

ECONOMETRICS I

Take Home Final Examination

Fall 2005

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Today is Thursday, December 8, 2005. This exam is due by 4PM, Monday, December 19, 2005. Feel free to turn it in early, of course. You may submit your answers to me electronically as an attachment to an e-mail if you wish. Your submission for this examination is to be a single authored project – you are assumed to be working alone. Please submit this page, with your signature, with your answers to the questions.

NOTE: In the empirical results below, a number of the form .nnnnnnE+aa means multiply the number .nnnnnn by 10 to the aa power. E-aa implies multiply 10 to the minus aa power. Thus, .123456E-02 is 0.00123456.

This test comprises 100 points. The 8 parts below are each worth 10% of the exam. The remaining 20% is given for the correct entry in the signature line below.

This course is governed by the Stern honor code:

I will not lie, cheat or steal to gain an academic advantage, or tolerate those who do.

Signature _____

Part 1.

In the first assigned problem set for this course, we used a data set on production by a sample of Spanish dairy farms. The data are a panel, $T=6$ years (1993 – 1998) on $N=247$ farms, a total of 1,482 observations. The variables in the data set are:

MILK	= liters of output;	YIT	= $\log(\text{milk})$
COWS		X1	= $\log(\text{COWS})$
LAND		X2	= $\log(\text{LAND})$
LABOR		X3	= $\log(\text{LABOR})$
FEED		X4	= $\log(\text{FEED})$

The data have been transformed so that the means of X1, X2, X3 and X4 are all zero. (Note, means of the logs). We are also using additional variables in this exercise,

X11, X12, X13, X14	= $X1*X1, X1*X2, X1*X3, X1*X4$
X22, X23, X24	
X33, X34	
X44	and so on. Squares and cross products
T	= Time = -2, -1, 0, 1, 2, 3
T2	= Time ²

The Cobb-Douglas model of production specifies $\log Y = \alpha + \sum_k \beta_k \log X_k$. The production function displays constant returns to scale if $\sum_k \partial \log Y / \partial \log X_k = 1 = \sum_k \beta_k$

The translog model of production specifies $\log Y = \alpha + \sum_k \beta_k \log X_k + \sum_k \sum_{m \geq k} \gamma_{km} \log X_k \log X_m$

Technical change in a logarithmic production function is represented by $\log Y = *** + \delta_0 t + \delta_1 t^2$.

The rate of technical change is then $\partial \log Y / \partial t = \delta_0 + 2\delta_1 t$. A constant rate of technical change is imposed by constraining $\delta_1 = 0$.

Results for various specifications of the production function, estimated by ordinary least squares, are presented below. Your responses to this part of the test (and a few others) will be based on these results.

- Test the null hypothesis that the technology is Cobb-Douglas against the alternative that it is translog. Show exactly how you are doing the test and on what statistic you are basing your conclusion.
- Test the hypothesis of constant returns to scale in the Cobb-Douglas model using (1) a simple t-test, (2) a Wald test, (3) an F test and (4) a likelihood ratio test. Explain in excruciatingly clear detail exactly how you are carrying out the tests. In your presentation, state clearly what assumptions about the model underlie the test statistics.
- Form a 95% confidence interval for the “cows elasticity,” β_1 . As usual, explain clearly what you are doing.
- I want to test the hypothesis that the rate of technical change is zero in the fourth year, but not necessarily in the other years. Show how to carry out the test. Do so. Now, I want to test the hypothesis that technical change is zero in every year. Again, explain the test procedure, and carry out the test.

Regressions for Part 1.

+-----+			
Ordinary	least squares regression		
LHS=YIT	Mean	=	11.57749
	Standard deviation	=	.6434377
WTS=none	Number of observs.	=	1482
Model size	Parameters	=	5
	Degrees of freedom	=	1477
Residuals	Sum of squares	=	29.09570
	Standard error of e	=	.1403538
Fit	R-squared	=	.9525473
	Adjusted R-squared	=	.9524188
Model test	F[4, 1477] (prob)	=	7412.19 (.0000)
Diagnostic	Log likelihood	=	809.6761
	Restricted(b=0)	=	-1448.908
	Chi-sq [4] (prob)	=	4517.17 (.0000)
+-----+			

+-----+					
Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
+-----+					
Constant	11.5774868	.00364586	3175.515	.0000	
X1	.59517558	.01958331	30.392	.0000	0
X2	.02305014	.01122274	2.054	.0400	0
X3	.02319244	.01303099	1.780	.0751	0
X4	.45175783	.01078465	41.889	.0000	0

Covariance Matrix

	1	2	3	4	5
+-----+					
1	.1329230D-04	0	0	0	0
2	0	.00038	-.00011	-.5457328D-04	-.00017
3	0	-.00011	.00013	.1001475D-04	.2018432D-04
4	0	-.5457328D-04	.1001475D-04	.00017	-.1404125D-04
5	0	-.00017	.2018432D-04	-.1404125D-04	.00012

+-----+			
Linearly restricted regression			
LHS=YIT	Mean	=	11.57749
	Standard deviation	=	.6434377
WTS=none	Number of observs.	=	1482
Model size	Parameters	=	4
	Degrees of freedom	=	1478
Residuals	Sum of squares	=	30.19235
	Standard error of e	=	.1429260
Fit	R-squared	=	.9507588
	Adjusted R-squared	=	.9506588
Model test	F[3, 1478] (prob)	=	9512.50 (.0000)
Diagnostic	Log likelihood	=	782.2604
	Restricted(b=0)	=	-1448.908
	Chi-sq [3] (prob)	=	4462.34 (.0000)
Restrictns.	F[1, 1477] (prob)	=	55.67 (.0000)
Not using OLS or no constant. Rsqd & F may be < 0.			
Note, with restrictions imposed, Rsqd may be < 0.			
+-----+			

+-----+					
Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
+-----+					
Constant	11.5774868	.00371268	3118.366	.0000	
X1	.56677311	.01956182	28.973	.0000	0
X2	-.00282763	.01086898	-.260	.7947	0
X3	-.04325035	.00968783	-4.464	.0000	0
X4	.47930487	.01031870	46.450	.0000	0

Matrix Cov.Mat. has 5 rows and 5 columns.

	1	2	3	4	5
+-----+					
1	.1378397D-04	0	0	0	0
2	0	.00038	-.00013	-.9174458D-04	-.00016
3	0	-.00013	.00012	-.2164277D-04	.3420965D-04
4	0	-.9174458D-04	-.2164277D-04	.9385403D-04	.1953333D-04
5	0	-.00016	.3420965D-04	.1953333D-04	.00011

```

+-----+
| Ordinary   least squares regression
| Model was estimated Dec 07, 2005 at 10:35:58AM
| LHS=YIT      Mean           = 11.57749
|              Standard deviation = .6434377
| WTS=none     Number of observs. = 1482
| Model size   Parameters      = 7
|              Degrees of freedom = 1475
| Residuals    Sum of squares   = 28.94244
|              Standard error of e = .1400785
| Fit          R-squared        = .9527973
|              Adjusted R-squared = .9526053
| Model test   F[ 6, 1475] (prob) =4962.20 (.0000)
| Diagnostic   Log likelihood    = 813.5895
|              Restricted(b=0)    = -1448.908
|              Chi-sq [ 6] (prob) =4525.00 (.0000)
| Info criter. LogAmemiya Prd. Crt. = -3.926392
|              Akaike Info. Criter. = -3.926392
| Autocorrel   Durbin-Watson Stat. = .8081626
|              Rho = cor[e,e(-1)] = .5959187
+-----+

```

```

+-----+
|Variable | Coefficient | Standard Error |b/St.Er.|P[|Z|>z] | Mean of X|
+-----+
|Constant | 11.5799428 | .00543433 | 2130.888 |.0000 |
|X1        | .59987916 | .01962113 | 30.573 |.0000 | 0
|X2        | .02390494 | .01120536 | 2.133 |.0329 | 0
|X3        | .02737137 | .01309626 | 2.090 |.0366 | 0
|X4        | .44479994 | .01106948 | 40.183 |.0000 | 0
|T         | .00745280 | .00267869 | 2.782 |.0054 | .50000000
|T2        | -.00195232 | .00146170 | -1.336 |.1817 | 3.16666667
+-----+
Estimated Asymptotic covariance matrix.

```

```

          1          2          3          4          5          6          7
1 2.95E-05 -6.16E-08 6.64E-07 -1.32E-07 5.48E-07 3.34E-06 -5.67E-06
2 -6.16E-08 0.000385 -0.00011 -5.17E-05 -0.00017 4.39E-06 -6.74E-07
3 6.64E-07 -0.00011 0.000126 1.04E-05 1.94E-05 8.41E-07 -3.43E-07
4 -1.32E-07 -5.17E-05 1.04E-05 0.000172 -1.80E-05 3.89E-06 -5.73E-07
5 5.48E-07 -0.00017 1.94E-05 -1.80E-05 0.000123 -6.46E-06 8.48E-07
6 3.34E-06 4.39E-06 8.41E-07 3.89E-06 -6.46E-06 7.18E-06 -2.19E-06
7 -5.67E-06 -6.74E-07 -3.43E-07 -5.73E-07 8.48E-07 -2.19E-06 2.14E-06

```

```

+-----+
| Ordinary least squares regression
| Model was estimated Dec 07, 2005 at 10:40:26AM
| LHS=YIT      Mean      = 11.57749
|               Standard deviation = .6434377
| WTS=none     Number of observs. = 1482
| Model size   Parameters = 15
|               Degrees of freedom = 1467
| Residuals    Sum of squares = 28.28359
|               Standard error of e = .1388520
| Fit          R-squared = .9538718
|               Adjusted R-squared = .9534316
| Model test   F[ 14, 1467] (prob) =2166.83 (.0000)
| Diagnostic   Log likelihood = 830.6529
|               Restricted(b=0) = -1448.908
|               Chi-sq [ 14] (prob) =4559.12 (.0000)
| Info criter. LogAmemiya Prd. Crt. = -3.938623
|               Akaike Info. Criter. = -3.938623
| Autocorrel   Durbin-Watson Stat. = .8282196
|               Rho = cor[e,e(-1)] = .5858902
+-----+

```

Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
Constant	11.5688643	.00726864	1591.613	.0000	
X1	.60693062	.02186356	27.760	.0000	0
X2	.01351936	.01169178	1.156	.2476	0
X3	.02384747	.01355872	1.759	.0786	0
X4	.45379253	.01199340	37.837	.0000	0
X11	.47329234	.14309910	3.307	.0009	.11926376
X22	-.08046399	.04929507	-1.632	.1026	.10386698
X33	-.04840054	.09251157	-.523	.6008	.05873631
X44	.17968695	.04555562	3.944	.0001	.28859636
X12	-.08379987	.06166837	-1.359	.1742	.15316763
X13	.18430294	.07247569	2.543	.0110	.09495984
X14	-.28574011	.07559639	-3.780	.0002	.33056414
X23	-.00815564	.04326168	-.189	.8505	.05287598
X24	.05222215	.03095889	1.687	.0916	.19230806
X34	-.05821369	.04040862	-1.441	.1497	.14262419

Part 2.

- Ordinary least squares was used to compute the regressions analyzed in Part 1. Show (algebraically) how the least squares coefficient estimator, \mathbf{b} , and the estimated asymptotic covariance matrix are computed. (Theoretically, not for each regression.)
- Show how each of the values in the box above the coefficient estimates in the first regression is computed, and interpret the value given. (Again, theoretically.)
- What are the finite sample properties of the least squares estimator? Make your assumptions explicit.
- What are the asymptotic properties of the least squares estimator? Again, be explicit about all assumptions, and explain your answer carefully.

Part 3.

As noted, the Spanish dairy data are a panel, with six years of data. The results below show the six separate regressions when the Cobb-Douglas model is fit separately in each year.

- Theory 1 states that the coefficient vectors are the same in the all periods. Is there an optimal way that I could combine these six estimators to form a single efficient estimator of the model parameters? How should I do that? Show the computations explicitly. (Show the theoretical result, not the numbers.)
- Use a Chow test to test the hypothesis that the six coefficient vectors are the same. Explain the computations in full detail so that I know exactly how you obtained your result.
- The test in the preceding question could be done with a Wald test. Is there any particular reason to use the Wald test or the Chow test – i.e., one and not the other in this setting? What assumptions would justify each? Do the regression results suggest that one or the other test might be appropriate? Explain. (You need not carry out the Wald test. This question asks about the test, in principle.)
- The residual vectors from the six regressions described here are collected after we compute the least squares coefficients. This produces 6 sets of 247 observations. The correlations of these six residuals are

	E93	E94	E95	E96	E97	E98
E93	1.00000	.71657	.65440	.63681	.60213	.56930
E94	.71657	1.00000	.79441	.69008	.63613	.57305
E95	.65440	.79441	1.00000	.77181	.68184	.62782
E96	.63681	.69008	.77181	1.00000	.67485	.66965
E97	.60213	.63613	.68184	.67485	1.00000	.74868
E98	.56930	.57305	.62782	.66965	.74868	1.00000

These are obviously not zero. Treating each year as a separate equation, suggest how the information here could be used to construct a more efficient estimator than equation by equation least squares. Show the estimator you propose to use in complete detail.

Regressions for Part 3

```
--> include;new;year93=1$
```

LHS=YIT	Mean	=	11.37137
	Standard deviation	=	.5741918
WTS=none	Number of observs.	=	247
Model size	Parameters	=	5
	Degrees of freedom	=	242
Residuals	Sum of squares	=	4.828938
	Standard error of e	=	.1412597
Fit	R-squared	=	.9404609
	Adjusted R-squared	=	.9394767

Variable	Coefficient	Standard Error	t-ratio	P[T >t]	Mean of X
Constant	11.5451542	.01020751	1131.045	.0000	
X1	.67127435	.04640328	14.466	.0000	-.10034736
X2	-.02051104	.02652538	-.773	.4401	-.03237374
X3	.02515508	.03348824	.751	.4533	.00367797
X4	.38883467	.02677635	14.522	.0000	-.27565075

```
--> include;new;year94=1$
```

LHS=YIT	Mean	=	11.47123
	Standard deviation	=	.5935680
WTS=none	Number of observs.	=	247
Model size	Parameters	=	5
	Degrees of freedom	=	242
Residuals	Sum of squares	=	4.736023
	Standard error of e	=	.1398940
Fit	R-squared	=	.9453566
	Adjusted R-squared	=	.9444534

Variable	Coefficient	Standard Error	t-ratio	P[T >t]	Mean of X
Constant	11.5692455	.00921237	1255.839	.0000	
X1	.62146662	.04577958	13.575	.0000	-.05794187
X2	.00884371	.02699075	.328	.7435	-.02396359
X3	.03644040	.03192103	1.142	.2548	-.00850757
X4	.42051389	.02749437	15.295	.0000	-.14620599

```
--> include;new;year95=1$
```

LHS=YIT	Mean	=	11.55324
	Standard deviation	=	.6240203
WTS=none	Number of observs.	=	247
Model size	Parameters	=	5
	Degrees of freedom	=	242
Residuals	Sum of squares	=	5.064954
	Standard error of e	=	.1446705
Fit	R-squared	=	.9471259
	Adjusted R-squared	=	.9462520

Variable	Coefficient	Standard Error	t-ratio	P[T >t]	Mean of X
Constant	11.5855294	.00923661	1254.306	.0000	
X1	.60484401	.04837102	12.504	.0000	-.03079058
X2	.04572600	.02771820	1.650	.1003	-.00774750
X3	-.00435116	.03291678	-.132	.8949	-.00850784
X4	.43360824	.02709758	16.002	.0000	-.03078562

```
--> include;new;year96=1$
```

LHS=YIT	Mean	=	11.63496
	Standard deviation	=	.6479655
WTS=none	Number of observs.	=	247
Model size	Parameters	=	5
	Degrees of freedom	=	242
Residuals	Sum of squares	=	4.729280
	Standard error of e	=	.1397944
Fit	R-squared	=	.9542115
	Adjusted R-squared	=	.9534547

Variable	Coefficient	Standard Error	t-ratio	P[T >t]	Mean of X
Constant	11.5821246	.00900367	1286.378	.0000	
X1	.58637409	.04892288	11.986	.0000	.02860376
X2	.03589870	.02837273	1.265	.2070	-.01317060
X3	.00405593	.03224129	.126	.9000	.00181355
X4	.47513989	.02686625	17.685	.0000	.07687632

```
--> include;new;year97=1$
```

LHS=YIT	Mean	=	11.68610
	Standard deviation	=	.6578417
WTS=none	Number of observs.	=	247
Model size	Parameters	=	5
	Degrees of freedom	=	242
Residuals	Sum of squares	=	4.391745
	Standard error of e	=	.1347134
Fit	R-squared	=	.9587467
	Adjusted R-squared	=	.9580648

Variable	Coefficient	Standard Error	t-ratio	P[T >t]	Mean of X
Constant	11.5788492	.00884093	1309.687	.0000	
X1	.56735320	.04922175	11.526	.0000	.06587095
X2	.03422551	.02785101	1.229	.2203	.02440856
X3	.06784837	.03052456	2.223	.0272	-.00173056
X4	.46078641	.02686617	17.151	.0000	.15009731

```
--> include;new;year98=1$
```

LHS=YIT	Mean	=	11.74802
	Standard deviation	=	.6852052
WTS=none	Number of observs.	=	247
Model size	Parameters	=	5
	Degrees of freedom	=	242
Residuals	Sum of squares	=	4.697669
	Standard error of e	=	.1393264
Fit	R-squared	=	.9593270
	Adjusted R-squared	=	.9586547

Variable	Coefficient	Standard Error	t-ratio	P[T >t]	Mean of X
Constant	11.5844443	.00944006	1227.157	.0000	
X1	.53130880	.05112715	10.392	.0000	.09460511
X2	.04776864	.02783742	1.716	.0874	.05284686
X3	.03305216	.03153783	1.048	.2957	.01325446
X4	.48898533	.02854318	17.131	.0000	.22566872

Part 4.

We now return to the panel data set examined in question 1. The results below show OLS, fixed effects and random effects estimates for the Cobb-Douglas model.

- Test the hypothesis of ‘no effects’ vs. ‘some effects’ using the results given below.
- Explain in precise detail the difference between the fixed and random effects model.
- Carry out the Hausman test for fixed effects vs. random effects and report your conclusion. Carefully explain what you are doing in this test. Hint: Transcribing and entering the matrices may be a pain. Here are some matrix commands for LIMDEP – these should also be easily transportable to Stata or Matlab as well – that should make life a little easier. Also, do note, if you carry out your test using either Stata or LIMDEP’s automatic procedures for the panel data models, you will get a somewhat different answer for the statistic. The reason is that you are not using all the internal digits of the computed matrices when you do it using this ‘hint’ while you probably are using the internal procedures.

```
matrix:vfe=[
    .00061      ,-.7143086D-04 , -.2953098D-04,      -.00020/
    -.7143086D-04,      .00026      , -.5828137D-05,-.1381531D-04/
    -.2953098D-04, -.5828137D-05,      .00054      , .6528495D-05/
    -.00020      ,-.1381531D-04 , .6528495D-05,      .00014]$
matrix:vre=[
    .00043      ,-.9201395D-04 ,-.6666412D-04,      -.00017/
    -.9201395D-04,      .00018      ,-.4320561D-05, .8595354D-06/
    -.6666412D-04, -.4320561D-05,      .00030      , -.2862694D-05/
    -.00017      , .8595354D-06 ,-.2862694D-05,      .00012      ]$
matrix:bfe=[
    .66200103/
    .03735244/
    .03039947/
    .38251038]$
matrix:bre=[
    .65025754/
    .03004298/
    .03506960/
    .39954471]$
```

- In the context of the fixed effects model, test the hypothesis that there are no effects – i.e., that all individuals have the same constant term. (The statistics you need to carry out the test are given in the results.)
- In the second set of results, we have added a set of year dummy variables, YEAR93, etc., to the model to allow for time variation as well as for variation across farms. Test the hypothesis that there is no separate time variation using these and the first set of results.

Regression Results for Part 4.

--> regress ; lhs=yit;rhs=cobddgls;panel;pds=6;pri\$

+-----+-----+-----+-----+-----+-----+					
OLS Without Group Dummy Variables					
Ordinary least squares regression					
Model was estimated Dec 07, 2005 at 11:22:00AM					
LHS=YIT	Mean	=	11.57749		
	Standard deviation	=	.6434377		
WTS=none	Number of observs.	=	1482		
Model size	Parameters	=	5		
	Degrees of freedom	=	1477		
Residuals	Sum of squares	=	29.09570		
	Standard error of e	=	.1403538		
Fit	R-squared	=	.9525473		
	Adjusted R-squared	=	.9524188		
Model test	F[4, 1477] (prob)	=	7412.19 (.0000)		
Diagnostic	Log likelihood	=	809.6761		
	Restricted(b=0)	=	-1448.908		
	Chi-sq [4] (prob)	=	4517.17 (.0000)		
	Akaike Info. Criter.	=	-3.923810		
+-----+-----+-----+-----+-----+-----+					
+-----+-----+-----+-----+-----+-----+					
Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
+-----+-----+-----+-----+-----+-----+					
X1	.59517558	.01958331	30.392	.0000	-.188987D-14
X2	.02305014	.01122274	2.054	.0400	-.272590D-14
X3	.02319244	.01303099	1.780	.0751	.124737D-14
X4	.45175783	.01078465	41.889	.0000	.779238D-14
Constant	11.5774868	.00364586	3175.515	.0000	
+-----+-----+-----+-----+-----+-----+					
Least Squares with Group Dummy Variables					
Model was estimated Dec 07, 2005 at 11:22:00AM					
LHS=YIT	Mean	=	11.57749		
	Standard deviation	=	.6434377		
WTS=none	Number of observs.	=	1482		
Model size	Parameters	=	251		
	Degrees of freedom	=	1231		
Residuals	Sum of squares	=	8.161094		
	Standard error of e	=	.8142264E-01		
Fit	R-squared	=	.9866899		
	Adjusted R-squared	=	.9839868		
Model test	F[250, 1231] (prob)	=	365.02 (.0000)		
Diagnostic	Log likelihood	=	1751.644		
	Restricted(b=0)	=	-1448.908		
	Chi-sq [250] (prob)	=	6401.11 (.0000)		
+-----+-----+-----+-----+-----+-----+					
+-----+-----+-----+-----+-----+-----+					
Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
+-----+-----+-----+-----+-----+-----+					
X1	.66200103	.02467845	26.825	.0000	-.188987D-14
X2	.03735244	.01613309	2.315	.0206	-.272590D-14
X3	.03039947	.02320776	1.310	.1902	.124737D-14
X4	.38251038	.01201690	31.831	.0000	.779238D-14

Matrix Cov.Mat. has 4 rows and 4 columns.

```

+-----+
1 |      .00061   -.7143086D-04  -.2953098D-04   -.00020
2 | -.7143086D-04   .00026   -.5828137D-05  -.1381531D-04
3 | -.2953098D-04  -.5828137D-05   .00054   .6528495D-05
4 |   -.00020   -.1381531D-04  .6528495D-05   .00014
+-----+

Test Statistics for the Classical Model
+-----+
Model      Log-Likelihood      Sum of Squares      R-squared
(1) Constant term only      -1448.90832      .6131518321D+03      .0000000
(2) Group effects only      412.25944      .4974526192D+02      .9188696
(3) X - variables only      809.67611      .2909570093D+02      .9525473
(4) X and group effects      1751.64437      .8161093811D+01      .9866899
+-----+

Hypothesis Tests
+-----+
Likelihood Ratio Test
Chi-squared      d.f.      Prob.
(2) vs (1)      3722.336      246      .00000
(3) vs (1)      4517.169      4      .00000
(4) vs (1)      6401.105      250      .00000
(4) vs (2)      2678.770      4      .00000
(4) vs (3)      1883.937      246      .00000

F Tests
num. denom. Prob value
(2) vs (1)      56.859      246      1235      .00000
(3) vs (1)      7412.185      4      1477      .00000
(4) vs (1)      365.021      250      1231      .00000
(4) vs (2)      1568.114      4      1231      .00000
(4) vs (3)      12.836      246      1231      .00000
+-----+

Random Effects Model: v(i,t) = e(i,t) + u(i)
Estimates: Var[e] = .662965D-02
Var[u] = .130695D-01
Corr[v(i,t),v(i,s)] = .663456
Lagrange Multiplier Test vs. Model (3) = 1582.16
( 1 df, prob value = .000000)
(High values of LM favor FEM/REM over CR model.)
Fixed vs. Random Effects (Hausman) = ???
( 4 df, prob value = .014557)
(High (low) values of H favor FEM (REM).)
Sum of Squares .296237D+02
+-----+

+-----+
|Variable | Coefficient | Standard Error | b/St.Er. | P[|Z|>z] | Mean of X |
+-----+
X1      .65025754      .02082134      31.230      .0000      -.188987D-14
X2      .03004298      .01334238      2.252      .0243      -.272590D-14
X3      .03506960      .01732942      2.024      .0430      .124737D-14
X4      .39954471      .01084665      36.836      .0000      .779238D-14
Constant 11.5774868      .00757539      1528.302      .0000
Matrix Cov.Mat. has 5 rows and 5 columns.
+-----+
1 |      .00043   -.9201395D-04  -.6666412D-04   -.00017      0
2 | -.9201395D-04   .00018   -.4320561D-05   .8595354D-06      0
3 | -.6666412D-04  -.4320561D-05   .00030   -.2862694D-05      0
4 |   -.00017   .8595354D-06  -.2862694D-05   .00012      0
5 |      0      0      0      0      .5738657D-04
+-----+

```

+-----+-----+-----+-----+-----+-----+-----+						
OLS Without Group Dummy Variables						
Model was estimated Dec 07, 2005 at 11:24:55AM						
LHS=YIT	Mean	=	11.57749			
	Standard deviation	=	.6434377			
WTS=none	Number of observs.	=	1482			
Model size	Parameters	=	10			
	Degrees of freedom	=	1472			
Residuals	Sum of squares	=	28.91153			
	Standard error of e	=	.1401463			
Fit	R-squared	=	.9528477			
	Adjusted R-squared	=	.9525594			
Model test	F[9, 1472] (prob)	=	3305.11 (.0000)			
Diagnostic	Log likelihood	=	814.3815			
	Restricted(b=0)	=	-1448.908			
	Chi-sq [9] (prob)	=	4526.58 (.0000)			
+-----+-----+-----+-----+-----+-----+-----+						
Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X	
+-----+-----+-----+-----+-----+-----+-----+						
X1	.60073303	.01964659	30.577	.0000	-.188987D-14	
X2	.02358238	.01121736	2.102	.0355	-.272590D-14	
X3	.02727236	.01310340	2.081	.0374	.124737D-14	
X4	.44443130	.01107955	40.113	.0000	.779238D-14	
YEAR93	-.03446509	.01301167	-2.649	.0081	.16666667	
YEAR94	-.01746941	.01281015	-1.364	.1727	.16666667	
YEAR95	-.00345211	.01269074	-.272	.7856	.16666667	
YEAR96	-.00541532	.01264493	-.428	.6685	.16666667	
YEAR97	-.00999072	.01262022	-.792	.4286	.16666667	
Constant	11.5892856	.00901935	1284.935	.0000		
+-----+-----+-----+-----+-----+-----+-----+						
Least Squares with Group Dummy Variables						
Model was estimated Dec 07, 2005 at 11:24:55AM						
LHS=YIT	Mean	=	11.57749			
	Standard deviation	=	.6434377			
WTS=none	Number of observs.	=	1482			
Model size	Parameters	=	256			
	Degrees of freedom	=	1226			
Residuals	Sum of squares	=	7.379538			
	Standard error of e	=	.7758349E-01			
Fit	R-squared	=	.9879646			
	Adjusted R-squared	=	.9854613			
Model test	F[255, 1226] (prob)	=	394.67 (.0000)			
Diagnostic	Log likelihood	=	1826.239			
	Restricted(b=0)	=	-1448.908			
	Chi-sq [255] (prob)	=	6550.29 (.0000)			
+-----+-----+-----+-----+-----+-----+-----+						
Panel:Groups	Empty	0,	Valid data	247		
	Smallest	6,	Largest	6		
	Average group size			6.00		
+-----+-----+-----+-----+-----+-----+-----+						
Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X	
+-----+-----+-----+-----+-----+-----+-----+						
X1	.63796531	.02379854	26.807	.0000	-.188987D-14	
X2	.04127557	.01544463	2.672	.0075	-.272590D-14	
X3	.02819226	.02217322	1.271	.2036	.124737D-14	
X4	.30816028	.01322571	23.300	.0000	.779238D-14	
YEAR93	-.09400525	.00892438	-10.534	.0000	.16666667	
YEAR94	-.06108644	.00813834	-7.506	.0000	.16666667	
YEAR95	-.03263851	.00761097	-4.288	.0000	.16666667	
YEAR96	-.02205545	.00721130	-3.058	.0022	.16666667	
YEAR97	-.01870213	.00704283	-2.655	.0079	.16666667	

Test Statistics for the Classical Model							
Model	Log-Likelihood	Sum of Squares	R-squared				
(1) Constant term only	-1448.90832	.6131518321D+03	.0000000				
(2) Group effects only	412.25944	.4974526192D+02	.9188696				
(3) X - variables only	814.38155	.2891152515D+02	.9528477				
(4) X and group effects	1826.23878	.7379537558D+01	.9879646				
Hypothesis Tests							
Likelihood Ratio Test			F Tests				
	Chi-squared	d.f.	Prob.	F	num.	denom.	Prob value
(2) vs (1)	3722.336	246	.00000	56.859	246	1235	.00000
(3) vs (1)	4526.580	9	.00000	3305.109	9	1472	.00000
(4) vs (1)	6550.294	255	.00000	394.667	255	1226	.00000
(4) vs (2)	2827.959	9	.00000	782.048	9	1226	.00000
(4) vs (3)	2023.714	246	.00000	14.542	246	1226	.00000

Random Effects Model: v(i,t) = e(i,t) + u(i)							
Estimates: Var[e]		=	.601920D-02				
Var[u]		=	.136218D-01				
Corr[v(i,t),v(i,s)]		=	.693539				
Lagrange Multiplier Test vs. Model (3) = 1621.50							
(1 df, prob value = .000000)							
(High values of LM favor FEM/REM over CR model.)							
Baltagi-Li form of LM Statistic		=	1621.50				
Fixed vs. Random Effects (Hausman)		=	.00				
<-- Not computable							
(9 df, prob value = 1.000000)							
(High (low) values of H favor FEM (REM).)							
Sum of Squares			.310345D+02				
R-squared			.949385D+00				

Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X		
X1	.66230811	.02020885	32.773	.0000	-.188987D-14		
X2	.03781071	.01301123	2.906	.0037	-.272590D-14		
X3	.05518575	.01709150	3.229	.0012	.124737D-14		
X4	.35287326	.01159807	30.425	.0000	.779238D-14		
YEAR93	-.06688085	.00805543	-8.303	.0000	.16666667		
YEAR94	-.04042410	.00756342	-5.345	.0000	.16666667		
YEAR95	-.01774170	.00725994	-2.444	.0145	.16666667		
YEAR96	-.01371576	.00708616	-1.936	.0529	.16666667		
YEAR97	-.01431767	.00700864	-2.043	.0411	.16666667		
Constant	11.6030002	.00908962	1276.512	.0000			

Part 5

This question is based on the original Mroz data used in his 1975 study of female labor supply. In the following model, we analyze the number of children in the family household using a Poisson regression model. The model is

$$\text{Prob}[N_{\text{kids}} = K_i] = \exp(-\lambda_i) \lambda_i^{K_i} / K_i!$$

where

$$\lambda_i = \exp(\beta_1 + \beta_2 \text{AGE}_i + \beta_3 \text{AGE}_i^2 + \beta_4 \text{WE}_i + \beta_5 \text{INCOME}_i)$$

K_i = the number of kids

Age = age in years

Age2 = age²

WE = wife's education in years

INCOME = family income in \$10,000

Maximum likelihood Poisson regression results appear below.

- Test the hypothesis that the number of children is unrelated to AGE using a Wald test.
- Compute the marginal effect of an additional year in age on the expected number of kids.
- Prove that the sample mean of the estimated λ_i s (that is, the estimates of λ_i when you plug in the data and the maximum likelihood estimates of the parameters) equals the sample mean of K_i . (Note, this is a common result in 'loglinear' models such as this.)
- Carry out a likelihood ratio test of the hypothesis that the four coefficients on AGE, AGE2, WE and INCOME are all zero.
- Show exactly how to compute a Lagrange multiplier test statistic for testing the hypothesis that the coefficient on HA, the husband's age, is zero. Note that HA is not in the model, and I want to know if it has been inappropriately omitted. When I do this test, the actual test value that is computed is 5.873. Should the hypothesis that the coefficient on HA in this model is zero be rejected? Explain your answer precisely.

```

+-----+
| Poisson Regression                               |
| Model estimated: Nov 30, 2004 at 04:42:02PM.    |
| Dependent variable          NKIDS               |
| Number of observations      753                 |
| Iterations completed        7                  |
| Log likelihood function     -1083.397           |
| Number of parameters        5                  |
| Restricted log likelihood    -1279.522          |
+-----+
+-----+-----+-----+-----+-----+-----+
|Variable | Coefficient | Standard Error | b/St.Er. | P[|Z|>z] | Mean of X |
+-----+-----+-----+-----+-----+-----+
| Constant | -7.64180956 | 1.14268278    | -6.688   | .0000    | 42.5378486 |
| AGE      | .49624655   | .05663388     | 8.762    | .0000    | 1874.54847 |
| AGE2     | -.00686403  | .00069963     | -9.811   | .0000    | 12.2868526 |
| WE       | -.03430021  | .01448182     | -2.369   | .0179    | 2.30805950 |
| INCOME   | .01193400   | .02569902     | .464     | .6424    |              |
+-----+-----+-----+-----+-----+-----+
Matrix Cov.Mat. has 5 rows and 5 columns.
      1      2      3      4      5
+-----+-----+-----+-----+-----+
1 | 1.30572   -.06373   .00078   -.00319   .00284
2 | -.06373   .00321   -.3948059D-04 .3794861D-04 -.00012
3 | .00078    -.3948059D-04 .4894781D-06 -.3460068D-06 .1216040D-05
4 | -.00319   .3794861D-04 -.3460068D-06 .00021   -.00014
5 | .00284    -.00012   .1216040D-05 -.00014   .00066

```

Part 6

In homework 8, you examined a model that might be used for lifetimes of electric or electronic parts, the exponential regression model,

$$f(y_i|x_i, \alpha, \gamma) = \theta_i \exp(-\theta_i y_i), \theta_i = \exp(\alpha + \gamma x_i).$$

The regression aspect of this model emerges when we note that $E[y_i|x_i] = 1/\theta_i$.

a. Though we estimated the parameters of the model by using maximum likelihood in homework 8, we could also have used nonlinear least squares. The nonlinear least squares estimates, with the MLEs are shown below. They are, of course, similar. Noting that neither is actually more difficult to compute than the other, is there a statistical reason to prefer one estimator or the other? Explain.

User Defined Optimization					
Nonlinear least squares regression					
Model was estimated Dec 07, 2005 at 02:29:21PM					
LHS=YI	Mean	=	.5717307		
	Standard deviation	=	.6128446		
WTS=none	Number of observs.	=	50		
Model size	Parameters	=	2		
	Degrees of freedom	=	48		
Residuals	Sum of squares	=	17.78242		
	Standard error of e	=	.5963626		
Fit	R-squared	=	.3373985E-01		
	Adjusted R-squared	=	.5306506E-01		
Model test	F[1, 48] (prob)	=	1.68 (.2016)		
Diagnostic	Log likelihood	=	-45.10161		
	Restricted(b=0)	=	-45.95966		
	Chi-sq [1] (prob)	=	1.72 (.1902)		

Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	

AL	.17403168	.26918380	.647	.5179	
CL	.77328951	.54264646	1.425	.1541	
Maximum Likelihood Estimates					

Exponential (Loglinear) Regression Model					
Maximum Likelihood Estimates					
Model estimated: Dec 08, 2005 at 03:58:46PM.					
Dependent variable		YI			
Weighting variable		None			
Number of observations		50			
Iterations completed		5			
Log likelihood function		-21.31650			
Number of parameters		2			
Info. Criterion: AIC =		.93266			
Finite Sample: AIC =		.93777			
Info. Criterion: BIC =		1.00914			
Info. Criterion:HQIC =		.96178			
Restricted log likelihood		-22.04564			
Chi squared		1.458283			
Degrees of freedom		1			
Prob[ChiSq > value] =		.2272035			

Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X

Parameters in conditional mean function					
Constant	.25008087	.30543726	.819	.4129	
XI	.61138675	.51150525	1.195	.2320	.52927082

b. The model is heteroscedastic. $\text{Var}[y_i|x_i] = 1/\theta_i^2$. Could you improve on the nonlinear least squares estimator with this knowledge? Show how to do generalized least squares in this model.

c. As a model for lifetimes, the exponential model has a number of shortcomings. The most oft noted is its property of ‘lack of memory.’ Regardless of how long the part has lasted (y_i), the probability it will fail in

the next interval Δy_i is the same. A common alternative to the model is the gamma model we examined in class. A slightly simpler one is the Weibull model, which adds a scale parameter, P , to the model;

$$f(y_i|x_i, P, \alpha, \gamma) = P \theta_i y_i^{P-1} \exp(-\theta_i y_i^P), \theta_i = \exp(\alpha + \gamma x_i), P > 0.$$

The exponential model is the special case, with $P = 1$. Estimates of the parameters of this model using the data from Assignment 8 are shown below:

+-----+-----+-----+-----+-----+-----+					
Weibull (Loglinear) Regression Model					
Maximum Likelihood Estimates					
Model estimated: Dec 07, 2005 at 02:39:27PM.					
Dependent variable		YI			
Weighting variable		None			
Number of observations		50			
Iterations completed		9			
Log likelihood function		-21.31393			
Number of parameters		3			
Info. Criterion: AIC =		.97256			
Finite Sample: AIC =		.98299			
Info. Criterion: BIC =		1.08728			
Info. Criterion:HQIC =		1.01624			
Restricted log likelihood		-22.04564			
Chi squared		1.463416			
Degrees of freedom		2			
Prob[ChiSqd > value] =		.4810867			
+-----+-----+-----+-----+-----+-----+					
+-----+-----+-----+-----+-----+-----+					
Variable	Coefficient	Standard Error	b/St.Er.	P[Z >z]	Mean of X
+-----+-----+-----+-----+-----+-----+					
Parameters in conditional mean function					
Constant	.24858901	.31142980	.798	.4247	
XI	.61583268	.50392591	1.222	.2217	.52927082
Scale parameter for Weibull model					
P_scale	1.00772871	.16215101	6.215	.0000	

- (1) Derive the log likelihood and the likelihood equations for estimation of α , γ and P for the Weibull model.
- (2) Sketch a proof of the consistency of the maximum likelihood estimator. Note, this can be in general terms, as your results will include this model, since it does satisfy the regularity conditions.
- (3) Sketch a proof of the asymptotic normality of this maximum likelihood estimator.
- (4) Derive the asymptotic covariance matrix for the MLE of (α, γ, P) . Derive the BHHH estimator of the asymptotic covariance matrix.
- (4) Test the hypothesis of the exponential model (null) against the Weibull (alternative) using the results given here. Use a Wald test and a likelihood ratio test. Be explicit about how you are doing your computations.

Part 7

In the Bertschek and Lechner paper discussed in class, the model analyzed by the authors was

$$y_{it}^* = \beta'x_{it} + \varepsilon_{it}, t = 1, \dots, 5 \text{ and } i = 1, \dots, N$$

$$y_{it} = 1 \text{ if } y_{it}^* > 0$$

where $\varepsilon_{i1}, \varepsilon_{i2}, \dots, \varepsilon_{i5}$ have a 5-variate normal distribution with means zero, variances 1, and full correlation matrix, R . They would like to do maximum likelihood estimation of β and the $4(5)/2 = 10$ free correlation coefficients in R . However, this will involve doing 5 dimensional integration of the normal distribution, which is (was) not technologically possible for them. They proposed, instead, to develop a GMM estimator that will make it unnecessary to estimate R altogether. The idea is this: Note, first of all, this is a probit model. In every period, for every observation,

$$\text{Prob}[y_{it} = 1 | x_{it}] = \Phi(\beta'x_{it}), \text{Prob}[y_{it} = 0 | x_{it}] = 1 - \Phi(\beta'x_{it})$$

where $\Phi(\beta'x_{it})$ is the standard normal CDF. They could just pool the data and use maximum likelihood to estimate β . This would be consistent, but would waste a large amount of information. They do have the following moment equations:

$$E[(y_{it} - \Phi(\beta'x_{it})) x_{is}] = 0, s, t = 1, \dots, 5.$$

Note that this is actually 25 sets of moment equations, because, for example, in period 1, $(y_{i1} - \Phi(\beta'x_{i1}))$ is uncorrelated with (orthogonal to) $x_{i1}, x_{i2}, \dots, x_{i5}$. The same is true for periods 2 through 5. So, suppose there are K regressors. In each period, there are $5K$ moment conditions, and there are 5 periods. So, this provides $5 \times 5 \times K$ moment conditions for estimating K parameters. The parameter vector is vastly overidentified.

Explain how to use this model to obtain GMM estimators of the model parameters. Be precise and detailed on the computations that you will do. Include in your description exactly what computations you will do to obtain the estimator and also how you will estimate the asymptotic covariance matrix for your estimator

Part 8

This question involves some “library” research. (You can do it on the web, of course.) Locate an empirical (applied) paper (study) in any field (political science, economics, finance, management, accounting, pharmacology, environment, etc.) that is an application of a discrete choice model – Poisson, probit or other binary choice, multinomial logit, ordered probit, or something else if you prefer. Report (a) what empirical issue the study was about; (b) what the model was; (c) what estimation technique the author used; (d) (briefly) what results they obtained. In part (d), describe the actual statistics that the author reported, and what conclusion they drew. This entire essay should not exceed one double spaced page.