

# **ECONOMETRICS I**

## **Take Home Final Examination**

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**Fall 2016**

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Today is Thursday, December 8, 2016. This exam is due by 3PM, Friday, December 16, 2016. You may submit your answers to me electronically as an attachment to an e-mail if you wish. Please do not include a copy of the exam questions with your submission; submit only your answers to the questions. Your submission for this examination is to be a single authored project. You are assumed to be working alone.

<p>NOTE: In the empirical results below, a number of the form .nnnnnnE+aa means multiply the number .nnnnnn by 10 to the aa power. E-aa implies multiply 10 to the minus aa power. Thus, .123456E-04 is 0.0000123456.</p>
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1. Properties of the least squares estimator
  - a. Show (algebraically) how the ordinary least squares coefficient estimator,  $\mathbf{b}$ , and the estimated asymptotic covariance matrix are computed.
  - b. What are the *finite sample* properties of this estimator? Make your assumptions explicit.
  - c. What are the *asymptotic properties* of the least squares estimator? Again, be explicit about all assumptions, and explain your answer carefully.
  - d. How would you compare the properties of the least absolute deviations (LAD) estimator to those of the ordinary least squares (OLS) estimator? Which is a preferable estimator?
2. The paper, Farsi, M, M. Filippini, and W. Greene, "Efficiency Measurement in Network Industries, Application to the Swiss Railroads," *Journal of Regulatory Economics*, 28, 1, 2005, pp. 69-90 is an analysis of an unbalanced panel of data on 50 railroads for 13 years, 605 observations in total. The variables in the data set are

ct	= total cost
q	= total output, sum of freight, passenger and mail
pe	= price of electricity
pk	= price of capital
pl	= price of labor
narrow	= dummy for narrow gauge track
tunnel	= dummy variable for long tunnels on routes
rack	= dummy variable for a certain track configuration

I propose first to analyze the cost data with a loglinear model. My first model is

$$\ln c_{it} = \beta_1 + \beta_2 \ln q_{it} + \beta_3 \ln p_{e_{it}} + \beta_4 \ln p_{l_{it}} + \beta_5 \ln p_{k_{it}} + \beta_6 \text{narrow}_i + \beta_7 \text{tunnel}_i + \beta_8 \text{rack}_i + \varepsilon_{it}, \quad \varepsilon_{it} \sim N[0, \sigma^2],$$

where "i" indicates the railroad and "t" indicates the year. Note that some variables are time invariant. For this application, I intend to ignore any panel data aspects of the data set, and treat the whole thing as a cross section of 605 observations. The ordinary least squares results are shown as Regression 1 on page 4.

- a. Show how each of the values in the box above the coefficient estimates is computed, and interpret the value given.
- b. Using the results given, form a confidence interval for the true value of the coefficient on the RACK dummy variable.
- c. An expanded, now nonlinear model appears as follows:

$$\ln c_{it} = \beta_1 + \beta_2 \ln q_{it} + \beta_3 \ln^2 q_{it} + \beta_4 \ln p_{e_{it}} + \beta_5 \ln p_{l_{it}} + \beta_6 \ln p_{k_{it}} + \beta_7 \text{narrow}_i + \beta_8 \text{tunnel}_i + \beta_9 \text{rack}_i + \beta_{10} \ln q_{it} \times \text{tunnel}_i + \varepsilon_{it}, \quad \varepsilon_{it} \sim N[0, \sigma^2].$$

The second set of results given includes the quadratic specification ( $\beta_3$ ) and an interaction of log output with tunnel ( $\beta_{10}$ ). Test the hypothesis of the linear model as a restriction on the nonlinear model. Do the test in three ways: 1. Use a Wald test to test the hypothesis that the two coefficients in the quadratic terms ( $\beta_3$  and  $\beta_{10}$ ) are zero. 2. Use an F test. 3. Use a likelihood ratio test assuming that the disturbances are normally distributed. The estimated least squares regression for this model is shown on page 5 with the estimated asymptotic covariance matrix.

- d. I am interested economies of scale for railroads. In the loglinear equation (regression 1), that quantity is

$$\Delta = \partial \ln ct / \partial \ln q = \beta_2$$

In the second model (regression 2), the measure is a linear function of  $\ln q$  and tunnel;

$$\Delta = \beta_2 + 2\beta_3 \ln q + \beta_{10} \text{Tunnel}.$$

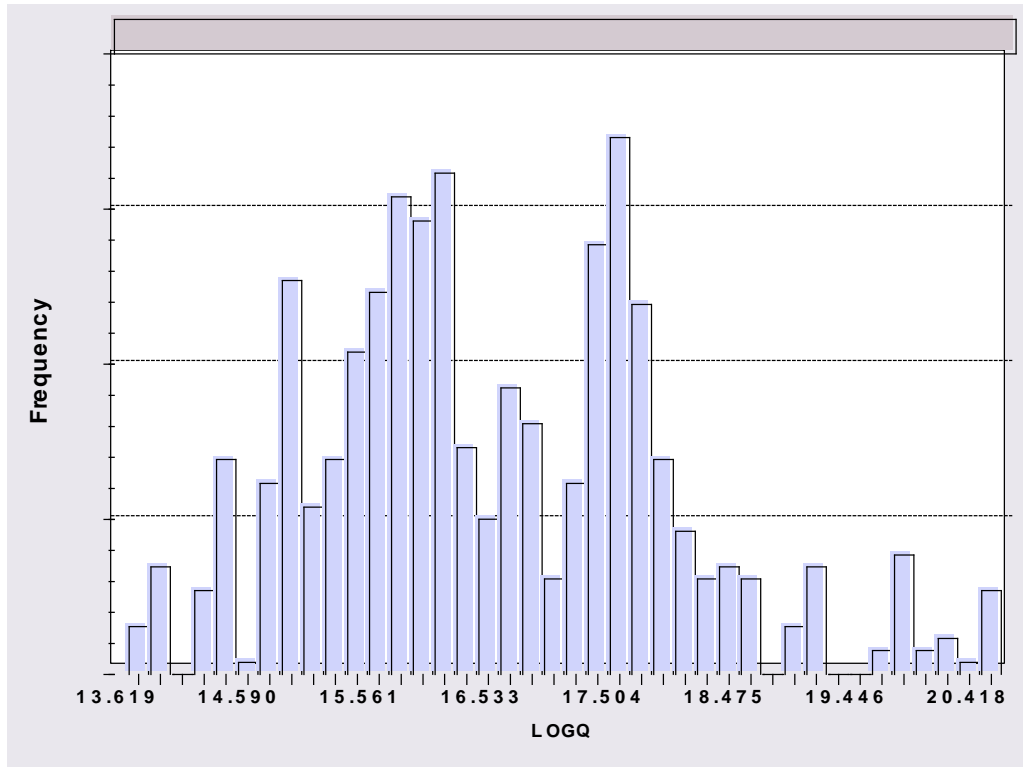
Estimate this value for the average sized railroad with long tunnels (tunnel = 1). (The average of  $\ln q$  is shown in the regression results.) Form a confidence interval for

$$\Delta \text{ (given average } \ln q = 16.4616 \text{ and tunnel} = 1 \text{)}.$$

- e. The efficient scale for a production model is that point where economies of scale equal 1. Assuming that tunnel = 0, the scale elasticity is

$$\Delta = \beta_2 + 2\beta_3 \ln q.$$

Solving for  $\Delta = 1$ , I obtain  $\ln q = (1 - \beta_2)/(2\beta_3)$ . Using the results of regression 2 (and the delta method), form a confidence interval for this function of the parameters. The histogram below shows the distribution of  $\ln q$  in the sample. Locate the point of constant returns to scale in the graph and comment on the efficient size of firm compared to the values found in the sample.



3. The third set of results (on page 6) is computed using White's heteroscedasticity consistent, robust estimator of the covariance matrix.
  - a. How is the White estimator computed?
  - b. Looking at these results, would you conclude that there is evidence of heteroscedasticity in these data?
  
4. The fourth regression is reported (on page 6) with a correction of the standard errors to accommodate the "clustering" in the data – these data are a panel.
  - a. How is the cluster estimator computed?
  - b. Why is it computed; what problem is it intended to solve?
  - c. Compare the results to regression 2 (which is the same model with conventional standard errors). What do you conclude about the effect of "clustering" in these data?

## REGRESSION 1

Ordinary	least squares regression .....				
LHS=LNCT	Mean	=	11.30622		
	Standard deviation	=	1.10169		
	Number of observs.	=	605		
Model size	Parameters	=	8		
	Degrees of freedom	=	597		
Residuals	Sum of squares	=	48.49299		
	Standard error of e	=	.28500		
Fit	R-squared	=	.93385		
	Adjusted R-squared	=	.93308		
Model test	F[ 7, 597] (prob)	=	1204.0(.0000)		
Diagnostic	Log likelihood	=	-95.00555		
	Restricted(b=0)	=	-916.54939		
	Chi-sq [ 7] (prob)	=	1643.1(.0000)		
Info criter.	Akaike Info. Criter.	=	-2.49736		
-----+					
	LNCT	Coefficient	Standard Error	z	Prob. z> Z
					Mean of X
-----+					
Constant	-7.78982***	1.98172	-3.93	.0001	
LNQ	.77200***	.01125	68.64	.0000	16.4616
LNPL	.15615	.17391	.90	.3692	13.2194
LNPE	-.41790**	.18157	-2.30	.0214	-1.85956
LNPK	.34916***	.02891	12.08	.0000	10.1795
NARROW	-.13716***	.02699	-5.08	.0000	.67603
RACK	.48192***	.03098	15.56	.0000	.23471
TUNNEL	-.14985***	.03790	-3.95	.0001	.18843
-----+					

## REGRESSION 2

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-----
Ordinary      least squares regression .....
LHS=LNCT      Mean          =          11.30622
               Standard deviation =          1.10169
               Number of observs. =           605
Model size    Parameters     =           10
               Degrees of freedom =          595
Residuals     Sum of squares  =          38.53713
               Standard error of e =          .25450
Fit           R-squared       =          .94743
               Adjusted R-squared =          .94664
Model test    F[ 9, 595] (prob) = 1191.5(.0000)
Diagnostic     Log likelihood =         -25.49196
               Restricted(b=0)   =         -916.54939
               Chi-sq [ 9] (prob) = 1782.1(.0000)
Info criter.  LogAmemiya Prd. Crt. =         -2.72055
               Akaike Info. Criter. =         -2.72055
               Bayes Info. Criter. =         -2.64773
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```

	Coefficient	Standard Error	z	Prob. z> Z	Mean of X
Constant	14.1048***	3.08775	4.57	.0000	
LNQ	-1.81538***	.28753	-6.31	.0000	16.4616
LNQSQ	.07876***	.00890	8.85	.0000	272.853
LNPL	.14406	.15695	.92	.3587	13.2194
LNPE	-.30467*	.16469	-1.85	.0643	-1.85956
LNPK	.30383***	.02608	11.65	.0000	10.1795
NARROW_T	.00122	.02783	.04	.9651	.67603
RACK	.42526***	.02867	14.83	.0000	.23471
TUNNEL	2.05124***	.75377	2.72	.0065	.18843
LNQTUNL	-.13334***	.04345	-3.07	.0021	3.40727

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-----
9.53418
-0.723063  0.0826737
0.0225023 -0.00255569  7.91249e-005
-0.313888  0.00430309 -0.000147886  0.0246349
-0.282299  0.00320699 -0.000112702  0.023335  0.0271221
-0.0147959  0.000883942 -2.83063e-005  0.000107397  0.000194245  0.00068034
0.0345858 -0.00400616  0.000122773 -5.22415e-005  0.000303053 -0.000125345  0.0007747
-0.0276479  0.00216755 -6.61934e-005  0.000655545  0.000603249  0.000244045 -0.00031242
1.65435 -0.181774  0.00568441 -0.0178937 -0.0196452 -0.00127023 0.00082181
-0.0952809  0.0105601 -0.000330446  0.000968367  0.00105588 7.36731e-005 -0.00047073
0.000300395 -0.0327144 0.00188776
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## REGRESSION 3

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Ordinary	least squares regression .....				
LHS=LNCT	Mean	=	11.30622		
	Standard deviation	=	1.10169		
	Number of observs.	=	605		
Model size	Parameters	=	10		
	Degrees of freedom	=	595		
Residuals	Sum of squares	=	38.53713		
	Standard error of e	=	.25450		
Fit	R-squared	=	.94743		
	Adjusted R-squared	=	.94664		
Model test	F[ 9, 595]	(prob) =	1191.5(.0000)		
White heteroscedasticity robust covariance matrix.					
Br./Pagan LM Chi-sq [ 9]	(prob) =	58.95	(.0000)		
-----					
	LNCT	Coefficient	Standard Error	z	Prob. z> Z
					Mean of X
-----					
Constant	14.1048***	2.94550	4.79	.0000	
LNQ	-1.81538***	.24881	-7.30	.0000	16.4616
LNQSQ	.07876***	.00775	10.16	.0000	272.853
LNPL	.14406	.17091	.84	.3993	13.2194
LNPE	-.30467*	.18467	-1.65	.0990	-1.85956
LNPK	.30383***	.02510	12.10	.0000	10.1795
NARROW_T	.00122	.02850	.04	.9659	.67603
RACK	.42526***	.02564	16.59	.0000	.23471
TUNNEL	2.05124***	.57610	3.56	.0004	.18843
LOGQTUNL	-.13334***	.03355	-3.97	.0001	3.40727
-----					

## REGRESSION 4

+-----					
	Covariance matrix for the model is adjusted for data clustering.				
	Sample of 605 observations contained		50 clusters defined by		
	variable ID		which identifies groups by a cluster ID.		
+-----					
Ordinary	least squares regression .....				
LHS=LNCT	Mean	=	11.30622		
Model size	Parameters	=	10		
+-----					
	LNCT	Coefficient	Standard Error	z	Prob. z> Z
					Mean of X
+-----					
Constant	14.1048*	7.65472	1.84	.0654	
LNQ	-1.81538**	.79951	-2.27	.0232	16.4616
LNQSQ	.07876***	.02490	3.16	.0016	272.853
LNPL	.14406	.33121	.43	.6636	13.2194
LNPE	-.30467	.45559	-.67	.5037	-1.85956
LNPK	.30383***	.07350	4.13	.0000	10.1795
NARROW_T	.00122	.09298	.01	.9895	.67603
RACK	.42526***	.08414	5.05	.0000	.23471
TUNNEL	2.05124	1.78394	1.15	.2502	.18843
LNQTUNL	-.13334	.10411	-1.28	.2003	3.40727
+-----					

5. Munnell, A., "Why has Productivity Declined? Productivity and Public Investment," New England Economic Review, 1990, pp. 3-22, examined the productivity of public capital in a panel of data using the lower 48 states and 17 years. These data are examined at length in Chapter 10 of the 7<sup>th</sup> edition of your text. In this exercise, we will use a very simple version of her model,

$$\log \text{GSP}_{it} = \beta_1 \beta_2 \log \text{PublicK}_{it} + \beta_3 \log \text{PrivateK}_{it} + \beta_4 \log \text{Labor}_{it} + \varepsilon_{it}$$

where GSP is gross state product. Ordinary least squares regression results appear below. KP is public capital; PC is private capital.

- Test the hypothesis that the marginal products of (coefficients on) private and public capital are the same.
- Test the hypothesis of constant returns to scale (that is, the hypothesis that the three coefficients sum to 1.0)
- Test the two hypotheses simultaneously.

```

+-----+
| Ordinary least squares regression |
| LHS=LOGGSP Mean = 10.50885 |
| Standard deviation = 1.021132 |
| WTS=none Number of observs. = 816 |
| Model size Parameters = 4 |
| Degrees of freedom = 812 |
| Residuals Sum of squares = 6.469532 |
| Standard error of e = .8926031E-01 |
| Fit R-squared = .9923871 |
| Adjusted R-squared = .9923589 |
| Model test F[ 3, 812] (prob) =***** (.0000) |
| Diagnostic Log likelihood = 815.7689 |
| Restricted(b=0) = -1174.417 |
| Chi-sq [ 3] (prob) =3980.37(.0000) |
+-----+
+-----+-----+-----+-----+-----+
|Variable| Coefficient | Standard Error |b/St.Er.|P[|Z|>z]| Mean of X|
+-----+-----+-----+-----+-----+
Constant| 1.64886431 | .05833603 | 28.265 | .0000 |
LOGKP | .15078348 | .01735707 | 8.687 | .0000 | 9.67920583
LOGPC | .30553817 | .01037855 | 29.439 | .0000 | 10.5594618
LOGEMP | .59815198 | .01390006 | 43.032 | .0000 | 6.97849785
+-----+
Asymptotic Covariance Matrix
      1      2      3      4
1| .00340
2| -.00059 .00030
3| -.00020 -.00009078 .00011
4| .00064 -.00020 -.000008636 .00019

```

- The three sets of results below show the least squares estimates for two of the states, then the results for these two states combined. (Presumably, these two are representative of the 48 in the data set.)
  - Theory 1 states that the coefficient vectors are the same for the two states. Is there an optimal way that I could combine these two estimators to form a single efficient estimator of the model parameters? How should I do that? Describe the computations in detail.
  - Use a Chow test to test the hypothesis that the two coefficient vectors are the same. Explain the computations in full detail so that I know exactly how you obtained your result.
  - Use a Wald test to test the hypothesis that the coefficients are the same. Again, document your computations.

```

| LHS=LOGGSP   Mean                = 10.53753 |
|              Standard deviation   = .1584103 |
| WTS=none     Number of observs.   = 17      |
| Model size   Parameters           = 4        |
| Residuals    Sum of squares       = .9063141E-02 |
|              Standard error of e  = .2640388E-01 |
| Fit          R-squared             = .9774269 |
|              Adjusted R-squared   = .9722177 |
| Model test   F[ 3, 13] (prob) = 187.64 (.0000) |
+-----+-----+-----+-----+-----+
|Variable| Coefficient | Standard Error | t-ratio | P[|T|>t] | Mean of X |
+-----+-----+-----+-----+-----+
Constant| 6.20005924 | 3.26548334    | 1.899   | .0800    | 9.79136043
LOGKP   | -1.11530796 | .69214743     | -1.611  | .1311    | 10.8133466
LOGPC   | .49706712   | .25806357     | 1.926   | .0762    | 7.13004557
LOGEMP  | 1.38609118  | .30315422     | 4.572   | .0005    |
      1          2          3          4
+-----+-----+-----+-----+
1| 10.66338   -2.24221   .68603   .54315
2| -2.24221   .47907    -.14467  -.12400
3| .68603     -.14467    .06660   .00145
4| .54315     -.12400    .00145   .09190

| LHS=LOGGSP   Mean                = 11.54882 |
| Residuals    Sum of squares       = .2267268E-02 |
|              Standard error of e  = .1320626E-01 |
| Fit          R-squared             = .9972928 |
|              Adjusted R-squared   = .9966681 |
| Model test   F[ 3, 13] (prob) =1596.37 (.0000) |
+-----+-----+-----+-----+
|Variable| Coefficient | Standard Error | t-ratio | P[|T|>t] | Mean of X |
+-----+-----+-----+-----+
Constant| 1.81099728 | 1.00540552    | 1.801   | .0949    | 10.5943058
LOGKP   | .48580792   | .28675566     | 1.694   | .1140    | 11.3937222
LOGPC   | -.24199469  | .13905321     | -1.740  | .1054    | 8.07221529
LOGEMP  | .91031345   | .09078688     | 10.027  | .0000    |
      1          2          3          4
+-----+-----+-----+-----+
1| 1.01084    -.28472    .12393   .07353
2| -.28472    .08223    -.03730  -.02000
3| .12393     -.03730    .01934   .00631
4| .07353     -.02000    .00631   .00824

| LHS=LOGGSP   Mean                = 11.04318 |
| WTS=none     Number of observs.   = 34      |
| Residuals    Sum of squares       = .3297875E-01 |
|              Standard error of e  = .3315557E-01 |
| Fit          R-squared             = .9966796 |
|              Adjusted R-squared   = .9963475 |
| Model test   F[ 3, 30] (prob) =3001.65 (.0000) |
+-----+-----+-----+-----+
|Variable| Coefficient | Standard Error | t-ratio | P[|T|>t] | Mean of X |
+-----+-----+-----+-----+
Constant| 2.63395245 | .82374079     | 3.198   | .0033    | 10.1928331
LOGKP   | -.02655970  | .19357706     | -.137   | .8918    | 11.1035344
LOGPC   | .05983913   | .04814774     | 1.243   | .2236    | 7.60113043
LOGEMP  | 1.05451622  | .17031130     | 6.192   | .0000    |
      1          2          3          4
+-----+-----+-----+-----+
1| .67855     -.14905    -.01781  .13662
2| -.14905    .03747     .00111   -.03227
3| -.01781    .00111     .00232   -.00254
4| .13662     -.03227    -.00254   .02901

```



7. We now return to the panel data set examined in question 2. The results below show OLS, fixed effects and random effects estimates.

- Test the hypothesis of ‘no effects’ vs. ‘some effects’ using the results given below.
- Explain in precise detail the difference between the fixed and random effects model.
- In the context of the fixed effects model, test the hypothesis that there are no effects – i.e., that all individuals have the same constant term. (The statistics you need to carry out the test are given in the results.)
- The variables narrow\_t, rack and tunnel are time invariant. Explain why it is necessary to omit these variables from the equation to compute the fixed effects regression..
- Since there are time invariant variables in the model, the Hausman statistic cannot be computed. What I did instead was compute the group means of the time varying variables (logq, lnpe, lnpl, lnpg) and add them to the model. I then used this regression to compute the Wu statistic to test the hypothesis of fixed versus random effects. The value of the statistic is 106.6.
  - How is the statistic computed?
  - What should I conclude on the basis of the test?

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OLS Without Group Dummy Variables.....					
Ordinary least squares regression .....					
LHS=LNCT	Mean	=	11.30622		
	Standard deviation	=	1.10169		
	Number of observs.	=	605		
Model size	Parameters	=	8		
	Degrees of freedom	=	597		
Residuals	Sum of squares	=	48.49299		
	Standard error of e	=	.28500		
Fit	R-squared	=	.93385		
	Adjusted R-squared	=	.93308		
Model test	F[ 7, 597] (prob)	=	1204.0(.0000)		
Diagnostic	Log likelihood	=	-95.00555		
	Restricted(b=0)	=	-916.54939		
	Chi-sq [ 7] (prob)	=	1643.1(.0000)		
Info criter.	LogAmemiya Prd. Crt.	=	-2.49736		
	Akaike Info. Criter.	=	-2.49736		
	Bayes Info. Criter.	=	-2.43911		
-----					
	LNCT				
	Coefficient	Standard Error	z	Prob. z> Z	Mean of X
-----					
LOGQ	.77200***	.01125	68.64	.0000	16.4616
LNPL	.15615	.17391	.90	.3692	13.2194
LNPE	-.41790**	.18157	-2.30	.0214	-1.85956
LNPK	.34916***	.02891	12.08	.0000	10.1795
NARROW_T	-.13716***	.02699	-5.08	.0000	.67603
RACK	.48192***	.03098	15.56	.0000	.23471
TUNNEL	-.14985***	.03790	-3.95	.0001	.18843
Constant	-7.78982***	1.98172	-3.93	.0001	
-----					

Least Squares with Group Dummy Variables.....

Ordinary least squares regression .....

LHS=LNCT Mean = 11.30622

Standard deviation = 1.10169

Number of observs. = 605

Model size Parameters = 57

Degrees of freedom = 548

Residuals Sum of squares = 3.35773

Standard error of e = .07828

Fit R-squared = .99542

Adjusted R-squared = .99495

Model test F[ 56, 548] (prob) = 2126.7(.0000)

Diagnostic Log likelihood = 712.71615

Restricted(b=0) = -916.54939

Chi-sq [ 56] (prob) = 3258.5(.0000)

Info criter. LogAmemiya Prd. Crt. = -5.00497

Akaike Info. Criter. = -5.00553

Bayes Info. Criter. = -4.59050

Estd. Autocorrelation of e(i,t) = .682787

Panel:Groups Empty 0, Valid data 50

Smallest 1, Largest 13

Average group size in panel 12.10

These 3 variables have no within group variation.

NARROW\_T RACK TUNNEL

F.E. estimates are based on a generalized inverse.

		Standard		Prob.	Mean
LNCT	Coefficient	Error	z	z> Z	of X
LOGQ	.31912***	.02940	10.85	.0000	16.4616
LNPL	.45701***	.06676	6.85	.0000	13.2194
LNPE	-.30126***	.08677	-3.47	.0005	-1.85956
LNPK	.30843***	.01886	16.35	.0000	10.1795
NARROW_T	.000	.....(Fixed Parameter).....			.67603
RACK	.000	.....(Fixed Parameter).....			.23471
TUNNEL	.000	.....(Fixed Parameter).....			.18843

Note: \*\*\*, \*\*, \* ==> Significance at 1%, 5%, 10% level.

Fixed parameter ... is constrained to equal the value or had a nonpositive st.error because of an earlier problem.

Test Statistics for the Classical Model							
Model		Log-Likelihood		Sum of Squares		R-squared	
(1)	Constant term only	-916.54938		733.08869		.00000	
(2)	Group effects only	306.82066		12.84646		.98248	
(3)	X - variables only	-95.00554		48.49299		.93385	
(4)	X and group effects	712.71616		3.35773		.99542	
Hypothesis Tests							
Likelihood Ratio Test				F Tests			
	Chi-squared	d.f.	Prob	F	num	denom	P value
(2) vs (1)	2446.74	49	.0000	635.03	49	555	.00000
(3) vs (1)	1643.09	7	.0000	1204.01	7	597	.00000
(4) vs (1)	3258.53	56	.0000	2126.72	56	548	.00000
(4) vs (2)	811.79	7	.0000	221.23	7	548	.00000
(4) vs (3)	1615.44	49	.0000	150.33	49	548	.00000

```

-----
Random Effects Model: v(i,t)      = e(i,t) + u(i)
Estimates:  Var[e]                =      .006127
            Var[u]                =      .075101
            Corr[v(i,t),v(i,s)] =      .924567
Lagrange Multiplier Test vs. Model (3) =2638.43
( 1 degrees of freedom, prob. value =  .000000)
(High values of LM favor FEM/REM over CR model)
Fixed vs. Random Effects (Hausman) =      .00
( 7 degrees of freedom, prob. value = 1.000000)
(High (low) values of H favor F.E.(R.E.) model)
Sum of Squares                97.926721
R-squared                      .866509
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```

	LNCT	Coefficient	Standard Error	z	Prob. z> Z	Mean of X
LOGQ		.50994***	.02267	22.50	.0000	16.4616
LNPL		.33111***	.06530	5.07	.0000	13.2194
LNPE		-.38440***	.08547	-4.50	.0000	-1.85956
LNPK		.29346***	.01849	15.87	.0000	10.1795
NARROW_T		-.08171	.08643	-.95	.3444	.67603
RACK		.30482***	.09890	3.08	.0021	.23471
TUNNEL		.36262***	.11075	3.27	.0011	.18843
Constant		-5.24152***	.72709	-7.21	.0000	

```

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Random Effects Model: v(i,t)      = e(i,t) + u(i)
Estimates:  Var[e]                =      .006172
            Var[u]                =      .072825
            Corr[v(i,t),v(i,s)] =      .921867
Lagrange Multiplier Test vs. Model (3) =2838.51
( 1 degrees of freedom, prob. value =  .000000)
(High values of LM favor FEM/REM over CR model)
Baltagi-Li form of LM Statistic =      1739.15
Sum of Squares                47.156449
R-squared                      .935674
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```

	LNCT	Coefficient	Standard Error	z	Prob. z> Z	Mean of X
LOGQ		.31912***	.02951	10.81	.0000	16.4616
LNPL		.45701***	.06700	6.82	.0000	13.2194
LNPE		-.30126***	.08709	-3.46	.0005	-1.85956
LNPK		.30843***	.01893	16.29	.0000	10.1795
NARROW_T		-.16251*	.08937	-1.82	.0690	.67603
RACK		.50776***	.10847	4.68	.0000	.23471
TUNNEL		-.20105	.12896	-1.56	.1190	.18843
LOGQB		.45832***	.04880	9.39	.0000	16.4616
LNPLB		.16415	.95577	.17	.8636	13.2194
LNPEB		.32844	.87197	.38	.7064	-1.85956
LNPKB		.05097	.11268	.45	.6511	10.1795
Constant		-13.2818	11.06214	-1.20	.2299	

8. You may use any data set you wish for this exercise. You are going to fit a binary choice model, so choose one that contains an interesting binary variable that you will explain with your binary choice model. The data set may be one of the ones that we used during the semester, or one of the data sets on the data page for your text, or any other data set that you find interesting.

You will need a statistical package to do this part of the exam. You may use NLOGIT, Stata, R, MatLab, Gauss, or any other package you are familiar with.

Your assignment is to estimate a binary choice model using your data. Your report should document the data set, lay out the model you are estimating, and present the results in a professional looking format in a table of results. The entire writeup need not be more than a page or two, but it should document your empirical work as if you were submitting for review by a referee or a colleague in your department.

You may fit a logit or probit model. (We are not interested in the ‘linear probability model.’) In addition to the main assignment above, also do the following:

- In a ‘technical appendix,’ write down the log likelihood for your model. Derive the likelihood equations (first order conditions) for estimation of the model parameters. Derive an asymptotic covariance matrix for your estimator.
  - Using an alternative specification for your equation, carry out a hypothesis test using a likelihood ratio test.
  - Report and interpret the partial effects for your model. Indicate whether you have used the partial effects at the means of the data or the average partial effects.
9. The probit model below examines the probability that an individual reports Health Satisfaction greater than 6 in the 0 – 10 scale for HSAT in the GSOEP. Age10 = AGE/10. RICH is a dummy variable that equals one if the individual’s income is in the top 20% of the incomes in the sample. The results agree with my expectations (with one exception; I would not expect HEALTHY to increase with age, as it . However, I am concerned that RICH may be endogenous in this model – the unobservables that influence health are likely to influence the ability to obtain a high income as well. How can I consistently estimate the model in this case of an endogenous variable in a binary choice model? Once I do, how do I estimate the interesting impact of income on health (i.e., the “treatment effect”)?

Binomial Probit Model						
Dependent variable		HEALTHY				
Log likelihood function		-17476.18255				
Restricted log likelihood		-18279.94994				
Chi squared [ 5](P= .000)		1607.53478				
Significance level		.00000				
McFadden Pseudo R-squared		.0439699				
Estimation based on N = 27326, K = 6						
Inf.Cr.AIC = 34964.4 AIC/N = 1.280						
-----+-----						
HEALTHY	Coefficient	Standard Error	z	Prob.  z >Z*	95% Confidence Interval	
-----+-----						
Index function for probability.....						
Constant	1.73753***	.12904	13.46	.0000	1.48461	1.99044
RICH	.21776***	.02357	9.24	.0000	.17156	.26395
FEMALE	-.14399***	.01572	-9.16	.0000	-.17479	-.11319
MARRIED	.04220**	.01923	2.20	.0282	.00452	.07989
AGE10	-.40707***	.06182	-6.58	.0000	-.52824	-.28591
AGESQ100	.01652**	.00690	2.40	.0166	.00300	.03004
-----+-----						