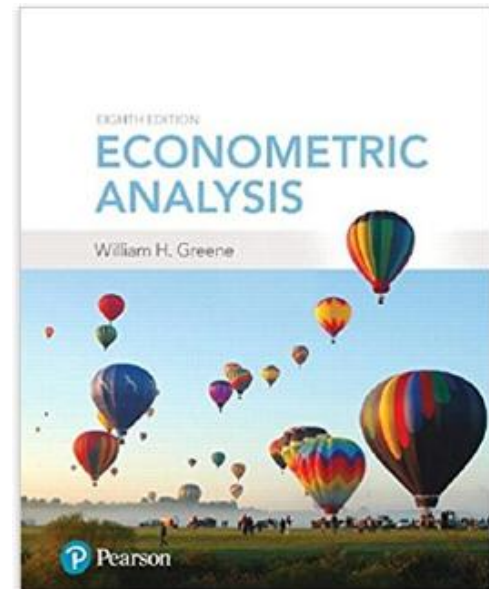


Econometrics I

Professor William Greene
Stern School of Business
Department of Economics



Econometrics I

Part 10 – Interval Estimation and Prediction

Interval Estimation

- **b** = point estimator of β
- We acknowledge the sampling variability.
 - Estimated sampling variance

logG	Coefficient	Standard Error	t	Prob. t >T*	95% Confidence Interval	
Constant	-23.9398*	12.19453	-1.96	.0593	-48.8804	1.0008
YEAR	.00902	.00677	1.33	.1931	-.00483	.02287
logPG	-.47878***	.06162	-7.77	.0000	-.60482	-.35275
logY	1.27365***	.15010	8.49	.0000	.96666	1.58064
logPN	1.13036***	.18906	5.98	.0000	.74369	1.51703
logPD	.79402***	.22581	3.52	.0015	.33217	1.25586
logPS	-1.19347***	.20592	-5.80	.0000	-1.61462	-.77233

Note: ***, **, * ==> Significance at 1%, 5%, 10% level.

- Point estimate is only the best single guess
- Form an interval, or range of plausible values
- Plausible \Leftrightarrow likely values with acceptable degree of probability.
- To assign probabilities, we require a distribution for the variation of the estimator.
- The role of the normality assumption for ε

Robust Inference: Confidence Interval for β_k

b_k = the point estimate

Std.Err[b_k] = $\text{sqr}\{\text{Estimated Asymptotic Variance of } b_k\} = v_k$

The matrix may be any robust or appropriate covariance matrix.

- $b_k \sim \text{Asy.N}[\beta_k, v_k^2]$
- $(b_k - \beta_k)/v_k \sim \text{Asy.N}[0, 1]$

Consider a range of plausible values of β_k given the point estimate b_k . $b_k \pm$ sampling error.

- Measured in standard error units,
- $|(b_k - \beta_k)/v_k| < z^*$
- Larger $z^* \rightarrow$ greater probability (“confidence”)
- Given normality, e.g., $z^* = 1.96 \rightarrow 95\%$, $z^* = 1.645 \rightarrow 90\%$
- Plausible range for β_k then is $b_k \pm z^* v_k$

Critical Values for the Confidence Interval

Assume normality of $\boldsymbol{\varepsilon}$:

- $b_k \sim N[\beta_k, v_k^2]$ for the true β_k .

- $(b_k - \beta_k)/v_k \sim N[0, 1]$

$v_k = [s^2(\mathbf{X}'\mathbf{X})^{-1}]_{kk}$ $(b_k - \beta_k)/\text{Est.}(v_k) \sim t[n-K]$.

Use critical values from $t[n-K]$ distribution instead of standard normal. Will be the same as normal if $n \geq 100$.

Based on asymptotic results or any robust covariance matrix estimator, use $N[0, 1]$.

Confidence Interval

Ordinary LHS=logG	least squares regression				
	Mean	=	5.39299		
	Standard deviation	=	.24878		
	No. of observations	=	36		
Regression	Sum of Squares	=	2.15567	DegFreedom	6
Residual	Sum of Squares	=	.105227E-01	Mean square	.35928
Total	Sum of Squares	=	2.16619		29
	Standard error of e	=	.01905		35
Fit	R-squared	=	.99514	Root MSE	.01710
Model test	F[6, 29]	=	990.14814	R-bar squared	.99414
				Prob F > F*	.00000

	logG	Coefficient	Standard Error	t	Prob. t >T*	95% Confidence Interval	
Constant		-23.9398*	12.19453	-1.96	.0593	-48.8804	1.0008
YEAR		.00902	.00677	1.33	.1931	-.00483	.02287
logPG		-4.7878***	.06162	-7.77	.0000	-6.0482	-.35275
logY		1.27365***	.15010	8.49	.0000	.96666	1.58064
logPN		1.13036***	.18906	5.98	.0000	.74369	1.51703
logPD		.79402***	.22581	3.52	.0015	.33217	1.25586
logPS		-1.19347***	.20592	-5.80	.0000	-1.61462	-.77233

Note: ***, **, * ==> Significance at 1%, 5%, 10% level.

Critical $t_{[.975,29]} = 2.045$

Confidence interval based on t: $1.27365 \pm 2.045 * .1501$

Confidence interval based on normal: $1.27365 \pm 1.960 * .1501$

Specification and Functional Form: Interaction Effects

Population

$$y = \beta_1 + \beta_2 x + \beta_3 z + \beta_4 xz + \varepsilon$$

$$\delta_x = \frac{\partial E[y | x, z]}{\partial x} = \beta_2 + \beta_4 z$$

Estimators

$$\hat{y} = b_1 + b_2 x + b_3 z + b_4 xz$$

$$\hat{\delta}_x = b_2 + b_4 z$$

Estimator of the variance of $\hat{\delta}_x$

$$\text{Est.Var}[\hat{\delta}_x] = \text{Var}[b_2] + z^2 \text{Var}[b_4] + 2z \text{Cov}[b_2, b_4]$$

Interaction Effect

```

-----
Ordinary least squares regression .....
LHS=LOGY Mean = -1.15746
          Standard deviation = .49149
          Number of observs. = 27322
Model size Parameters = 4
          Degrees of freedom = 27318
Residuals Sum of squares = 6540.45988
          Standard error of e = .48931
Fit R-squared = .00896
     Adjusted R-squared = .00885
Model test F[ 3, 27318] (prob) = 82.4(.0000)
  
```

```

-----+-----
Variable| Coefficient      Standard Error  b/St.Er.  P[|Z|>z]  Mean of X
-----+-----
Constant| -1.22592***      .01605        -76.376   .0000
      AGE| .00227***        .00036         6.240    .0000      43.5272
      FEMALE| .21239***        .02363         8.987    .0000      .47881
      AGE_FEM| -.00620***       .00052        -11.819   .0000      21.2960
  
```

Do women earn more than men (in this sample?) The $+.21239$ coefficient on FEMALE would suggest so. But, the female "difference" is $+.21239 - .00620 \cdot \text{Age}$. At average Age, the effect is $.21239 - .00620(43.5272) = -.05748$.

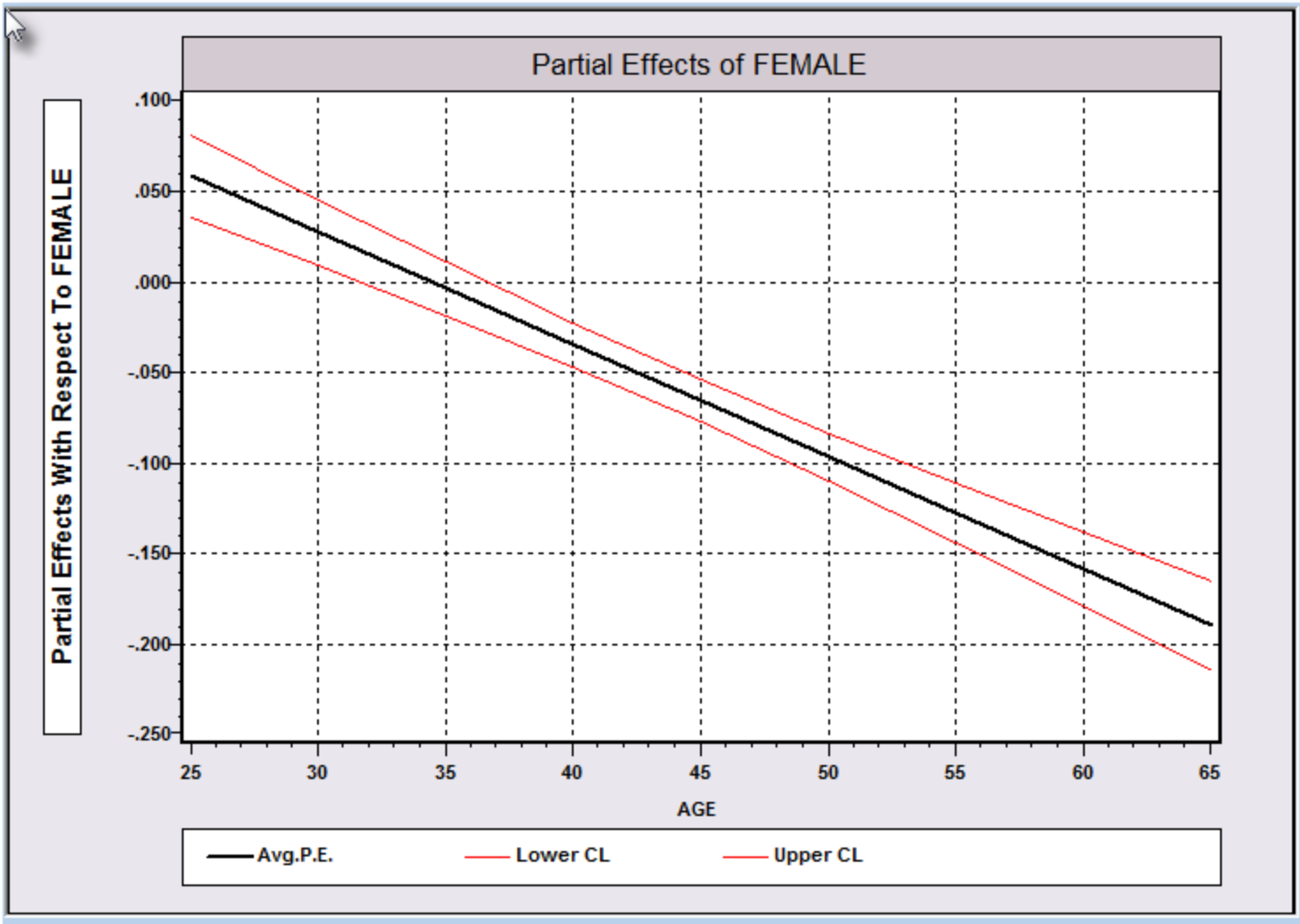
 Partial Effects Analysis for Linear Regression Function

Effects on function with respect to FEMALE

Results are computed by average over sample observations

Partial effects for binary var FEMALE computed by first difference

df/dFEMALE (Delta method)	Partial Effect	Standard Error	t	95% Confidence Interval	
APE. Function	-.05750	.00595	9.67	-.06916	-.04585
AGE = 25.00	.05737	.01143	5.02	.03498	.07977
AGE = 30.00	.02637	.00929	2.84	.00817	.04457
AGE = 35.00	-.00463	.00746	.62	-.01926	.01000
AGE = 40.00	-.03563	.00624	5.71	-.04786	-.02341
AGE = 45.00	-.06664	.00599	11.12	-.07838	-.05490
AGE = 50.00	-.09764	.00683	14.30	-.11102	-.08426
AGE = 55.00	-.12864	.00843	15.26	-.14517	-.11211
AGE = 60.00	-.15964	.01046	15.27	-.18014	-.13915
AGE = 65.00	-.19065	.01270	15.01	-.21555	-.16575



Bootstrap Confidence Interval For a Coefficient

```
regr;lhs=logg;rhs=logi,one,logpg,logpnc,logpuc,logppt$
calc;bi = b(1) ; bbar=0 $
matrix ; bir=init(1000,1,0.0)$
proc$
draw      ; n=52 ; replacement $
regress   ; quietly
           ; lhs=logg;rhs=logi,one,logpg,logpnc,logpuc,logppt$
matrix    ; bir(i)=b(1) $
calc      ; bbar=bbar + 1/1000 *b(1) $
endproc $
exec;i=1,1000$
matrix ; ones = init(1000,1,1.0)
       ; brep = bi*ones + bir-bbar*ones $
quantiles ; rhs = brep $
```

Bootstrap CI for Least Squares

Ordinary least squares regression					
LHS=LOGG	Mean	=	-9.94245		
	Standard deviation	=	.23881		
	No. of observations	=	52	DegFreedom	Mean square
Regression	Sum of Squares	=	2.79360	5	.55872
Residual	Sum of Squares	=	.114979	46	.00250
Total	Sum of Squares	=	2.90858	51	.05703
	Standard error of e	=	.05000	Root MSE	.04702
Fit	R-squared	=	.96047	R-bar squared	.95617
Model test	F[5, 46]	=	223.52968	Prob F > F*	.00000

LOGG	Coefficient	Standard Error	t	Prob. t >T*	95% Confidence Interval	
LOGI	1.28609***	.14566	8.83	.0000	.99289	1.57930
Constant	-20.9548***	1.52568	-13.73	.0000	-24.0259	-17.8838
LOGPG	-.02797	.04338	-.64	.5222	-.11529	.05934
LOGPNC	-.15576	.21002	-.74	.4621	-.57852	.26699
LOGPUC	.02850	.10199	.28	.7811	-.17680	.23381
LOGPPT	-.18283	.11908	-1.54	.1315	-.42252	.05686

Percentiles		BREP:1
Sample size		1000
Min.		.77187
01th		.94799
*025		1.01719
05th		1.06692
10th		1.12893
20th		1.19682
25th		1.21584
30th		1.23299
40th		1.25748
Med.		1.28525
60th		1.31333
70th		1.34672
75th		1.36097
80th		1.37399
90th		1.43407
95th		1.49031
*975		1.55168
99th		1.64709
Max.		1.96372

Bootstrap CI for Least Absolute Deviations

```

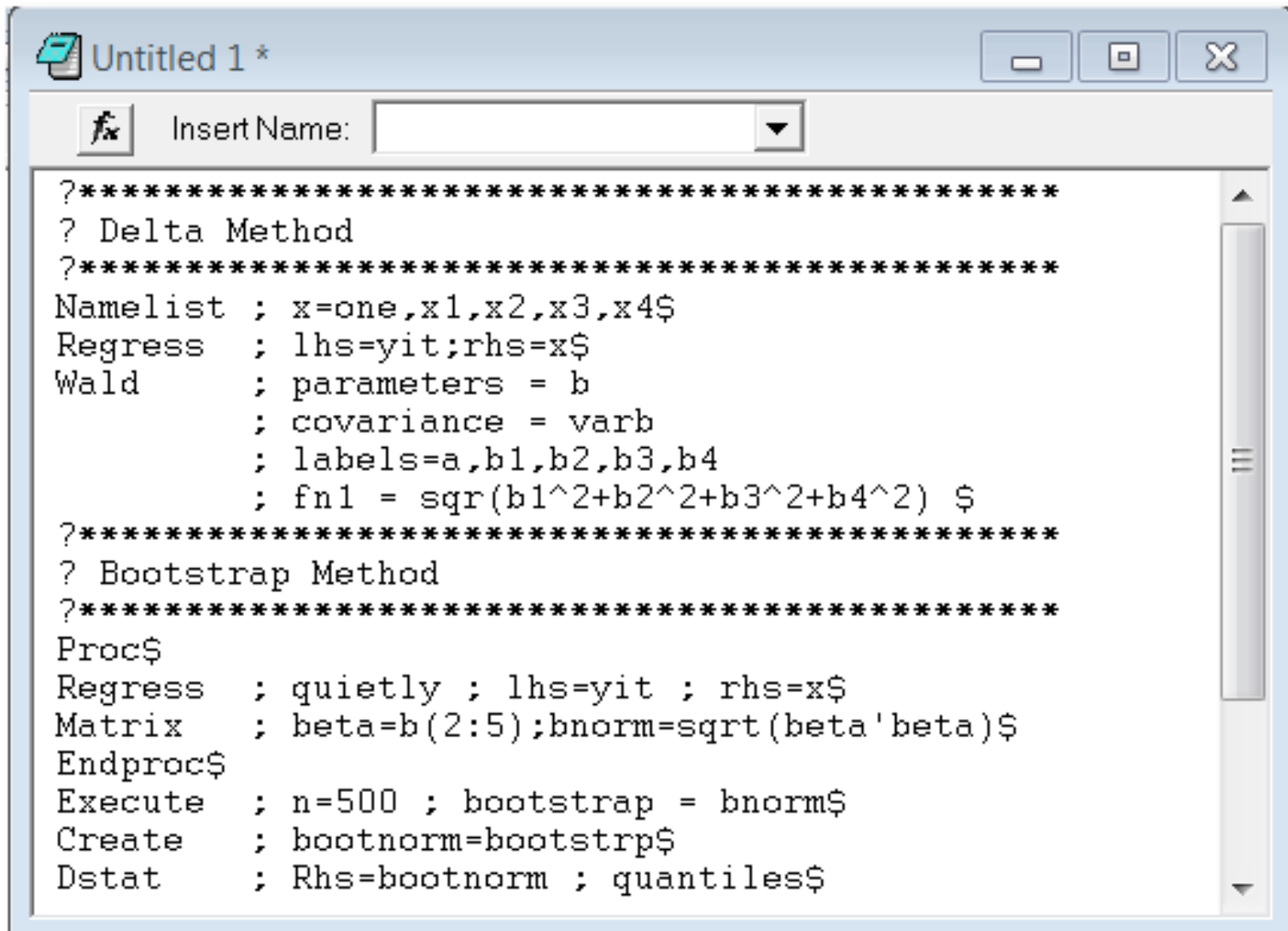
Least absolute deviations estimator.....
Nonlinear least squares regression .....
LHS=LOGG Mean = -9.95023
Standard deviation = .24202
Number of observs. = 52
Model size Parameters = 6
Degrees of freedom = 46
Residuals Sum of squares = .119690
Standard error of e = .05101
Fit R-squared = .95993
Adjusted R-squared = .95558
Model test F[ 5, 46] (prob) = 220.4(.0000)
Not using OLS or no constant. Rsqrd & F may be < 0
Sum of absolute deviations = 1.6665854
    
```

LOGG	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
Covariance matrix based on 50 replications.						
LOGI	1.20636***	.16869	7.15	.0000	.87573	1.53698
Constant	-19.6510***	1.75404	-11.20	.0000	-23.0889	-16.2132
LOGPG	.04376	.03901	1.12	.2619	-.03269	.12022
LOGPNC	-.50766	.34833	-1.46	.1450	-1.19038	.17505
LOGPUC	.04250	.17786	.24	.8111	-.30610	.39110
LOGPPT	-.01514	.16300	-.09	.9260	-.33461	.30433

Percentiles	BREP:1
Sample size	1000
Min.	.11903
01th	.53445
*025	.72084
05th	.83721
10th	.93752
20th	1.04461
25th	1.08954
30th	1.11451
40th	1.16544
Med.	1.22374
60th	1.26602
70th	1.31469
75th	1.34340
80th	1.37143
90th	1.45886
95th	1.54635
*975	1.60223
99th	1.70422
Max.	2.16943

Bootstrapped Confidence Intervals

$$\text{Estimate Norm}(\beta) = (\beta_1^2 + \beta_2^2 + \beta_3^2 + \beta_4^2)^{1/2}$$



```
Untitled 1 *
Insert Name:
?*****
? Delta Method
?*****
Namelist ; x=one,x1,x2,x3,x4$
Regress ; lhs=yit;rhs=x$
Wald ; parameters = b
; covariance = varb
; labels=a,b1,b2,b3,b4
; fn1 = sqr(b1^2+b2^2+b3^2+b4^2) $
?*****
? Bootstrap Method
?*****
Proc$
Regress ; quietly ; lhs=yit ; rhs=x$
Matrix ; beta=b(2:5);bnorm=sqrt(beta'beta)$
Endproc$
Execute ; n=500 ; bootstrap = bnorm$
Create ; bootnorm=bootstrp$
Dstat ; Rhs=bootnorm ; quantiles$
```

Variable	Mean	Standard Deviation	Minimum	Maximum	Cases	Missing Values
BOOTNORM	.748701	.012426	.715269	.791033	500	982

Descriptive Statistics for 1 variables
 DSTAT results are matrix LASTDSTA in current project.

Percentiles BOOTNORM
 Sample size 500

Min.	.715269
01th	.722066
*025	.725308
05th	.727917
10th	.732863
20th	.738778
25th	.740387
30th	.742011
40th	.745226
Med.	.748275
60th	.751571
70th	.754829
75th	.757144
80th	.75874
90th	.765108
95th	.770278
*975	.77354
99th	.777276
Max.	.791033

WALD procedure. Estimates and standard errors for nonlinear functions and joint test of nonlinear restrictions.

Wald Statistic = 4841.03525
 Prob. from Chi-squared[1] = .00000
 Functions are computed at means of variables

WaldFcns	Function	Standard Error	z	Prob. z >Z*	95% Confidence Interval
Fncn(1)	.74792***	.01075	69.58	.0000	.72685 .76899

HEALTH ECONOMICS

Health Econ. **19**: 1063–1074 (2010)

Published online in Wiley InterScience (www.interscience.wiley.com).10.1002/hec.1647

QUANTILE REGRESSION ANALYSIS OF THE RATIONAL ADDICTION MODEL: INVESTIGATING HETEROGENEITY IN FORWARD-LOOKING BEHAVIOR

AUDREY LAPORTE^{a,*}, ALFIA KARIMOVA^b and BRIAN FERGUSON^c

^a*Department of Health Policy, Management and Evaluation, University of Toronto, Toronto, Ont., Canada*

^b*Department of Economics, University of Toronto, Toronto, Ont., Canada*

^c*Department of Economics, University of Guelph, Guelph, Ont., Canada*

SUMMARY

The time path of consumption from a rational addiction (RA) model contains information about an individual's tendency to be forward looking. In this paper, we use quantile regression (QR) techniques to investigate whether the tendency to be forward looking varies systematically with the level of consumption of cigarettes. Using panel data,

Coefficient on MALE dummy variable in quantile regressions

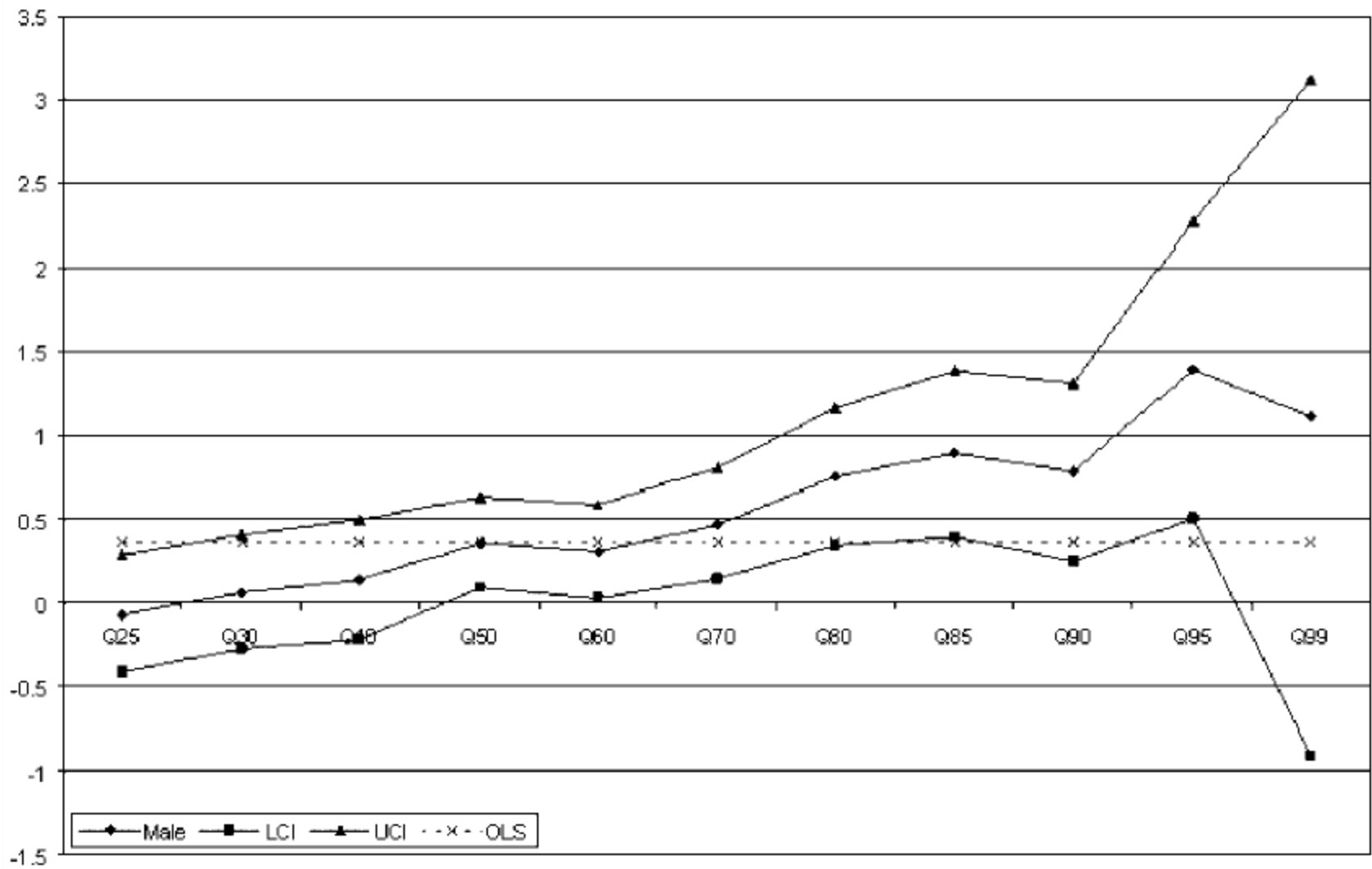


Figure 4. Male

Forecasting

- **Objective: Forecast**
- Distinction: Ex post vs. Ex ante forecasting
 - Ex post: RHS data are observed
 - Ex ante: RHS data must be forecasted
- Prediction vs. model validation.
 - Within sample prediction
 - “Hold out sample”

Prediction Intervals

Given \mathbf{x}^0 predict y^0 .

Two cases:

Estimate $E[y|\mathbf{x}^0] = \boldsymbol{\beta}'\mathbf{x}^0$;

Predict $y^0 = \boldsymbol{\beta}'\mathbf{x}^0 + \varepsilon^0$

Obvious predictor, $\mathbf{b}'\mathbf{x}^0$ + estimate of ε^0 . Forecast ε^0 as 0, but allow for variance.

Alternative: When we predict y^0 with $\mathbf{b}'\mathbf{x}^0$, what is the 'forecast error?'

Est. $y^0 - y^0 = \mathbf{b}'\mathbf{x}^0 - \boldsymbol{\beta}'\mathbf{x}^0 - \varepsilon^0$, so the variance of the forecast error is
 $\mathbf{x}^{0'}\text{Var}[\mathbf{b} - \boldsymbol{\beta}]\mathbf{x}^0 + \sigma^2$

How do we estimate this? Form a confidence interval. Two cases:

If \mathbf{x}^0 is a vector of constants, the variance is just $\mathbf{x}^{0'}\text{Var}[\mathbf{b}]\mathbf{x}^0$. Form confidence interval as usual.

If \mathbf{x}^0 had to be estimated, then we use a random variable. What is the variance of the product? (Ouch!) One possibility: Use bootstrapping.

Forecast Variance

Variance of the forecast error is

$$\sigma^2 + \mathbf{x}^0' \text{Var}[\mathbf{b}]\mathbf{x}^0 = \sigma^2 + \sigma^2[\mathbf{x}^0'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}^0]$$

If the model contains a constant term, this is

$$\text{Var}[e^0] = \sigma^2 \left[1 + \frac{1}{n} + \sum_{j=1}^{K-1} \sum_{k=1}^{K-1} (x_j^0 - \bar{x}_j)(x_k^0 - \bar{x}_k)(\mathbf{Z}'\mathbf{M}^0\mathbf{Z})^{jk} \right]$$

In terms squares and cross products of deviations from means. Interpretation: Forecast variance is smallest in the middle of our “experience” and increases as we move outside it.

Butterfly Effect

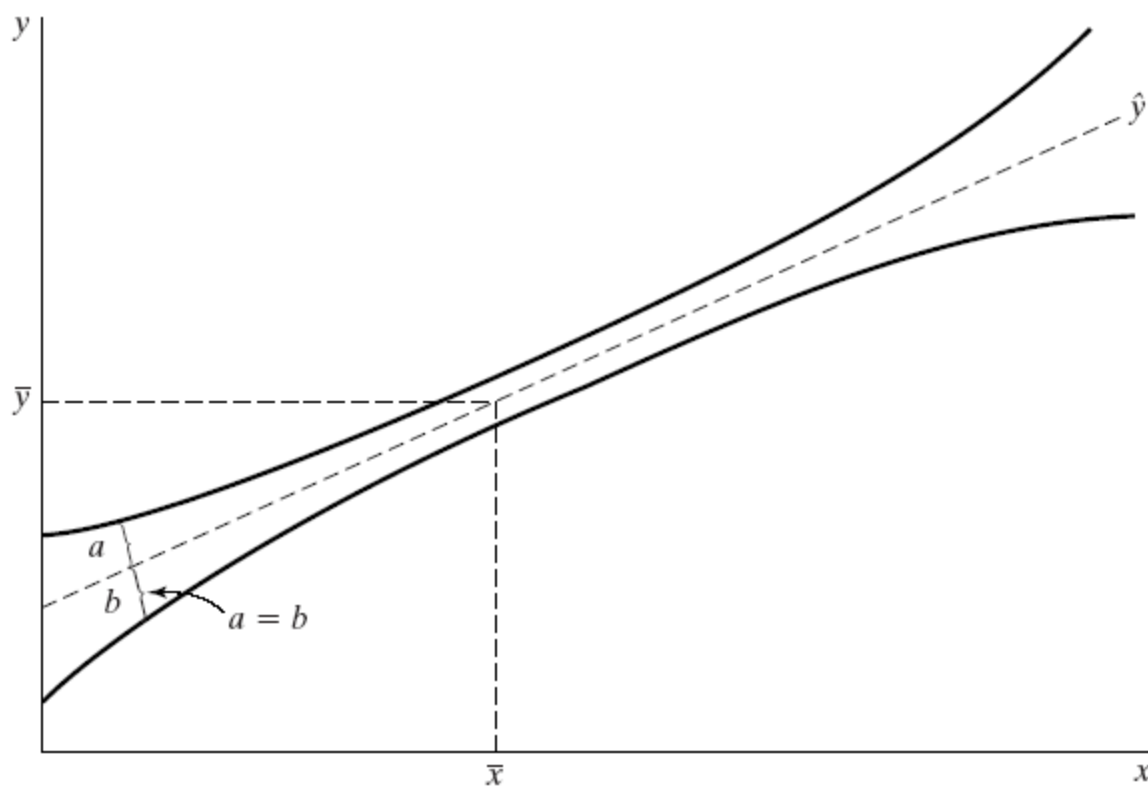
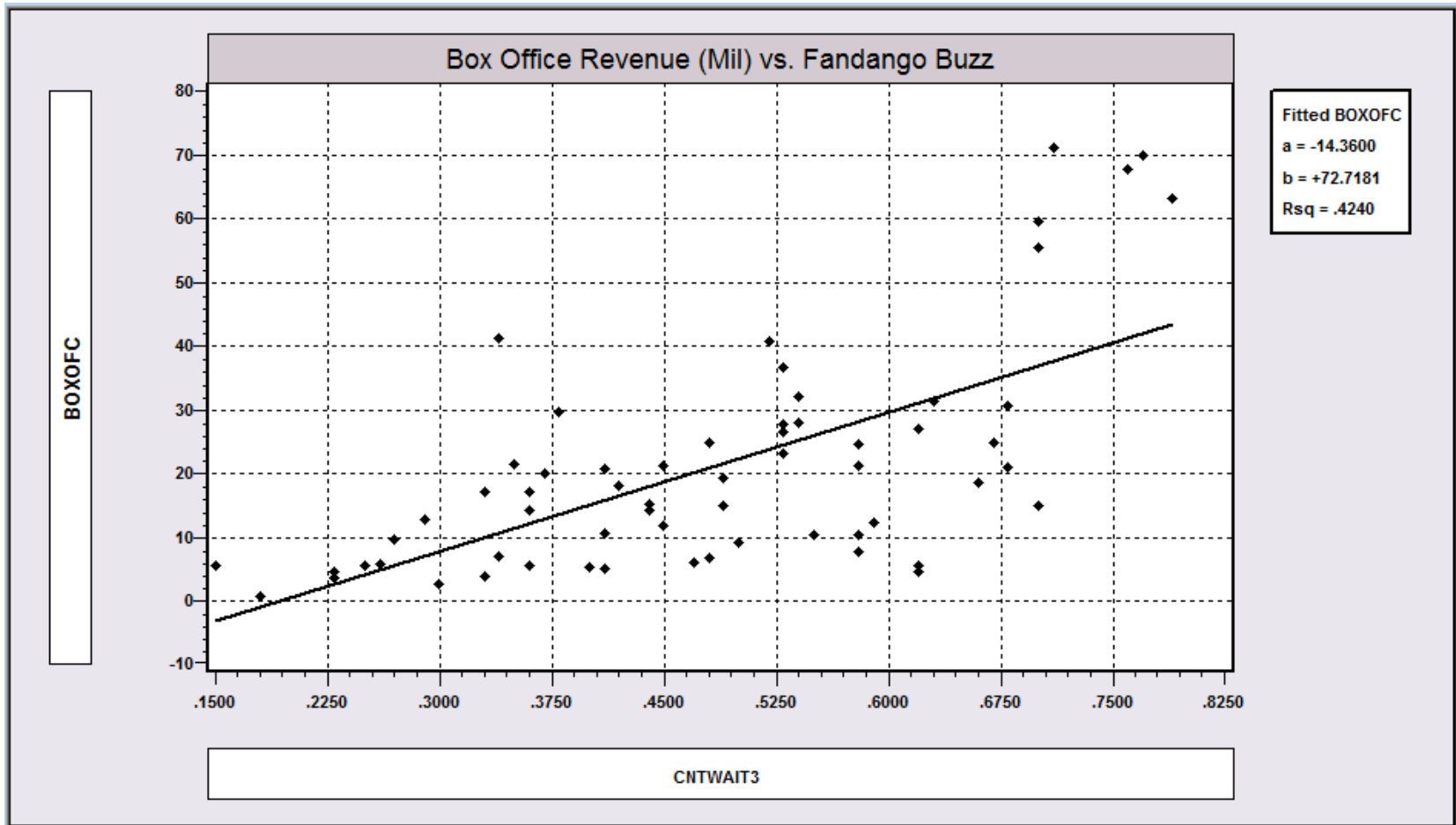


FIGURE 6.1 Prediction Intervals.

Internet Buzz Data



A Prediction Interval

Prediction includes a range of uncertainty

Point estimate: $\hat{y} = a + bx^*$

The range of uncertainty around the prediction:

$$a + bx^* \pm 1.96 \sqrt{S_e^2 \left(1 + \frac{1}{n} + \frac{(x^* - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \right)}$$

The usual 95%

Due to ε

Due to estimating α and β with a and b

Slightly Simpler Formula for Prediction

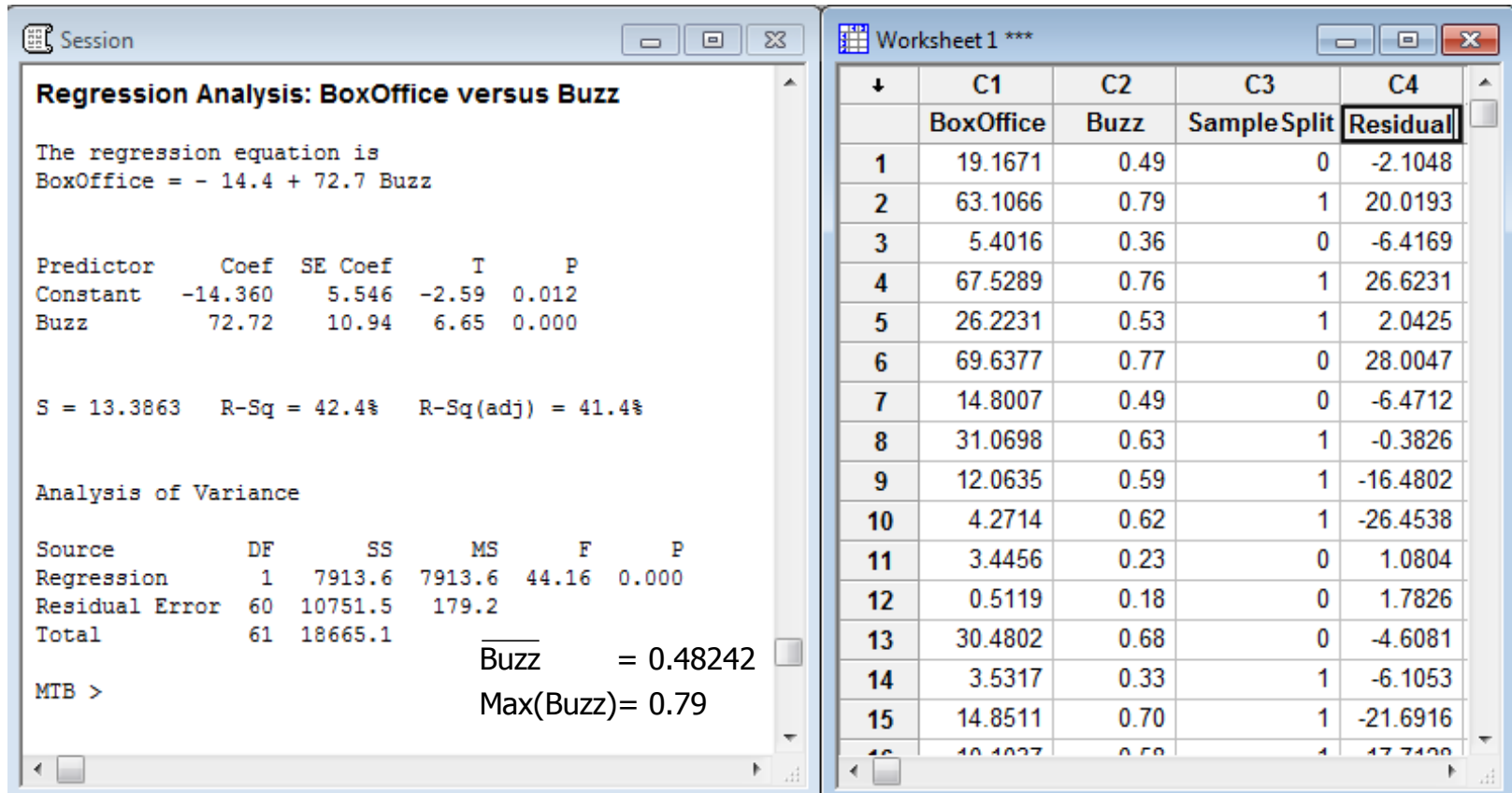
Prediction includes a range of uncertainty

Point estimate: $\hat{y} = a + bx^*$

The range of uncertainty around the prediction:

$$a + bx^* \pm 1.96 \sqrt{S_e^2 \left(1 + \frac{1}{n}\right) + (x^* - \bar{x})^2 (\text{SE}(b))^2}$$

Prediction from Internet Buzz Regression



Prediction Interval for Buzz = .8

Predict Box Office for Buzz = .8

$$a+bx = -14.36 + 72.72(.8) = 43.82$$

$$\sqrt{s_e^2 \left(1 + \frac{1}{N}\right) + (.8 - \overline{\text{Buzz}})^2 \text{SE}(b)^2}$$

$$= \sqrt{13.3863^2 \left(1 + \frac{1}{62}\right) + (.8 - .48242)^2 10.94^2}$$

$$= 13.93$$

$$\text{Interval} = 43.82 \pm 1.96(13.93)$$

$$= 16.52 \text{ to } 71.12$$

Semi- and Nonparametric Estimation

Application: Stochastic Frontier Model

Production Function Regression: $\log Y = b'x + v - u$

where u is “inefficiency.” $u > 0$. v is normally distributed.

Save for the constant term, the model is consistently estimated by OLS.

If the theory is right, the OLS residuals will be skewed to the left, rather than symmetrically distributed if they were normally distributed.

Application: Spanish dairy data used in Assignment 2

y_{it} = log of milk production

x_1 = log cows, x_2 = log land, x_3 = log feed, x_4 = log labor

Regression Results

```

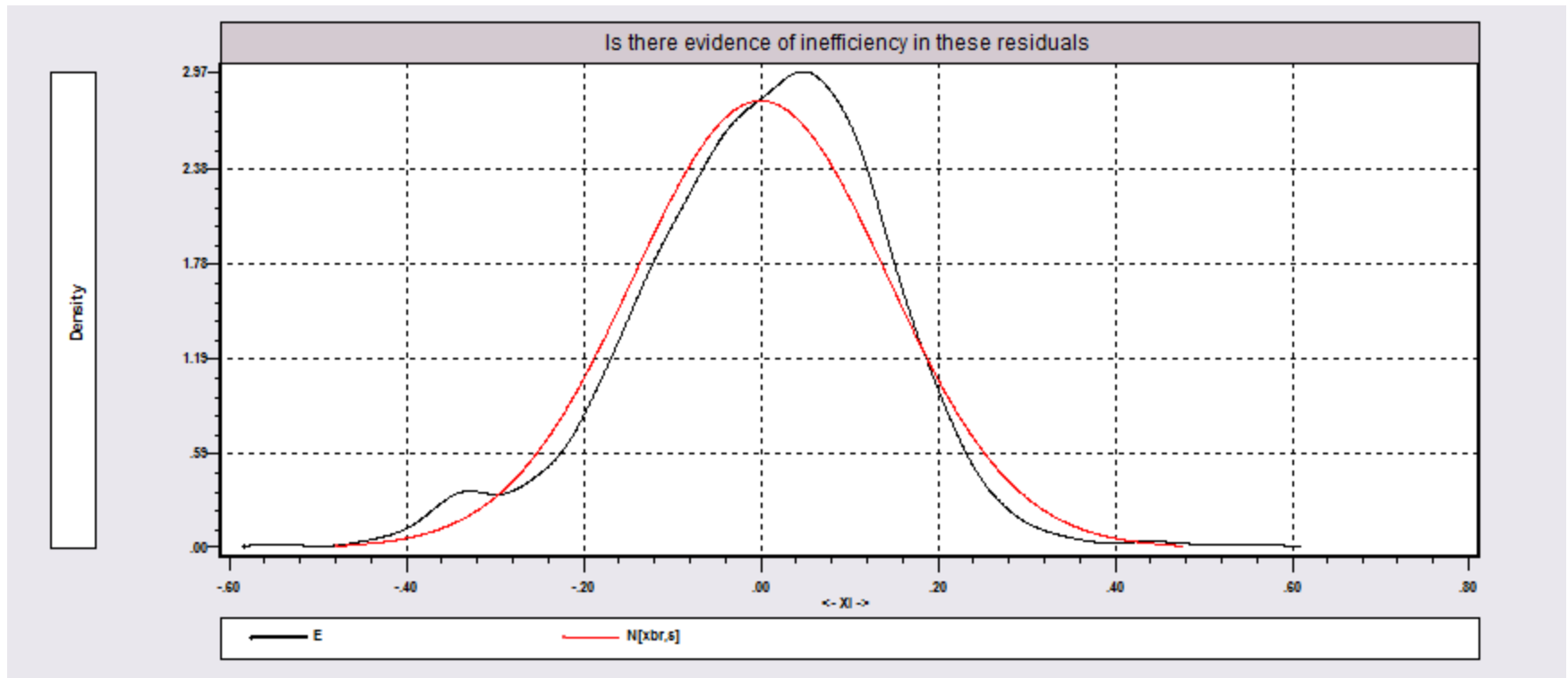
-----
Ordinary least squares regression
LHS=YIT Mean = 11.57749
Standard deviation = .64344
No. of observations = 1482 Degrees of freedom
Regression Sum of Squares = 584.056 4
Residual Sum of Squares = 29.0957 1477
Total Sum of Squares = 613.152 1481
Standard error of e = .14035
Fit R-squared = .95255 R-bar squared = .95242
Model test F[ 4, 1477] = 7412.18529 Prob F > F* = .00000
Diagnostic Log likelihood = 809.67609 Akaike I.C. = -3.92381
Restricted (b=0) = -1448.90834 Bayes I.C. = -3.90592
Chi squared [ 4] = 4517.16885 Prob C2 > C2* = .00000
Model was estimated on Aug 29, 2011 at 09:44:07 AM
-----

```

YIT	Coefficient	Standard Error	z	Prob. z > Z*	95% Confidence Interval	
Constant	11.5775***	.00365	3175.52	.0000	11.5703	11.5846
X1	.59518***	.01958	30.39	.0000	.55679	.63356
X2	.02305**	.01122	2.05	.0400	.00105	.04505
X3	.02319*	.01303	1.78	.0751	-.00235	.04873
X4	.45176***	.01078	41.89	.0000	.43062	.47290

Note: ***, **, * ==> Significance at 1%, 5%, 10% level.

Distribution of OLS Residuals



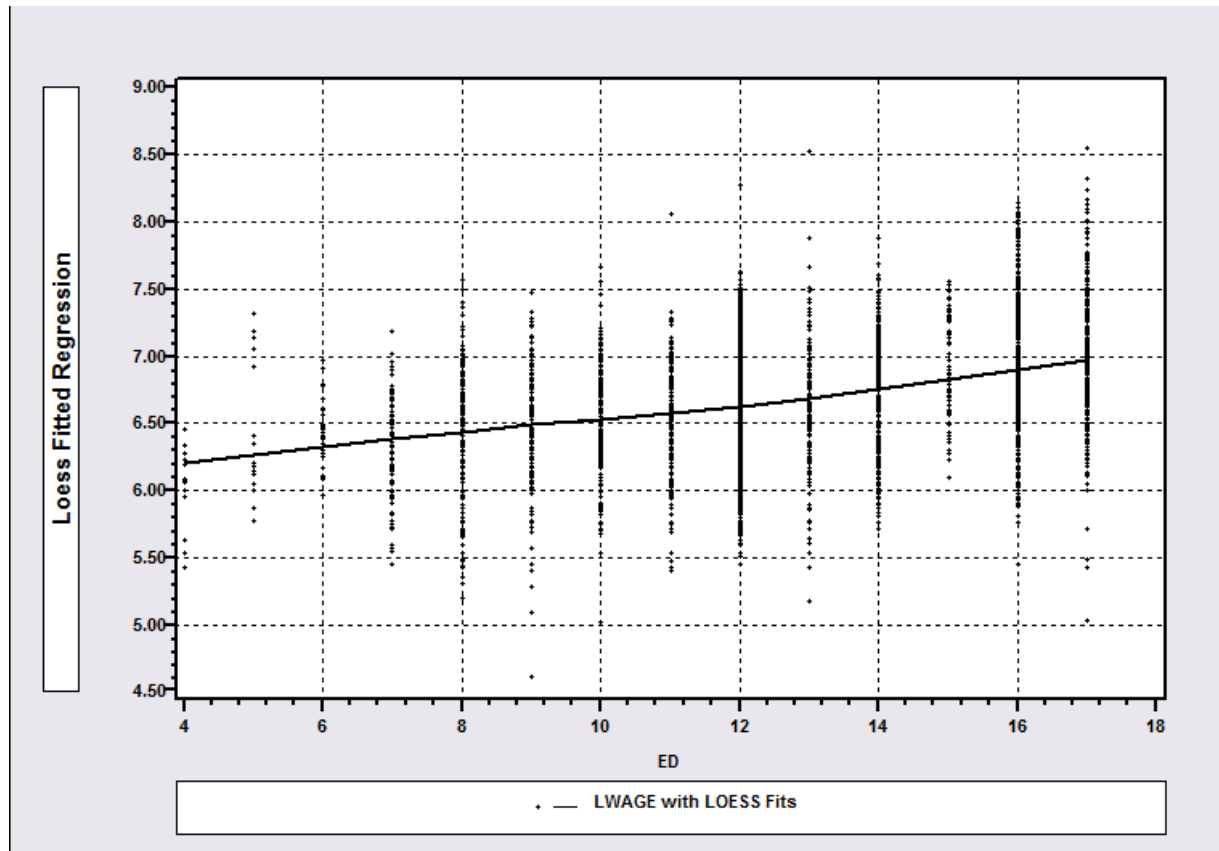
A Nonparametric Regression

- $y = \mu(x) + \varepsilon$
- Smoothing methods to approximate $\mu(x)$ at specific points, x^*
- For a particular x^* , $\mu(x^*) = \sum_i w_i(x^*|\mathbf{x})y_i$
 - E.g., for ols, $\mu(x^*) = a + bx^*$
 - $w_i = 1/n + (x_i - \bar{x}) / \sum_i (x_i - \bar{x})^2$
- We look for weighting scheme, local differences in relationship. OLS assumes a fixed slope, b .

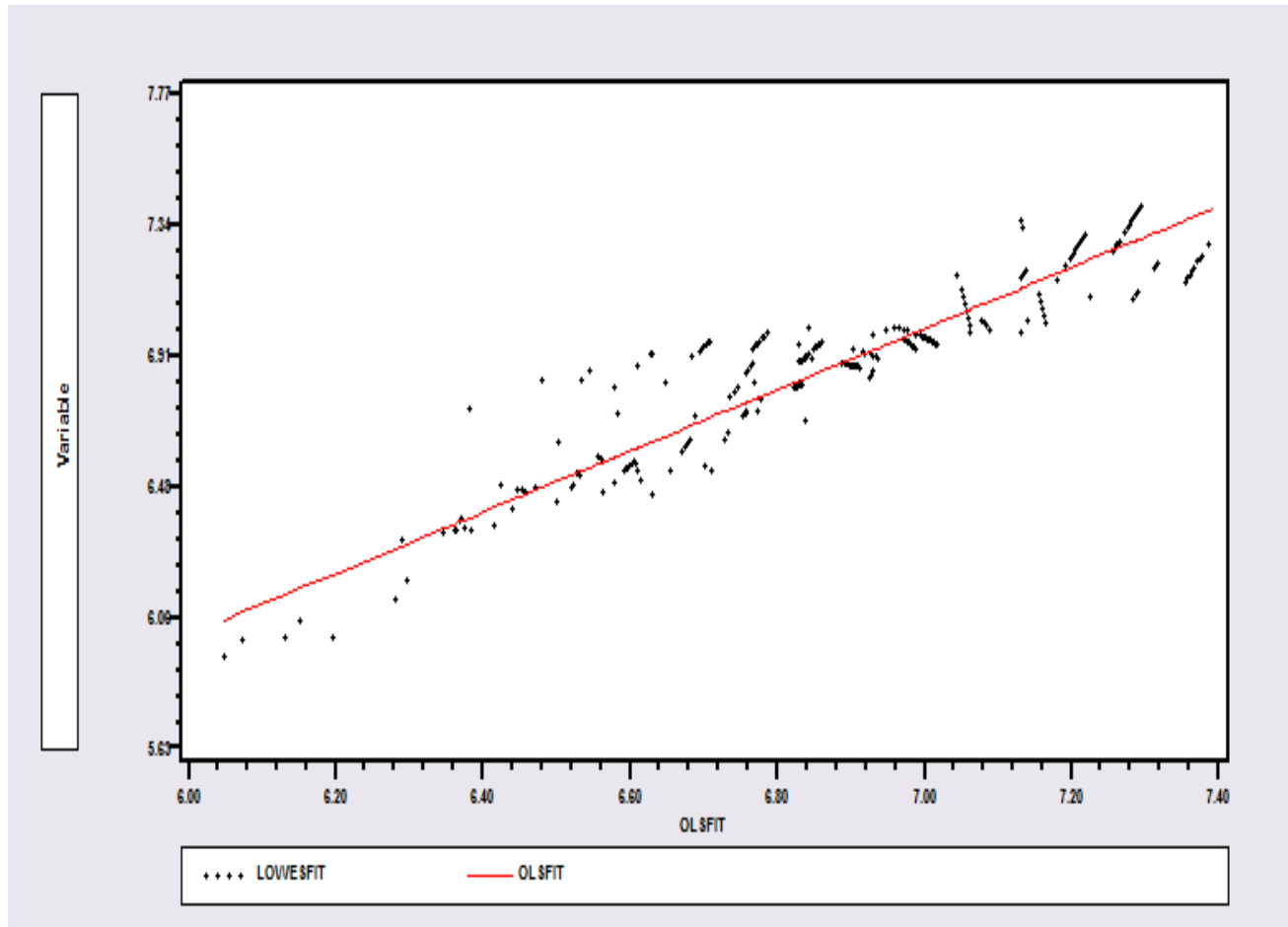
Nearest Neighbor Approach

- Define a neighborhood of x^* . Points near get high weight, points far away get a small or zero weight
- Bandwidth, h defines the neighborhood:
e.g., Silverman $h = .9 \text{Min}[s, (\text{IQR}/1.349)]/n^{.2}$
Neighborhood is $+ \text{ or } - h/2$
- LOWESS weighting function: (tricube)
 $T_i = [1 - [\text{Abs}(x_i - x^*)/h]^3]^3$.
- Weight is $w_i = 1[\text{Abs}(x_i - x^*)/h < .5] * T_i$.

LOWESS Regression



OLS Vs. Lowess



Smooth Function: Kernel Regression

$$\hat{\mu}(x^* | \mathbf{x}, B) = \frac{\sum_{i=1}^n \frac{1}{B} K\left[\frac{x_i - x^*}{B}\right] y_i}{\sum_{i=1}^n \frac{1}{B} K\left[\frac{x_i - x^*}{B}\right]}$$

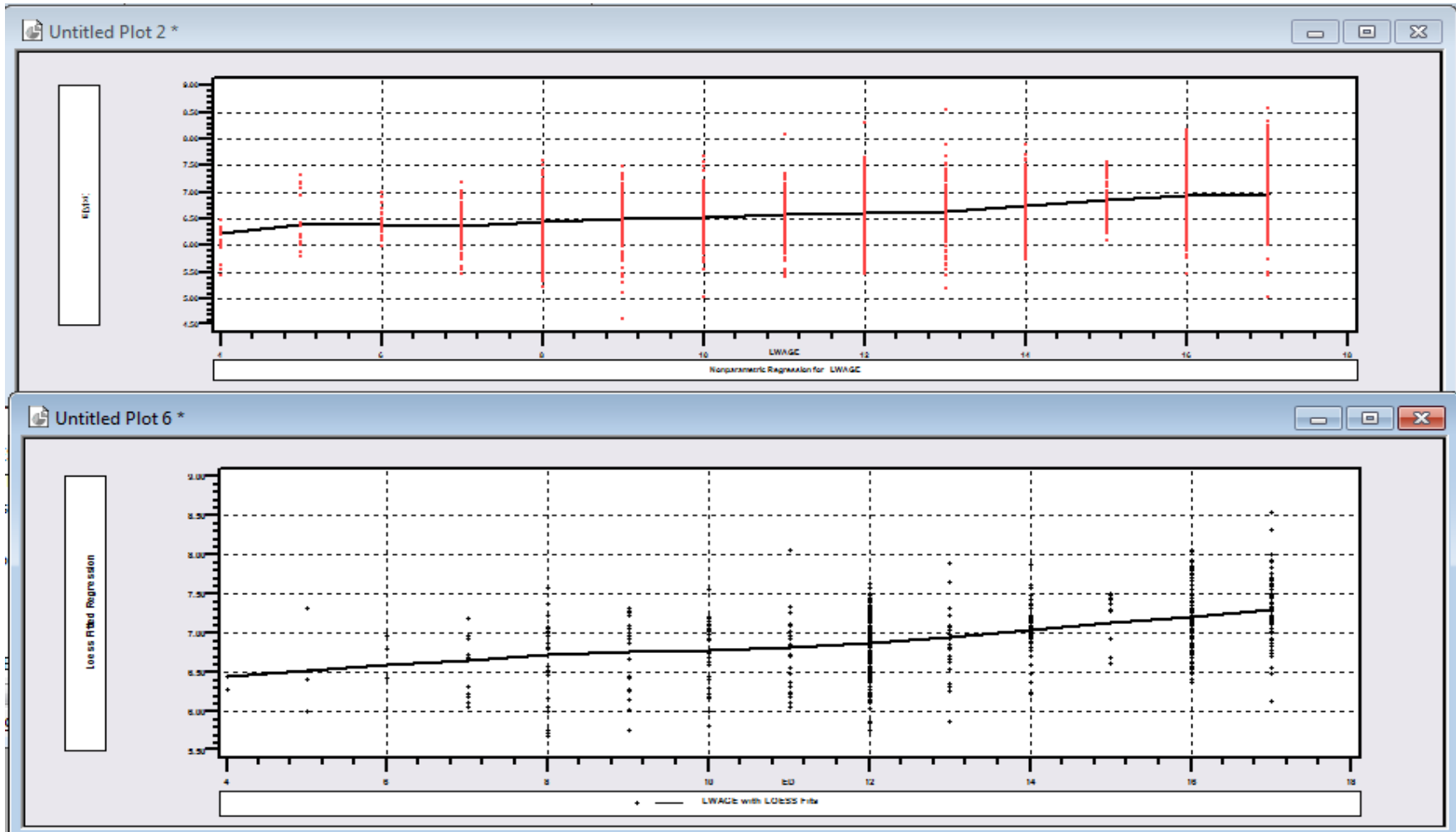
Kernel Functions:

Normal: $K(t) = \phi(t)$

Logistic: $K(t) = \Lambda(t)[1-\Lambda(t)]$

Epanechnikov: $K(t) = .75(1-.2t^2)/\sqrt{5}$, if $|t| \leq 5$ and 0 otherwise

Kernel Regression vs. Lowess (Lwage vs. Educ)



Locally Linear Regression

$$\mu(\mathbf{x}^*) = \beta(\mathbf{x}^*)' \mathbf{x}^*.$$

$$\beta(\mathbf{x}^*) = \left[\sum_{i=1}^n w_i(\mathbf{x}^*, \mathbf{x}_i) \mathbf{x}_i \mathbf{x}_i' \right]^{-1} \left[\sum_{i=1}^n w_i(\mathbf{x}^*, \mathbf{x}_i) \mathbf{x}_i y_i \right]$$

$$w_i(\mathbf{x}^*, \mathbf{x}_i) = K[(\mathbf{x}^* - \mathbf{x}_i)'(\mathbf{x}^* - \mathbf{x}_i), h]$$

```

-----
Ordinary least squares regression .....
LHS=LWAGE Mean = 6.95074
Standard deviation = .43840
No. of observations = 595 Degrees of freedom
Regression Sum of Squares = 38.4740 4
Residual Sum of Squares = 75.6913 590
Total Sum of Squares = 114.165 594
Standard error of e = .35818
Fit R-squared = .33700 R-bar squared = .33251
Model test F[ 4, 590] = 74.97447 Prob F > F* = .00000
Diagnostic Log likelihood = -230.85370 Akaike I.C. = -2.04509
Restricted (b=0) = -353.12149 Bayes I.C. = -2.00821
Chi squared [ 4] = 244.53559 Prob C2 > C2* = .00000
Model was estimated on Aug 29, 2011 at 09:12:14 AM

```

OLS vs. LOWESS

LWAGE	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
Constant	5.57608***	.38522	14.48	.0000	4.82106	6.33109
ED	.07657***	.00548	13.97	.0000	.06583	.08732
FEM	-.44938***	.04704	-9.55	.0000	-.54157	-.35719
UNION	.10451***	.03239	3.23	.0013	.04103	.16798
LOGWKS	.10534	.09720	1.08	.2785	-.08517	.29585

Note: ***, **, * ==> Significance at 1%, 5%, 10% level.

Matrix - LOCLBETA

[595, 5] Cell: 0.305729

	1	2	3	4	5
1	4.78094	0.0473557	-0.541554	0.364504	0.360682
2	5.62988	0.0774372	-0.46378	0.102727	0.084603
3	6.97646	0.0351276	-0.365454	0.0817198	-0.134264
4	4.09146	0.0978699	-0.519258	0.212737	0.42492
5	4.24776	0.0951743	-0.408972	-0.0647883	0.387792
6	6.55862	0.048071	-0.414722	0.0847438	-0.066336
7	6.99921	0.0350738	-0.36576	0.0812928	-0.139938
8	5.80921	0.0136297	-0.583179	0.228573	0.198185
9	4.24905	0.0951543	-0.408957	-0.0648068	0.38754
10	4.25924	0.0949729	-0.408828	-0.0650166	0.385654