

Econometric Analysis of Panel Data

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URL for course web page:
people.stern.nyu.edu/wgreene/Econometrics/PanelDataEconometrics.htm

Final Examination: Spring 2018

This is a ‘take home’ examination. Today is Tuesday, May 1, 2018. Your answers are due on Monday, May 14, 2018. You may use any resources you wish – textbooks, computer, the web, etc. – but please work alone and submit only your own answers to the questions.

The five parts of the exam are weighted as follows:

Part I.	Literature	20
Part II.	The Mundlak Estimator	20
Part III.	Panel Data Regressions	50
Part IV.	Binary Choice Models	50
Part V.	A Loglinear Model	60

Note, in parts of the exam in which you are asked to report the results of computation, please filter your response so that you present the numerical results as part of an organized discussion of the question. Do not submit long, unannotated pages of computer output. Some of the parts require you to do some computations. Use Stata, R, *NLOGIT*, MatLab or any other software you wish to use.

Part I. Literature

Locate a published study in a field that interests you that uses a panel data based methodology. Describe in no more than one page the study, the estimation method(s) used and the conclusion(s) reached by the author(s).

Part II. The Mundlak Approach in Estimation

Many recent studies have revived Mundlak’s approach to modeling common effects in linear regression and nonlinear models. Describe in detail the standard common effects models. How is the Mundlak estimator motivated? How is it employed? Show how the estimator provides a constructive test for fixed vs. random effects.

Part III. Panel Data Regressions

The course website contains an abbreviated version of the WHO health outcomes data set,

<http://people.stern.nyu.edu/wgreene/Econometrics/WHO-balanced-panel.csv>

and as an nlogit project,

<http://people.stern.nyu.edu/wgreene/Econometrics/WHO-balanced-panel.lpj>

The csv file is a text, comma delimited file that should be directly readable by other programs such as Stata and R. The original data set contained 840 observations as an unbalanced panel for 191 countries. It also contained data for some internal political districts such as the 24 states of Mexico and the provinces of Canada and Australia. This panel retains the data for the 140 countries that contain all 5 years of data. The variables in the file are

COUNTRY	= Country name (text)
ID, STRATUM	= Country ID. Ignore STRATUM
YEAR	= 1993, ..., 1997
COMP and LOGCOMP	= WHO health outcome measure and its log
DALE and LOGDALE	= WHO life expectancy and its log
EDUC, LOGEDUC, LOGEDUC2	= Education, log and square of log
HLTHEXP, LOGHEXP, LOGHEXP2	= Health expenditure, log and square of log
PUBTHE	= Share of health expenditure paid by government
LOGED_EX	= LOGHEXP * LOGEDUC
GINI	= Gini coefficient income distribution
TROPICS	= Dummy variable for tropical country
POPDEN, LOGPOPDN	= Population density, people per square kilometer and log
GDPC, LOGGDPC	= Per capita GDP and log
T93,...,T97	= Year dummy variables
GEFF	= World bank measure of government effectiveness
VOICE	= World Bank measure of political efficacy
OECD	= OECD member dummy variable
MEANLCMP	= Country mean of log COMP
MEANLHC	= Country mean of log EDUC
MEANLHC2	= Country mean of log EDUC squared
MEANLEXP	= Country mean of log HEXP

Note that COMP, DALE, EDUC and HLTHEXP are time varying, but all other measured variables are time invariant.

The WHO model originally specified was

$$y_{it} = \alpha + \beta_1 x_{1,it} + \beta_2 x_{2,it} + \gamma_{11} x_{1,it}^2 + \gamma_{22} x_{2,it}^2 + \gamma_{12} x_{1,it} x_{2,it} + \varepsilon_{it}$$

where

$$y = \log \text{COMP}, x_1 = \log \text{EDUC}, x_2 = \log \text{HEXP}.$$

Call this Model A. This is a translog production function. The authors found that the values of γ_{ki} implied a nonconcave production function, and fixed γ_{22} and γ_{12} both to zero in their final presentation. Call this restricted model Model B.

- a. Fit the “pooled” model and report your results.
- b. Using the pooled model, test the null hypothesis of Model B against the alternative Model A.
- c. Using the formulation of Model B, fit a random effects model and a fixed effects model. Use your estimation results to decide which is the preferable model. If you find that neither panel data model is preferred to the pooled model, show how you reached that conclusion. As part of the analysis, test the hypothesis that there are no “country effects.”
- d. Using the Mundlak approach, determine which model, fixed or random effects is preferred.
- e. Assuming that there are “latent individual (county) effects,” the asymptotic covariance matrix that is computed for the pooled estimator, $s^2(\mathbf{X}'\mathbf{X})^{-1}$, is inappropriate. What estimator can be computed for the covariance matrix of the pooled estimator that will give appropriate standard errors?
- f. The hypothesis of constant returns to scale in the translog model (Model A) would be

$$H_0: \beta_1 + \beta_2 = 1 \text{ and } \gamma_{11} + \gamma_{22} + 2\gamma_{12} = 0$$

Test this hypothesis in the context of Model A.

- g. The 2004 Health Economics paper by Greene argued that WHO did not handle the obvious heterogeneity across countries appropriately. Variables GINI, TROPICS, logPOPDN, logGDPC, GEFF, VOICE, OECD all capture dimensions of this heterogeneity. Extend the random effects model to include some (or all) of these variables and test the hypothesis that they significantly add to the explanatory power of the model.
- h. Are there “time effects” in the data. One approach find out would be to add the time variables (less one of them) to the preferred regression model and test for their joint significance. A second approach would be to use a CHOW test to test for homogeneity of the regression model over the 5 years. Test the homogeneity assumption using your preferred pooled model.

Part IV. Binary Choice Models

The course website describes the “German Manufacturing Innovation Data.” The actual data are not published on the course website. We will use them for purposes of this exercise, however. You can obtain them by downloading either a csv file,

<http://people.stern.nyu.edu/wgreene/Econometrics/probit-panel.csv>

or an nlogit project file,

<http://people.stern.nyu.edu/wgreene/Econometrics/probit-panel.lpj>

This data set contains 1,270 firms and 5 years of data for 6,350 observations in total – a balanced panel. The variables that you need for this exercise are described in the data sets area of the course home page,

<http://people.stern.nyu.edu/wgreene/Econometrics/PanelDataSets.htm>

(The csv file can easily be ported to other software such as R, SAS and Stata.) I am interested in a binary choice model for the innovation variable, IP. You will fit your model using at least three of the independent variables in the data set. With respect to the model you specify,

A. THEORY

- (a) If you fit a pooled **logit** model, there is the possibility that you might be ignoring unobserved heterogeneity (effects). Wooldridge argues that when one fits a probit model while ignoring unobserved heterogeneity, the raw coefficient estimator (MLE) is inconsistent, but the quantity of interest, the “Average Partial Effects” might well be estimated appropriately. Explain in detail what he has in mind here.
- (b) Suppose we were to estimate a “fixed effects” probit model by “brute force,” just by including the 1,270 dummy variables needed to create the empirical model. What would the properties of the resulting estimator likely be? What is “the incidental parameters problem?”
- (c) How would I proceed to use Chamberlain’s estimator to obtain a consistent slope estimator for the fixed effects logit model.
- (d) Describe in detail how to fit a random effects logit model using quadrature and using simulation for the part of the computations where they would be necessary, under the assumption that the effects are uncorrelated with the other included exogenous variables.
- (e) Using the random effects logit model that you described in part (d), describe how you would test the hypothesis that the same logit model applies to the four different sectors in the data set (CONSGOOD, FOOD, RAWMTL, INVGGOOD).

B. PRACTICE

- (a) Fit a pooled probit model using your specification. Provide all relevant estimation results. (Please condense and organize the results in a readable form.)
- (b) Fit a random effects probit model.
- (c) Use the Mundlak (correlated random effects) approach to approximate a fixed effects model. Recall this means adding the group means of the time varying variables to the model, then using a random effects model.
- (d) Note the difference between the estimates in (b) and (c). Which do you think is appropriate? Explain.

Tip for nlogit users: You can use

CREATE ; new variable = GroupMean(variable,pds=5)\$

To obtain the group means you need for a variable.

Part V. A Loglinear Model

This semester, we have examined several ‘loglinear models,’ including the logit model for binary choice, Poisson and negative binomial models for counts and the exponential model for a continuous nonnegative random variable. We will now examine one more loglinear model. The nonnegative, continuous random variable $y|x$ has a Weibull distribution:

$$f(y|x) = \lambda_i P y_i^{P-1} \exp(-\lambda_i y_i^P), y \geq 0, P > 0,$$

$$\lambda_i = \exp(\alpha + \beta' \mathbf{x}_i).$$

(We examined a version of this model in Assignment 5.) Estimation and analysis is based on a sample of N observations on y_i, \mathbf{x}_i . The conditional mean function is

$$E[y_i|x_i] = \frac{1}{\lambda_i} \Gamma\left(\frac{P+1}{P}\right) = \exp(-\alpha - \beta' \mathbf{x}_i) \Gamma\left(\frac{P+1}{P}\right) \text{ (Note the minus sign.)}$$

The variables used in the regressions are described below.

Variable	Mean	Std.Dev.	Minimum	Maximum	Cases	Missing
INCOME	.352135	.176857	.001500	3.067100	27326	0
logINCOM	-1.157442	.491452	-6.502290	1.120732	27326	0
AGE	43.52569	11.33025	25.0	64.0	27326	0
EDUC	11.32063	2.324885	7.0	18.0	27326	0
HSAT	6.785662	2.293725	0.0	10.0	27326	0
MARRIED	.758618	.427929	0.0	1.0	27326	0
HHKIDS	.402730	.490456	0.0	1.0	27326	0

The data set is a panel. There are 7,293 groups with group sizes ranging from 1 to 7. This exercise will examine a variety of regression formulations. I have done the estimation for you; the results appear below. Some of the questions will involve a small amount of ancillary computation.

A. I propose to estimate the parameters (P, α, β) by maximum likelihood. The results are shown in regression 1 below. Derive the log likelihood function, likelihood equations and Hessian. Show precisely how to use Newton’s method to estimate the parameters. How will you obtain asymptotic standard errors for your estimator? Test the hypothesis of ‘the regression model.’ That is, test the hypothesis that all of the coefficients in β are equal to zero using the likelihood ratio test.

B. There are several interesting special cases of the Weibull model. If $P = 1$, the model reduces to the exponential model discussed in class. We considered three different ways to test a parametric restriction such as this, Wald, Likelihood ratio and LM tests. Using the results of regressions 1, 2 and 3 below, carry out the three tests. Do the results of the three tests agree?

C. The conditional mean function shown above suggests a nonlinear least squares approach. Note that the conditional mean function can be written

$$E[y|x] = \exp\left[\log \Gamma\left(\frac{P+1}{P}\right) - \alpha - \beta' \mathbf{x}\right] = \exp(\delta - \beta' \mathbf{x})$$

Thus, the constant term in the conditional mean function is not $-\alpha$. The nonlinear least squares results are shown in regression 4. How do the two results compare to the MLE? We now have two possible estimators of β . In theoretical terms, which is better, MLE or NLS? Why? Do the empirical results support your argument?

D. The likelihood equations for estimation of (P, α, β) imply that $E[y^P|x] = 1/\lambda$. Prove this result.

E. Derive the partial effects for the Weibull conditional mean function, $\partial E[y|x]/\partial x$. Compute the partial effects at the means of the data. Hint: $\Gamma((P+1)/P)$ for the P in regression 1 equals .88562. How would you obtain standard errors for your estimated partial effects? Explain in detail.

F. Regression 5 presents *linear* least squares results for the regression of $-y$ on $(1, x)$. (The minus sign on y changes the sign of the coefficients so they will be comparable to the earlier results.) How do these results compare to the MLEs in part A? How do they compare to the results in part E? Why would they resemble the results in part E?

G. The log of a Weibull distributed variable has a type 1 extreme value distribution. The expected value of $\log y$ is $-(\alpha + \beta'x) + \gamma$, where γ is the Euler-Mascheroni constant, 0.57721566.... Regression 6 presents the results of linear regression of $-\log y$ on x . Which other result should these resemble? Do they?

H. Since these are panel data, it is appropriate to rebuild the model to accommodate the unobserved heterogeneity. Explain the difference between fixed and random effects models. How would they appear in the loglinear model formulated here?

I. Regressions 7 and 8 show FEM and REM.

- (1) What is the incidental parameters problem? Would the result apply to the model shown in (7)?
- (2) Show how the parameters of the random effects model in regression 8 are computed. I.e., describe how the maximum simulated likelihood estimator is computed.
- (3) Regression 9 presents estimates of a random effects model that also contains the group means of the regressors. As noted earlier, this Mundlak style treatment helps to distinguish the FE and RE specifications. Based on the results given, which appears to be the preferable model, FE or RE?

J. Some have argued that marital status might be endogenous in an income equation when there are households that have two working people. (You probably thought people married for love.) To investigate in the present model, I will use a control function approach. Regression 10 presents a probit equation for marital status based on age, education, gender and whether the household head has a white collar job. The variable GENRES is the generalized residual from this model,

$GENRES = q\phi(\beta'x)/\Phi(q\beta'x)$ where $q = 2\text{Married} - 1$. The expected value of GENRES is zero, and since it is the derivative of $\log L$ with respect to the constant term, it will sum to zero in the sample. I am going to use GENRES as a control function. What is a control function, and why will I use it in the INCOME model?

K. Regression 11 presents estimates of the Weibull INCOME model that includes the control function. Regression 12 is similar to 11, but regression 12 includes normal heterogeneity in the model in the form of what appears to be a random effect – a random constant. But, this is not a panel data model look closely at the results and note that the ‘panel’ has one period. The implied two equation model underlying 12 is

$$\text{MARRIED}_i^* = \gamma'z + u_i, \text{ MARRIED}_i = 1[\text{MARRIED}_i^* > 0], u_i \sim N[0,1].$$

$$\text{INCOME}_i^* \sim \text{Weibull}(\lambda_i, P) \text{ where } \lambda_i = \exp(\beta'x_i + \varepsilon_i)$$

where (ε_i, u_i) have a bivariate normal distribution with means $(0,0)$, standard deviations $(\sigma_\varepsilon, 1)$ and correlation ρ . The endogeneity issue turns on ρ . The coefficient on GENRES in the model in regression 12 will approximate $\sigma_\varepsilon \rho$. So, based on the estimated model, marry for money (endogenous, ρ not equal to zero) or marry for love (exogenous, ρ equal to zero)?

L. In this model, the argument in parts J and K about MARRIED could also be made about health satisfaction, HSAT. But, HSAT is an ordered outcome, coded 0,1,2 (bad, middling, good) in our data. How would you proceed to deal with endogeneity of HSAT in this model?

1. Weibull, MLE

Weibull (Loglinear) Regression Model

Dependent variable INCOME
 Log likelihood function 12133.14495
 Restricted log likelihood 1195.24508 (Log likelihood when $\beta = 0$)
 Chi squared [7] (P= .000) 21875.79975
 Significance level .00000
 McFadden Pseudo R-squared -9.1511775
 Estimation based on N = 27326, K = 8
 Inf.Cr.AIC = -24250.3 AIC/N = -.887

INCOME	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
Parameters in conditional mean function						
Constant	1.67075***	.01433	116.62	.0000	1.64267	1.69883
AGE	.00086***	.00022	3.91	.0001	.00043	.00130
EDUC	-.05084***	.00073	-69.23	.0000	-.05228	-.04940
HSAT	-.01233***	.00077	-15.96	.0000	-.01385	-.01082
MARRIED	-.16990***	.00371	-45.79	.0000	-.17717	-.16262
FEMALE	-.02041***	.00334	-6.11	.0000	-.02696	-.01386
HHKIDS	.06403***	.00375	17.07	.0000	.05668	.07139
Scale parameter for Weibull model						
P_scale	2.13722***	.00495	431.40	.0000	2.12751	2.14693

***, **, * ==> Significance at 1%, 5%, 10% level.

	Constant	AGE	EDUC	HSAT	MARRIED	HHKIDS	P_scale
Constant	0.000203282	-2.18842e-006	-5.43762e-006	-4.95755e-006	1.02554e-005	-1.15906e-005	-2.77586e-005
AGE	-2.18842e-006	4.7703e-008	-7.89778e-009	4.31199e-008	-3.07207e-007	1.75484e-007	1.31682e-007
EDUC	-5.43762e-006	-7.89778e-009	5.32837e-007	-5.61965e-008	-3.2693e-007	2.5195e-007	9.33686e-007
HSAT	-4.95755e-006	4.31199e-008	-5.61965e-008	5.90819e-007	-2.22175e-008	-3.58706e-007	-5.27993e-007
MARRIED	1.02554e-005	-3.07207e-007	-3.2693e-007	-2.22175e-008	1.33275e-005	-4.0085e-006	-6.66408e-007
HHKIDS	-1.15906e-005	1.75484e-007	2.5195e-007	-3.58706e-007	-4.0085e-006	1.35887e-005	3.46969e-006
P_scale	-2.77586e-005	1.31682e-007	9.33686e-007	-5.27993e-007	-6.66408e-007	3.46969e-006	2.40568e-005

2. Exponential, MLE

```

Exponential (Loglinear) Regression Model
Dependent variable      INCOME
Log likelihood function  1558.04494
Restricted log likelihood 1195.24508
Chi squared [ 5] (P= .000) 725.59973
Significance level      .00000
McFadden Pseudo R-squared -.3035360
Estimation based on N = 27326, K = 6
Inf.Cr.AIC = -3104.1 AIC/N = -.114

```

INCOME	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
Parameters in conditional mean function						
Constant	1.85106***	.04834	38.29	.0000	1.75632	1.94580
AGE	.00158**	.00064	2.48	.0133	.00033	.00283
EDUC	-.05438***	.00268	-20.27	.0000	-.05963	-.04912
HSAT	-.01101***	.00275	-4.00	.0001	-.01641	-.00561
MARRIED	-.26249***	.01568	-16.75	.0000	-.29322	-.23177
HHKIDS	.06619***	.01399	4.73	.0000	.03877	.09360

	Constant	AGE	EDUC	HSAT	MARRIED	HHKIDS
Constant	0.00233667	-2.04546e-005	-8.41197e-005	-5.58263e-005	-1.67755e-005	-0.000163745
AGE	-2.04546e-005	4.06572e-007	1.57594e-007	3.21119e-007	-3.44467e-006	3.49705e-006
EDUC	-8.41197e-005	1.57594e-007	7.19294e-006	-7.79848e-007	1.52206e-006	-7.72587e-008
HSAT	-5.58263e-005	3.21119e-007	-7.79848e-007	7.58913e-006	-6.64156e-007	-7.8368e-007
MARRIED	-1.67755e-005	-3.44467e-006	1.52206e-006	-6.64156e-007	0.00024571	-8.04925e-005
HHKIDS	-0.000163745	3.49705e-006	-7.72587e-008	-7.8368e-007	-8.04925e-005	0.000195638

3. Constrained Weibull, MLE

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Weibull (Loglinear) Regression Model
Dependent variable      INCOME
LM Stat. at start values 21526.22099
LM statistic kept as scalar LMSTAT
Log likelihood function  1558.04494
Restricted log likelihood 1195.24508
Chi squared [ 6] (P= .000) 725.59973
Significance level      .00000
McFadden Pseudo R-squared -.3035360
Estimation based on N = 27326, K = 7
Inf.Cr.AIC = -3102.1 AIC/N = -.114

```

INCOME	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
Parameters in conditional mean function						
Constant	1.85106***	.09976	18.56	.0000	1.65553	2.04658
AGE	.00158	.00130	1.22	.2233	-.00096	.00412
EDUC	-.05438***	.00574	-9.47	.0000	-.06563	-.04312
HSAT	-.01101**	.00556	-1.98	.0477	-.02190	-.00011
MARRIED	-.26249***	.02881	-9.11	.0000	-.31896	-.20603
HHKIDS	.06619**	.02827	2.34	.0192	.01078	.12160
Scale parameter for Weibull model						
P scale	1.0***	.00672	148.81	.0000	.98683	.10132D+01

***, **, * ==> Significance at 1%, 5%, 10% level.

4. Nonlinear Least Squares, y on $\exp(-b'x)$

```

Nonlinear least squares regression .....
LHS=INCOME Mean = .35214
Standard deviation = .17686
Fit R-squared = .11070
Adjusted R-squared = .11073
Model test F[ 5, 27320] (prob) = 680.2(.0000)
-----+-----
UserFunc| Coefficient Standard Error z Prob. 95% Confidence
| | | | | | | Interval
-----+-----
B_ONE| 1.92270*** .02202 87.33 .0000 1.87955 1.96585
B_AGE| -.00022 .00030 -.75 .4535 -.00081 .00036
B_EDUC| -.05378*** .00103 -52.09 .0000 -.05580 -.05175
B_HSAT| -.01072*** .00132 -8.12 .0000 -.01330 -.00813
B_MARR| -.25986*** .00821 -31.66 .0000 -.27594 -.24377
B_KIDS| .05581*** .00664 8.40 .0000 .04279 .06883

```

5. Linear Least Squares, $-y$ on $b'x$

Ordinary	least squares regression				
LHS=MINCOME	Mean	=	-.35214		
	Standard deviation	=	.17686		
-----	No. of observations	=	27326	DegFreedom	Mean square
Regression	Sum of Squares	=	93.8115	5	18.76231
Residual	Sum of Squares	=	760.870	27320	.02785
Total	Sum of Squares	=	854.682	27325	.03128
-----	Standard error of e	=	.16688	Root MSE	.16687
Fit	R-squared	=	.10976	R-bar squared	.10960
Model test	F[5, 27320]	=	673.68410	Prob F > F*	.00000
-----+-----					
		Standard		Prob.	95% Confidence
MINCOME	Coefficient	Error	z	z >Z*	Interval
-----+-----					
Constant	-.03873***	.00815	-4.75	.0000	-.05470 -.02275
AGE	.00012	.00010	1.14	.2535	-.00008 .00032
EDUC	-.02088***	.00044	-47.07	.0000	-.02175 -.02001
HSAT	-.00366***	.00046	-8.03	.0000	-.00455 -.00277
MARRIED	-.08630***	.00260	-33.21	.0000	-.09140 -.08121
HHKIDS	.02024***	.00238	8.51	.0000	.01558 .02489

6. Linear Least Squares, $-\log y$ on $b'x$

Ordinary least squares regression						
LHS=MLINCOME						
	Mean	=	1.15744			
	Standard deviation	=	.49145			
-----	No. of observations	=	27326	DegFreedom	Mean square	
Regression	Sum of Squares	=	968.991	5	193.79827	
Residual	Sum of Squares	=	5630.67	27320	.20610	
Total	Sum of Squares	=	6599.66	27325	.24152	
-----	Standard error of e	=	.45398	Root MSE	.45393	
Fit	R-squared	=	.14682	R-bar squared	.14667	
Model test	F[5, 27320]	=	940.30851	Prob F > F*	.00000	

MLINCOME	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
Constant	2.03085***	.02217	91.60	.0000	1.98740	2.07430
AGE	.00190***	.00028	6.73	.0000	.00135	.00246
EDUC	-.05651***	.00121	-46.82	.0000	-.05887	-.05414
HSAT	-.01175***	.00124	-9.47	.0000	-.01418	-.00932
MARRIED	-.34733***	.00707	-49.14	.0000	-.36118	-.33348
HHKIDS	.06628***	.00647	10.25	.0000	.05361	.07896

7. Fixed Effects Weibull, MLE

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FIXED EFFECTS Weibul Model
Dependent variable      INCOME
Log likelihood function      34910.40335
Estimation based on N = 27326, K = 7299
Inf.Cr.AIC = -55222.8 AIC/N = -2.021
Unbalanced panel has 7293 individuals
Skipped 0 groups with inestimable ai
Weibull loglinear regression model

```

		Standard		Prob.	95% Confidence	
INCOME	Coefficient	Error	z	z >Z*	Interval	

	Index function for probability					
AGE	-.04322***	.00055	-78.85	.0000	-.04429	-.04214
EDUC	-.07959***	.00616	-12.91	.0000	-.09167	-.06750
HSAT	-.00339***	.00088	-3.85	.0001	-.00511	-.00166
MARRIED	-.18215***	.00836	-21.80	.0000	-.19853	-.16578
HHKIDS	.07732***	.00550	14.06	.0000	.06654	.08810
	Scale parameter for Weibull distribution					
P scale	5.77115***	.02935	196.61	.0000	5.71362	5.82868

***, **, * ==> Significance at 1%, 5%, 10% level.

8. Random Effects Weibull, Maximum Simulated Likelihood

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-----
Random Coefficients WeiblReg Model
Dependent variable      INCOME
Log likelihood function      19489.51857
Restricted log likelihood      1558.04494
Chi squared [ 1](P=.000) 35862.94726
Significance level      .00000
McFadden Pseudo R-squared -11.5089579
Estimation based on N = 27326, K = 8
Inf.Cr.AIC = -38963.0 AIC/N = -1.426
Unbalanced panel has 7293 individuals
Simulation based on 100 Halton draws
Weibull loglinear regression model

```

		Standard		Prob.	95% Confidence	
INCOME	Coefficient	Error	z	z >Z*	Interval	

	Nonrandom parameters					
AGE	-.01369***	.00015	-91.51	.0000	-.01398	-.01339
EDUC	-.06413***	.00057	-111.59	.0000	-.06525	-.06300
HSAT	-.00478***	.00059	-8.05	.0000	-.00594	-.00361
MARRIED	-.19181***	.00307	-62.53	.0000	-.19782	-.18580
HHKIDS	.08751***	.00271	32.26	.0000	.08219	.09282
	Means for random parameters					
Constant	2.43436***	.01030	236.40	.0000	2.41418	2.45455
	Scale parameters for dists. of random parameters					
Constant	.50166***	.00138	364.84	.0000	.49896	.50435
	Scale parameter for Weibull distribution					
P_scale	4.15999***	.01130	368.03	.0000	4.13783	4.18214

***, **, * ==> Significance at 1%, 5%, 10% level.

9. Random Effects Weibull, Maximum Simulated Likelihood with Group Means

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Random Coefficients  WeibReg Model
Dependent variable      INCOME
Log likelihood function    21443.98658
Restricted log likelihood   1735.72267
Chi squared [ 1](P= .000) 39416.52782
Significance level        .00000
McFadden Pseudo R-squared -11.3545005
Estimation based on N = 27326, K = 13
Inf.Cr.AIC = -42862.0 AIC/N = -1.569
Unbalanced panel has 7293 individuals
Simulation based on 100 Halton draws
Weibull loglinear regression model
-----
```

INCOME	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
Nonrandom parameters						
AGE	-.04140***	.00041	-100.59	.0000	-.04221	-.04060
EDUC	-.07436***	.00382	-19.47	.0000	-.08185	-.06687
HSAT	-.00775***	.00076	-10.19	.0000	-.00924	-.00626
MARRIED	-.20973***	.00473	-44.37	.0000	-.21899	-.20046
HHKIDS	.08242***	.00454	18.17	.0000	.07353	.09132
gmnAGE	.04656***	.00043	107.94	.0000	.04571	.04741
gmnEDUC	.03803***	.00384	9.90	.0000	.03050	.04556
gmnHSAT	-.01136***	.00083	-13.69	.0000	-.01299	-.00974
gmnMARRI	.01497**	.00587	2.55	.0107	.00347	.02646
gmnHHKID	-.01393**	.00603	-2.31	.0210	-.02575	-.00210
Means for random parameters						
Constant	1.41275***	.00930	151.96	.0000	1.39453	1.43097
Scale parameters for dists. of random parameters						
Constant	.46708***	.00118	395.64	.0000	.46477	.46940
Scale parameter for Weibull distribution						
P_scale	4.38591***	.01205	363.95	.0000	4.36229	4.40953

Binomial Probit Model						
Dependent variable		MARRIED				
	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
Index function for probability						
Constant	.20370***	.05761	3.54	.0004	.09079	.31662
AGE	.02234***	.00076	29.28	.0000	.02084	.02383
EDUC	-.03308***	.00367	-9.02	.0000	-.04027	-.02589
FEMALE	-.12946***	.01727	-7.50	.0000	-.16330	-.09562
WHITEC	-.03858**	.01861	-2.07	.0382	-.07506	-.00210

Weibull (Loglinear) Regression Model						
Dependent variable		INCOME				
Log likelihood function		12160.11190				
Restricted log likelihood		1195.24508				
Chi squared [7](P= .000)		21929.73366				
Significance level		.00000				
-----+-----						
INCOME	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
-----+-----						
Parameters in conditional mean function						
Constant	1.19668***	.03721	32.16	.0000	1.12374	1.26962
AGE	-.00528***	.00051	-10.39	.0000	-.00627	-.00428
EDUC	-.04215***	.00107	-39.51	.0000	-.04425	-.04006
HSAT	-.01251***	.00077	-16.34	.0000	-.01401	-.01101
MARRIED	.67313***	.06091	11.05	.0000	.55375	.79251
HHKIDS	.05201***	.00371	14.03	.0000	.04475	.05927
GENRES	-.49197***	.03508	-14.02	.0000	-.56073	-.42321
Scale parameter for Weibull model						
P_scale	2.13826***	.00492	434.41	.0000	2.12862	2.14791
-----+-----						

Random Coefficients WeibReg Model						
Dependent variable INCOME						
Log likelihood function		13158.01422				
Restricted log likelihood		1563.62291				
Sample is 1 pds and 27326 individuals						
Simulation based on 10 Halton draws						
INCOME	Coefficient	Standard Error	z	Prob. z >Z*	95% Confidence Interval	
Nonrandom parameters						
AGE	-.00331***	.00046	-7.24	.0000	-.00421	-.00241
EDUC	-.04512***	.00097	-46.49	.0000	-.04702	-.04321
HSAT	-.01211***	.00072	-16.84	.0000	-.01352	-.01070
MARRIED	.45855***	.05526	8.30	.0000	.35024	.56686
HHKIDS	.06500***	.00369	17.60	.0000	.05776	.07224
GENRES	-.39467***	.03169	-12.45	.0000	-.45678	-.33256
Means for random parameters						
Constant	1.31805***	.03473	37.95	.0000	1.24999	1.38612
Scale parameters for dists. of random parameters						
Constant	.25368***	.00194	130.76	.0000	.24988	.25748
Scale parameter for Weibull distribution						
P_scale	2.64003***	.00733	359.95	.0000	2.62566	2.65441