). The first and fifth bars show the average integration of migrants in top and bottom quintile counties, respectively. Syrian migrants in top quintile counties make 8.31 native local friends on average, versus 3.49 in bottom quintile counties. In the second bar we multiply the bottom quintile averages of general friendliness and relative friending, thereby removing any within-quintile covariance. Doing so somewhat increases the value from the first bar, consistent with the small negative correlation between the two components in Table 3. The third and fourth bars replace the bottom-quintile averages of general friendliness and relative friending with the corresponding top-quintile averages, respectively. We view this as a counterfactual in which we hold one of the two integration components of low-integration regions fixed and adjust the other to the levels of high-integration regions. We interpret the difference between the second and fourth bars (2.68), compared to the second and third bars (1.43), as relative friending explaining about 1.9x as much of the difference between high and low-integration places as general friendliness.

**Figure A13:** Decomposition of Difference Between High- and Low-Integration Regions



**Note:** Figure shows how much of the difference between high and low friending integration counties is driven by general friendliness versus relative friending. The first and fifth bars show the average friending integration of Syrian migrants in top and bottom quintile counties, respectively. The second bar replaces each county observation from the first bar with the bottom quintile averages of general friendliness and relative friending. The third and fourth bars replace the bottom-quintile averages of general friendliness and relative friending with the corresponding top-quintile averages, respectively.

# I Individual-level Correlates of Natives Behavior Towards Migrants

This appendix explores the relationship between observable native characteristics and behaviors toward Syrian migrants. In particular we focus on their (i) friending of local Syrian migrants; (ii) general friendliness; (iii) relative friending; and (iv) joining of pro-immigration organizations on Facebook.

**Table A12: Natives - Measures of Friending**

	N Local SY Friends		General Friendliness		Relative Friending		In Pro Imm. Group (0/100)	
Age 25 - 34	-0.073*** (0.000)	-0.073*** (0.000)	-19.097*** (0.098)	-14.407*** (0.092)	-0.059*** (0.001)	-0.061*** (0.001)	0.359*** (0.018)	0.146*** (0.018)
Age 35 - 44	-0.116*** (0.000)	-0.114*** (0.000)	-55.586*** (0.103)	-52.328*** (0.097)	-0.081*** (0.001)	-0.080*** (0.001)	0.951*** (0.018)	0.858*** (0.018)
Age 45 - 54	-0.132*** (0.000)	-0.131*** (0.000)	-62.533*** (0.108)	-62.415*** (0.102)	-0.098*** (0.001)	-0.095*** (0.001)	1.116*** (0.019)	1.152*** (0.019)
Age 55+	-0.139*** (0.000)	-0.141*** (0.000)	-82.666*** (0.108)	-84.728*** (0.102)	-0.098*** (0.001)	-0.095*** (0.001)	2.105*** (0.020)	2.157*** (0.020)
Female	-0.015*** (0.000)	-0.015*** (0.000)	-19.519*** (0.056)	-18.725*** (0.053)	-0.008*** (0.001)	-0.009*** (0.001)	0.882*** (0.010)	0.843*** (0.010)
Has College	0.006*** (0.000)	0.006*** (0.000)	4.131*** (0.060)	7.619*** (0.056)	-0.000 (0.001)	-0.002*** (0.001)	1.931*** (0.011)	1.788*** (0.011)
Prev Quarters in NUTS3 FEs	X	X	X	X	X	X	X	X
Personal Usage Controls	X	X	X	X	X	X	X	X
County FEs		X		X		X		X
N	17,768,822	17,768,822	17,768,822	17,768,822	17,515,164	17,515,164	17,768,141	17,768,141
R-Squared	0.020	0.031	0.170	0.263	0.001	0.002	0.035	0.042
Sample Mean	0.086	0.086	122.510	122.510	0.074	0.074	4.835	4.835

**Note:** Table shows results from regressing various measures on various outcomes for users in the German native Facebook sample. The outcome is their number of local friends in the Syrian migrant sample in columns 1 and 2; their number of local friends in the German native sample in columns 3 and 4; their relative friending to Syrians and Germans defined by equation 6 in columns 5 and 6; and the number of groups registered with *ProAsyl* they are in in columns 7 and 8. Columns 1, 3, 5, and 7 include controls for age, gender, and whether they list a college on Facebook, as well as fixed effects the number of quarters on Facebook in their current county. They also include linear controls for  $\log(0.5 + \text{minutes on FB in the last 28 days})$ ,  $\log(91 - \text{days on Facebook out of the last 90})$ ,  $\log(1081 - \text{days on Facebook out of the last 1080})$ . Columns 2, 4, 6, and 8 add county fixed effects. In columns 7 and 8 the personal usage controls also include fixed effects for each number of Facebook groups a user is in. Significance levels: \*( $p < 0.10$ ), \*\*( $p < 0.05$ ), \*\*\*( $p < 0.01$ ).

Equation 1 is our multivariate regression of interest. Each observation is a German native user. In all specifications we include controls for the amount of time each user spends on Facebook and for the number of quarters they have been on Facebook in their current county. In certain specifications we also include county fixed effects.  $Y_i$  represents measures of the four outcomes listed above. Friending of local Syrian migrants is measured by the user's number of local Syrian migrant friends. Individual-level general friendliness is measured by the user's number of local native friends. We construct individual-level relative friending by replacing each term in the numerator of equation 6— $NLocalFriends_c^{DE \rightarrow SY}$  and  $NLocalFriends_c^{DE \rightarrow DE}$ —with its individual-level analog.<sup>27</sup> We identify pro-immigration Facebook pages and groups using a combination of string, url, and manual matching. Our outcome measure is

<sup>27</sup>A user must have at least one local native friend for this individual-level measure. The county-level average of this measure will equal the county-level measure in equation 6 if each observation in the former is weighted by the user's number of local native friends.

whether a user “likes” one of these page or is in one of these groups. In total, we identify 8,171 groups and pages, and measure 2.1 million user-page or user-group connections.

Table A12 presents results. Columns 1 and 2 show that younger natives and male natives are more likely to befriend migrants than older and female natives, respectively. Columns 3 and 4 show that these patterns are driven in part by general friendliness: a native being younger, male, or college educated is associated with having a larger network of local native friends. Columns 5 and 6 show that our individual-level measure of relative friending is also higher for younger and male German natives, while it is somewhat lower for college educated Germans compared to college educated Germans. Because Syrian migrants in Germany are more likely to be young and male than the average German native (see Table 1), one possible explanation for this finding is that homophily plays a strong role in shaping which natives befriend Syrian migrants. For example, younger German natives might be more likely to connect with younger Syrian migrants because younger people in general are more likely to connect, rather than because of particular behaviors toward migrants.

Columns 7 and 8 show that older, female, and college-educated natives are more likely than others to join pro-immigration groups on Facebook, conditional on Facebook usage. (For these analyses we include fixed effects for each number of total Facebook groups as user is in, holding constant a user’s overall propensity to join Facebook groups. Our results remain qualitatively unchanged without this control.) These are *opposite* the relationships presented for relative friending in columns 5 and 6, suggesting that is not necessarily those who are most supportive of pro-immigration groups that are most likely to disproportionately befriend Syrian migrants. This is again consistent with a story in which homophily, above specific attitudes or behaviors toward migrants, contribute to the demographic differences we observe in prior columns.

## J Data Description of County-Level Covariates

Variable	Description	Data Source	Link to Data
Average Age	Average age of populaton, 2014	German Statistical Office	<a href="https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en">https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en</a>
% Female Age	Share of population that is female, 2014	German Statistical Office	<a href="https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en">https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en</a>
Pop. Density 2018	Population density, 2018.	Regionalatlas Deutschland	<a href="https://www-genesis.destatis.de/gis/genView?GenMLURL=https://www-genesis.destatis.de/regatlas/A1002-1.xml&amp;CONTEXT=REGATLAS01">https://www-genesis.destatis.de/gis/genView?GenMLURL=https://www-genesis.destatis.de/regatlas/A1002-1.xml&amp;CONTEXT=REGATLAS01</a>
% Empty Flats	Share of flats that that are vacant, 2017	Thünen-Landatlas	<a href="https://karten.landatlas.de/app/landatlas/">https://karten.landatlas.de/app/landatlas/</a>
Average Income	Average income, 2018	Statistische Ämter des Bundes und der Länder ( <i>Federal and state statistical offices</i> )	<a href="https://www.statistikportal.de/de/vgrd/ergebnisse-kreisebene">https://www.statistikportal.de/de/vgrd/ergebnisse-kreisebene</a>
% Unemployed	Unemployment rate, 2014	Bundesagentur für Arbeit ( <i>Federal Employment Agency</i> )	<a href="https://statistik.arbeitsagentur.de/SiteGlobals/Forms/Suche/Einzelheftsuche_Formular.html?topic_f=beschaeftigung-eu-heft-eu-heft">https://statistik.arbeitsagentur.de/SiteGlobals/Forms/Suche/Einzelheftsuche_Formular.html?topic_f=beschaeftigung-eu-heft-eu-heft</a>
Train. Positions per Applicant	Number of training positions (Lehrstellen) per applicant (Auszubildender)	Bundesagentur für Arbeit ( <i>Federal Employment Agency</i> )	<a href="https://statistik.arbeitsagentur.de/SiteGlobals/Forms/Suche/Einzelheftsuche_Formular.html?gtp=15084_list%253D4&amp;topic_f=analyse">https://statistik.arbeitsagentur.de/SiteGlobals/Forms/Suche/Einzelheftsuche_Formular.html?gtp=15084_list%253D4&amp;topic_f=analyse</a>
Syrians Employed / in Train.	Number of Syrians employed or in training divided by Syrian population	Bundesagentur für Arbeit ( <i>Federal Employment Agency</i> )	<a href="https://statistik.arbeitsagentur.de/SiteGlobals/Forms/Suche/Einzelheftsuche_Formular.html?topic_f=beschaeftigung-eu-heft-eu-heft">https://statistik.arbeitsagentur.de/SiteGlobals/Forms/Suche/Einzelheftsuche_Formular.html?topic_f=beschaeftigung-eu-heft-eu-heft</a>
All Crimes 2014	Reported crimes (total) per population, 2014	Polizeiliche Kriminalstatistik ( <i>Police Crime Statistics</i> )	<a href="https://www.bka.de/DE/AktuelleInformationen/StatistikenLagebilder/PolizeilicheKriminalstatistik/pks_node.html">https://www.bka.de/DE/AktuelleInformationen/StatistikenLagebilder/PolizeilicheKriminalstatistik/pks_node.html</a>
Thefts 2014	Theft crimes per population, 2014	Polizeiliche Kriminalstatistik ( <i>Police Crime Statistics</i> )	<a href="https://www.bka.de/DE/AktuelleInformationen/StatistikenLagebilder/PolizeilicheKriminalstatistik/pks_node.html">https://www.bka.de/DE/AktuelleInformationen/StatistikenLagebilder/PolizeilicheKriminalstatistik/pks_node.html</a>
Violent crimes 2014	Violent crimes per population, 2014	Polizeiliche Kriminalstatistik ( <i>Police Crime Statistics</i> )	<a href="https://www.bka.de/DE/AktuelleInformationen/StatistikenLagebilder/PolizeilicheKriminalstatistik/pks_node.html">https://www.bka.de/DE/AktuelleInformationen/StatistikenLagebilder/PolizeilicheKriminalstatistik/pks_node.html</a>
% Christian	Number of Christians per population, 2011	Zensus Datenbank ( <i>Census Results</i> )	<a href="https://ergebnisse2011.zensus2022.de/datebank/online/">https://ergebnisse2011.zensus2022.de/datebank/online/</a>
% AfD 2014	Vote share Alternative für Deutschland (AfD), European elections, 2014, demeaned by state	Der Bundeswahlleiter ( <i>Federal Returning Officer</i> )	<a href="https://www.bundeswahlleiter.de/europawahlen/2014/ergebnisse.html">https://www.bundeswahlleiter.de/europawahlen/2014/ergebnisse.html</a>
% Voted 2014	Log turnout, European elections, 2014	Der Bundeswahlleiter ( <i>Federal Returning Officer</i> )	<a href="https://www.bundeswahlleiter.de/europawahlen/2014/ergebnisse.html">https://www.bundeswahlleiter.de/europawahlen/2014/ergebnisse.html</a>

% Syrians 2010	Number of Syrians divided by population, 2010	German Statistical Office	<a href="https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en">https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en</a>
% Syrians 2019	Number of Syrians divided by population, 2019	German Statistical Office	<a href="https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en">https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en</a>
% Foreign 2010	Number of foreigners divided by population, 2010	German Statistical Office	<a href="https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en">https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en</a>
% Foreign 2019	Number of foreigners divided by population, 2019	German Statistical Office	<a href="https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en">https://www-genesis.destatis.de/genesis/online?operation=sprachwechsel&amp;language=en</a>
Integr. Courses per Syrian	Number of integration courses completed 2015-2019 per Syrian	Federal Office for Migration and Refugees	<a href="https://www.bamf.de/DE/Themen/Statistik/integrationskurszahlen/integrationskurszahlen-node.html">https://www.bamf.de/DE/Themen/Statistik/integrationskurszahlen/integrationskurszahlen-node.html</a>
Pro-Immigr. Groups per Syrian	Number of groups affiliated with ProAsyl activist group per Syrian	ProAsyl	not publicly available, data received directly from organisation
Integr. Sports Clubs per Syrian	Number of sports clubs that are part of Integration through Sport initiative	German Olympic Sports Confederation	<a href="https://integration.dosb.de">https://integration.dosb.de</a>
Unemp. General Schools Teachers per Pop. 2014	Number of unemployed general school teachers divided by population, 2014	Bundesagentur für Arbeit (Federal Employment Agency)	not publicly available, data received directly from organisation
Unemp. Higher Ed. School Teachers per Pop. 2014	Number of unemployed university and research institute teachers divided by population, 2014	Bundesagentur für Arbeit (Federal Employment Agency)	not publicly available, data received directly from organisation
Unemp. Driving and Sports Teachers per Pop. 2014	Number of driving and sports teachers divided by population, 2014	Bundesagentur für Arbeit (Federal Employment Agency)	not publicly available, data received directly from organisation
Unemp. Other School Teachers per Pop. 2014	Number of teachers in other education centers divided by population, 2014	Bundesagentur für Arbeit (Federal Employment Agency)	not publicly available, data received directly from organisation

---

## K High School Matching Procedure

We assign users to high schools using a three-step process. On Facebook, users can provide the high school that they attended in their profile. Some of these high schools (such as "Hogwarts" and "the School of Hard Knocks") are obviously incorrect, so we begin by filtering out such schools. We are left with a list of plausible high school names, which we then need to disambiguate, since many high schools share the same name. For this, we use a listing of high schools from the websites of German state governments (see [DatenSchule Project](#).) For each user in our sample, we are able to observe the counties in which they lived during high school age. We use this information and their self-reported high school name to match them to a high school in the administrative data. To do this, we make use of a fuzzy string matching algorithm, applied to the list of high schools that are in the regions in which they lived between the ages of 13 and 18.<sup>28</sup> Using this methodology, we are able to match 1.2 million of the 2.2 million users to high schools from the administrative data.

In the second step, we consider the users who report a high school that we are unable to find in the administrative data. In some cases, simple misspellings or inconsistencies in the school's name prevent a match from being formed between the two data sets. In other cases, these discrepancies are due to variations in states' criteria for including schools in the lists provided on their websites (e.g., states differ in their inclusion of vocational high schools in the lists we use). For this reason, we create a listing of school names that are reported by 50 or more users in a single county, but which are not included in the administrative data. We allow users to be assigned to these well-attested schools as we would any other. We call these schools the "non-canonical schools", and include them in all regressions, though our results are robust to excluding them. This process adds another 81 thousand users to our sample. For users who attend a school which we cannot find in the administrative data, and which appears in the self-reported data fewer than 50 times in the same county, we discard their self-reported school.

Finally, for users without a validated self-reported high school, we attempt to impute the school they attended using information on their social network. Intuitively, this approach takes advantage of the fact that most users will attend the same school as their friends who live in the same area and are the same age. To do this, we find the modal high school among a user's friends in the county they live in (as well as counties bordering it) and who are no more than 3 years different in age from the user. If this modal high school is attended by at least 10 friends, and there are at least 5 times as many friends attending this high school as the next most common school, we assign the user to this high school. We repeat this process 10 times, adding 137 thousand more users to our sample.<sup>29</sup>

We are able to assign 63% of native users to high schools using this methodology. In the cohorts we use for our regression, the median cohort has 31 students, with an inter-quartile range of 15 to 52 students. The match rate is lower (24%) for Syrian migrant students, since they have relatively few local friends and are less likely to list a high school on their profile. Any mistakes we make in assigning Syrians to high schools are likely to bias our analyses away from finding an effect of exposure.

---

<sup>28</sup>If we are unable to find a high school that matches in one of the regions that they lived in, we consider the regions that neighbor the regions the user lived in.

<sup>29</sup>To get a sense for the predictive power of the above imputation methodology, we can examine how accurate it is in determining the high school attended by users who self-report the school they attended. The imputation method is able to assign a school to 25% of such users, agreeing with the self-reported school in more than 90% of cases.