

Econometric Analysis of Panel Data

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URL for course web page:

www.stern.nyu.edu/~wgreene/Econometrics/PanelDataEconometrics.htm

Assignment 1

Part I. Mathematical Statistics

The density $f(y)$ for a nonnegative random variable, y , is exponential with parameter λ , so

$$f(y) = 1/\lambda \exp(-y/\lambda), y \geq 0, \lambda > 0.$$

For this random variable, the mean is $E[y] = \lambda$. We make this a regression model by formulating the conditional mean function

$$\lambda(x) = \exp(\alpha + \beta x).$$

(This makes it a 'loglinear model.'). Now, the regression function is

$$E[y|x] = \exp(\alpha + \beta x).$$

Suppose, further, that x is distributed uniformly with density

$$f(x) = 1, 0 \leq x \leq 1.$$

Note that with this assumption, the joint density of y and x is

$$f(y, x) = f(y|x) f(x) = [1/\exp(\alpha + \beta x)] \exp[-y/\exp(\alpha + \beta x)].$$

1. Derive the parameters of the linear projection,

$$P(x) = \delta_0 + \delta_1 x,$$

where $\delta_0 = E[y] - \delta_1 E[x]$ and

$$\delta_1 = \text{Cov}[x, y] / \text{Var}[x].$$

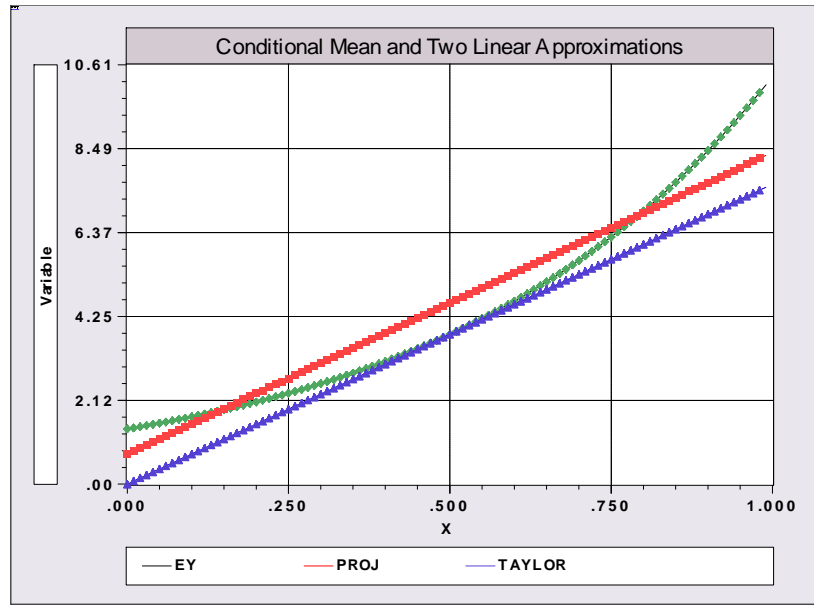
Suppose $\alpha = 1/3$ and $\beta = 2$. What are the values of δ_0 and δ_1 ?

Hint: $E[y] = E_x E[y|x] = \int_0^1 \exp(\alpha + \beta x) \times 1 dx = \exp(\alpha) \int_0^1 \exp(\beta x) dx$ and

$$\text{Cov}[y, x] = \text{Cov}[x, E[y|x]] = E_x \{x \times E[y|x]\} - E[x] E[y] = \exp(\alpha) \int_0^1 x \exp(\beta x) dx - E[x] E[y].$$

Find help at http://en.wikipedia.org/wiki/List_of_integrals_of_exponential_functions

2. Consider the linear Taylor series approximation to the conditional mean function. What are the values of θ_0 and θ_1 in the Taylor series: $E^*[y|x] = \theta_0 + \theta_1 x$ when the expansion point is $E[x] = 1/2$ and as before, $\alpha = 1/3$ and $\beta = 2$.



1. $f(x) = 1, 0 \leq x \leq 1$. $f(y|x) = [1/\lambda(x)]\exp[-y/\lambda(x)], y \geq 0, \lambda(x) = \alpha + \beta x, \alpha = 1/3, \beta = 2$.

$E[x] = 1/2, \text{Var}[x] = 1/12$. (Standard results)

$E[y] = E_x E[y|x] = E[\lambda(x)] = E[\exp(\alpha + \beta x)]$

$E[y] = \exp(\alpha) E[\exp(\beta x)]$

$$= \exp(\alpha) \int_0^1 \exp(\beta x) dx$$

$$= \exp(\alpha) \left[\frac{\exp(\beta x)}{\beta} \Big|_0^1 \right] = \exp(\alpha) \left[\frac{\exp(\beta)}{\beta} - \frac{1}{\beta} \right] = \frac{\exp(\alpha)}{\beta} [\exp(\beta) - 1] = 4.5832$$

$\text{Cov}(x, y) = \text{Cov}(x, E[y|x]) = E[x \exp(\alpha + \beta x)] - E[x]E[y]$

$$= \int_0^1 x \exp(\alpha + \beta x) dx - \frac{1}{2} 4.5832$$

$$= \exp(\alpha) \int_0^1 x \exp(\beta x) dx - \frac{1}{2} 4.5832$$

$$= \exp(\alpha) \left[\frac{\exp(\beta x)}{\beta^2} (\beta x - 1) \Big|_0^1 \right] - \frac{1}{2} 4.5832$$

$$= \exp(\alpha) \left[\frac{\exp(\beta)}{\beta^2} (\beta - 1) - \frac{1}{\beta^2} (-1) \right] - \frac{1}{2} 4.5832$$

$$= \frac{\exp(\alpha)}{\beta^2} [\exp(\beta)(\beta - 1) + 1] - \frac{1}{2} 4.5832 = .697808.$$

The slope is $\text{Cov}(x, y) / \text{Var}[x] = .697808 / (1/12) = 8.37369$

The constant term is $E[y] - \text{slope} \times E[x] = 4.5832 - 8.37369(1/2) = 0.27147$

2. If $E[y|x] = \lambda(x)$, The Taylor series approximation would be

$\lambda(E[x]) + \partial \lambda(x) / \partial x |_{x=E[x]} \times (x - E[x])$

$= \lambda(E[x])[1 - \beta E[x]]$

$E^*(y|x) = \lambda(1/2) + \partial \lambda(x) / \partial x |_{1/2} [x - 1/2]$

$$= \exp(1/3 + 2(1/2)) - (1/2)\beta \lambda(1/2) + \beta \lambda(1/2)x$$

The slope is $2\exp(1/3 + 2(1/2)) = 2\exp(4/3) = 7.587$.
The constant is $\lambda(1/2)[1 - \frac{1}{2}\beta] = 0$.

Part II. Linear Regression Analysis

Data for this exercise are on the course website – please use the “Cornwell and Rupert Returns to Schooling Data.” We begin with the linear regression model (using the variable names in the data set)

$$(*) \text{ LWAGE}_{it} = \beta_1 + \beta_2 \text{OCC}_{it} + \beta_3 \text{SMSA}_{it} + \beta_4 \text{MS}_{it} + \beta_5 \text{FEM}_i + \beta_6 \text{ED}_i + \beta_7 \text{EXP}_{it} + \varepsilon_{it}$$

The dependent variable is log wage. The RHS variables are defined in the data set. Although this is a panel data set, we are going to ignore that aspect and “pool” the data.

1. Compute the linear least squares regression results and report the coefficients, standard errors, ‘t-ratios,’ R^2 , adjusted R^2 , residual standard deviation, and F statistic for testing the joint significance of all the variables in the equation.

2. Test the hypothesis that neither education (ED) nor experience (EXP) is a significant determinant of the expected log wage. Use an F (Wald), likelihood ratio (assuming normality of ε), and a Lagrange multiplier (also assuming normality) test. In each case, document in minute detail exactly how you are computing your results and what conclusion you reach.

3. The model contains a dummy variable for sex, $\text{FEM} = 1$ for female, 0 for male. What is the value of the coefficient on FEM in your estimated model? How do you interpret this value? I.e., what is the economic meaning of the value you computed for this coefficient? Test the hypothesis that this coefficient equals zero.

GENDER DIFFERENCE (PROPORTIONAL). -39%. HYPOTHESIS THAT IT EQUALS ZERO IS REJECTED BASED ON T RATIO OF -15.164.

+-----+-----+-----+-----+-----+-----+					
Ordinary	least squares regression				
LHS=LWAGE	Mean	=	6.676346		
	Standard deviation	=	.4615122		
WTS=none	Number of observs.	=	4165		
Model size	Parameters	=	7		
	Degrees of freedom	=	4158		
Residuals	Sum of squares	=	556.3030		
	Standard error of e	=	.3657745		
Fit	R-squared	=	.3727592		
	Adjusted R-squared	=	.3718541		
Model test	F[6, 4158] (prob)	=	411.84 (.0000)		
Diagnostic	Log likelihood	=	-1717.476		
	Restricted(b=0)	=	-2688.806		
	Chi-sq [6] (prob)	=	1942.66 (.0000)		
Info criter.	LogAmemiya Prd. Crt.	=	-2.009797		
	Akaike Info. Criter.	=	-2.009797		
	Bayes Info. Criter.	=	-1.999151		
+-----+-----+-----+-----+-----+-----+					
Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
Constant	5.66098***	.04685914	120.808	.0000	
OCC	-.11220***	.01464317	-7.662	.0000	.5111645
SMSA	.15504***	.01233744	12.567	.0000	.6537815

MS	.09569***	.02133490	4.485	.0000	.8144058
FEM	-.39478***	.02603413	-15.164	.0000	.1126050
ED	.05688***	.00267743	21.244	.0000	12.845378
EXP	.01044***	.00054206	19.256	.0000	19.853782

Note: ***, **, * = Significance at 1%, 5%, 10% level.

--> MATRIX ; List ; Wald = b2'*<v2>*b2 \$

Matrix WALD has 1 rows and 1 columns.

```

1
+-----+
1| 668.77622

```

--> CALC ; Logl1 = Logl \$

--> Regress ; Lhs = Lwage ; Rhs = One,OCC,SMSA,MS,FEM ; Res = e0 \$

Ordinary	least squares regression
LHS=LWAGE	Mean = 6.676346
	Standard deviation = .4615122
WTS=none	Number of observs. = 4165
Model size	Parameters = 5
	Degrees of freedom = 4160
Residuals	Sum of squares = 645.7792
	Standard error of e = .3939992
Fit	R-squared = .2718733
	Adjusted R-squared = .2711731
Model test	F[4, 4160] (prob) = 388.32 (.0000)
Diagnostic	Log likelihood = -2028.070
	Restricted(b=0) = -2688.806
	Chi-sq [4] (prob) = 1321.47 (.0000)
Info criter.	LogAmemiya Prd. Crt. = -1.861613
	Akaike Info. Criter. = -1.861613
	Bayes Info. Criter. = -1.854009
Autocorrel	Durbin-Watson Stat. = .7730622
	Rho = cor[e,e(-1)] = .6134689
Model was estimated Feb 10, 2009 at 06:49:39AM	

Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
Constant	6.61346***	.02495768	264.987	.0000	
OCC	-.28539***	.01249653	-22.838	.0000	.5111645
SMSA	.19827***	.01316629	15.059	.0000	.6537815
MS	.15351***	.02274471	6.749	.0000	.8144058
FEM	-.40738***	.02799278	-14.553	.0000	.1126050

Note: ***, **, * = Significance at 1%, 5%, 10% level.

--> Create ; e02 = e0*e0 \$

--> CALC ; Logl0 = Logl \$

--> Matrix ; List ; LMtest = e0'XPart2 * <XPart2'[E02]XPart2> * Xpart2'e0 \$

Matrix LMTEST has 1 rows and 1 columns.

```

1
+-----+
1| 411.91215

```

--> CALC ; list ; LR = 2*(Logl1 - Logl0) \$

Listed Calculator Results

LR = 621.187281

Part III. Structural Change

The implication of the specification of FEM in the model in Part II is that the extent of the difference between men and women is captured in a shift of the regression function (based on a change in the intercept alone). Consider, instead, the hypothesis that different regression functions apply to men and women. Fit the model separately for men and women, then use a Chow test to test the null hypothesis that the same equation applies to men and women. (Note, for purposes of this exercise, your model will not contain the FEM variable.) The model is

$$(**) \quad \text{LWAGE}_{it} = \beta_1 + \beta_2 \text{OCC}_{it} + \beta_3 \text{SMSA}_{it} + \beta_4 \text{MS}_{it} + \beta_5 \text{ED}_i + \beta_6 \text{EXP}_{it} + \varepsilon_{it}$$

Completely document your analysis. Include in your results a table that shows the results of the three regressions, male, female and pooled, so that the reader can easily see the comparison of the estimated coefficients. What is the result of the test?

Looking ahead to our work in panel data modeling, repeat this analysis for the 7 years of data in the sample. That is, compute the regression in (**) using the full pooled data set, then again for each of the 7 years. (There are 595 observations for each of the 7 years.) Using a Chow (F) test, test the null hypothesis that the same model applies to all 7 years. To investigate whether a structural change might be explained by a simple shift of the function, fit the model

$$(***) \quad \text{LWAGE}_{it} = \beta_1 + \beta_2 \text{OCC}_{it} + \beta_3 \text{SMSA}_{it} + \beta_4 \text{MS}_{it} + \beta_5 \text{ED}_i + \beta_6 \text{EXP}_{it} + \gamma_1 T_{2,t} + \dots \gamma_6 T_{6,t} + \varepsilon_{it}$$

where $T_{2,t}, \dots, T_{6,t}$ are 6 dummy variables for the 6 years, omitting the first. Test the null hypothesis that the 6 dummy variable coefficients all equal zero and report all results. Interpret your findings.

```
--> NAMELIST ; XPart3 = One,OCC,SMSA,MS,ED,EXP $
--> REGRESS ; Lhs = Lwage ; Rhs = XPart3 $
```

```
+-----+
| Ordinary   least squares regression
| LHS=LWAGE  Mean                = 6.676346
|            Standard deviation  = .4615122
| WTS=none   Number of observs.  = 4165
| Model size Parameters          = 6
|            Degrees of freedom  = 4159
| Residuals  Sum of squares      = 587.0679
|            Standard error of e = .3757074
| Fit        R-squared           = .3380713
|            Adjusted R-squared  = .3372755
| Model test F[ 5, 4159] (prob) = 424.83 (.0000)
| Diagnostic Log likelihood      = -1829.572
|            Restricted(b=0)     = -2688.806
|            Chi-sq [ 5] (prob) =1718.47 (.0000)
| Info criter. LogAmemiya Prd. Crt. = -1.956450
|            Akaike Info. Criter. = -1.956450
|            Bayes Info. Criter.  = -1.947325
| Autocorrel Durbin-Watson Stat. = .7555981
|            Rho = cor[e,e(-1)]   = .6222010
| Model was estimated Feb 10, 2009 at 06:56:48AM
+-----+
```

```
+-----+
```

Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
Constant	5.40421***	.04487911	120.417	.0000	
OCC	-.09787***	.01500945	-6.521	.0000	.5111645
SMSA	.14884***	.01266551	11.752	.0000	.6537815
MS	.32660***	.01534942	21.277	.0000	.8144058
ED	.05882***	.00274699	21.412	.0000	12.845378
EXP	.01024***	.00055661	18.398	.0000	19.853782

Note: ***, **, * = Significance at 1%, 5%, 10% level.

--> CALC ; SSPool = sumsqdev \$

--> REGRESS ; For[FEM=0] ; Lhs = Lwage ; Rhs = XPart3 \$

```
*****
* Setting up an iteration over the values of FEM      *
* The model command will be executed for 1 values    *
* of this variable. In the current sample of 4165    *
* observations, the following counts were found:      *
* Subsample Observations Subsample Observations    *
* FEM = 0 3696 FEM = **** *
*-----*
* Actual subsamples may be smaller if missing values *
* are being bypassed. Subsamples with 0 observations *
* will be bypassed. *
*****
```

```
*****
* Subsample analyzed for this command is FEM = 0 *
*****
```

Ordinary least squares regression	
LHS=LWAGE	Mean = 6.729774
	Standard deviation = .4382202
WTS=none	Number of observs. = 3696
Model size	Parameters = 6
	Degrees of freedom = 3690
Residuals	Sum of squares = 503.8896
	Standard error of e = .3695341
Fit	R-squared = .2898728
	Adjusted R-squared = .2889105
Model test	F[5, 3690] (prob) = 301.25 (.0000)
Diagnostic	Log likelihood = -1561.981
	Restricted(b=0) = -2194.572
	Chi-sq [5] (prob) = 1265.18 (.0000)
Info criter.	LogAmemiya Prd. Crt. = -1.989402
	Akaike Info. Criter. = -1.989402
	Bayes Info. Criter. = -1.979313
Autocorrel	Durbin-Watson Stat. = .7604481
	Rho = cor[e,e(-1)] = .6197759
Model was estimated Feb 10, 2009 at 06:56:48AM	

Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
Constant	5.63563***	.05029666	112.048	.0000	
OCC	-.08984***	.01576766	-5.697	.0000	.5262446
SMSA	.15905***	.01299522	12.239	.0000	.6360931
MS	.09638***	.02199861	4.381	.0000	.9145022
ED	.05707***	.00286496	19.920	.0000	12.846591
EXP	.01083***	.00058000	18.677	.0000	20.214015

Note: ***, **, * = Significance at 1%, 5%, 10% level.

--> CALC ; SSMen = sumsqdev \$

--> REGRESS ; For[FEM=1] ; Lhs = Lwage ; Rhs = XPart3 \$

```
*****
* Setting up an iteration over the values of FEM      *
* The model command will be executed for 1 values    *
* of this variable. In the current sample of 4165    *
* observations, the following counts were found:      *
* Subsample Observations Subsample Observations    *
*****
```

```

* FEM      =      1      469      FEM      =****      *
* -----*
* Actual subsamples may be smaller if missing values *
* are being bypassed.  Subsamples with 0 observations *
* will be bypassed. *
*****

```

```

*****
* Subsample analyzed for this command is FEM      =      1      *
*****

```

```

+-----+
| Ordinary least squares regression |
| LHS=LWAGE Mean = 6.255308 |
| Standard deviation = .4227426 |
| WTS=none Number of observs. = 469 |
| Model size Parameters = 6 |
| Degrees of freedom = 463 |
| Residuals Sum of squares = 47.90453 |
| Standard error of e = .3216605 |
| Fit R-squared = .4272321 |
| Adjusted R-squared = .4210467 |
| Model test F[ 5, 463] (prob) = 69.07 (.0000) |
| Diagnostic Log likelihood = -130.4956 |
| Restricted(b=0) = -261.1765 |
| Chi-sq [ 5] (prob) = 261.36 (.0000) |
| Info criter. LogAmemiya Prd. Crt. = -2.255805 |
| Akaike Info. Criter. = -2.255806 |
| Bayes Info. Criter. = -2.202707 |
| Autocorrel Durbin-Watson Stat. = .8105792 |
| Rho = cor[e,e(-1)] = .5947104 |
| Model was estimated Feb 10, 2009 at 06:56:48AM |
+-----+

```

Variable	Coefficient	Standard Error	t-ratio	P[T >t]	Mean of X
Constant	5.26320***	.11046622	47.645	.0000	
OCC	-.25591***	.03915105	-6.536	.0000	.3923241
SMSA	.12504***	.03921038	3.189	.0015	.7931770
MS	-.00121	.09661168	-.013	.9900	.0255864
ED	.06611***	.00754249	8.765	.0000	12.835821
EXP	.00851***	.00154633	5.503	.0000	17.014925

Note: ***, **, * = Significance at 1%, 5%, 10% level.

```

--> CALC ; SSWomen = Sumsqdev $
--> CALC ; K = Col(XPart3) $
--> CALC ; List ; Chow = ((SSPool - (SSMen+SSWomen))/ K) /
      ((SSMen + SSWomen) / (N - 2*K)) $
+-----+
| Listed Calculator Results |
+-----+
CHOW      =      44.247081
--> calc;LIST;FTB(.95,K,(N-2*K))$
+-----+
| Listed Calculator Results |
+-----+
Result    =      2.100770

```

```

CREATE ; T = Trn(-7,0)$
CALC ; SST = 0 $
Procedure
REGRESS ; For[T=j] ; Lhs = Lwage ; Rhs = Xpart3 ; quietly $
CALC ; SST = SST + Sumsqdev $
ENDPROC
Exec ; J = 1,7 $
CALC ; List ; Chow = ((SSPool - SST)/(6*K)) / (SST/(N-7*K)) $
--> CALC ; List ; Chow = ((SSPool - SST)/(6*K)) / (SST/(N-7*K)) $
+-----+
| Listed Calculator Results |
+-----+
CHOW      =      34.789262
--> CALC ; List ; FTB(.95,(6*K),(n-7*K)) $
+-----+
| Listed Calculator Results |
+-----+
Result    =      1.419543

CALC ; List ; FTB(.95,(6*K),(n-7*K)) $
CREATE ; T1 = t=1 ; T2 = t=2 ; T3 = t=3 ; T4 = t=4 ; T5 = t=5 ; T6 = T=6 $
REGRESS ; Lhs = Lwage ; Rhs = XPart3 $
CALC ; SS0 = Sumsqdev $
REGRESS ; Lhs = Lwage ; Rhs = XPart3,t1,t2,t3,t4,t5,t6 $
CALC ; List ; FStat = ((SS0 - Sumsqdev)/6) / (sumsqdev/(n-k-6)) $
--> CALC ; List ; FStat = ((SS0 - Sumsqdev)/6) / (sumsqdev/(n-k-6)) $
+-----+
| Listed Calculator Results |
+-----+
FSTAT     =      203.701855

```

Part IV. A Nonlinear Regression

1. The model (*) above omits a well known phenomenon with respect to the association of wages and experience – earnings often do not increase uniformly with experience, but rather increase more rapidly in the earlier years of employment than in the later years. Test this theory by adding EXP^2 to your model. The equation is

$$\begin{aligned}
 (****) \text{ LWAGE}_{it} = & \beta_1 + \beta_2 \text{OCC}_{it} + \beta_3 \text{SMSA}_{it} + \beta_4 \text{MS}_{it} + \beta_5 \text{FEM}_i + \beta_6 \text{ED}_i \\
 & + \gamma \text{EXP}_{it} + \delta \text{EXP}_{it}^2 + \varepsilon_{it}
 \end{aligned}$$

Refit the model by least squares and discuss your results. Use the entire sample. Does squared experience help to explain the variation in log wages? Test the null hypothesis that it does not. What do you find?

2. Partial Effect. As part of your analysis, derive and statistically analyze the partial effect of experience,

$$\theta(\text{EXP}_{it}) = \partial E[\text{LWAGE}|\mathbf{x}]/\partial \text{EXP} = \gamma + 2\delta \text{EXP}_{it}$$

at the sample mean value of EXP_{it} . Compute an asymptotic standard error for the estimator of θ then test the hypothesis that θ equals zero.

3. Examining the Regression. Obtain the sample mean values of OCC, SMSA, MS ED. Then, using your estimated coefficients, compute

$$\hat{\alpha} = \hat{\beta}_1 + \hat{\beta}_2 \overline{\text{OCC}} + \hat{\beta}_3 \overline{\text{SMSA}} + \hat{\beta}_4 \overline{\text{MS}} + \hat{\beta}_6 \overline{\text{ED}}.$$

We are interested in what the regression model implies about the trajectory of wages as a function of experience. Thus, we want to plot

$$\widehat{WAGE}(Exp | Female) = \exp(\hat{\alpha} + \hat{\beta}_5 + \hat{\gamma}Exp + \hat{\delta}Exp^2)$$

and

$$\widehat{WAGE}(Exp | Male) = \exp(\hat{\alpha} + \hat{\gamma}Exp + \hat{\delta}Exp^2)$$

What do you find? Interpret the figure.

HINT: Here are NLOGIT commands that you can use to do this computation

```
regress;lhs=lwage;rhs=one,occ,smsa,ms,fem,ed,exp,expsq$
calc;occbar=xbr(occ);smsabar=xbr(smsa);msbar=xbr(ms);edbar=xbr(ed)$
calc;ahat=b(1)+b(2)*occbar+b(3)*smsabar+b(4)*msbar+b(6)*edbar$
samp;1-51$
create;fitexp=trn(1,1)$
create;wagef=exp(ahat + b(5) + b(7)*fitexp+b(8)*fitexp^2)$
create;wagem=exp(ahat + b(7)*fitexp+b(8)*fitexp^2)$
plot;lhs=fitexp;rhs=wagef,wagem;fil
;title =Wage Trajectories for Men and Women Based on Experience
;Vaxis =Projected Weekly Earnings
;Grid$
```

+-----+-----+-----+-----+			
Ordinary	least squares regression		
LHS=LWAGE	Mean	=	6.676346
	Standard deviation	=	.4615122
WTS=none	Number of observs.	=	4165
Model size	Parameters	=	8
	Degrees of freedom	=	4157
Residuals	Sum of squares	=	529.2850
	Standard error of e	=	.3568246
Fit	R-squared	=	.4032224
	Adjusted R-squared	=	.4022174
Model test	F[7, 4157] (prob)	=	401.25 (.0000)
Diagnostic	Log likelihood	=	-1613.797
	Restricted(b=0)	=	-2688.806
	Chi-sq [7] (prob)	=	2150.02 (.0000)
Info criter.	LogAmemiya Prd. Crt.	=	-2.059103
	Akaike Info. Criter.	=	-2.059103
	Bayes Info. Criter.	=	-2.046936
Autocorrel	Durbin-Watson Stat.	=	.7637078
	Rho = cor[e,e(-1)]	=	.6181461
Model was estimated Feb 10, 2009 at 07:00:17AM			
+-----+-----+-----+-----+			

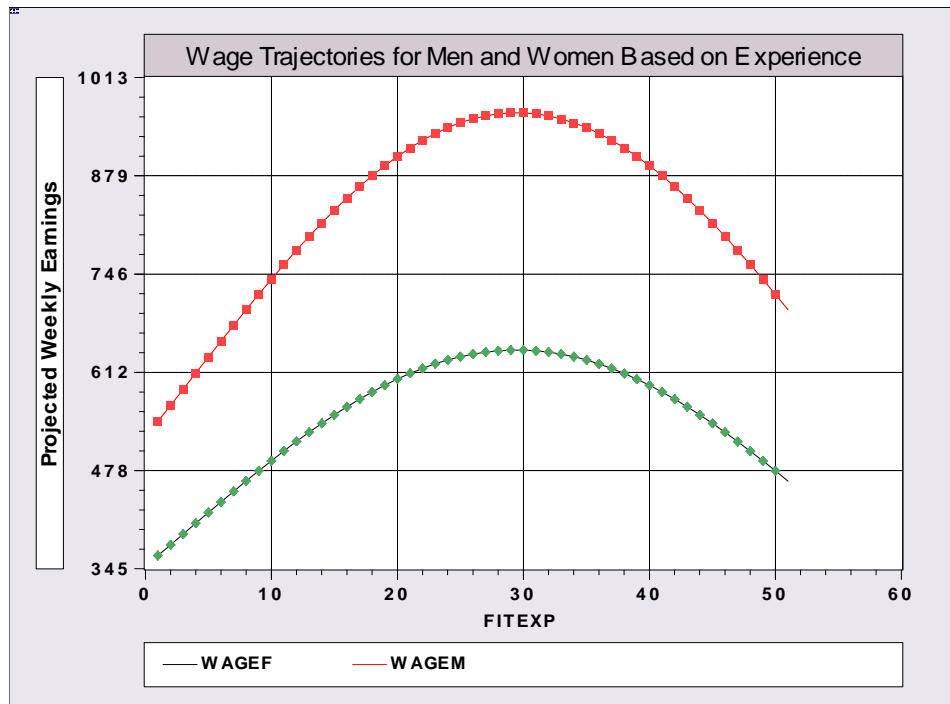
+-----+-----+-----+-----+-----+-----+					
Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
+-----+-----+-----+-----+-----+-----+					
Constant	5.41799***	.04866099	111.342	.0000	
OCC	-.10684***	.01428962	-7.477	.0000	.5111645
SMSA	.16053***	.01204145	13.331	.0000	.6537815
MS	.07571***	.02085803	3.630	.0003	.8144058
FEM	-.40722***	.02541147	-16.025	.0000	.1126050
ED	.05683***	.00261192	21.758	.0000	12.845378
EXP	.04152***	.00219803	18.888	.0000	19.853782
EXPSQ	-.00070***	.483146D-04	-14.567	.0000	514.40504
+-----+-----+-----+-----+-----+-----+					
Note: nnnnn.D-xx or D+xx => multiply by 10 to -xx or +xx.					
Note: ***, **, * = Significance at 1%, 5%, 10% level.					
+-----+-----+-----+-----+-----+-----+					

```
--> CALC ; EXpBar = Xbr(EXP) $
--> Calc ; list ; Theta = b(7)+2*b(8)*Expbar $
```

```

+-----+
| Listed Calculator Results |
+-----+
THETA = .013570
--> Calc ; list ; stheta = sqrt(Varb(7,7) + 4*expbar*expbar*varb(8,8) + 4*Expb...
+-----+
| Listed Calculator Results |
+-----+
STHETA = .000571
--> Calc ; list ; ttheta = theta/stheta $
+-----+
| Listed Calculator Results |
+-----+
TTHETA = 23.772215

```



Part V. Nonlinear Function of Parameters

Based on the regression model in part IV,

$$\begin{aligned}
 \text{(***) } LWAGE_{it} = & \beta_1 + \beta_2 OCC_{it} + \beta_3 SMSA_{it} + \beta_4 MS_{it} + \beta_5 FEM_i + \beta_6 ED_i + \\
 & EXP_{it} + \delta EXP_{it}^2 + \varepsilon_{it}
 \end{aligned}$$

In Part IV, you plotted the trajectory of WAGE (not log WAGE) using the mean values of the variables in the model. Note that the figure shows a parabola with a maximum at about 29 years. We are interested in exploring the computation of the peak earning year.

TIP. For purposes of this exercise, you will find it convenient to use

$$\begin{aligned}
 EXP100 &= EXP/100 \\
 \text{and} \\
 EXP100SQ &= EXP100^2
 \end{aligned}$$

for your regression. This scaling will make the standard errors that you need to use much more convenient but will, of course, not change the model.

$$(\text{*****}) \text{ LWAGE}_{it} = \beta_1 + \beta_2 \text{OCC}_{it} + \beta_3 \text{SMSA}_{it} + \beta_4 \text{MS}_{it} + \beta_5 \text{FEM}_i + \beta_6 \text{ED}_i + \gamma \text{EXP100}_{it} + \delta \text{EXP100}_{it}^2 + \epsilon_{it}$$

Refit the model, using EXP100 and EXP100SQ. Our prediction of WAGE is

$$\widehat{\text{WAGE}}(\text{Exp} | \text{Female}) = \exp(\hat{\alpha} + \hat{\beta}_5 + \hat{\gamma} \text{Exp100} + \hat{\delta} \text{Exp100}^2)$$

1. Prove that the maximum of this function occurs at $\text{EXP100}^* = -\hat{\gamma}/2\hat{\delta}$. We are interested in estimating and forming a confidence interval for EXP100*.
2. Compute the value of EXP100* using the results of your regression.
3. Use the delta method to obtain estimated asymptotic standard error for EXP100*.

HINTS: You can use the following NLOGIT commands to do the regression.

```
sample;all$
create;exp100=exp/100 ; exp100sq=exp100^2$
regress;lhs=lwage;rhs=one,occ,smsa,ms,fem,ed,exp100,exp100sq$
```

The rest of the computations can be done with a hand calculator or with the CALC command. After you compute the regression, go into the project window, open the Matrices list, then double click on VARB to show the asymptotic covariance matrix.

```
--> calc ; list ; max = -b(7)/(2*b(8)) $
+-----+
| Listed Calculator Results |
+-----+
MAX      =      .294943
--> calc ; g1 = -1/(2*b(8)) ; g2 = -max/b(8) $
--> calc ; smax = sqrt(g1*g1*varb(7,7)+g2*g2*varb(8,8)+2*g1*g2*varb(7,8))$
--> calc ; list ; max/smax $
+-----+
| Listed Calculator Results |
+-----+
Result    =      46.619582
--> wald ; start = b ; var = varb ; labels = 8_c ; fn1=-c7/(2*c8)$
```

```
+-----+
| WALD procedure. Estimates and standard errors |
| for nonlinear functions and joint test of    |
| nonlinear restrictions.                      |
| Wald Statistic          =      2173.38527    |
| Prob. from Chi-squared[ 1] =      .00000    |
+-----+
+-----+-----+-----+-----+-----+
|Variable| Coefficient | Standard Error |b/St.Er.|P[|Z|>z]|
+-----+-----+-----+-----+-----+
|Fncn(1) |      .29494*** |      .00632660 |  46.620 | .0000 |
+-----+-----+-----+-----+-----+
```

4. We will now use the method of Krinsky and Robb as an alternative to the delta method. The method proceeds as follows. We are interested in analyzing a nonlinear function of $\hat{\gamma}$ and $\hat{\delta}$. We have estimated the 2×2 asymptotic covariance matrix for this pair of estimators; call it Σ . Our estimators are asymptotically normally distributed with mean (γ, δ) and asymptotic covariance matrix Σ . What we will do is draw a large random sample from this population, (γ_r, δ_r) , $r = 1, \dots, R$, then compute from this sample, a sample of values $\text{EXP}_r = -\gamma_r / (2\delta_r)$. We will then use the empirical standard deviation from the sample of draws as our estimator of the asymptotic standard deviation of the estimator of EXP.

How to draw a random sample from this population: We need a sample of draws of the form

$$\gamma_r = \hat{\gamma} + w_{\gamma,r} \text{ and } \delta_r = \hat{\delta} + w_{\delta,r} \text{ then } \text{EXP}_r^* = -\gamma_r / (2\delta_r)$$

where $(w_{\gamma,r}, w_{\delta,r})$ have bivariate normal distribution with mean vector $(0, 0)$ and covariance matrix Σ . Here is how to do that. We will use the Cholesky decomposition of Σ . L is a lower triangular matrix such that $LL' = \Sigma$. Let $v_{1,r}$ and $v_{2,r}$ be samples of independent draws from the standard normal distribution. Then,

$$w_{\gamma,r} = L_{11}v_{1,r} \text{ and } w_{\delta,r} = L_{21}v_{1,r} + L_{22}v_{2,r}.$$

- Let σ_{11} = the asymptotic variance of $\hat{\gamma}$, σ_{22} = the asymptotic variance of $\hat{\delta}$ and let σ_{12} = the asymptotic covariance.
Show that $L_{11} = \text{sqr}(\sigma_{11})$, $L_{21} = \sigma_{12}/L_{11}$ and $L_{22} = \text{sqr}(\sigma_{22} - \sigma_{12}^2/\sigma_{11})$.
- Compute the random sample of draws on v_1 and v_2 .
- Compute the random sample of draws on γ and δ .
- Compute the random sample of draws on EXP^*
- Compute the standard error for your estimate of EXP^* using the sample standard deviation.
- Compare your result to the results using the delta method in part 3.

HINT: This set of NLOGIT commands does the computation after the regression.

```
calc      ;sgg=varb(7,7);sgd=varb(8,7);sdd=varb(8,8)$
calc      ;list; L11=sqr(sgg) ; L21=sgd/sqr(sgg) ; L22=sqr(sdd-sgd^2/sgg)$
create    ;u1=rnn(0,1);u2=rnn(0,1)$
create    ;gr=b(7) + L11*u1 ; dr=b(8) + L21*u1 + L22*u2 $
create    ;expr = -gr/(2*dr)$
dstat     ;rhs=expr$
? This command uses the delta method after the regression
wald      ;start=b;var=varb;labels=8_b ; fn1=-b7/(2*b8)$
```

```

--> create ;u1=rnn(0,1);u2=rnn(0,1)$
--> create ;gr=b(7) + L11*u1 ; dr=b(8) + L21*u1 + L22*u2 $
--> create ;expr = -gr/(2*dr)$
--> dstat ;rhs=expr$
Descriptive Statistics
All results based on nonmissing observations.
=====
Variable      Mean      Std.Dev.      Minimum      Maximum      Cases Missing
=====
All observations in current sample
-----+-----
EXPR | .295373      .635183E-02  .277729      .327965      4165      0
--> wald ;start=b;var=varb;labels=8_b ; fn1=-b7/(2*b8)$

+-----+
| WALD procedure. Estimates and standard errors |
| for nonlinear functions and joint test of    |
| nonlinear restrictions.                      |
| Wald Statistic      =      2173.38527        |
| Prob. from Chi-squared[ 1] =      .00000    |
+-----+

+-----+-----+-----+-----+-----+
|Variable| Coefficient | Standard Error |b/St.Er.|P[|Z|>z]|
+-----+-----+-----+-----+-----+
|Fncn(1) | .29494***   | .00632660      |46.620  |.0000 |
+-----+-----+-----+-----+-----+
| Note: ***, **, * = Significance at 1%, 5%, 10% level. |
+-----+

```