

Econometric Analysis of Panel Data

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Final Examination: Spring 2013

This is a 'take home' examination. Today is Tuesday, May 7, 2013. Your answers are due on Friday, May 17, 2013. You may use any resources you wish – textbooks, computer, the web, etc. – but please work alone and submit only your own answers to the questions.

The six parts of the exam are weighted as follows:

Part I.	Literature	20
Part II.	The Hausman and Taylor Estimator	20
Part III.	Panel Data Regressions	20
Part IV.	Binary Choice Models	20
Part V.	A Loglinear Model	20

Note, in parts of the exam in which you are asked to report the results of computation, please filter your response so that you present the numerical results as part of an organized discussion of the question. Do not submit long, unannotated pages of computer output. Some of the parts require you to do some computations. Use Stata, R, *NLOGIT*, MatLab or any other software you wish to use.

Part I. Literature

Locate a published study in a field that interests you that uses a panel data based methodology. Describe in no more than one page the study, the estimation method(s) used and the conclusion(s) reached by the author(s).

Part II. The Hausman and Taylor Estimator

Write out a full statement of the procedure that Hausman and Taylor devised for estimation of the parameters in a panel data model in which some independent variables are correlated with the time invariant part of the disturbance in a random effects model. Now, show how the Arellano/Bond/Bover estimator uses the Hausman and Taylor result.

Part III. Panel Data Regressions

The course website contains Munnell's data set on statewide production, in ASCII text form,

http://people.stern.nyu.edu/wgreene/Econometrics/productivity.csv

and as an nlogit project,

http://people.stern.nyu.edu/wgreene/Econometrics/productivity.lpj

The data in the file are a panel on the following variables for the lower 48 states, 17 years (816 observations),

STATE	= state name
YR	= year, 1970,,1986
P_CAP	= public capital
HWY	= highway capital
WATER	= water utility capital
UTIL	= utility capital
PC	= private capital
GSP	= gross state product
EMP	= employment
UNEMP	= unemployment rate

The basic model of interest is

$$Y_{it} \ = \ \alpha + \beta_1 X \mathbf{1}_{it} + \beta_2 X \mathbf{2}_{it} + \beta_3 X \mathbf{3}_{it} + \beta_4 X \mathbf{4}_{it} + \beta_5 X \mathbf{5}_{it} + c_i + \epsilon_{it}$$

Where Y is logGSP, X1 is logPC, X2 is logHWY, X3 is logWATER, X4 is logUTIL and X5 is logEMP.

This is a Cobb-Douglas production function.

a. Fit the "pooled" model and report your results

b. Fit a random effects model and a fixed effects model. Use your estimation results to decide which is the preferable model. If you find that neither panel data model is preferred to the pooled model, show how you reached that conclusion. As part of the analysis, test the hypothesis that there are no "state effects."

c. Assuming that there are "latent individual (state) effects," the asymptotic covariance matrix that is computed for the pooled estimator, $s^2(\mathbf{X'X})^{-1}$, is inappropriate. What estimator can be computed for the covariance matrix of the pooled estimator that will give appropriate standard errors?

d. The hypothesis of constant returns would be that

H₀:
$$\beta_1 + \beta_2 + \beta_3 + \beta_4 + \beta_5 = 1$$

Test this hypothesis in the context of the model in a. and in the context of your preferred model in part b. Do you reach the same conclusion in both cases?

e. A theory of production posed by Professor Axehandle claims that much of state production is committed years in advance, so that a model with lagged production,

$$Y_{it} = \alpha + \beta_1 X 1_{it} + \beta_2 X 2_{it} + \beta_3 X 3_{it} + \beta_4 X 4_{it} + \beta_5 X 5_{it} + \gamma Y_{i,t-1} + c_i + \varepsilon_{it}$$

is more appropriate than the earlier model. What method should Professor A use to obtain consistent estimates of the parameters of her model? Explain your suggestion. (You need not actually estimate Professor A's model. Just advise her on how to do the estimation.)

Part IV. Binary Choice Models

The course website describes the "German Manufacturing Innovation Data." The actual data are not published on the course website. We will use them for purposes of this exercise, however. You can obtain them by downloading either a csv file,

http://people.stern.nyu.edu/wgreene/Econometrics/probit-panel.csv

or an nlogit project file,

http://people.stern.nyu.edu/wgreene/Econometrics/probit-panel.lpj

This data set contains 1,270 firms and 5 years of data for 6,350 observations in total - a balanced panel. The variables that you need for this exercise are described in the data sets area of the course home page,

http://people.stern.nyu.edu/wgreene/Econometrics/PanelDataSets.htm

I am interested in a binary choice model for the innovation variable, IP. You will fit your model using at least three of the independent variables in the data set. With respect to the model you specify,

A. THEORY

(a) If you fit a pooled **logit** model, there is the possibility that you might be ignoring unobserved heterogeneity (effects). Wooldridge argues that when one fits a probit model while ignoring unobserved heterogeneity, the raw coefficient estimator (MLE) is inconsistent, but the quantity of interest, the "Average Partial Effects" might well be estimated appropriately. Explain in detail what he has in mind here.

(b) Suppose we were to estimate a "fixed effects" probit model by "brute force," just by including the 1,270 dummy variables needed to create the empirical model. What would the properties of the resulting estimator likely be? What is "the incidental parameters problem?"

(c) How would I proceed to use Chamberlain's estimator to obtain a consistent slope estimator for the fixed effects logit model.

(d) Describe in detail how to fit a random effects logit model using quadrature and using simulation for the part of the computations where they would be necessary, under the assumption that the effects are uncorrelated with the other included exogenous variables.

(e) Using the random effects logit model that you described in part (d), describe how you would test the hypothesis that the same logit model applies to the four different sectors in the data set (CONSGOOD,FOOD,RAWMTL,INVGOOD).

B. PRACTICE

(a) Fit a pooled probit model using your specification. Provide all relevant estimation results. (Please condense and organize the results in a readable form.)

(b) Fit a random effects probit model.

(c) Use the Mundlak (correlated random effects) approach to approximate a fixed effects model. Recall this means adding the group means of the time varying variables to the model, then using a random effects model.(d) Note the difference between the estimates in (b) and (c). Which do you think is appropriate? Explain.

Tip for nlogit users: You can use **SETPANEL ; Group=firm ; Pds=ti \$** To define the panel, then **CREATE ; new variable = GroupMean(variable,pds=ti)\$** To obtain the group means you need for a variable.

Part V. A Loglinear Model

The random variable $y|\mathbf{x}$ has a gamma distribution:

$$f(y | \mathbf{x}) = \frac{\lambda_i^P}{\Gamma(P)} \exp(-\lambda_i y_i) y_i^{P-1}, y \ge 0, P > 0,$$
$$\lambda_i = \exp(\alpha + \beta' \mathbf{x}_i).$$

Estimation and analysis is based on a sample of N observations on y_i, x_i .

(a) I propose to estimate the parameters P,α,β by maximum likelihood. Derive the log likelihood function, likelihood equations and Hessian. Show precisely how to use Newton's method to estimate the parameters. How will you obtain asymptotic standard errors for your estimator? Note, the derivatives of the gamma function, $\Gamma(P)$ are dlog $\Gamma(P)/dP = \psi(P)$ and d²log $\Gamma(P)/dP^2 = \Psi'(P)$. Just use these symbols in your results.

(b) The conditional mean function is

$$E[y_i | \mathbf{x}_i] = P/\lambda_i = P \times \exp(-\alpha - \beta' \mathbf{x}_i) = \exp[(\log P - \alpha) - \beta' \mathbf{x}_i] = \exp(\theta - \beta' \mathbf{x}_i)$$

Show how this can be used to construct a nonlinear least squares estimator of β . Note that it is not possible to obtain separate estimates of P and α by this method – the constant term will be α – logP for the reason shown in the equation above.

(c) Another conditional mean is

$$E[\log y_i | x_i] = \psi(P) - \log \lambda_i = -[\alpha - \psi(P)] - \beta' x_i = -(\delta + \beta' x_i)$$

How can you use this result to obtain an estimator of β .

(d) Which of the three estimators suggested above will be the most efficient? Explain.

(e) Maximum likelihood estimates of the model using our German health care data, in which the dependent variable is household income, HHNINC, appear on the next page.

- (1) The Erlang form of the gamma model has an integer value of P. The second set of results below are obtained with P = 3. Test the hypothesis of the Erlang form vs. the more general form using both Wald and Likelihood ratio tests.
- (2) Test the hypothesis that $\beta=0$ using a likelihood ratio test.
- (3) Show how to use a Wald test to test the hypothesis that $\beta = 0$.

(f) Derive the functional form of the partial effects in the model, $\partial E[y|x]/\partial x$. Then, using the general model with unrestricted P, compute these partial effects at the means of the variables. (The means are given in the final set of results after the least squares results.)

(g) Nonlinear least squares, OLS, and OLS results for the log of income are shown. What is the relationship between thes results, the parameters computed in (g) and the partial effects? Note, compared to the MLEs, the OLS results seem to have thw wrong signs. Explain.

Gamma (Lo Dependent Log like Restricte Chi squat Estimatic Inf.Cr.A	oglinear) Regressi t variable lihood function ed log likelihood red [5 d.f.] on based on N = 2 IC = -28546.8 AIC	on Model INCC 14279.415 1195.245 26168.340 7326, K = /N = -1.0	DME 554 508 092 6 045				
INCOME	Coefficient	Standard Error	Z	Prob. $ z > Z^*$	95% Cor Inte	nfidence erval	
	Parameters in con	ditional me	ean funct	ion			
Constant	3.53885***	.02233	158.48	.0000	3.49509	3.58262	
AGE	.00040	.00026	1.55	.1213	00011	.00091	
EDUC	05435***	.00119	-45.81	.0000	05667	05202	
MARRIED	23530***	.00646	-36.44	.0000	24796	22264	
HSAT	01074***	.00122	-8.82	.0000	01313	00836	
	Scale parameter f	or gamma mo	odel				
P_scale	5.11950***	.04245	120.61	.0000	5.03630	5.20270	

🛄 Estimated	l covariance ma	trix for coefficie	ent vector			
[6, 6] Ce	ell: 0.00049860	4	✓ ×			
	Constant	AGE	EDUC	MARRIED	HSAT	P_scale
Constant	0.000498604	-3.41396e-006	-1.65021e-005	-1.62231e-005	-1.10514e-005	0.000351957
AGE	-3.41396e-006	6.71176e-008	3.12955e-008	-4.04454e-007	6.56027e-008	-4.67543e-014
EDUC	-1.65021e-005	3.12955e-008	1.40767e-006	3.09253e-007	-1.51836e-007	-6.21876e-013
MARRIED	-1.62231e-005	-4.04454e-007	3.09253e-007	4.17027e-005	-1.9308e-007	2.24721e-012
HSAT	-1.10514e-005	6.56027e-008	-1.51836e-007	-1.9308e-007	1.48273e-006	1.98232e-013
P_scale	0.000351957	-4.67543e-014	-6.21876e-013	2.24721e-012	1.98232e-013	0.00180184
P_scale	0.000351957	-4.67543e-014	-6.21876e-013	2.24721e-012	1.98232e-013	0.00180184

Gamma (Loglinear) Regression ModelDependent variableINCOMELog likelihood function12505.22870Restricted log likelihood1195.24508Chi squared [5 d.f.]22619.96724Estimation based on N = 27326, K = 5Inf.Cr.AIC = -25000.5 AIC/N = -.915

INCOME	Coefficient	Standard Error	z	Prob. z >Z*	95% Cor Inte	nfidence erval	
	Parameters in co	nditional me	ean funct	ion			
Constant	3.00441***	.02708	110.93	.0000	2.95132	3.05749	
AGE	.00040	.00034	1.19	.2356	00026	.00106	
EDUC	05435***	.00155	-35.07	.0000	05739	05131	
MARRIED	23530***	.00844	-27.89	.0000	25183	21877	
HSAT	01074***	.00159	-6.75	.0000	01386	00763	
	Scale parameter	for gamma mo	odel				
P scale	3.0.	(Fixed H	Parameter)			

🔲 Estimated	l covariance ma	trix for coefficio	ent vector			
[6, 6] C	ell: 0.00073354	8	✓ ×			
	Constant	AGE	EDUC	MARRIED	HSAT	P_scale
Constant	0.000733548	-5.82592e-006	-2.81609e-005	-2.76847e-005	-1.88592e-005	0
AGE	-5.82592e-006	1.14536e-007	5.34058e-008	-6.90201e-007	1.11951e-007	0
EDUC	-2.81609e-005	5.34058e-008	2.40218e-006	5.2774e-007	-2.59109e-007	0
MARRIED	-2.76847e-005	-6.90201e-007	5.2774e-007	7.11657e-005	-3.29491e-007	0
HSAT	-1.88592e-005	1.11951e-007	-2.59109e-007	-3.29491e-007	2.53028e-006	0
P_scale	0	0	0	0	0	0

Nonlinear	least squares	regression				
LHS=INCOM	E Mean	=		35214		
	Standard devia	ation =		17686		
	Number of obse	ervs. =	1	27326		
Model size	e Parameters	=		5		
	Degrees of fre	eedom =	1	27321		
Residuals	Sum of squares	s =	763	1.973		
	Standard error	cofe =	•	16699		
Fit	R-squared	=	•	10847		
	Adjusted R-squ	ared =		10850		
Model test	F[4, 27321]	(prob) =	831.0(.0	1000)		
+-		Ctondowd		Dreb	05% Com	fidongo
UcerFund	Coefficient	Frror	7		. 95% COI. * Thte	
				12122		
B1	-1.97224***	02139	-92.21	.0000	-2.01416	-1.93032
B2	.00132***	.00027	4.88	.0000	.00079	.00185
B3	.05381***	.00103	52.16	.0000	.05179	.05584
в4	.23520***	.00764	30.77	.0000	.22022	.25018
в5	.01043***	.00132	7.89	.0000	.00784	.01301
+-						
Ordinary	least squares	regression				
	- No. of observa	ations =	:	27326	DegFreedom	Mean square
Regression	n Sum of Squares	s =	91	. 7939	- 4	22.94847
Residual	Sum of Squares	s =	76:	2.888	27321	.02792
Total	Sum of Squares	3 =	854	4.682	27325	.03128
	- Standard error	ofe =		16710	Root MSE	.16709
Fit	R-squared	=		10740	R-bar squared	l .10727
Model test	F[4, 27321]	=	821.8	34443	Prob F > F*	.00000
+-						
		Standard		Prob	. 95% Con	fidence
INCOME	Coefficient	Error	z	z >Z	* Inte	erval
+-						
Constant	.02071***	.00788	2.63	.0086	.00527	.03616
AGE	.00025*** .	9450D-04	2.65	.0081	.00007	.00044
EDUC	.02094***	.00044	47.15	.0000	.02007	.02181
MARRIED	.07801***	.00241	32.34	.0000	.07328	.08274
HSAT	.00358***	.00046	7.84	.0000	.00268	.00447
+-	1 t					
Ordinary	least squares	regression				
THS=IOGING	COM Mean	=	-1	15/44		
	Standard devia	ation =		49145	DeerDeeeder	Maan
Democratics	- NO. OI ODSErva	$\alpha = 1000$	0.47	2/320	Degrreedom	Mean square
Regression	1 Sum of Granes	· =	94	1.342	4 07201	∠30.03541 20600
Residual	Sum of Gruares	, =	50	54.34 DD CC	2/321 27225	. 20009
IOLAL	Sum OI Squares		055	99.00 46405	Z/JZ5	.24152
 ₽i+	- Stanuard error	. or e =	•	±0405 14254	ROOL MSE	.45481
rit Model toot	- E[4 07001]	=	11//	14334 76500	r-bar squared	14342 00000
	- ri 4, 2/321]	=	±±44.	20209	FION L > L.	.00000
		Standard		Prob	95% Cor	fidence
	Coefficient	Error	7	2 2 2	* 55% COL	rval
+-						
Constant	-2.08985***	.02145	-97.43	.0000	-2.13190	-2.04781
AGE	00069***	.00026	-2.70	.0070	00120	00019
EDUC	.05670***	.00121	46.90	.0000	.05433	.05907
MARRIED	.32016***	.00657	48.77	.0000	.30729	.33303
HSAT	.01148***	.00124	9.24	.0000	.00904	.01391
+-						
Descriptiv	ve Statistics for	6 variab	les			
Variable	Mean	Std.Dev.	Minim	um	Maximum	Cases Missing
INCOME	. 352135	.176857	001	500	3.067100	27326
	-1 157442	491452	-6 502	290	1 120732	27326
TOGTICOUL	43 52569	11 33025	0.002.	5 0	£4 N	27326
EDIIC	11 32063	2.324885	2.	7.0	18 0	27326
MARRIED	.758618	427929	(1 0	27326
HOVL	6 785662	2 293725			10 0	27326
116A1	0.703002				±0.0	
+-						